



# Modeling attenuation of photosynthetically active radiation across the optical gradient in the Laurentian Great Lakes with application to Lake Erie

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## ABSTRACT

Representation of subsurface photosynthetically active radiation (PAR) in biophysical models of the Laurentian Great Lakes (LGL) is imperative to their utility as tools for research and management. Here we consolidated measured vertical profiles of subsurface PAR with concurrent water quality (WQ) data from four LGL. We estimated the diffuse attenuation coefficient of PAR (K<sub>d</sub>(PAR)) by fitting an exponential function to measured PAR over depth, and evaluated 68 regressions predicting K<sub>d</sub>(PAR) as a function of water quality variables (K<sub>d</sub>-WQ regressions). We compare four of the top cross-lake calibrated regressions against two published regressions trained on western Lake Erie (WLE) data. Then, as a case study, we demonstrate the utility of our cross-lake calibrated K<sub>d</sub>-WQ regressions with a simplified biophysical model of Lake Erie consisting of the Finite Volume Community Ocean Model with submodules for simulating suspended sediment and dissolved organic carbon (FVCOM-SS-DOC). Twenty-five K<sub>d</sub>-WQ regressions were identified as candidates for use in biophysical models based on their skill determined via cross-validation. WLE-trained K<sub>d</sub>-WQ regressions were less able to simulate K<sub>d</sub>(PAR) and PAR in more transparent waters compared to cross-lake calibrated K<sub>d</sub>-WQ regressions, which translated to considerable differences in primary production estimates for the central and eastern basins when using WQ data simulated by FVCOM-SS-DOC. A cross-lake calibrated K<sub>d</sub>-WQ regression was installed into FVCOM-SS-DOC, which then simulated spatial patterns of suspended sediments and K<sub>d</sub>(PAR). These calibrated K<sub>d</sub>-WQ regressions can be used in a variety of biophysical models across optically-distinct waters of the LGL to support adaptive management of nutrient inputs and fisheries.

## 1. Introduction

Light is a fundamental energy source that drives the ecology of aquatic systems. By use of intracellular chlorophyll and other accessory pigments, primary producers are able to capture photosynthetically active radiation (PAR, 400–700 nm) to convert inorganic carbon into carbohydrates and other biomolecules. Light levels influence higher trophic levels in addition to primary producers. For instance, changing light levels can alter the zooplankton community and biomass by influencing bottom-up (i.e., reduction in quantity or quality of algal food, Faithfull et al. (2011)) and top-down controls (i.e., increased light, increased predation, Bramm et al. (2009)). Photosynthesis-irradiance (PI) curves describe the effect of light on primary productivity as a hyperbolic function where photosynthetic rates, as commonly measured by <sup>14</sup>C uptake, increase with PAR irradiance towards an asymptote or

saturation point after which photoinhibition may occur (Jassby and Platt, 1976). High turbidity increases the attenuation of PAR, which can decrease phytoplankton biomass and primary production (Bramm et al., 2009; Brothers et al., 2017; Ge et al., 2020), alter phytoplankton stoichiometry (Urabe et al., 2002), reduce bottom dissolved oxygen (Moriarty et al., 2021; Pilla and Couture, 2021), and affect predation rates of certain visually-foraging zooplankton and piscivorous fish (Hansen et al., 2013; Jokela et al., 2013; Pangle et al., 2009). The abundance of research in this field makes it apparent that adequately representing PAR is imperative for any aquatic ecosystem model seeking to simulate primary production or higher order ecological processes such as phytoplankton competition and succession (Huisman et al., 2004; Liu et al., 2021), biological drivers of hypoxia (Brothers et al., 2017; Lashaway and Carrick, 2010; Moriarty et al., 2021), secondary production (Bramm et al., 2009; Faithfull et al., 2011), or nutrient cycling and algal

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stoichiometry (Urabe et al., 2002).

Biophysical models are useful tools for improving our understanding of how physical, biogeochemical and trophic interactions influence aquatic ecosystems. Here, we use the term “biophysical model” to refer to a numerical ocean model linked to a module that simulates advection–diffusion and source–sink terms on one or more state variables representing ecological or biogeochemical quantities. Biophysical models may be formulated with varying degrees of complexity for different applications, ranging from a few to dozens of state variables. Using a biophysical model of Lake Michigan, Rowe et al. (2017) showed that varying mixed layer depth, which affected light limitation of phytoplankton and interaction with benthic filter feeders, was an important driver of primary production. Using a biophysical model of Chesapeake Bay, Moriarty et al. (2021) found that increased light attenuation due to sediment resuspension effected a spatial shift in primary production that also reduced dissolved oxygen up to 2.2 mg L<sup>-1</sup> in some areas due to changes in photosynthesis and the remineralization rates of organic material. Similar models developed for Lake Michigan (Ji et al., 2002) and the Changjiang Estuary (Ge et al., 2020) have demonstrated how sediment plumes can regulate phytoplankton growth through increased light attenuation and nutrient availability. In Lake Erie, the case study in this manuscript, several biophysical models have provided insights on ecosystem dynamics (e.g. Leon et al., 2011) and key management issues related to harmful algal blooms (Verhamme et al., 2016), hypoxia (Bocaniov et al., 2016; Leon et al., 2006; Rowe et al., 2019), and *Cladophora* (Valipour et al., 2016). Accurate modeling of subsurface light is critical to developing scenarios aimed at resolving nutrient effects on the eutrophication issues facing Lake Erie (Annex 4 Objectives and Targets Task Team, 2015) due to the interactions between light, algal growth, nutrients, and associated eutrophication endpoints.

Attenuation of incident PAR occurs through scattering and absorption by water, suspended particles, and dissolved substances (Abdelrhman, 2017). Changes in subsurface light in coastal zones and lakes are often driven by terrestrial sources of chromophoric dissolved organic material (CDOM) (Gerea et al., 2016; Pilla and Couture, 2021) and suspended sediments (SS), or internal sources such as resuspension of sediment, algal particles, and/or autochthonous CDOM (Niu et al., 2018; Valipour et al., 2017). The intensity of subsurface PAR in aquatic systems has commonly been modeled as an exponential function of depth (Leon et al., 2011; Morel, 1988; Verhamme et al., 2016; Weiskerger et al., 2018):

$$I_z = I_0 \exp(K_d(PAR) \cdot z) \quad (1)$$

$$K_d(PAR) = K_a[SS] + K_b[DOC] + K_c[CHL] + K_w \quad (2)$$

where  $I_z$  is irradiance at depth,  $z$  is the vertical coordinate (negative downward from the surface),  $I_0$  is surface incident PAR, and  $K_d(PAR)$  is the diffuse attenuation coefficient for PAR (Equation (1)), which can be related to the attenuation due to water ( $K_w$ ) and concentrations of optically-active absorbing and scattering constituents such as SS, dissolved organic carbon (DOC), and chlorophyll-*a* (CHL) with coefficients ( $K_a, K_b, K_c, K_w$ ) determined by regression (Equation (2)). We recognize that the optical properties of bulk DOC vary depending on the source of DOC and residence time in lakes (Smith et al., 2004); however, CDOM absorption is not often simulated directly in aquatic biogeochemical models, therefore we chose to use DOC to represent CDOM absorption for compatibility with commonly used aquatic biogeochemical models. Absorption and scattering processes (i.e., inherent optical properties, [IOPs]) are wavelength-dependent, and also vary for direct-beam versus diffuse attenuation depending on solar zenith angle (Lee et al., 2005). Spectrally-resolved light attenuation models have been developed that can account for the variable absorption and backscattering of different wavelengths within the PAR spectrum (e.g., Abdelrhman, 2017; Gregg and Rousseaux, 2016; Morel, 1988; Shuchman et al., 2013). Use of

complex spectral light attenuation models versus simpler models of diffuse attenuation at a single or binned wavelength may have a negligible or significant impact in aquatic biogeochemical models, depending on latitude, water transparency, model formulations, and uncertainty in other model processes (Gregg and Rousseaux, 2016). In order for spectral light attenuation models to be used for modeling subsurface PAR, IOPs must be defined for the wavelength bands representing the contiguous PAR spectrum; and their calibration requires region-specific spectral observations of subsurface light, while  $K_d(PAR)$  may be estimated from more widely available PAR sensor data.

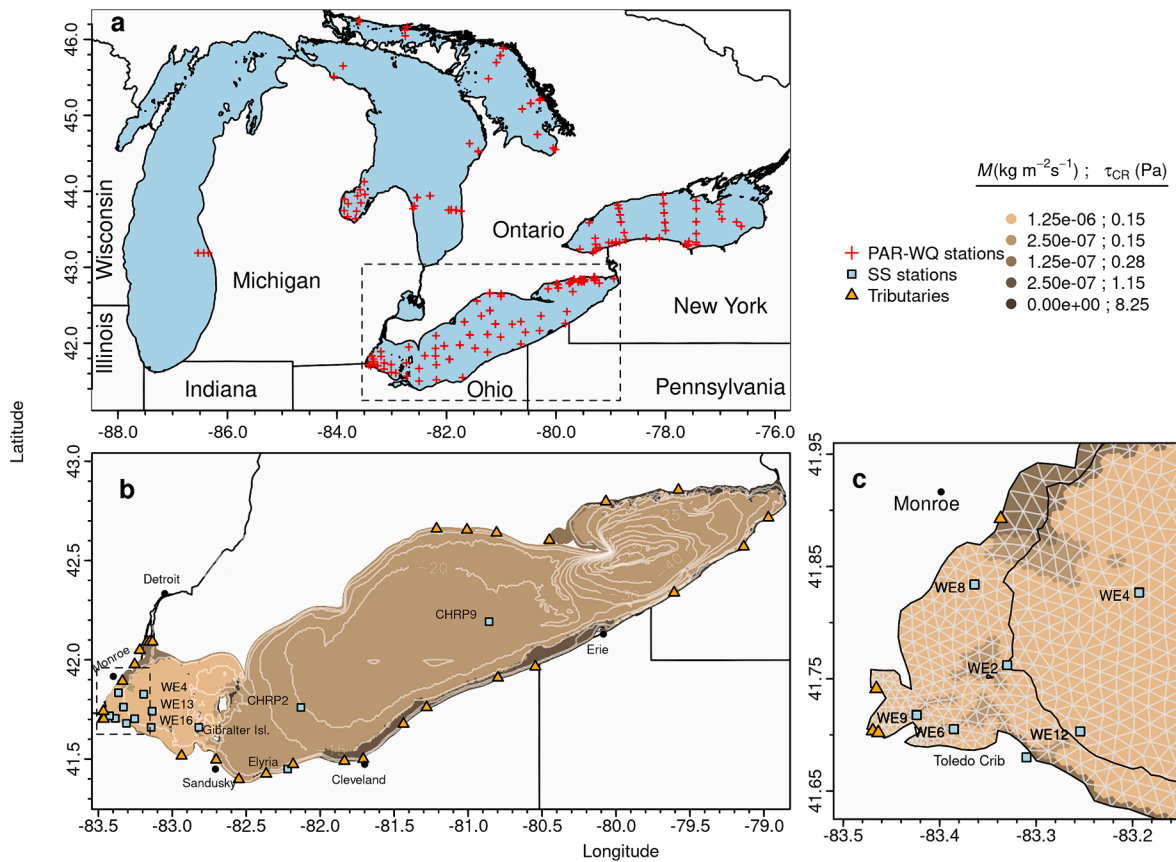
The availability of concurrently measured water quality samples and PAR profiles provides the necessary data for modeling  $K_d(PAR)$  as a function of water quality variables ( $K_d$ -WQ regressions; Eq. (2)), and  $K_d$ -WQ regressions can be used in biophysical models to simulate light attenuation. In the Lake Erie Ecosystem Model (LEEM),  $K_d(PAR)$  was estimated as a power function ( $y = 0.2273x^{0.734}$ ) of total suspended solids (TSS) concentration (Limnotech, 2021; Verhamme et al., 2016), whereas Leon et al. (2011) estimated  $K_d(PAR)$  from a multiple linear function of SS, algae, DOC, and detritus in their Lake Erie application of the coupled Estuary and Lake Computer Model-Computational Aquatic Ecosystem Dynamics Model (ELCOM-CAEDYM). In their Lake Michigan biophysical model, Rowe et al. (2017) simulated  $K_d(PAR)$  as a linear function of chlorophyll and detritus. Lake Erie-specific  $K_d$ -WQ regressions have been described in the literature and used in other applications. For their *Cladophora* growth model, Higgins et al. (2005) estimated  $K_d(PAR)$  from simple regressions on TSS or turbidity data. Weiskerger et al. (2018) evaluated 23 different regressions using  $K_d(PAR)$  and water quality data from the western basin of Lake Erie, 15 of which showed excellent predictive skill ( $R^2 > 0.8$ ). While  $K_d$ -WQ regressions have been sufficient in modeling light attenuation in lake-specific or basin-specific applications, there has been little investigation into relationships that could be reliably used in locations with different optical properties within and across the Laurentian Great Lakes (LGL).

In this manuscript we describe the development of  $K_d$ -WQ regressions for use in mechanistic ecosystem models of the LGL. We calibrated and assessed various  $K_d$ -WQ regressions using a large database of subsurface PAR observations and concurrently-collected water quality data collected from four of the five LGL to identify regressions that can skillfully predict  $K_d(PAR)$  across a gradient of optical conditions. We then demonstrated the utility of these regressions in a case study by using them to predict  $K_d(PAR)$  and primary productivity using simulated output from a 3D biophysical model of Lake Erie consisting of a hydrodynamic model, SS model, and simplified ecological model. Using the 3D lake model's output, we evaluate the sensitivity of primary production rates to  $K_d(PAR)$  prediction between cross-lake calibrated regressions and regressions trained on data exclusively from Lake Erie's western basin. Finally, we install a cross-lake calibrated  $K_d$ -WQ regression into the 3D lake model and evaluate the model's ability to recreate spatial patterns of  $K_d(PAR)$  in Lake Erie.

