# Estimates of Cetacean Abundance in the Northern Gulf of Mexico from Vessel Surveys 

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## Introduction

The Southeast Fisheries Science Center (SEFSC) initiated annual, vessel-based visual sampling surveys of northern Gulf of Mexico marine mammals in 1990. The primary goal of these surveys was to meet Marine Mammal Protection Act requirements for estimating abundance and monitoring trends of marine mammal stocks in United States waters. The surveys were designed to collect: 1) marine mammal sighting data to estimate abundance and to determine distribution and diversity; and 2) environmental data to evaluate factors which may affect the distribution, abundance and diversity of marine mammals. The analyses for abundance estimation from the 1991-I 994 surveys are presented in this report.

Survey Methods

The Gulf of Mexico surveys were conducted during the spring-summer period (April-June), lasting from 15 to 55 days. The 1990 and 1991 surveys were the shortest; surveys during 1992-I 994 were all approximately 50 days in length. The 1990 and 1991 surveys were conducted in one leg which sampled the off-shelf waters of the northern Gulf between $83^{\circ}-96^{\circ} \mathrm{W}$ with a survey track similar to that shown in Figure 1. The 1992-I 994 surveys were conducted in three separate legs, with one or two legs similar to the track shown in Figure 1, and one or two legs sampling the area between $87^{\circ}-96^{\circ} \mathrm{W}$ with a survey track similar to that shown in Figure 2. There was a major difference in sampling between the two tracks. The track shown in Figure 1 was based on a pre-determined track for sampling ichthyoplankton stations and was transited 24 hours a day. Daylight transects on this track could be latitudinal or longitudinal. or a combination of both. The track shown in Figure 2 was designed specifically to collect marine mammal sightings along transects perpendicular to the depth gradient. This resulted in visual sampling on only longitudinal transects.

Visual marine mammal sighting data were collected by two teams of three observers during daylight hours, weather permitting (i.e., no rain, Beaufort sea state $<6$ ), utilizing standard vessel survey data collection methods for cetaceans developed by the Southwest Fisheries Science Center (e.g., see Holt and Sexton 1987). Each
team had at least two members experienced in shipboard marine mammal observation and identification techniques. Two observers searched for marine mammals using highpower (25X), large format "Bigeye" binoculars mounted on the ship's flying bridge. The third observer maintained a search of the area near the trackline without visual aids and with handheld binoculars, and recorded data. Sighting data were recorded with a computer data acquisition program and included species, herd-size, bearing and reticle (a measure of radial distance) of a sighting, and data on environmental conditions (i.e., Beaufort sea state, sun position, etc.) which could affect the observers' ability to sight animals. The reticle relative to a sighting was measured using an eyepiece with a graduated scale in the binoculars. The bearing of a sighting relative to the trackline was measured using a 360 " graduated scale attached to the base of the binoculars. Ancillary data also collected included, but were not limited to, time of day, position, behavior, and associated animals. If necessary, the vessel was diverted from the trackline to identify species and obtain herd size estimates.

In general, environmental stations were located every 30 minutes of latitude or longitude along the cruise track. The stations included CTD/STD hydrocasts to just off the bottom, or to 500 m when depth exceeded 500 m . An XBT sample was obtained halfway between the environmental stations. A thermo-salinograph operated throughout the entire cruise; surface water salinity and temperature were recorded every minute of time. Analysis of environmental data and possible relationships with cetacean abundance and distribution are not included in this report.

The sighting and effort data were summarized by survey for the line transect distance sampling analysis. The basic sample unit considered in the analyses was one day's survey effort with associated herd size and sighting distance for each sighting. Effort and sighting data were pooled across environmental conditions which may have had different sighting rates because of effects on observers' abilities to sight animals (i.e., sighting rates tended to decrease as wind and wave height increased).