## 2. Methods

### 2.1. Study area

The geographic scope of our empirical analyses on PAR and WQ spans four of the five LGL including Lakes Michigan, Huron, Ontario, and Erie (Fig. 1a). The LGL are located at temperate latitudes situated along or near the Canadian-US border with trophic statuses ranging from *meso*-eutrophic in some nearshore areas to oligotrophic in the offshore. The LGL are all dimictic, stratifying in the summer and winter and mixing in the spring and late fall. The lakes span a range of water transparency making them a suitable testing ground for light attenuation model development. We also present a biophysical modeling analysis that focuses on Lake Erie (Fig. 1b,c). Lake Erie is the shallowest, warmest, and most productive of the LGL. It is often disaggregated into



**Fig. 1.** Sample locations (red crosses) with photosynthetically active radiation (PAR) vertical profiles and associated water quality (WQ) data used to develop light attenuation models (a). Spatial domain (b) and grid (c) of the Lake Erie Finite Volume Community Ocean Model's Suspended Sediment and Dissolved Organic Carbon model (FVCOM-SS-DOC) with suspended sediment sample locations used for calibration and assessment (labeled blue squares), 5 m depth contours, tributary pourpoints (orange triangles), and calibrated erodibility parameters ( $M$ ,  $\text{kg m}^{-2} \text{s}^{-1}$ ;  $\tau_{\text{CR}}$ , Pa). Darker colors indicate less erodible sediment. Cities are indicated as black points.

three subbasins: the western, central, and eastern. Lake Erie has received considerable attention by policymakers, managers, scientists, and the public in recent decades due to the recurring effects of anthropogenic eutrophication such as harmful algal blooms of *Microcystis* in the western basin (Stumpf et al., 2012; Wynne et al., 2011), central basin hypoxia (Stow et al., 2023; Zhou et al., 2013), and nuisance *Cladophora glomerata* in the eastern basin (Higgins et al., 2005).

## 2.2. Water quality and PAR profile data

We compiled individual profiles of PAR measured using spherical sensors (quantum scalar irradiance) from submersible instruments deployed from research vessels with coincident surface water quality data from various sampling programs across Lakes Erie, Michigan, Ontario, and Huron (Fig. 1a, Electronic Supplementary Material (ESM) Table S1). Water quality data included TSS, DOC, CHL, volatile suspended solids (VSS), and non-volatile suspended solids (NVSS). NVSS is the mineral component of TSS and measured as the residual mass following combustion of TSS at 550° C. VSS is calculated as the difference between gravimetrically-measured TSS and NVSS. The cross-lake dataset included PAR profiles,  $K_d(\text{PAR})$  estimates, and water quality data from the Western Lake Erie Monitoring program run by NOAA's Great Lakes Environmental Research Laboratory's (NOAA-GLERL) and the University of Michigan's Cooperative Institute for Great Lakes Research's (CIGLR) (Boegehold et al., 2023; CIGLR and NOAA-GLERL, 2019), the International Field Year for Lake Erie (IFYLE) (IFYLE, n.d.), Environment and Climate Change Canada's (ECCC) Nearshore Nuisance Benthic Algae Dataset in Lake Erie (ECCC, 2012), the Saginaw Bay

Multi-stressor study (Stow et al., 2020), the EPA Great Lakes National Program Office's (GLNPO) Lake Erie Dissolved Oxygen Monitoring Program for 2018, and data from the most recent Cooperative Science and Monitoring Initiative (CSMI) field years for Lakes Huron (2017) (Wick et al., in press), Ontario (2018) (Sullivan and Gurdak, 2022), and Michigan (2021).  $K_d(\text{PAR})$  data produced by the color-producing agent algorithm (CPA-A) from the Visible Infrared Imaging Radiometer Suite (VIIRS) satellite (NOAA CoastWatch, 2023) in Lake Erie for 2018 were also used to assess  $K_d$ -WQ regressions (ESM Table S1). PAR profiles and water quality data were used in four distinct ways:

1. To estimate  $K_d(\text{PAR})$  from PAR profiles using equation (1)
2. To develop regressions capable of predicting  $K_d(\text{PAR})$  from water quality variables (e.g., equation (2);
3. To assess prediction of subsurface PAR using PAR profiles across lakes Erie, Ontario, Huron, and Michigan; and
4. To assess the performance of selected  $K_d$ -WQ regressions when used with a Lake Erie biophysical model for 2018

We applied quality assurance criteria to PAR profiles and concurrently-collected water quality data. PAR profiles were often noisy near the surface due to the instrument interacting with the water surface. We removed PAR measurements within the top 0.5 m of the profile. To focus our analysis on the euphotic zone, we removed all PAR measurements  $< 0.5 \mu\text{E m}^{-2}\text{s}^{-1}$ . To estimate  $K_d(\text{PAR})$ , we fit two regressions sequentially. First, we used equation (1) by fitting a linear least-squares regression of log-transformed PAR over depth on the top 5 m of the profile to estimate incident surface PAR ( $I_0$ ). If  $I_0$  estimates exceeded



3000  $\mu\text{E m}^{-2}\text{s}^{-1}$ , the highest intensity expected at the water surface, then they were forced to 3000  $\mu\text{E m}^{-2}\text{s}^{-1}$  for the sake of estimating a realistic euphotic zone depth. We then fit equation (1) in the same manner previously described over PAR measurements within the euphotic zone, which ranged from the surface to the depth at which irradiance is equal to 10% of  $I_0$  (Jerome et al., 1983; Lee et al., 2005), to estimate  $K_d(\text{PAR})$ .  $K_d(\text{PAR})$  was averaged across casts for sampling events with replicate casts at the same location and time. Following estimation of  $K_d(\text{PAR})$  for each PAR profile, we screened all remaining profiles for quality using a combination of automated screening and visual inspection of the fitted functions. Profiles that lacked concurrent water quality data were removed except the EPA GLNPO DO survey data and NOAA GLERL's and CIGLR's Western Lake Erie Monitoring Program data for 2018, which were used to assess  $K_d(\text{PAR})$  prediction using simulated water quality variables and thus did not require concurrent water quality data. Nineteen additional profiles and water quality samples were removed due to influential outliers in TSS, CHL, and  $K_d(\text{PAR})$ . Twenty-six instances of small-magnitude negative NVSS estimates in the water quality dataset were set to zero so those samples and concurrent profiles in oligotrophic waters would be retained for our analyses. Of the initial 1301 sampling events and 1453 PAR profiles in our cross-lake dataset, we retained 995 PAR profiles from 866 sampling events (129 profiles were replicates); 742 sampling events had concurrent water quality samples sufficient for fitting models and cross-validation, and the other 124 profiles were sufficient for model assessment (ESM Table S1).

Lake Erie made up the majority of the dataset's 866  $K_d(\text{PAR})$  estimates ( $n = 710$ , 82% of  $K_d(\text{PAR})$  estimates in dataset), followed by Lakes Huron ( $n = 92$ , 10.6%), Ontario ( $n = 57$ , 6.6%), and Michigan ( $n = 7$ , 0.8%) (Fig. 1, ESM Table S1). Lakes Michigan and Huron data were grouped together ("Mich-Huron") in subsequent analyses due to the small amount of Lake Michigan observations and the central and eastern basins of Lake Erie were aggregated into one basin for the purpose of our analyses. Saginaw Bay observations were distinguished from Michigan-Huron observations.

### 2.3. $K_d$ -WQ regression fitting

We fit 68 regressions to predict  $K_d(\text{PAR})$  from CHL, DOC, NVSS, VSS, and/or TSS in R (Version 4.2.3) using ordinary least squares linear regressions, segmented regressions (R package "segmented"), regression trees (R package "rpart"), and power law regressions on up to 742 paired water quality and  $K_d(\text{PAR})$  data from the compiled cross-lake dataset. Power laws and linear regressions have been commonly used to relate  $K_d(\text{PAR})$  to water quality (Limnotech, 2021; Morel, 1988; Weiskerger et al., 2018). We adjusted predicted  $K_d(\text{PAR})$  estimates from the power law models using a retransformation bias correction factor ( $\epsilon_{adj}$ ) equal to the natural exponent of variance in the residuals (Stow et al., 2006). Segmented regressions fit ordinary least square regressions over two intervals of independent variable data separated by a fitted threshold that maximizes the difference in the predictor's coefficient while ensuring continuity in the predicted relationship. In comparison to ordinary linear regression, segmented regressions can be useful in  $K_d(\text{PAR})$  modeling because they may be able to better capture non-linearity in the relationship between  $K_d(\text{PAR})$  and water quality as concentrations of optically active constituents approach zero by applying two different slopes based on the concentration of an important predictor (Smith and Baker, 1978). In that sense, it serves as an alternate approach to the power law that is also potentially capable of resolving non-linearity. We evaluated both univariate and multivariate segmented regressions. For multivariate segmented regressions, the segmented predictor variable (i. e., the independent variable with two slopes) was that which produced the largest  $r^2$  when related to  $K_d(\text{PAR})$  via simple linear regression. We

used 10-fold cross-validation to evaluate model skill in predicting  $K_d(\text{PAR})$  outside of the training dataset. Training and testing folds ranged between 543–668 and 58–75 observations, respectively. Model-predicted  $K_d(\text{PAR})$  ( $\hat{K}_d$ ) and observed  $K_d(\text{PAR})$  ( $K_d$ ) were used to calculate root mean square error (RMSE) and the squared Pearson correlation coefficient ( $r_{adj}^2$ ) adjusted for the number of observations ( $n$ ) and number of independent variables ( $p$ ) (ESM Equations S1 and S2).  $K_d$ -WQ regressions scoring in the top 50% of RMSE and  $r_{adj}^2$  were identified as candidate models that could be used in mechanistic ecosystem models.

We then calculated the relative contribution of each water quality variable to  $K_d(\text{PAR})$  using multivariate linear regressions selected from the list of candidate models. This was done by dividing the product of a given predictor's coefficient and concentration by the difference of the predicted  $K_d(\text{PAR})$  and attenuation due to water. Contribution of a given water quality variable to  $K_d(\text{PAR})$  ( $K_{wq}$ ) for an individual sampling event was calculated as:

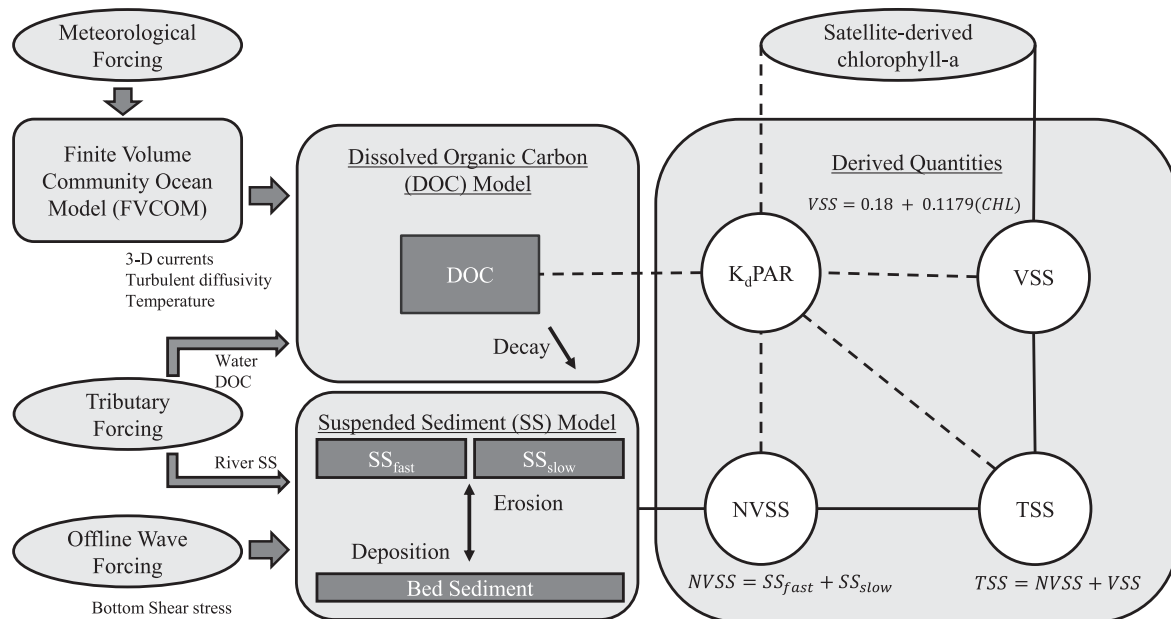
$$\% \text{ Contribution to } K_d(\text{PAR}) = 100 \left( \frac{K_{wq}[WQ]}{(K_d(\text{PAR}) - K_w)} \right) \quad (3)$$

### 2.4. Assessment of subsurface PAR prediction with observed water quality

Four of the top performing  $K_d$ -WQ regressions from the cross-validation were fit to the entire cross-lake dataset and further evaluated against two published  $K_d$ -WQ regressions trained on western Lake Erie data including the top performing multiple linear regression from Weiskerger et al. (2018) which used DOC, NVSS, and CHL as predictors (hereafter referred to as Weiskerger et al., 2018), and a power law function with TSS as the predictor used in the Lake Erie Ecosystem Model (hereafter referred to as LEEM-Kd; Limnotech, 2021; Verhamme et al., 2016). Each  $K_d$ -WQ regression predicted the  $K_d(\text{PAR})$  from water quality observations and then was input into equation (1) to iteratively calculate PAR at 20 discrete depths across a given profile's euphotic zone. For sampling events with replicate casts, we compared the models' simulated PAR profile against all measured replicate profiles individually. Because the cross-lake dataset lacked PAR measurements at the surface-water interface, we initialized each PAR profile with  $I_0$  estimated from fitting equation (1) to the top 3 m of each PAR profile. The estimated intercept from that regression was input as  $I_0$ . In total 683 or 754 profiles of the 995 in the cross-lake dataset were used for this analysis depending on if the  $K_d$ -WQ regression used NVSS + VSS or TSS, respectively. We excluded 163–234 profiles due to the absence of data for any required predictor for a given regression; and an additional 78 profiles were excluded due to noisy data in the top 3 m of the profile that reduced our confidence in the resulting estimates of  $I_0$ . Model skill in simulating subsurface PAR was determined by comparing simulated PAR for each model at each depth ( $I_{mod}^i$ ) against the inverse distance weighted value of the two nearest PAR measurements ( $I_{obs}^i$ ). Model performance was evaluated and compared against each other using several skill metrics including the Nash-Sutcliffe Efficiency (NSE, ESM Equation S3), RMSE, average bias (ESM Equation S4), and relative euclidean distance (relED, Equation S5, Abdelrhman, 2017). RelED was used to convey the ability of each model to recreate an entire profile, as opposed to individual PAR values.

### 2.5. Lake Erie case study: FVCOM-SS-DOC model structure and assessment

To demonstrate the utility of our cross-lake regressions, we evaluated the skill of four selected cross-lake calibrated  $K_d$ -WQ regressions, Weiskerger et al. (2018), and LEEM-Kd in predicting  $K_d(\text{PAR})$  given simulated water quality data from a biophysical model of Lake Erie. The



biophysical model consisted of a hydrodynamics model, a SS model, and a simplified ecological model that simulated DOC dynamics. The biophysical model's structure is detailed below.

We generated meteorological forcing conditions using station-based interpolation methods, using the same methodology as the NOAA Great

We used the FVCOM sediment transport module to simulate suspended sediments (FVCOM-SS, Fig. 2). FVCOM-SS is a customizable sediment transport module capable of simulating dynamics of suspended sediments and bedload. This module can be designed to represent spatially-variable sediment bed characteristics and it allows for an unlimited amount of cohesive or non-cohesive suspended sediment state variables. FVCOM-SS also has the ability to resolve sediment stratigraphy and fluid-mud layer effects on bottom boundary layer dynamics, however we did not utilize these features in our application nor did we represent bedload. Our application was entirely focused on SS and was driven by offline outputs from the WaveWatch III model for Lake Erie (Hu et al., 2021). Bottom shear stress ( $\tau_b$ ) is calculated from wave model output and takes into account wave-current interaction in FVCOM's bottom boundary layer module. Erosion rate ( $E$ ,  $\text{kg m}^{-2} \text{s}^{-1}$ ) was modeled as a third-order polynomial (Van Prooijen and Winterwerp, 2010; Equation (4)) with the tunable erosion parameter  $M$  (Table S2), which can be modulated by sediment porosity ( $\phi$ , ESM Table S2), the bed fraction (*BedFrac*, Table S2) of the given suspended sediment variable,  $\tau_b$ , and critical shear stress ( $\tau_{\text{CR}}$ , ESM Table S2) (Ge et al., 2015; Ge et al., 2020):

$$E = \begin{cases} 0, & \frac{\tau_b}{\tau_{CR}} < 0.52 \\ M(1 - \varphi)BedFrac \left( \frac{\tau_b}{\tau_{CR}} \right)^3 + 0.904 \left( \frac{\tau_b}{\tau_{CR}} \right)^2 - 0.823 \left( \frac{\tau_b}{\tau_{CR}} \right) + 0.204, & 0.52 < \frac{\tau_b}{\tau_{CR}} < 1.7 \\ M(1 - \varphi)BedFrac \left( \frac{\tau_b}{\tau_{CR}} - 1 \right), & \frac{\tau_b}{\tau_{CR}} > 1.7 \end{cases} \quad (4)$$

Erosion began once  $\tau_b: \tau_{CR} > 0.52$ , representing incipient mobilization of sediments, and then increased with  $\tau_b$  according to the third-order polynomial formulation until  $\tau_b: \tau_{CR} > 1.7$  at which point  $E$  increased linearly with  $\tau_b$  (Van Prooijen and Winterwerp, 2010). Suspended sediments were deposited back to the bed at a constant settling rate for suspended sediment state variables included in the model ( $W_{SET}$ , ESM Table S2) and horizontally transported via advection simulated by FVCOM. We included two, non-cohesive SS variables to represent fast-settling ( $SS_{fast}$ ) and slow-settling sediments ( $SS_{slow}$ ). The former is intended to be characteristic of coarser sediments while the latter represents finer sediments. We set  $W_{SET} = 0.076$  mm/s for  $SS_{fast}$ , which was calibrated within a range of 0.03 – 0.95 mm/s reported in the literature (Hawley and Eadie, 2007; Hawley and Lesht, 1992).  $W_{SET}$  for  $SS_{slow}$  was set to 0.0076 mm/s and the initial bed fractions were calibrated to 0.9 and 0.1 for  $SS_{fast}$  and  $SS_{slow}$ , respectively. Initial bed fractions were uniformly applied to the entire model domain. In reality, settling velocity will vary with particle diameter and kinematic fluid viscosity (Cheng, 1997) and likely sediment composition as well, but our simplification of two settling rates was sufficient to achieve our objective of simulating aggregate NVSS dynamics in observations following calibration of initial bed fractions. We report on the calibration of erodibility parameters in the results and ESM. For other parameters we did not calibrate, we used Lake Erie-specific values from Valipour et al. (2017) when available (ESM Table S2).

We calibrated the SS model given the influence of SS on  $K_d(PAR)$  in Lake Erie (Weiskerger et al., 2018). Because our focus was to resolve subsurface PAR, we aimed to design FVCOM-SS so it could adequately simulate total suspended sediment concentrations near the surface and, thus, calibrated it against NVSS observations collected within the top 1 m of the water column. We calibrated the model against various data types including bias-corrected satellite suspended mineral data, *in-situ* grab sample data of NVSS, and continuous turbidity data converted to an NVSS equivalent (ESM Figure S1). Specifically, we used *in-situ* grab samples from NOAA GLERL and CIGLR's Western Lake Erie Monitoring program (Boegehold et al., 2023; CIGLR and NOAA-GLERL, 2019), continuous turbidity data measured at several locations in the western (CIGLR and NOAA-GLERL, 2019; Ohio State University's Gibraltar Island Buoy: [https://seagull-erddap.glos.org/erddap/tabledap/obs\\_56.html](https://seagull-erddap.glos.org/erddap/tabledap/obs_56.html)) and central basins (GLERL's Coastal Hypoxia Research Program [CHRP] Mooring data (NOAA-GLERL and CIGLR, 2020; Stow et al., 2023); City of Elyria Water Quality Station: [https://seagull-erddap.glos.org/erddap/tabledap/obs\\_19.html](https://seagull-erddap.glos.org/erddap/tabledap/obs_19.html)), and the CPA-A product for suspended mineral solids from the VIIRS satellite. We applied a bias-correction to the satellite SS data, which we developed from regressing satellite detected values on corresponding grab samples of NVSS (ESM Figure S2). *In-situ* SS monitoring stations are shown in Fig. 1b-c.

Suspended sediments were also input through 29 pourpoints representing 26 rivers in the model domain (Triangles in Fig. 1, ESM Table S3). Suspended sediment concentrations input from tributaries were taken from the Heidelberg University's National Center for Water Quality Research Tributary Monitoring Program for tributaries covered by that program. We assumed SS coming in from the rivers were 100% NVSS, which is a simplification that we assumed would not greatly impact model skill since wind-wave driven resuspension is the main

driver of high turbidity events in Lake Erie (Niu et al., 2018). Extended data gaps in Heidelberg-monitored tributaries were filled using concentration-flow regression trees derived from the tributary-specific Heidelberg data from 2014 to 2019. Regression trees or flow-weighted mean concentrations from nearby tributaries were also used for other tributaries identified as priorities for eutrophication management on the US side of Lake Erie that lacked intensive monitoring during 2014–2019 (i.e. Cattaraugus Creek, Vermilion River, and Grand River, OH) (Annex 4 Objectives and Targets Task Team, 2015). Suspended sediment concentrations for the Detroit River were set to the average bias-corrected concentration detected by VIIRS near the river mouth. Suspended sediment concentrations for all other tributaries were set to flow-weighted mean concentrations calculated by dividing annual load estimates from loading models like the SWAT model (Liu et al., 2016) and the LOADEST model (Robertson et al., 2018) by annual flows. For rivers that lacked sufficient observations of SS to create a tributary-specific concentration estimate, we borrowed estimates from nearby tributaries with sufficient concentration data or loading estimates.

We simulated water quality using the General Ecological Module (GEM) provided with the FVCOM code. GEM is a biophysical modeling framework that can be configured with any number of state variables from functional groups including nutrients, phytoplankton, zooplankton, particulate organic matter, and bacteria (e.g., Rowe et al., 2017; Rowe et al., 2019). We configured GEM to only include DOC (mg L<sup>-1</sup>) as a state variable and linked it to FVCOM-SS (FVCOM-SS-DOC, Fig. 2) to calculate light attenuation as a function of the water quality variables used in the  $K_d$ -WQ regressions. Local DOC source/sink terms include advection and diffusion by FVCOM and a prescribed constant decay rate ( $2.7 \times 10^{-12}$  s<sup>-1</sup>) simulated in the ecological module. DOC was initialized at 2.1 mg L<sup>-1</sup> across the model grid and input through all rivers at 5.68 mg L<sup>-1</sup> (Prater et al., 2018), except the Maumee River and Detroit River. The Maumee River's DOC concentration was defined using the average of *in-situ* measurements of DOC at NOAA GLERL's WE2 and WE6 stations and the Detroit River's DOC was set to 2.37 mg L<sup>-1</sup> (Prater et al., 2018, ESM Table S3). Simulated DOC was compared against surface observations collected by ECCC, the Great Lakes Intake Program, and NOAA GLERL's weekly monitoring program in western Lake Erie from May thru October of 2018. For simplicity in this application focused on calibrating and assessing model components contributing to simulation of subsurface PAR, CHL was prescribed as spatially-variable and temporally-constant concentrations representing the average bias-corrected CHL values from VIIRS between April 1st, 2018 and September 30th, 2018. In future work, the suspended sediment and light attenuation model components can be incorporated into ecological models with dynamic nutrient and chlorophyll concentrations, which was beyond the scope of this study. The bias correction was empirically derived by relating satellite-detected CHL and corresponding *in-situ* extracted CHL ( $CHL = 0.77[\text{satellite CHL}] - 0.495$ ,  $r^2 = 0.6$ ). VSS was estimated from an empirical relationship to CHL derived from Lake Erie data within our cross-lake dataset ( $VSS = 0.18 + 0.1179 \cdot CHL$ ,  $r^2 = 0.82$ ) and added to NVSS simulated by the SS model to estimate TSS.

## 2.6. Lake Erie case study: Assessment of $K_d(\text{PAR})$ prediction with simulated water quality from FVCOM-SS-DOC and the sensitivity of primary production to $K_d$ -WQ regression

Following calibration of the SS model, we ran FVCOM-SS-DOC for 2018 and used the outputs to predict light attenuation with each of the six aforementioned  $K_d$ -WQ regressions at stations in the western and central basins with  $K_d(\text{PAR})$  observations. Model skill was reported as NSE.