## Analytical Methods

The abundance of cetacean species sighted during the surveys was estimated using distance sampling analysis for line transect surveys (Buckland et al. 1993). Northern Guff of Mexico abundance was estimated for the 1991-1994 surveys combined, and separately for each survey, for every species of cetacean sighted on effort. The 1990 survey was a pilot survey, and was not used in the analysis due to inconsistencies with the other surveys. Abundance estimates were made using program DISTANCE (Laake et al. 1993). The formula used to estimate density ( $\hat{D}$ ) was

$$
\bar{D}=\frac{n \cdot \hat{s} \cdot \hat{f}(0)}{2 \cdot L}
$$

$$
\begin{aligned}
& \text { where } \quad \begin{array}{l}
n=\text { number of on-effort group sightings } \\
\hat{S}=\text { mean group size or expected group size } \\
\hat{f}(0)=\text { sighting probability density function at } 0 \text { perpendicular distance } \\
L \quad=\text { length of transects sampled within a stratum. }
\end{array} .
\end{aligned}
$$

Abundance was estimated as the density times the size of the survey area, and the lognormal $95 \%$ confidence intervals were computed for each abundance estimate. The northern Gulf of Mexico survey area was considered to be the waters between the 100 m isobath and the EEZ boundary, a total of approximately $398,960 \mathrm{~km}^{2}$.

The parameter $\hat{f}(0)$ was estimated using a hazard-rate model and a half\&normal model (Buckland 1985, Buckland et al. 1993). The $\hat{f}(0)$ parameter was estimated using a maximum likelihood algorithm applied to exact sighting distances. Model selection of $\hat{f}(0)$ was determined using Akaike's Information Criterion (AIC, see Buckland et al. 1993).

The length of line sampled was determined using LORAN positions (latitude and longitude) collected at regular intervals (usually every two minutes) along the transect

In some cases, the LORAN readings were known to be in error and $I$ for these cases was determined using the elapsed time and average vessel speed of 18 km per hour.

No attempt was made to estimate the probability of sighting animals on the trackline ( $g_{0}$ ), which was assumed equal to one for all species. The resulting estimates do not account for animals that were not sighted due to observer error or that may have been unavailable for sighting (e.g., on the transect, beneath the water's surface). This effect could result in conservative estimates of abundance. However, it is not clear that biases due to assuming $g_{i}=1$ would not be countered by other effects, such as underestimation of average group size or attraction of animals to the survey vessel.

Variance of $\hat{D}$ was estimated as

$$
\operatorname{vär}(\hat{D})=\hat{D}^{2}\left[\frac{\nu \hat{a} r(n)}{n^{2}}+\frac{\nu \hat{a} r(\hat{S})}{\hat{S}^{2}}+\frac{v \hat{a} r[\hat{f}(0)]}{\hat{f}(0)^{2}}\right]
$$

and coefficient of variation (cv) estimated as

$$
c \hat{v}(\dot{D})=\frac{\sqrt{v a \hat{r}(\hat{D})}}{\hat{D}}
$$

The variance of $n$ was based on the variation in the number of on-effort group sightings between sampling units within each stratum. The sampling unit for the ship surveys was a day's visual sighting effort. The variance of $\hat{S}$ was based on the variation in group size within each stratum. The variance of $\hat{f}(0)$ was based on the variation between expected versus actual perpendicular sighting distance (PSD) distributions pooted across strata.

The group sizes for some species tended to be inversely related to PSD, a feature which can result from size bias (i.e., larger groups are easier to see at distance than are small groups). Therefore, the arithmetic mean of group size could overestimate the true mean group size and could have lead to positively biased abundance estimates. The program DISTANCE used a regression of the group size by sighting
distance to generate a mean "expected group size." The expected group size was used in the density estimation if it was significantly different from the arithmetic mean group size ( $p<0 . t 5$, Student's test, see Buckland et al. $t 993$ ). The sample size, herd size estimate, and other parameters used in the distance sampiing analysis are given in Table I.