Using the same biophysical model outputs, we estimated mean daily primary production (PP,  $\text{mg C m}^{-2} \text{ d}^{-1}$ ) over the year. Mean daily primary production was estimated using  $K_d(\text{PAR})$  predicted by each of the six  $K_d$ -WQ regressions and the phytoplankton production-irradiance curve from the Great Lakes Primary Production Model (Equation (5)) (Fahnenstiel et al., 1989; Fahnenstiel et al., 2016; Lang and Fahnenstiel, 1996) with the satellite-derived average CHL concentrations used in the biophysical model and PAR irradiance ( $I$ ) at each of the 20 vertical sigma layer depths ( $z$ ) as simulated by each of the four cross-lake and two published western Lake Erie  $K_d$ -WQ regressions (LEEM-Kd, Weiskerger et al., 2018). In Lake Erie P-I curves, Depew et al. (2006) determined a photoinhibition parameter significantly greater than zero in only 4% of cases, thus we neglected photoinhibition. Primary production was integrated across depths and time ( $t$ ) to estimate cumulative primary production between April 1st 00:00 GMT and September 30th 23:00 GMT, 2018 ( $\text{mg C m}^{-2}$ ) and then divided by the total number of days in the model run (182.9583) to estimate PP.

$$PP = \frac{\sum_t \sum_z \text{CHL}_{z,t} P_{\max}^B \left(1 - \exp\left(-\frac{\alpha I_{z,t}}{P_{\max}^B}\right)\right) \Delta z \Delta t}{182.9583} \quad (5)$$

Maximum photosynthetic rate at light saturation ( $P_{\max}^B$ ,  $\text{mg C mg Chl}^{-1} \text{ h}^{-1}$ ) and the initial linear slope for the P-I relationship at low irradiances ( $\alpha$ ,  $\text{mg C mg Chl}^{-1} \text{ h}^{-1}$ ) were empirically estimated as functions of temperature at each sigma layer and time step ( $T_{z,t}$ ) and seasonality (Equations 6, 7, ESM Figures S3-S6), respectively, using 12 seasonal

mean temperature,  $P_{\max}^B$ , and  $\alpha$  values representing nearshore and offshore for six seasons as reported by Depew et al. (2006) for the eastern basin of Lake Erie:

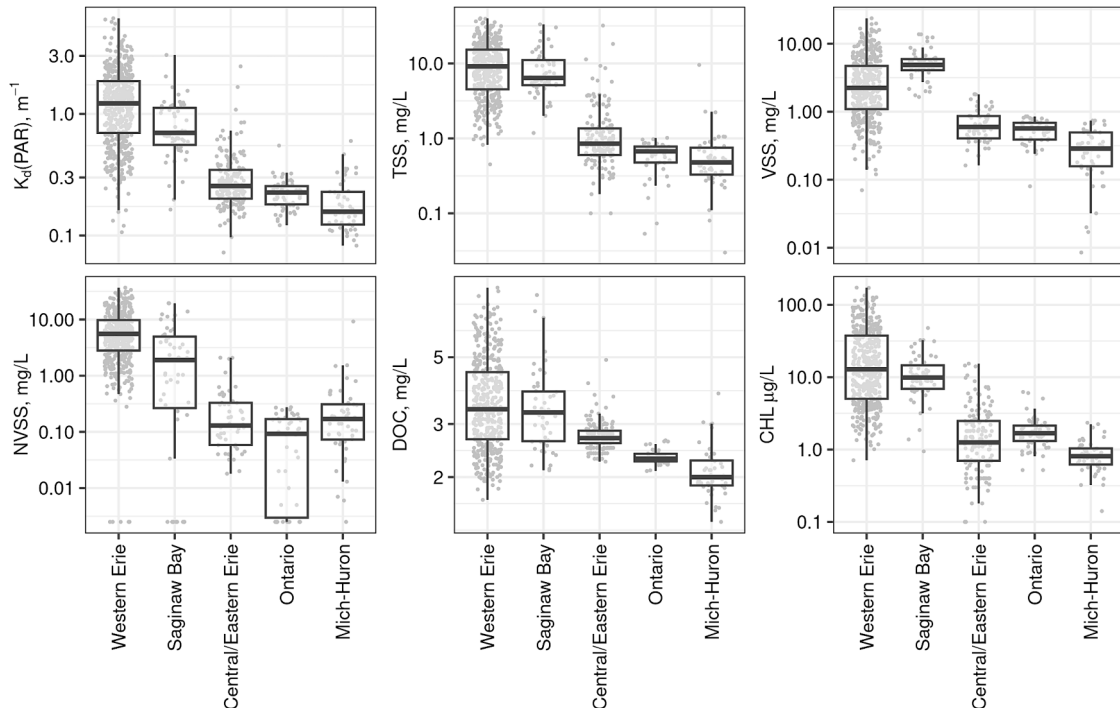
$$P_{\max}^B = 1.84T_{z,t}^{0.31}, \text{ Adjusted } r^2 = 0.59 \quad (6)$$

$$\alpha = 9.55 + 0.14 \cos\left(2\pi \left(\frac{\text{Month}}{12}\right)\right) - 3.365 \sin\left(2\pi \left(\frac{\text{Month}}{12}\right)\right), \quad \text{Adjusted } r^2 = 0.52 \quad (7)$$

Month was assigned the average month number (0–11 representing December–November) for each season based on the number of sample events that occurred in each month. For example, the winter seasonal mean values reflected two sampling events in March (month number = 3), one in February (month number = 2), and one in December (month number = 0), which translated to an average month number of 2. These empirical relationships and fits are shown in ESM Figures S3-S6. To demonstrate the sensitivity of PP to light prediction, we report the percent difference in primary production estimates derived from the different  $K_d$ -WQ regressions relative to one of the top-performing  $K_d$ -WQ regressions that was skilled across the gradient of optical conditions in Lake Erie.

## 2.7. Lake Erie case study: Assessment of simulated spatial patterns of Lake Erie $K_d(\text{PAR})$ against satellite estimates

We then implemented one of the top performing cross-lake calibrated  $K_d$ -WQ regressions from our previous analyses in the FVCOM-SS-DOC code and ran it for the entire calendar year of 2018. In this implementation, FVCOM-SS-DOC predicts  $K_d(\text{PAR})$  from concurrently simulated and derived water quality variables at each vertical sigma layer. The model then estimates PAR at the bottom of each sigma layer by inputting predicted  $K_d(\text{PAR})$  into equation (1) paired with  $I_0$  values, which represents incoming PAR from the previous layer for non-surface layers. For the surface layer,  $I_0$  is provided by the meteorological forcing file created from interpolated observations. We assessed the FVCOM-SS-



**Fig. 3.** The distribution of light attenuation and water quality variables included in our cross-lake dataset by basin. Variables include the attenuation coefficient of photosynthetically active radiation ( $K_d(\text{PAR})$ ), total suspended solids (TSS), volatile suspended solids (VSS), non-volatile suspended solids (NVSS), dissolved organic carbon (DOC), and chlorophyll- $a$  (CHL).  $K_d(\text{PAR})$  was estimated by fitting an exponential function (equation (1)) to vertical profiles of PAR.



DOC's ability to simulate spatial patterns of  $K_d(\text{PAR})$  and compared it against estimates from the CPA-A collected by VIIRS.

### 3. Results

#### 3.1. Water quality and PAR profile data

$K_d(\text{PAR})$  estimates were derived by fitting equation (1) to vertical profiles of subsurface PAR. On average, equation (1) explained 98% of the variance of subsurface PAR in profiles retained for analyses ( $n = 866$ ). The range of  $K_d(\text{PAR})$  values varied by lake and basin (Fig. 3), with 90.3 % of estimates  $< 0.3 \text{ m}^{-1}$  originating in the more transparent waters of lakes Huron, Michigan, Ontario, and the central and eastern basins of Lake Erie. Conversely, 88.7 % of observations  $\geq 0.3 \text{ m}^{-1}$  were from the more turbid waters of western Lake Erie and Saginaw Bay.  $K_d(\text{PAR})$  estimates from the *in-situ* cross-lake dataset ranged from 0.072–6.1  $\text{m}^{-1}$  with a median of 0.70  $\text{m}^{-1}$  and mean ( $\pm \text{sd}$ ) of  $1.0 \pm 0.90 \text{ m}^{-1}$  ( $n = 866$ , Fig. 3). The coefficient of variation (CV) of  $K_d(\text{PAR})$  for each basin was as follows: 78% (central/eastern Erie), 66.2% (western Erie), 59% (Saginaw Bay), 54.2% (Mich-Huron), and 30% (Ontario). CV across the entire dataset was 90 %. Water quality variables also varied by lake and basin (Fig. 3). Western Lake Erie had the highest median concentrations of all water quality variables except VSS, which was highest in Saginaw Bay with a median concentration of 4.87  $\text{mg L}^{-1}$ . Lakes Michigan, Huron, Ontario, and the central and eastern basins of Lake Erie had similar concentrations for all water quality variables. Western Lake Erie and Saginaw Bay were the only basins with TSS, VSS, and NVSS medians greater than 1  $\text{mg L}^{-1}$  and average CHL greater than 10  $\text{ug L}^{-1}$ . Across the entire dataset, variability in water quality concentrations relative to their mean values was similar for TSS (CV = 1.04), VSS (CV = 1.14), NVSS (CV = 1.16), and CHL (CV = 1.52). Dissolved organic carbon had the smallest variability relative to its mean (CV = 0.37).

#### 3.2. $K_d$ -WQ regression fitting

We assessed a total of 68  $K_d$ -WQ regressions (ESM Figure S7). The number of  $K_d(\text{PAR})$  estimates used in the cross-validation varied between 610–742 depending on the number of available observations of the independent variables. The abundance of western Lake Erie data in our dataset may bias the evaluation of  $K_d$ -WQ regressions towards those most capable in that basin. To compensate, we report model RMSE at lower  $K_d(\text{PAR})$  values typical in non-western Lake Erie waters ( $K_d(\text{PAR}) < 0.3 \text{ m}^{-1}$ ) to show which regressions have utility in more transparent waters. There were 27  $K_d$ -WQ regressions that scored above the median  $r_{\text{adj}}^2$  and RMSE when predicting across the entire data range and also scored above the median RMSE when predicting across  $K_d(\text{PAR}) < 0.3 \text{ m}^{-1}$  (ESM Table S4, Figure S8). Two regression trees were initially included in the list of top regressions, but additional inspection of prediction-observation comparisons indicated these regressions were not capable of predicting  $K_d(\text{PAR}) < 0.3 \text{ m}^{-1}$  and were removed from the list of regressions reported in ESM Table S4.

The resulting 25  $K_d$ -WQ regressions with the greatest skill consisted of 10 multiple linear regressions, 10 segmented multiple linear regressions, and 5 power law regressions (ESM Table S4, Figure S8). The

linear regression using VSS, NVSS, DOC (all in  $\text{mg L}^{-1}$ ), and CHL ( $\text{ug L}^{-1}$ ) as predictors was most accurate in predicting  $K_d(\text{PAR})$  (Table 1) followed by several other multiple linear regressions and segmented regressions. The difference among the candidate regressions'  $r_{\text{adj}}^2$  and RMSE values was small. The segmented regressions on VSS + NVSS + DOC + CHL, NVSS + DOC + CHL, and power laws on TSS + DOC and TSS + DOC + CHL were among the most skilled of their respective forms (ESM Table S4, Figures S7–S8). Multivariate linear regressions, segmented regressions, and power laws that used TSS + DOC + CHL as predictors were selected for further evaluation because these models were among the most skilled models within each form and using TSS-based models instead of NVSS + VSS-based models allowed us to maximize use of the cross-lake dataset since there were 93 more observations of TSS than VSS and NVSS. To evaluate the differential effect of VSS and NVSS on light attenuation prediction, we also included the multiple linear regression on NVSS + VSS + DOC + CHL in further analyses as a comparison to the TSS-based models we were evaluating. Skill metrics from the cross-validation and model fits for the four selected  $K_d$ -WQ regressions are shown in Table 1 and Fig. 4, respectively. Formulas for the linear regression on TSS + DOC + CHL (equation (8)), VSS + NVSS + DOC + CHL (equation (9)), power law (equation (10)), and segmented regression (equation (11)) models are reported below and in ESM Table S4 along with cross-validation skill metrics:

$$K_d\text{PAR} = 0.082(\text{TSS}) + 0.102(\text{DOC}) + 0.004(\text{CHL}) - 0.068 \quad (8)$$

$$K_d\text{PAR} = 0.033(\text{VSS}) + 0.09(\text{NVSS}) + 0.099(\text{DOC}) + 0.008(\text{CHL}) - 0.03 \quad (9)$$

$$K_d\text{PAR} = 0.178(\text{TSS})^{0.511} * \text{DOC}^{0.439} * \text{CHL}^{0.058} * \epsilon_{\text{adj}}, \epsilon_{\text{adj}} = 1.06 \quad (10)$$

$$K_d\text{PAR} = \begin{cases} 0.092(\text{TSS}) + 0.097(\text{DOC}) + 0.004(\text{CHL}) - 0.084, & \text{TSS} < 8.027 \\ 0.079(\text{TSS}) + 0.097(\text{DOC}) + 0.004(\text{CHL}) + 0.025, & \text{TSS} \geq 8.027 \end{cases} \quad (11)$$

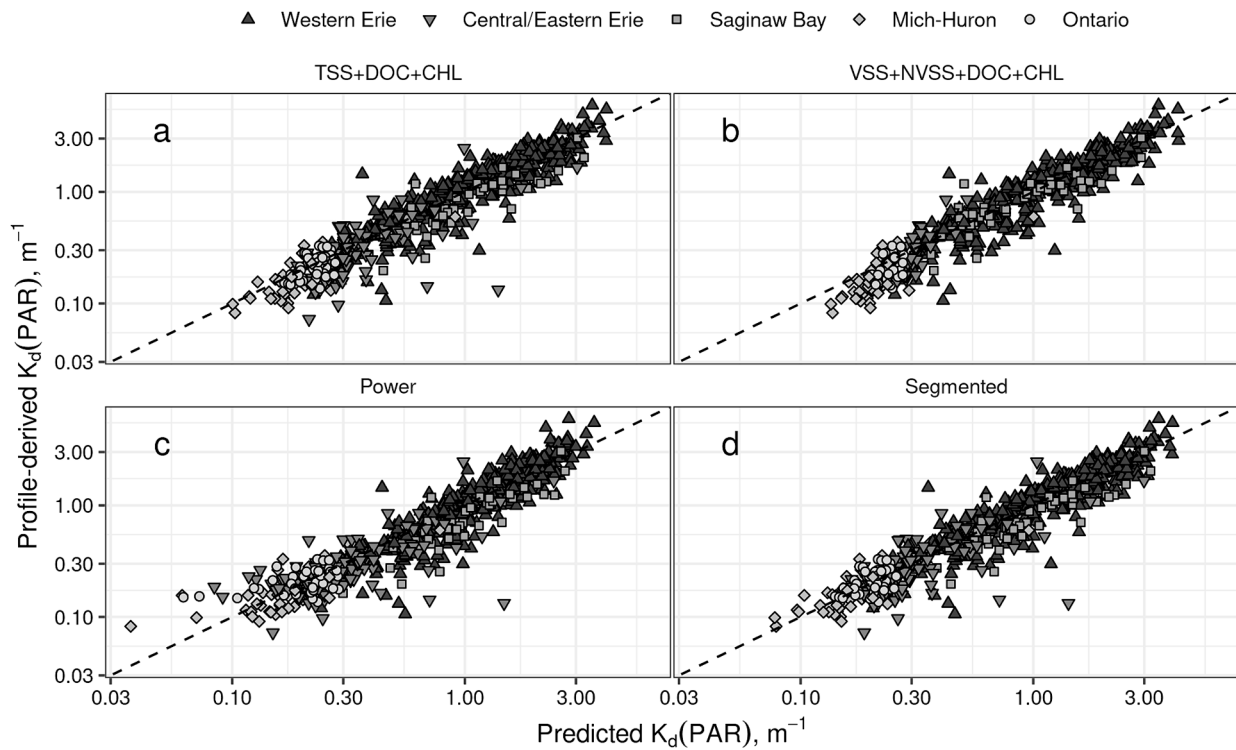
We calculated the relative contribution of each water quality variable to  $K_d(\text{PAR})$  using the multivariate linear regressions listed as equations 8–9. Attenuation due to water was negative in both equations (8) and (9), which likely indicates a negligible contribution to  $K_d(\text{PAR})$  and, therefore, we subtracted it from overall  $K_d(\text{PAR})$  in the denominator to investigate the water quality drivers of  $K_d(\text{PAR})$  and maintain their relative contributions  $\leq 100\%$ . In both equations (8) and (9), DOC was the primary factor affecting light attenuation in Lakes Michigan-Huron, Ontario, and the central and eastern basins of Lake Erie (Fig. 5). In equation (8), TSS was the main contributor to  $K_d(\text{PAR})$  in western Lake Erie and Saginaw Bay. Disaggregating TSS into NVSS and VSS in equation (9) demonstrated that NVSS was more impactful on  $K_d(\text{PAR})$  than VSS in western Lake Erie, but these constituents had relatively equal contributions in Saginaw Bay although NVSS' contribution was more variable in that basin. Regressions using VSS + NVSS in place of TSS appeared to have greater sensitivity to CHL; however, interpretation of this effect is limited owing to the fact that VSS + NVSS based regressions were based on a smaller sample size (610) than the TSS-based regressions (703). Most of the VSS + NVSS observations came from western Lake Erie and Saginaw Bay. DOC was also important in western

**Table 1**

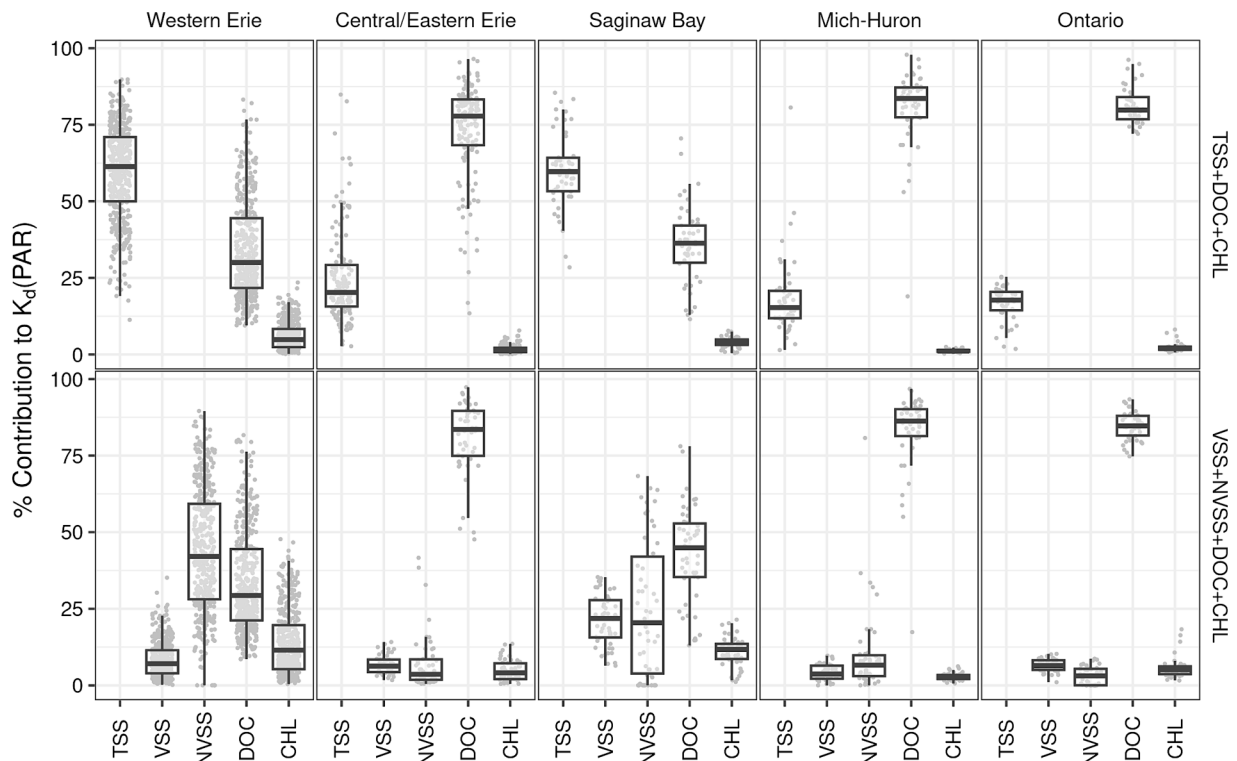
Skill of four photosynthetically active radiation attenuation ( $K_d(\text{PAR})$ ) regressions on water quality ( $K_d$ -WQ regressions) when predicting on withheld data as determined via 10-fold cross-validation. Skill metrics are reported for two ranges of data: the entire dataset and at  $K_d(\text{PAR}) < 0.3 \text{ m}^{-1}$ . Other candidate  $K_d$ -WQ regressions' skill metrics are reported in ESM Table S4.

Predictors	Form	Entire Dataset			$K_d(\text{PAR}) < 0.3 \text{ m}^{-1}$	
		RMSE, $\text{m}^{-1}$	Adjusted $r^2$	n	RMSE, $\text{m}^{-1}$	n
TSS + DOC + CHL	Multivariate Regression	0.34	0.8	703	0.15	191
	Power law	0.37	0.82	703	0.16	191
	Segmented	0.35	0.83	703	0.15	191
NVSS + VSS + DOC + CHL	Multivariate Regression	0.34	0.85	610	0.13	126

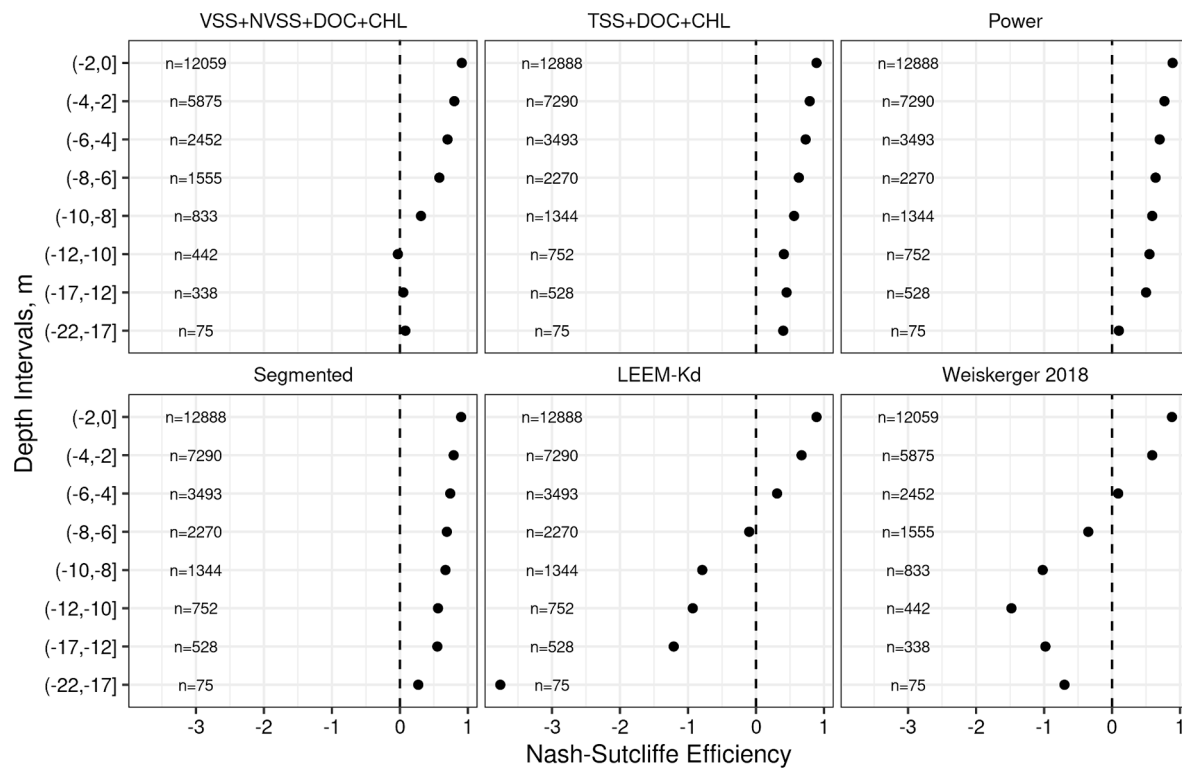




**Fig. 4.** Comparison of the attenuation coefficient for photosynthetically active radiation ( $K_d(\text{PAR})$ ) derived from PAR profiles from four Laurentian Great Lakes to  $K_d(\text{PAR})$  predicted by four selected  $K_d(\text{PAR})$ -Water Quality ( $K_d\text{-WQ}$ ) regressions including: (a) a multivariate linear regression on total suspended solids (TSS,  $\text{mg L}^{-1}$ ), dissolved organic carbon (DOC,  $\text{mg L}^{-1}$ ), and chlorophyll- $a$  (CHL,  $\mu\text{g/L}$ ); (b) a multivariate linear regression on volatile suspended solids (VSS,  $\text{mg L}^{-1}$ ), non-volatile suspended solids (NVSS,  $\text{mg L}^{-1}$ ), DOC, and CHL; (c) a power law regression on TSS, DOC, and CHL; and (d) a segmented linear regression on TSS, DOC, and CHL. The dashed 1:1 line indicates perfect agreement between predictions and profile-derived values.



**Fig. 5.** Percent contribution of total suspended solids (TSS), non-volatile suspended solids (NVSS), volatile suspended solids (VSS), dissolved organic carbon (DOC), and chlorophyll- $a$  (CHL) to the attenuation of photosynthetically active radiation ( $K_d(\text{PAR})$ ) for two different multivariate regressions (equations (8), (9), top row and bottom row, respectively). Percent contribution to  $K_d(\text{PAR})$  was calculated by dividing the product of a given predictor's coefficient and corresponding concentration by the difference of the predicted  $K_d(\text{PAR})$  and the intercept (equation (3)).



**Fig. 6.** Nash-Sutcliffe efficiency of six photosynthetically active radiation (PAR) models by depth interval using the attenuation coefficient of PAR ( $K_d(\text{PAR})$ ) predicted from different  $K_d(\text{PAR})$ -Water quality ( $K_d$ -WQ) regressions. Simulated and observed PAR data comparisons were aggregated by depth intervals across profiles for this skill metric calculation.  $K_d$ -WQ regressions used in PAR models include two published regressions calibrated with western Lake Erie data (LEEM-Kd and Weiskerger 2018) and four regressions that were calibrated with our cross-lake dataset including: a multivariate linear regression on total suspended solids (TSS), dissolved organic carbon (DOC), and chlorophyll- $a$  (CHL) ("TSS + DOC + CHL"); a multivariate regression on volatile suspended solids (VSS), non-volatile suspended solids (NVSS), DOC, and CHL ("VSS + NVSS + DOC + CHL"); a power law regression on TSS, DOC, and CHL ("Power"); and segmented regression on TSS, DOC, and CHL ("Segmented").

**Table 2**

Skill of the top two performing light attenuation-water quality regressions in simulating subsurface photosynthetically active radiation (PAR) profiles when implemented in equation (1) (i.e., PAR model) for each basin included in the cross-lake dataset. Model skill metrics were calculated for each individual profile and we report the median model skill of profiles ( $n_{\text{profile}}$ ) by basin and across the whole dataset here. Skill metrics for all PAR models are reported in ESM Table S3..

Basin	$K_d$ -WQ regression used in PAR model	reLED (%)	NSE	$n_{\text{profile}}$
Western Erie	Segmented:TSS + DOC + CHL	16.63	0.83	422
	VSS + NVSS + DOC + CHL	17.61	0.81	420
Central & Eastern Erie	Power:TSS + DOC + CHL	13.83	0.92	227
	Segmented:TSS + DOC + CHL	15.90	0.89	227
Saginaw Bay	VSS + NVSS + DOC + CHL	13.98	0.89	45
	Weiskerger 2018	15.84	0.80	45
	VSS + NVSS + DOC + CHL	15.86	0.87	32
Mich-Huron	TSS + DOC + CHL	19.90	0.86	32
	VSS + NVSS + DOC + CHL	15.33	0.91	28
Ontario	Power:TSS + DOC + CHL	18.20	0.84	28
	Segmented:TSS + DOC + CHL	17.23	0.85	754
Whole dataset	VSS + NVSS + DOC + CHL	17.37	0.84	683

Lake Erie and Saginaw Bay. Median DOC contribution to  $K_d(\text{PAR})$  in equation (8) for western Lake Erie and Saginaw Bay were 30% and 36.34%, respectively, and 29.3% and 44.89% in equation (9), respectively.