An exploratory analysis indicated that sightings made at small radial distances (generally $<0.247 \mathrm{~nm}$ ) resulted in a poor fit of the sighting probability density function. Exclusion of these sightings resulted in relatively better fits and more precise estimates of $\mathcal{K} 0$ ). It was felt that most of these sightings were of animals that were attracted to the vessel to bow ride. One requirement for unbiased estimates of abundance is that the sighting target(s) should not move in response to the observer or the observation platform (Buckland et al. t 993). To reduce the potential for bias due to attraction to the vessel, only sightings made at radial distances of $>=0.247 \mathrm{~nm}$ were included in the data summarized for the distance sampling analysis.

Examination of the bearing and reticle measurements indicated that most were rounded to the nearest 5 units ( 5 degrees for bearing, 0.5 for reticle readings). The bearing and reticle reading data for each sighting were smeared by adding a randomly selected value between -5 and 5 for the bearing, and between -0.5 and 0.5 for the radial distance. This was done to reduce the potential for artificial grouping of sighting distances due to rounding of measurements by observers.

The formula used for calculating radial sighting distances $(R)$, from Smith (1982), was

$$
R=h \cdot \tan \left[\arctan \left(\frac{Y_{1,2}}{\sqrt{h_{1,2}}}\right)-B \cdot r\right]
$$

where
$h_{1}=0.003508$, modeled height of binoculars above surface in nautical miles for $r<5.0$
$h_{2}=0.004818$, modeled height of binoculars above surface in nautical miles for $\boldsymbol{r} \boldsymbol{>}=5.0$

$$
\begin{array}{ll}
Y_{1} & =76.756, \text { modeled declination parameter for } r>5.0 \\
Y_{2}^{\prime} & =29.228, \text { modeled declination parameter for } r>5.0 \\
B & =0.0623, \text { a constant } \\
r & =\text { reticle measurement } .
\end{array}
$$

A non-linear model (SAS 1988) was used to produce least squares estimates of parameters $h$ and $Y$, using empirical data on reticle readings and distances. Estimates generated using the entire range of empirical data (for $r$ from 0.0 to 15.0) provided a good fit for distances greater than measured at $r=5.0$, but underestimated distances measured at $r=5.0$. Therefore, two addition sets of estimates were generated, one with using the ground truth measurements for $r<5.0$ (Model 1) and another for $r>=5.0$ (Model 2). A plot of these two sets of estimates against the mean empirical distances (Figure 3) indicated Model 1 predicted distances best fort $<4.0$ and Model 2 predicted distances best for $r>=4.0$. Thus, two sets of estimates were used for $h$ and $Y$; one set for $r<4.0$, and another set for $r>=4.0$. Perpendicular sighting distances ( $P$ ) were calculated as

$$
P=R \cdot \sin (b)
$$

where $\quad b=$ angle between sighting and trackline.

The sample sizes (number of groups sighted) of most species were considered insufficient to obtain accurate and precise estimates of $\hat{f}(0)$. Sightings of species with similar sighting characteristics (i.e., body size, group size, behavior) were pooled to estimate $\hat{f}(0)$ (Table II). For instance, $\hat{f}(0)$ for Cuvier's beaked whale (Ziphius cavirostris) was estimated by pooling with sightings of Blainville's beaked whale (Mesoplodon densirostris), unidentified beaked whales (family Ziphiidae), and dwarf and pygmy sperm whales (Kogia spp.) of group sizes less than five. Seven speci es did have sufficient sightings (30 or more, including non-GulfCet sightings) to estimate species $\hat{f}(0)$ without pooling; these were the sperm whale (Physeter macrocephalus), dwarf sperm whale (K. simus), Atlantic spotted dolphin (Stenella frontalis), pantropical spotted dolphin (S. attenuata), striped dolphin (S. coerulooalba), Risso's dolphin (Grampus
griseus), and bottlenose dolphin (Tursiops truncatus). The estimated values of $\hat{f}(0)$ and associated statistics are listed in Table III.