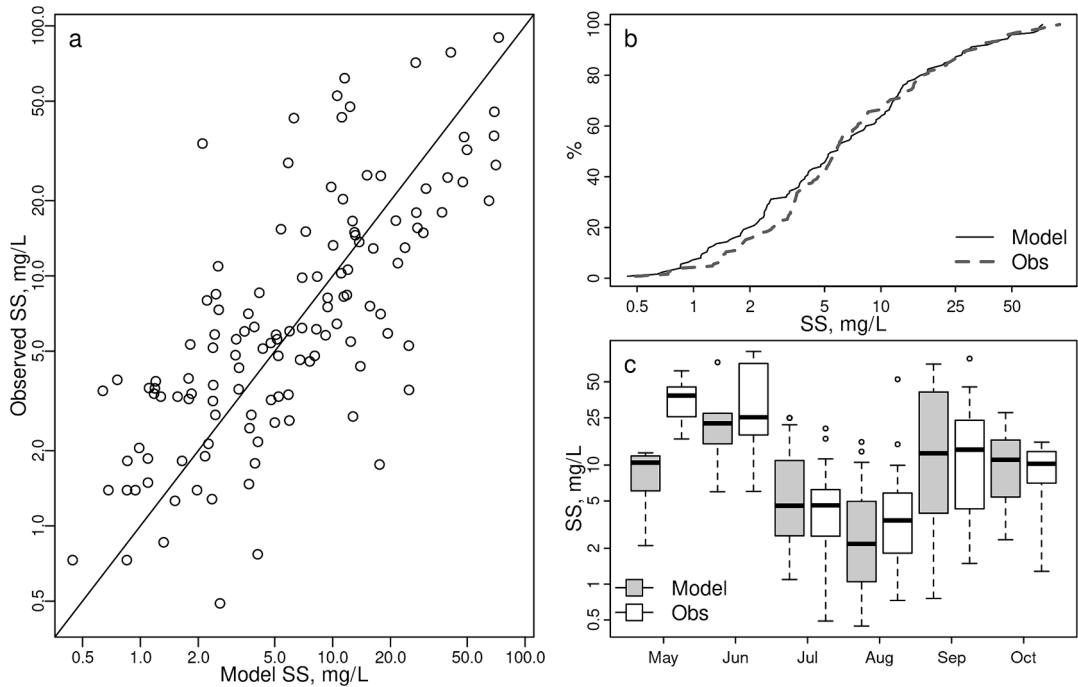
### 3.3. Assessment of subsurface PAR prediction with observed water quality

We estimated subsurface PAR (i.e. PAR model, e.g. ESM Figure S9) using equation (1) with  $I_0$  estimated from the observed PAR profile and  $K_d(\text{PAR})$  predicted from six  $K_d$ -WQ regressions. We used the four selected cross-lake  $K_d$ -WQ regressions (equations 8–11, Fig. 4) and the two published, western Lake Erie-trained  $K_d$ -WQ regressions (LEEM-Kd,

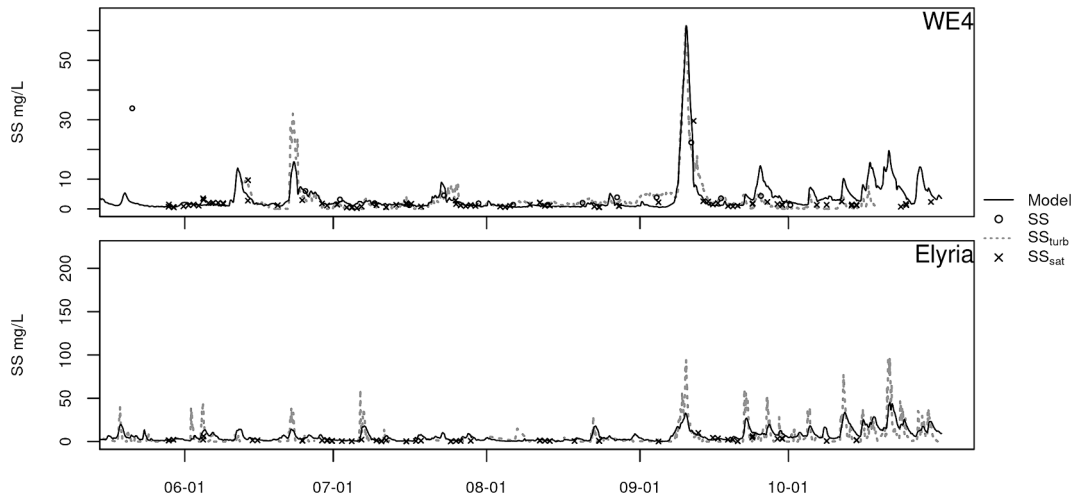
Weiskerger et al., 2018) to predict  $K_d(\text{PAR})$  from concurrent surface water quality measurements. The skill of each PAR model varied by depth and basin (Fig. 6 and ESM Figure S10, Tables 2 and ESM Table S5). All models evaluated demonstrated acceptable bias across depths in the western basin of Lake Erie (ESM Figure S10, Table S5). The four PAR models using cross-lake calibrated  $K_d$ -WQ regressions were more robust predictors of PAR across basins compared to those that used the LEEM-Kd and Weiskerger et al. (2018) regressions, but the most skilled PAR model differed by basin (Table 2). While most basins had several models that demonstrated similar median skill metrics in predicting subsurface PAR profiles, there was a larger gap in skill between PAR models assessed in Saginaw Bay. The best Saginaw Bay PAR model used the  $K_d$ -

**Table 3**  
Suspended sediment (SS) model skill metrics assessed using two different types of *in-situ* data ( $SS_{obs}$ ): non-volatile suspended solids (NVSS) from grab samples and NVSS converted from turbidity data.

Data type	Mean $SS_{obs}$ , mg L <sup>-1</sup>	$\sigma_{obs}$ , mg L <sup>-1</sup>	$\sigma_{mod}$ , mg L <sup>-1</sup>	Bias, mg L <sup>-1</sup>	RMSE	r	NSE	n	Number of stations
Grab Sample	12.14	15.99	15.80	-0.21	13.76	0.62	0.25	125	8
Converted Turbidity	6.54	11.63	12.15	1.31	9.59	0.68	0.32	26,755	13



**Fig. 7.** a)Suspended sediment (SS) model-simulated surface SS vs *in-situ* grab samples of non-volatile suspended solids (NVSS) with a 1:1 line, b) cumulative distribution of simulated surface SS and *in-situ* grab samples of NVSS, and c) monthly distribution of simulated surface SS compared to corresponding *in-situ* grab samples of NVSS. Simulated NVSS ranged from 0.44-72.94 mg/L and observed NVSS ranged from 0.49 – 89.87 mg/L.



**Fig. 8.** Suspended sediment (SS) model predicted surface SS (solid line) at the western basin site WE4 and the central basin site Elyria compared to non-volatile suspended solids (NVSS, circles) from grab samples, estimated from turbidity sensor data ( $SS_{turb}$ , dashed line), and from bias-corrected satellite data ( $SS_{sat}$ , X's).

WQ regression on VSS + NVSS + DOC + CHL followed by the Weiskerger et al. (2018) regression, but the next best model's reLED was 11.5% lower and produced a NSE that was 0.13 lower. Model skill tended to deteriorate with depth, as error in prediction of light transmission is cumulative over depth. Three of the four models calibrated with the cross-lake dataset maintained an NSE > 0 at all depth intervals

(Fig. 6). The PAR model using the regression on VSS + NVSS + DOC + CHL at 10–12 m was the only instance where a model using a cross-lake  $K_d$ -WQ regression produced a negative NSE (−0.03). Error in subsurface PAR predicted by the Weiskerger et al. (2018) and LEEM-Kd PAR models exceeded observed variance in PAR (i.e., NSE < 0) between 6 and 22 m.

### 3.4. Lake Erie case study: FVCOM-SS-DOC model evaluation

Prior to using the cross-lake calibrated  $K_d$ -WQ regressions in the Lake Erie FVCOM-SS-DOC model, we calibrated the SS model and compared simulated DOC against observations. Calibration and evaluation of FVCOM-SS-DOC is reported below.

We calibrated the two erodibility parameters ( $M$ ,  $\tau_{CR}$ ) based on data summarized by the Great Lakes Aquatic Habitat Framework (GLAHF) (Riseng et al., 2018), site-specific comparisons to observations, and spatial comparisons to satellite data. Calibration of the bed sediments resulted in five sediment types distinguished by their combination of  $M$  and  $\tau_{CR}$  (Fig. 1b). A majority of the sediment bed was classified under two of the five sediment types including western basin mud ( $M = 1.25 \cdot 10^{-6} \text{ kg m}^{-2} \text{ s}^{-1}$ ,  $\tau_{CR} = 0.15 \text{ Pa}$ ) and central and eastern basin mud ( $M = 2.5 \cdot 10^{-7}$ ,  $\tau_{CR} = 0.15 \text{ Pa}$ ). Regions of the model identified as hard/rock in the GLAHF data layer were set to be non-erodible ( $M = 0 \text{ kg m}^{-2} \text{ s}^{-1}$ ,  $\tau_{CR} = 8.25 \text{ Pa}$ ). The other two sediment types were implemented over smaller areas to refine the SS model's ability to simulate spatial patterns apparent in satellite imagery. More information on the regional bed parameterization can be found in the ESM.

The SS model compared well against *in-situ* data. Compared to NVSS grab samples, which were predominantly collected in the western basin, the model predicted SS concentrations with minimal bias across the entire range of observations (Table 3, Fig. 7a,b). The model was biased low in May and underestimated the observed variance in June 2018, but compared better to observations from July – October (Fig. 7c). All 13 stations had some turbidity data, and the model demonstrated apt skill compared against this dataset set as well (Table 3).

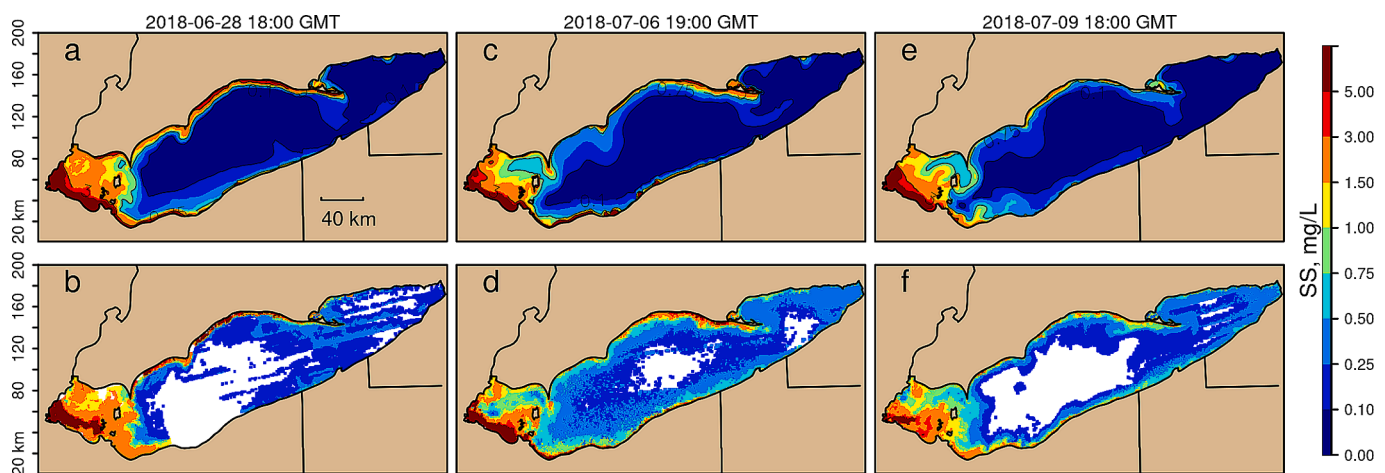
We assessed simulated SS using observed time-series of SS derived from turbidity sensor data, NVSS from grab samples, and bias-corrected satellite-derived SS (Fig. 8, ESM Figs. S11–S21). Model skill varied by station (ESM Table S6), with the best performance at the western basin sites, particularly WE4 and WE2 (Fig. 8 and ESM Fig. S11). At WE4, simulated SS was highly correlated to sensor observations ( $r = 0.86$ ), bias was low (Bias =  $0.4 \text{ mg L}^{-1}$ , RMSE =  $3.33 \text{ mg L}^{-1}$ ), and the model captured the variability of the observations well (NSE = 0.71,  $\sigma_{obs} = 6.17 \text{ mg L}^{-1}$ ,  $\sigma_{mod} = 6.32 \text{ mg L}^{-1}$ ) (Fig. 8). Resuspension events in late June, late July, and a large event in early September indicated in the sensor observations were all simulated by the model. Simulated SS compared well to the large resuspension event in September, considering the onset, rate of increase in surface SS, peak concentration, and rate of decline. Simulated SS also compared well to observations at WE2, which had turbidity data from May – October (Bias =  $-1.4 \text{ mg L}^{-1}$ , RMSE =  $11.54 \text{ mg L}^{-1}$ , NSE = 0.51,  $\sigma_{obs} = 16.57 \text{ mg L}^{-1}$ ,  $\sigma_{mod} = 12.21 \text{ mg L}^{-1}$ )

(ESM Figure S11). Simulated SS underestimated peak concentrations in large resuspension events at WE2, including one in late May and early September, but the timing and rate of change in surface SS was well simulated for most resuspension events aside from the late May event.

In the central basin, the SS model was less skilled in predicting site-specific time-series of SS relative to its performance in the western basin (ESM Table S6), but it still captured nearshore-offshore differences and simulated most resuspension events when they occurred (Fig. 8 and Figures S20–S21). Simulated SS underestimated peak concentrations at all central basin sites. The SS model simulated the lower levels of SS observed at the two offshore central basin stations, CHRP2 (ESM Figure S20) and CHRP9 (Figure S21), compared to the more nearshore site near Elyria, OH (Fig. 8). The SS model also underestimated converted turbidity data at the two offshore stations in the central basin during periods of resuspension (ESM Figures S20–S21). However, there was disagreement between the bias-corrected satellite data and the converted turbidity data at times and the SS model tended to show stronger agreement with the satellite data.

The SS model simulated common spatial patterns of SS and episodic resuspension events observed in VIIRS imagery (Fig. 9 and ESM Fig. S22). Compared to satellite data across the whole lake, the SS model was able to predict SS concentrations with low bias and positive NSE, although skill metrics varied by basin (Table 4). Common spatial patterns of SS in the satellite imagery that were simulated included a persistent difference in SS concentrations in the northern half of the western basin compared to the southern half (Fig. 9a–f), a SS plume extending from the western basin's eastern limit to Sandusky, OH (Fig. 9a–b,e–f), and elevated SS along the central basin's northern shoreline (Fig. 9a–f), which varied in extent and concentration. The SS model also simulated spatial patterns associated with a large resuspension event that was detected by all three types of observations in early September 2018 (ESM Figures S11–S22). The model's performance during this event is described further in the ESM.

We evaluated DOC simulated by FVCOM-SS-DOC using 270 observations from three sampling programs in Lake Erie. FVCOM-SS-DOC overestimated average observed DOC by  $0.15 \text{ mg L}^{-1}$  throughout the lake and had an RMSE of  $0.49 \text{ mg L}^{-1}$ . Simulated DOC in the central and eastern basins agreed well with observations but was biased high in the western basin and underestimated observed variability (ESM Figure S23). The observations indicated lower DOC in the western basin compared to the other two basins, which was not simulated by FVCOM-SS-DOC.



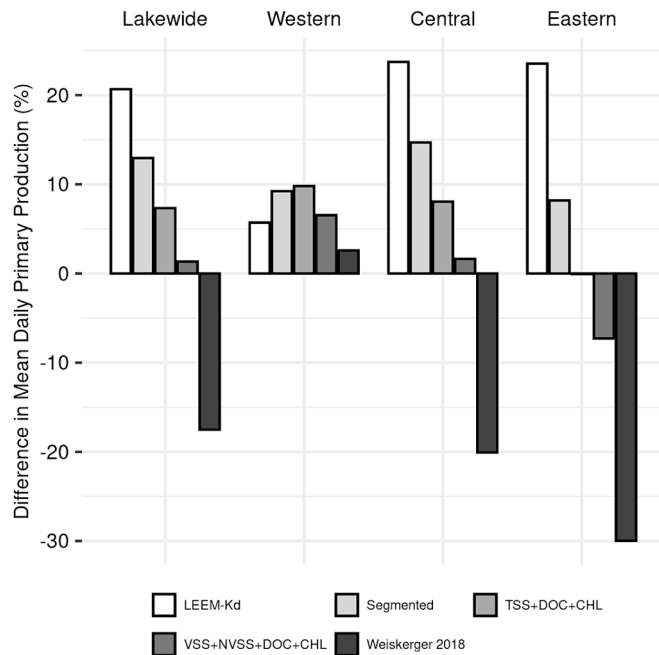
**Fig. 9.** Comparison of common spatial patterns of surface suspended sediments (SS) simulated by the Finite Volume Community Ocean Model's Suspended Sediment model (FVCOM-SS; a,c,e) to satellite-derived estimates from the color-producing agent algorithm (b,d,f) as represented in three examples from June 28, July 6th, and July 9th 2018.



**Table 4**

Suspended sediment (SS) model skill in comparison to bias-corrected satellite suspended mineral data ( $SS_{sat}$ ) across Lake Erie and its three basins. Satellite data spanned from 2018 to 05-28 to 2018-9-29.

Basin	Mean $SS_{sat}$ , mg L <sup>-1</sup>	Bias, mg L <sup>-1</sup>	RMSE, mg L <sup>-1</sup>	r	NSE	$\sigma_{sat}$ , mg L <sup>-1</sup>	$\sigma_{mod}$ , mg L <sup>-1</sup>	n
Western	3.04	0.85	4.75	0.65	0.08	4.96	6.00	85,805
Central	0.77	-0.21	1.5	0.55	0.19	1.66	1.42	155,437
Eastern	0.41	-0.2	0.51	0.42	-0.4	0.44	0.45	62,498
Lake wide	1.34	0.09	2.75	0.68	0.21	3.09	3.7	303,740



**Fig. 10.** Sensitivity of lake wide and basin-specific estimated daily primary production (PP) from April 1st – September 30th, 2018 using the Great Lakes primary production model (Fahnenstiel et al., 1989; Fahnenstiel et al., 2016; Lang and Fahnenstiel, 1996) and the attenuation coefficient of photosynthetically active radiation ( $K_d(PAR)$ ) predicted by four cross-lake calibrated  $K_d(PAR)$ -water quality ( $K_d$ -WQ) regressions and two western basin calibrated  $K_d$ -WQ regressions (LEEM-Kd and Weiskerger 2018). Water quality and suspended sediment data simulated by the Finite Volume Community Ocean Model's Suspended Sediment and Dissolved Organic Carbon Model (FVCOM-SS-DOC) were used as inputs to the primary production model. All primary production estimates were reported as the percent difference relative to the primary production prediction from the cross-lake calibrated power law regression on total suspended solids (TSS), DOC, and chlorophyll.

### 3.5. Lake Erie case study: Assessment of $K_d(PAR)$ and primary production prediction with simulated water quality from FVCOM-SS-DOC

We used the four selected cross-lake  $K_d$ -WQ regressions (equations 8–11, Fig. 4) and the two published, western Lake Erie-trained  $K_d$ -WQ regressions (LEEM-Kd, Weiskerger et al., 2018) to predict  $K_d(PAR)$  from water quality data simulated by FVCOM-SS-DOC. We also assessed the sensitivity of lake-wide mean estimates of daily primary production

rates to selection of  $K_d$ -WQ regression used with FVCOM-SS-DOC. All models produced positive NSE values in simulating  $K_d(PAR)$  when given simulated water quality and SS concentrations (ESM Figure S24). The power law on TSS + DOC + CHL produced the best fit to profile-derived  $K_d(PAR)$  values. The simulated water quality fields we prescribed for this analysis limited the skill of  $K_d$ -WQ regressions in the central basin. The lack of variation in the CHL and VSS fields prevented  $K_d$ -WQ regressions from being able to predict variation in  $K_d(PAR)$  below 0.25 m<sup>-1</sup>, which was characteristic of some central basin profiles used in this assessment.

The sensitivity of primary production to  $K_d$ -WQ regression was evaluated as the percent difference relative to the primary production prediction from the power law regression on TSS + DOC + CHL (Fig. 10). All cross-lake calibrated  $K_d$ -WQ regressions were able to predict  $K_d(PAR)$  across the optical gradient in the LGL, but the power law regression was chosen as the baseline model because it was among the best performing in Lake Erie when given observed or simulated inputs (Tables 2 and ESM S5, Figure S24). Weiskerger et al. (2018) produced the lowest lake wide primary production estimates that were 17.5% lower than the cross-lake calibrated power law, whereas LEEM-Kd produced the highest relative estimate (Percent difference = 20.67). The other cross-lake calibrated regressions differed from the power law's predicted primary production by 1.3–12.9%. Percent difference in primary production estimates in the central and eastern basins followed the ranking described for the lake wide estimates with LEEM-Kd and Weiskerger et al. (2018) being  $\geq 20\%$  different than primary production derived from the power law. Primary production estimates in western basin were comparable amongst all  $K_d$ -WQ regressions (2.5–9.8% different than power law).

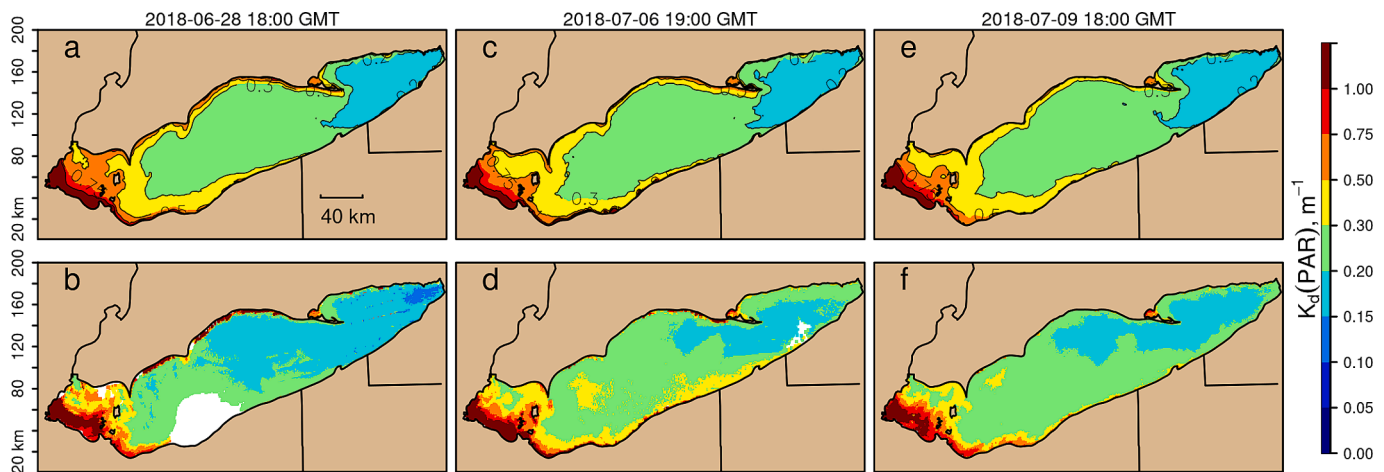
### 3.6. Lake Erie case study: Assessment of simulated spatial patterns of Lake Erie $K_d(PAR)$ against satellite estimates

We implemented the power law regression on TSS + DOC + CHL in the 3D Lake Erie FVCOM-SS-DOC code and assessed simulated  $K_d(PAR)$  against satellite-derived  $K_d(PAR)$  estimates in 2018. We found that satellite-derived  $K_d(PAR)$  agreed well with corresponding estimates from subsurface PAR profiles ( $r^2 = 0.93$ ; ESM Figure S25), supporting their use in model skill assessment. FVCOM-SS-DOC simulated  $K_d(PAR)$  with low bias across the entire lake and produced positive NSE values (Table 5). The model's ability to simulate variability in  $K_d(PAR)$  decreased from west to east, which was also reflected in model skill metrics. The model simulated recurrent and episodic spatial patterns indicated in the satellite imagery that reflected the dynamics of the SS model. This includes the north-south difference in the western basin, increased  $K_d(PAR)$  along the north shore of the central basin, and higher

**Table 5**

Skill assessment comparing the attenuation of photosynthetically active radiation ( $K_d(PAR)$ ) simulated by a power law on total suspended solids (TSS), dissolved organic carbon (DOC), and chlorophyll-*a* (CHL) implemented in the Finite Volume Community Ocean Model's Suspended Sediment and Dissolved Organic Carbon model (FVCOM-SS-DOC, a,c,e) and corresponding satellite-derived estimates of  $K_d(PAR)$  (b,d,f). Satellite data spanned from 2018 to 05-28 to 2018-9-29.

Basin	Mean $K_d(PAR)_{sat}$ , m <sup>-1</sup>	Bias, m <sup>-1</sup>	RMSE, m <sup>-1</sup>	r	NSE	$\sigma_{sat}$ , m <sup>-1</sup>	$\sigma_{mod}$ , m <sup>-1</sup>	n
Western	0.84	0.00	0.32	0.79	0.61	0.52	0.46	66,434
Central	0.36	-0.04	0.21	0.67	0.41	0.28	0.15	151,234
Eastern	0.26	-0.04	0.2	0.36	0.09	0.21	0.09	68,679
Lakewide	0.45	-0.03	0.24	0.81	0.64	0.40	0.34	286,347



**Fig. 11.** Comparison of the attenuation of photosynthetically active radiation ( $K_d(\text{PAR})$ ) simulated by a power law on total suspended solids (TSS), dissolved organic carbon (DOC), and chlorophyll- $a$  (CHL) implemented in the Finite Volume Community Ocean Model's Suspended Sediment and Dissolved Organic Carbon Model (FVCOM-SS-DOC; a,c,e) to corresponding satellite-derived estimates of  $K_d(\text{PAR})$  (b,d,f) for three dates in late June and early July.

attenuation along the sediment plume that occasionally extends from the western basin into the Sandusky subbasin (Fig. 11). The model also simulated a gradient of decreasing attenuation from west to east and from nearshore to offshore that was indicated in the satellite data (Fig. 11).

#### 4. Discussion

Previously published implementations of  $K_d$ -WQ regressions have been calibrated over a range of  $K_d(\text{PAR})$  values that limits their performance in other parts of the LGL with different levels of water transparency thereby limiting their utility to biophysical models (Higgins et al., 2005; Limnotech, 2021; Rowe et al., 2017; Weiskerger et al., 2018). In comparison to Weiskerger et al. (2018) and LEEM-Kd, the four cross-lake calibrated regressions we evaluated were more skilled in predicting  $K_d(\text{PAR})$  and PAR in transparent waters with euphotic zones deeper than 6 m. Moreover, the increased flexibility of the cross-lake calibrated regressions did not reduce  $K_d(\text{PAR})$  and PAR simulation skill in the western basin relative to the western basin trained regressions. Having  $K_d$ -WQ regressions that can work across the water transparency gradient within a given LGL is important to simulating higher order ecosystem processes, such as primary production, which we demonstrated by using the six evaluated  $K_d$ -WQ regressions with simulated data from our simplified Lake Erie biophysical model, FVCOM-SS-DOC. Furthermore, we were able to verify that FVCOM-SS-DOC could simulate spatial patterns in SS and  $K_d(\text{PAR})$  observed in Lake Erie when implemented with one of our top-performing cross-lake  $K_d$ -WQ regressions thereby highlighting how our calibration of 25  $K_d$ -WQ regressions can advance the development of biophysical models needed to support ecosystem management objectives in the LGL.