## Survey Results

The northern Gulf of Mexico surveys total transect kilometers sampled each survey during 1991-94 varied from 3,491-6,386 km, for a total over all the surveys of $22,041 \mathrm{~km}$. More than 850 sightings of at least 21 cetacean species were made during the 1991-I 994 surveys. Of these, 616 sightings, identified to species (or the genus Kogia or the family Ziphiidae), were used in the distance sampling analysis. The number of sightings by species and survey year are fisted in Table IV.

Analytical Results

The northern Gulf of Mexico abundance estimates for all species observed are listed in Table V with associated statistics. The all-surveys-combined (ASC) abundance estimates ranged from fewer than 1000 for most species to about 31,000 for pantropical spotted dolphins. The coefficient of variation (cv) of the ASC estimates was relatively large ( $>50 \%$ )for most species, but was about $30 \%$ or less for sperm whales, dwarf sperm whaies (K. simus), pantropical spotted dolphins, and grampus. The $\mathrm{ev}^{\prime}$ s of the abundance estimates by-survey were considerably larger for most species, although the cv's for the by-survey abundance estimates of the pantropical spotted dolphin ranged from about $29 \%$ to $48 \%$. The by survey abundance estimates and associated statistics are given in Table VI.

## Literature Cited

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Table I. Sighting and group-size statistics used for estimating cetacean species abundance for pooled 199194 shipboard surveys. Abbreviations are. as foltows: $\eta=$ number of sightings, $\boldsymbol{n} \boldsymbol{l}=$ groups encountered per 1000 km (survey effort given in column heading), $\mathrm{G}=$ mean group size, $\mathrm{S}=$ size-bias adjusted group size, CV $=\%$ coefficient of variation. Group sizes denoted with an * indicate which size estimate was used in density calculations.

| SPECIES | SPRING |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | n | $\begin{gathered} \mathrm{n} / \mathrm{L} \\ 22040 \mathrm{~km} \end{gathered}$ | $\begin{aligned} & \text { CV } \\ & n / L \end{aligned}$ | G | S | $\begin{gathered} \text { c v } \\ \text { she } \end{gathered}$ |
| Balaenoptera edeni | 3 | 0.000140 | 97.65 | -2.67 | 11.00 | 45.07 |
| Physeter macrocephalus | 60 | 0.002720 | 22.09 | 2.58 | -2.13 | 7.87 |
| Kogia simus | 34 | 0.001543 | 23.75 | -2.21 | 2.14 | 13.77 |
| Kogia brevicaps | 10 | 0.000454 | 34.00 | '2.24 | 2.13 | 11.75 |
| Kogia sp. | 61 | 0.002768 | 23.38 | 7.08 | 2.02 | 9.41 |
| Ziphius cavirostris | 5 | 0.000230 | 46.06 | '1.20 | 1.39 | 16.67 |
| Unidentified Ziphiidae | 12 | 0.000540 | 35.00 | '1.92 | 1.86 | 13.56 |
| Stenella frontalis | 20 | 0.000907 | 31.83 | 21.25 | 28.04 | 15.69 |
| Stenella attenuata | 158 | 0.007169 | 14.43 | 57.73 | ‘43.37 | 7.67 |
| Stenella coeruleoalba | 29 | 0.001316 | 22.42 | 36.74 | 41.67 | 11.63 |
| Stenella longirostris | 20 | 0.000907 | 33.57 | 70.80 | 72.34 | 24.62 |
| Stenella clymene | 23 | 0.001044 | 23.61 | 63.70 | '55.30 | 26.67 |
| Lagenodelphis hosei | 1 | 0.000045 | 89.11 | -34.00 | - | - |
| Steno bredanensis | 11 | 0.000499 | 27.96 | '14.36 | 13.66 | 12.18 |
| Orcinus orca | 7 | 0.000318 | 36.29 | '10.00 | 11.75 | 15.28 |
| Pseudorca crassidens | 5 | 0.000227 | 43.63 | -20.40 | 33.27 | 42.07 |
| Feresa attenuata | 4 | 0-000182 | 48.94 | 79.25 | 90.68 | 62.44 |
| Peponocephala electra | 9 | 0.000408 | 29.63 | '119.56 | 115.75 | 22.50 |
| Grampus griseus | 75 | 0.003403 | 23.34 | 10.73 | '10.10 | 9.44 |
| Globicephala macromynchus | 5 | 0.000227 | 44.67 | '16.60 | 14.19 | 21.16 |
| Tursiops truncatus | 107 | 0.004855 | 19.49 | 13.96 | *9.01 | 11.81 | Usmi=unldentified small whale.

SPECIES


Table lill. The estimated value of $f_{0}$ and associated statistics. ESW $=$ effective strip half-width ( $1 / f_{0}$ ), $\mathrm{CV}=\%$ coefficient of variation, $n\left(f_{0}\right)=$ number of sightings used to estimate $f_{0}(*$ indicates species which required pooling with other species to estimate $\left.f_{0}\right), \mathrm{n}(\mathrm{D})=$ number of sightings used for estimating density and abundance.

| SPECIES | $\underset{\left(\mathrm{km}^{-1}\right)}{f_{0}}$ | $\begin{aligned} & \text { ESW } \\ & (\mathrm{km}) \end{aligned}$ | $\begin{aligned} & C V \\ & \left(U_{0}\right) \end{aligned}$ | $n\left(f_{0}\right)$ | Spring <br> n(D) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Balaenoptera edeni | 0.4777 | 2.0933 | 21.07 | " 87 | 3 |
| Physeter macrocephalus | 0.4588 | 2.1795 | 19.69 | 69 | 60 |
| Kogia simus | 0.5025 | 1.9904 | 13.95 | 34 | 34 |
| Kogia breviceps | 0.4755 | 2.1028 | 12.80 | -62 | 10 |
| Kogia sp. | 0.4756 | 2.1028 | 12.80 | -63 | 61 |
| Ziphius cavirostris | 0.5601 | 1.7854 | 7.84 | 75 | 5 |
| Unidentified Ziphiidae | 0.5624 | 1.7780 | 7.72 | 76 | 12 |
| Steneilla frontatis | 0.8353 | 1.1972 | 26.55 | 58 | 20 |
| Stenelia attenuata | 10.5050 | 1.9802 | 11.77 | 170 | 158 |
| Stenella coeruleoalba | 0.5122 | 1.9524 | 36.42 | 31 | 29 |
| Stenella longirostris | 0.4928 | 2.0291 | 9.64 | 757 | 20 |
| Stenella clymene | 0.4840 | 2.0660 | 9.59 | 259 | 23 |
| Lagenodelphis hosei | 0.4112 | 2.4318 | 10.17 | '183 | 1 |
| Steno bredanensis | 0.5961 | 1.6775 | 7.30 | -284 | 11 |
| Orcinus orca | 0.4366 | 2.2905 | 14.41 | *83 | 7 |
| Pseudorca crassidens | 0.4128 | 2.4223 | 12.80 | -107 | 5 |
| Feresa attenuata | 0.4890 | 2.0449 | 15.73 | -104 | 4 |
| Peponocephala electra | 0.4072 | 2.4559 | 10.14 | '182 | 9 |
| Grampus griseus | 0.4024 | 2.4852 | 10.42 | 77 | 75 |
| Globicephala macromynchus | 0.4701 | 2.1271 | 16.83 | 73 | 5 |
| Tursiops truncatus | 0.6439 | 1.5529 | 11.97 | 140 | 107 |

Table IV. On-effort Cetacean sightings collected during 1991-1994 annual, spring-summer vessel surveys of the northern Gulf of Mexico.

| SPECIES | SURVEY |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1991 | 1992 | 1993 | 1996 | ALL |
| Balaencptera sp. | - | 2 | - | 1 | 3 |
| Balcenoptera physalus | - | - | - | 1 | 1 |
| Baleenoptere edeni | 3 | - | - | - | 3 |
| Physeter macroctophalus | 5 | 18 | 8 | 29 | 60 |
| Kogia simus |  | 16 | 12 | E | 34 |
| Kogia braviceps | 3 | 4 | 2 | 1 | 10 |
| Kogia sp. |  | 12 | 6 | - | 18 |
| Zuphius cavirostris | - | - | 3 | 2 | 5 |
| Mesapkocan densirostris | - | 1 | - | - | 1 |
| Ziphïidae | 1 | 7 | 7 | 12 | 27 |
| Stenelia tromelis | 1 | 7 | 9 | 6 | 23 |
| Stenella attonuats | 21 | 37 | 53 | 68 | 178 |
| Stenelita coenteratha | 3 | 7 | 8 | 11 | 29 |
| Stenelia longirostris | - | 6 | 5 | 9 | 20 |
| Stenella clymene | 3 | 6 | 9 | 7 | 25 |
| Stenella sp. | - | 1 | 4 | 5 | 10 |
| Lagenodelphis hosel | - | 1 | - | - | 1 |
| Steno bredanensis | 1 | 4 | 4 | 2 | 11 |
| Orcinus orca | - | 1 | 4 | 2 | 7 |
| Pseudorca crassidens | 1 | 1 | 1 | 2 | 5 |
| Faresa aftenuata | 1 | 2 | 1 | - | 4 |
| Peponocophala electra | - | 3 | 2 | 4 | 9 |
| Grampus griseus | 2 | 22 | 15 | 39 | 78 |
| Glabicephala marcomynchus | - | 3 | 1 | 1 | 5 |
| Tursiops touncatus | 8 | 46 | 47 | 20 | 121 |
| Unidentified large whaio | - | 1 | 6 | 3 | 10 |
| Unidentified small whale | 2 | 2 | 2 | 5 | 11 |
| Unidentified odomocate | 6 | 17 | 12 | 8 | 43 |
| Unidentified dolphin | 15 | 25 | 38 | 24 | 103 |
| All | 76 | 253 | 259 | 288 | 856 |

Table IV. On-effort cetacean sightings collected during 1991-1994 annual, spring-summer vessel surveys of the northern Gulf of Mexico.

| SPECIES | SURVEY |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1991 | 1992 | 1903 | 1994 | ALL |
| Balaenopiera sp. | - | 2 | - | 1 | 3 |
| Balaomoptara physalus | - | - | - | 1 | 1 |
| Balaenoptere edeni | 3 | - | - | - | 3 |
| Fhyseter macrocephalus | 5 | 18 | 8 | 29 | 60 |
| Kogie simus |  | 16 | 12 | 6 | 34 |
| Kogie brevicops | 3 | 4 | 2 | 1 | 10 |
| Kogie sp. |  | 12 | 8 | - | 18 |
| Ziphius cavirastris | - | - | 3 | 2 | 5 |
| Mesoplodon densirostris | - | 1 | - | - | 1 |
| Ziphiidae | 1 | 7 | 7 | 12 | 27 |
| Stenella frontalis | 1 | 7 | 9 | 6 | 23 |
| Stenella attenuata | 21 | 37 | 53 | 68 | 179 |
| Stenalta coerulearlbs | 3 | 7 | 8 | 11 | 29 |
| Stenelia longirostris | - | 6 | 5 | 9 | 20 |
| Stenella cilymeno | 3 | 6 | 9 | 7 | 25 |
| Stonells sp. | - | 1 | 4 | 5 | 10 |
| Lagenodelphis hossi | - | 1 | - | - | 1 |
| Steno bredanensis | 1 | 4 | 4 | 2 | 11 |
| Orcinus area | - | 1 | 4 | 2 | 7 |
| Pseudorca crassidens | 9 | 1 | $\dagger$ | 2 | 5 |
| Foresa attenuasa | 1 | 2 | 1 | - | 4 |
| Peponocephala electra | - | 3 | 2 | 4 | 9 |
| Grampus griseus | 2 | 22 | 45 | 38 | 78 |
| Globicephala marcartyynchus | - | 3 | 9 | 1 | 5 |
| Tursiops ouncstus | 8 | 46 | 47 | 20 | 121 |
| Unidentified large whale | - | 1 | 6 | 3 | 10 |
| Unidentified small whale | 2 | 2 | 2 | 5 | 11 |
| Unidentified odontocete | 6 | 17 | 12 | 8 | 43 |
| Unidentified doiphin | 15 | 26 | 38 | 24 | 103 |
| ALL | 76 | 253 | 259 | 268 | 856 |

Table V. Abundance estimates ( N ) and density estimates (D. per 1000km²) by species for the northern Gulf of Mexico for 1991-1994 surveys combined with \% coefficient of variation (CV) and log-normal 95\% confidence intervals (CI).

| SPECIES | SPRING |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $N$ | $D$ per $1000 \mathrm{~km}^{2}$ | cv | Cl |
| Balaenoptera edeni | 35 | 0.08660 | 109.59 | 6-205 |
| Physeter macrocephalus | 530 | 1.32840 | 30.64 | 295-953 |
| Kogia simus | 341 | 0.85490 | 30.80 | 189-615 |
| Kogia breviceps | 56 | 0.14024 | 38.18 | 27-115 |
| Kogia sp. | 547 | 1.37000 | 28.27 | 317-941 |
| Ziphius cavirostris | 30 | 0.07620 | 49.61 | 12-70 |
| Ziphiidae | 117 | 0.29350138 .32 |  | 57-242 |
| \|Stenella frontalis | 3213\|8.053501 44.39 |  |  | \| 1399-7378 |
| Stenella attenuata | $31320178.50500\|20.14\|$ |  |  | 21918-46298 |
| \|Stenella coeruleoalba | 4858 ${ }^{12.17700 \mid}$ |  | 44.33 | 2038-11330 |
| IStenella longirostris | 6316\| $15.83000 \mid 42.771$ |  |  | 2828-14104 |
| Stenella clymene | 5571 | 13.96500 | 36.87 | 2734-11355 |
| Lagenodelphis hosei | 127 | 0.31716 | 89.69 | 28-570 |
| Steno bredanensis | 852) 2.13670) 31.37 |  |  | 468-1554 |
| Orcinus orca | 277 | 0.69327141 .931 |  | 126-609 |
| Pseudorca crassidens | 381\| | 0.95525161 .95 \| |  | 115-1262 |
| Feresa attenuata | 518 | 1.29790 | 80.88 | 101-2659 |
| Peponocephala electra | 3965 | 9.93890 | 38.57 | 1883-8350 |
| Grampus griseus | 2749 | 6.89120 | 27.25 | 1627-4645 |
| Globicephala macrorhynchus | 353 | 0.88520 | 52.39 | 133-939 |
| Tursiops truncatus | 5618 | 14.08000 | 25.74 | 3419-9229 |

Table VI. Annual abundance estimates ( N ) by species for the northern Gulf of Mexico for 1991-1994 spring-summer surveys with \% coefficient of variation (CV), and log-normal $95 \%$ confidence intervals (CI).

| SPECIES | 1991 |  |  | 1992 |  |  | 1993 |  |  | 1994 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | N | CV | Cl | $N$ | CV | Cl | $N$ | CV | Cl | $N$ | CV | Cl |
| Gataenoplera edent | 218 | 101. 12 | 39-1235 | - |  |  |  |  |  |  |  |  |
| Physeter mancocophaius | 143 | 58.14 | 48-428 | 931 | 47.53 | 376-2290 | 229 | 52.30 | 88-614 | 771 | 41. 54 | 347.1709 |
| Kogia stmus | - | - |  | 541 | 42. 50 | 240-1223 | 502 | 49. 17 | 196-1275 | 154 | 64.00 | 47-504 |
| Kogia brevicops | 109 | 68.35 | 31-386 | 60 | 53.33 | 22-185 | 59 | 70.36 | 15-234 | 16 | 107.60 | 3- w |
| Kogia $¢ p$. | 109 | 68.35 | 31-386 | 1050 | 40. 10 | 460-2189 | 580 | 44. 61 | 247.1359 | 162 | 81.28 | 52-505 |
| Ziphlus cavirostia |  | , |  | - |  |  | 70 | 52.56 | 22-225 | 38 | 60. 23 | 8-159 |
| ZIphildet | 129 | 78.15 | 31-530 | 18 | 126. 66 | 2.129 | 53 | 78.43 | 12-223 | 267 | 47.96 | 144-719 |
| Stonalia trontais | ' |  |  | 4527 | 65. 16 | 1370-14978 | 4618 | 62.36 | 1464-14571 | 2186 | 85.35 | 483-8593 |
| SteneNa aftenuata | 19767 | 45. 20 | 8302-47058 | 15280 | 36. 45 | 7555-30904 | 29414 | 26.82 | 16762-51615 | 71847 | 31. 52 | 38889-132750 |
| Stenatita coenitoralibe | 3463 | 75. 73 | 865-14025 | 2574 | 52. 25 | 966-6891 | 4160 | 63.08 | 1313-13163 | 0147 | 59.69 | 2711-24479 |
| Stenalia fongtrostris | , |  |  | 2593 | 62. 69 | 012-8280 | 2336 | 62. 49 | 719.7587 | 15995 | 07.28 | 4670-54791 |
| Steneifla clymene | 1936 | 68.94 | 540-6944 | 3390 | 47.59 | 1385-6417 | 6466 | 46. 70 | 2693-15621 | 12255 | 62.27 | 3830-38132 |
| Lagenodatiph's hosei | , | . |  | 443 | 91.62 | 81-2157 |  |  |  |  |  |  |
| steno bredamensis | 545 | 114.67 | 83-3581 | 759 | 56. 26 | 230-2490 | 1192 | 49.03 | 4752987 | 527 | 65. 93 | 85-3250 |
| Oromus orca | - |  |  | 138 | 95. 52 | 27.705 | 641 | 50.36 | 248-1670 | 193 | 111.70 | 11.3360 |
| Psouderca crassidens | 661 | 88.17 | 139-3131 | 196 | 99.61 | 37-9053 | 77 | 108.08 | 13-458 | 744 | 113.58 | 41.13462 |
| Foresa-altanusta Paporocaphale alactra | $2347$ | $-60.86$ | -549-10032 | 356 3174 | 73.43 54.15 | $\begin{array}{r} 95-1341 \\ \hline 1941.8833 \end{array}$ | 153 | 112.99 69.95 | $\begin{array}{r} 24.954 \\ 231-2961 \end{array}$ | 10588 | 47.65 | 4245-23400 |
| Grampus griseus | 667 | 95.37 | 726217 | 2325 | 34. 40 | 1 192-4534 | 1408 | 4133 | 630-3118 | 6332 | 44.63 | 2683-14844 |
| Grabiosphala macrortymohus | $\checkmark$ |  |  | 808 | 62.05 | 267-2662 | 403 | 120.07 | 15-694 | 240 | 102.75 | 43-1338 |
| Tursfops inuneatus | 2392 | 52.71 | 645-6742 | 6937 | 40.14 | 4130-19338 | 6149 | 39. 69 | 2622-43167 | 5487 | 56. 96 | 1822-16523 |




Flgure 2. Map of the northern Gulf of Mexico with an example of on-elfort survey heck that occurred during a survey leg which was designed specifically to collect marine mammal sightings along transects perpendicular lo the depth gradient,


Figure 3. Estimated radial sighting distances, based on two sets of mooeied esumares or dinocurar neignt and declination, with mean empirical radial distances.