##### 4.1. The variable light environment in the Laurentian Great Lakes

Variability in  $K_d(\text{PAR})$  within and across the LGL is substantial when considered in context of what has been observed in marine and freshwater environments. We documented  $K_d(\text{PAR})$  ranges from 0.072–6.2  $\text{m}^{-1}$  with coefficients of variation ranging between 30–78% by basin and 90% across the entire dataset. Morris et al. (1995) found  $K_d(\text{PAR})$  in 59 inland lakes across the northeastern United States, Alaska, Colorado, and Argentina ranged from 0.1–5.73  $\text{m}^{-1}$  with a coefficient of variation of 8.7%. Song et al. (2017) used an empirical relationship of Landsat reflectance to estimate  $K_d(\text{PAR})$  for 178 lakes in northeast China, which ranged from 0.49  $\text{m}^{-1}$  to 15.8  $\text{m}^{-1}$ , although only four measurements in their calibration dataset ( $n = 167$ ) exceeded 8  $\text{m}^{-1}$ . The LGL's range in

$K_d(\text{PAR})$  indicated in our dataset is comparable to what has been observed in the Atlantic Ocean and is only slightly narrower than what has been measured in open water and coastal environments across the world's major oceans and seas when pooled together (range: 0.03–9.85  $\text{m}^{-1}$ ; Son and Wang, 2015). The variability of light attenuation in the LGL underscores the importance of fitting  $K_d$ -WQ regressions to a dataset that adequately represents the range of optical conditions across the lakes.

##### 4.2. $K_d(\text{PAR})$ partitioning using $K_d$ -WQ regressions

There was a dichotomy that emerged in the partitioning of  $K_d(\text{PAR})$  in each basin based on coefficients from equations (8) and (9). 75–83% of  $K_d(\text{PAR})$  on average was due to DOC and 16–25% was due to TSS in the basins that mostly comprised of open water samples (Mich-Huron, Central/Eastern Erie, Ontario). In the most opaque basins (Lake Erie's western basin and Saginaw Bay), TSS and DOC accounted for about 60% and 35% of  $K_d(\text{PAR})$ , respectively. The contribution of DOM to  $K_d(\text{PAR})$  relative to suspended particulates in the LGL, even in western Lake Erie and Saginaw bay, are notably higher than what has been reported in other systems where this has been evaluated including the eutrophic Lakes Taihu (% Contribution of CDOM to  $K_d(\text{PAR}) = 6.8\%$ , Zhang et al., 2007) and Donghu in China (17.8–23.9%, Yang et al., 2005), the oligotrophic Dom Helvécio lake in Brazil (4.8%, Brandao et al., 2017), Arhus Bay of the Kattegat between the North and Baltic seas (17%, Lund-Hansen, 2004), the Indian River Lagoon in Florida (5–25%, Christian and Sheng, 2003), and Japan's eastern Seto Inland Sea (13%, Yamaguchi et al., 2013). The contribution of CDOM to  $K_d(\text{PAR})$  reported by Chen et al. (2014) for the Caloosahatchee River Estuary in Florida demonstrated contributions of CDOM (range: <5–90%, mean contribution: 55%) within the range of average contributions of DOC by basin in the LGL (35–83%). It should be noted that our percent contribution refers to the percent contribution of DOC to light attenuation of non-water components since our estimated  $K_w$  was negative, whereas these other studies calculated contributions of CDOM based on regressions with a forced, positive constant for  $K_w$  (0.027–0.15  $\text{m}^{-1}$ ) that accounted for an additional 0.2–10% of  $K_d(\text{PAR})$  depending on the system. Choosing a forced or fitted intercept likely would not affect conclusions regarding the relative importance of each constituent to overall light attenuation since the attenuation due to water is often lower than attenuation due to the non-water components (Chen et al., 2014).

Our findings are consistent with relative contribution of dissolved and particulate components to IOPs in the LGL documented in previous studies. Perkins et al. (2013) found CDOM and total particulates (algae

and non-algae) contributed, on average, 39% and 61% to total absorption of the 440 nm wavelength, respectively, in Lake Erie's western basin, Green Bay, and Sandusky Bay, and 57% and 40%, respectively, in open waters of Lakes Superior, Michigan, Huron, Ontario, and Lake Erie's central and eastern basins. Effler et al. (2013) found that mineralogenic particles accounted for 80% of backscattering in the western basin and Sandusky Bay but were less influential on backscattering in most open water locations. While the contributions to  $K_d(\text{PAR})$  cannot be equated to the contributions to IOPs since the influence of absorption and backscattering on total attenuation are not equal (Lee et al., 2005), our study reinforces the relative importance of dissolved organic matter in offshore waters and total particulate matter in coastal waters as determining factors of the underwater light field in the LGL.

#### 4.3. Lake Erie case study

The calibration of the SS model in FVCOM-SS-DOC was a prerequisite to modeling light attenuation in Lake Erie because SS is a major driver of light attenuation (Weiskerger et al., 2018). The SS model simulated the recurrent patterns of resuspension seen along the south shore of the western basin and the north shore of the central basin. It also simulated episodic events like that observed in September 2018. By using three different datasets to calibrate our SS model, we were able to do an extensive quantitative comparison of our model to site-specific *in-situ* data and bias-corrected satellite data. Previous studies have used converted beam attenuation (Hawley et al., 2014), backscatter and turbidity data (Khazaei et al., 2021; Lin et al., 2021), and satellite true color imagery (Limnotech, 2021; Niu et al., 2018) to supplement intermittent grab sample data for quantitatively and qualitatively evaluating SS models. Our calibration approach is similar to that employed by Khazaei et al. (2021) for their sediment transport model of Green Bay, Lake Michigan and represents an improvement on previous calibration efforts in Lake Erie SS models by expanding the number of observations for quantitative model evaluation. As a result, our SS model demonstrated skill (Tables 3 and 4) that was comparable or better than other published SS models in the LGL including LEEM ( $r^2 = 0.2$ , Bias =  $-2.2 \text{ mg L}^{-1}$  of TSS; Limnotech, 2021), Niu et al. (2018)'s western Lake Erie SS model ( $r^2 = -0.07 - 0.78$ , RMSE =  $5.1\text{--}28.2 \text{ mg L}^{-1}$  of TSS), Lin et al. (2018)'s Lake Erie sediment transport model (near bed turbidity:  $r^2 = 0.39$ ), and Khazaei et al. (2021)'s Green Bay, Lake Michigan sediment transport model (Grab samples: Bias =  $-6.57\text{--}1.38 \text{ mg L}^{-1}$  of TSS,  $r = 0.48\text{--}0.85$ ). It should be noted that these aforementioned SS models simulate TSS, which includes inorganic and organic material, whereas we are simulating only the inorganic component.

Simulated DOC and SS concentrations from FVCOM-SS-DOC were sufficiently accurate to characterize general spatial patterns in  $K_d(\text{PAR})$  when used with the cross-lake calibrated  $K_d$ -WQ regressions, but our case study shows the challenge of modeling variation in  $K_d(\text{PAR})$  in more transparent waters for which SS is not the main driver. While it was well suited to simulate  $K_d(\text{PAR})$  in the turbid, western basin of Lake Erie, FVCOM-SS-DOC was limited in its ability to predict variation in  $K_d(\text{PAR})$  in the more transparent central and eastern basins, even though our  $K_d$ -WQ regressions showed improved performance at low  $K_d(\text{PAR})$  compared to the previously published  $K_d$ -WQ regressions. The lack of dynamic particulate organic matter (CHL and VSS) in our biophysical model and the model's limited design in regards to DOC sources and sinks likely constrained FVCOM-SS-DOC's skill in areas of the lake where NVSS was not the primary driver of light attenuation. Such is the case for the central and eastern basins of Lake Erie, where DOC accounted for 75% of  $K_d(\text{PAR})$  (Fig. 5) and VSS comprised 76% of TSS on average, compared to in the western basin where it comprised only 35%. The increased relative influence of organic matter on attenuation in more transparent waters highlights the tighter coupling of phytoplankton, phytoplankton-derivatives (e.g. POM, DOM), and light in offshore waters of the LGL relative to turbid coastal areas. Our calibrated  $K_d$ -WQ regressions can be incorporated into various biophysical model

configurations, and may be expected to accurately predict  $K_d(\text{PAR})$  if optically active constituents are accurately simulated.

#### 4.4. Ecological implications

Model-predicted primary production for Lake Erie was sensitive to selection of  $K_d$ -WQ regression. All model primary production estimates for the western basin in 2018 were in relatively close agreement with each other (Fig. 10), but there was notable difference amongst the western basin calibrated and cross-lake calibrated regressions in the other two basins. Despite being both trained on western basin data, the LEEM-Kd and Weiskerger et al. (2018) regressions generated the most divergent primary production estimates of the six regressions evaluated in all basins except the western, and they respectively overestimated and underestimated all cross-lake calibrated regressions' lake wide primary production estimates. In the central and eastern basins, the LEEM-Kd and Weiskerger et al. (2018) regressions resulted in primary production estimates that were -30–20% different than estimates from the cross-lake calibrated power law regression on TSS + DOC + CHL (Fig. 10). Conversely, primary production estimated from the cross-lake calibrated regressions tended to agree with each other—with an average percent difference relative to the power law of 5.2% and 8.1% in the central and eastern basins, respectively. The divergence of LEEM-Kd and Weiskerger et al. (2018) in the central and eastern basins demonstrates the uncertainty associated with extrapolating  $K_d$ -WQ regressions trained on basin-specific data across the wide-ranging gradient of water transparency that exists within Lake Erie and the other LGL. In ecological models, biases in primary production caused by biases in subsurface PAR may affect model predictions of hypoxia (e.g., Moriarty et al., 2021), phytoplankton dynamics, and other processes indirectly or directly influenced by light.

Biophysical models with skillfully-resolved light environments can be vital tools for disentangling the role of light and other factors affecting harmful algal blooms (HABs), hypoxia, nuisance algae, and fisheries in the LGL. In Lake Erie, predicting the extent, transport, and toxicity of the bloom-forming cyanobacterium, *Microcystis aeruginosa*, has been a primary focus for biophysical model development informed by Annex 4 of the Great Lakes Water Quality Agreement (Annex 4 Objectives and Targets Task Team, 2015; Wynne et al., 2011). Differences in absorbance spectra and the ability to access light via buoyancy can confer a circumstantial competitive advantage to *Microcystis* (Huisman et al., 2004; Tan et al., 2020), but the concurrent effects of other factors, like nutrients and turbulence, obscure the role each driver plays in affecting seasonal succession of phytoplankton (Huisman et al., 2004; Liu et al., 2021). Additionally, predicting central basin hypoxia and eastern basin blooms of the filamentous green algae *Cladophora glomerata* in Lake Erie and their response to nutrient load management have been priority objectives guiding the development of biophysical models for the lake (Annex 4 Objectives and Targets Task Team, 2015; Limnotech, 2021). Subsurface PAR regulates *Cladophora* growth potential and is highly correlated with maximum sustainable biomass, which when exceeded can lead to sloughing of biomass that may ultimately contaminate beaches (Auer et al., 2010; Higgins et al., 2005). In regards to hypoxia, PAR can influence benthic dissolved oxygen by regulating oxygen production from benthic photosynthesis (Brothers et al., 2014; Brothers et al., 2017) and by affecting the growth of algae that ultimately settles out and consumes hypolimnetic oxygen through decomposition (Carrick et al., 2005; Lashaway and Carrick, 2010).

#### 5. Conclusion

By consolidating a large database of coincident PAR and water quality observations across four of the five LGL, we were able to develop improved  $K_d$ -WQ regressions across a range of optical conditions observed in this system. We provide in the ESM the cross-lake dataset used for calibration and these calibrated regressions, which may be



applied in biophysical models configured with a variety of state variables. We demonstrated how these models perform in predicting sub-surface PAR when given incident PAR and concurrent water quality measurements as well as how they perform when given water quality concentrations simulated by a simplified biophysical model of Lake Erie. The former demonstrates the overall skill of several  $K_d$ -WQ regressions when provided accurate inputs and the latter demonstrates how the  $K_d$ -WQ regressions can be applied in a biophysical model to simulate spatial and temporal patterns of  $K_d$ (PAR). These advancements are important for future development of other ecological processes in LGL biophysical models that are influenced by light. Accurately resolving light attenuation for biophysical and biogeochemical models is a fundamental development step in order to use these tools to study the ecology of coastal and open waters in the LGL.

## CRediT authorship contribution statement

**Peter J. Alsip:** Conceptualization, Data curation, Formal analysis, Investigation, Methodology, Software, Validation, Visualization, Writing – original draft, Writing – review & editing. **Mark D. Rowe:** Conceptualization, Data curation, Funding acquisition, Methodology, Project administration, Resources, Software, Supervision, Writing – review & editing. **Alexander Kain:** Data curation, Investigation, Methodology, Software, Validation, Writing – review & editing. **Casey Godwin:** Conceptualization, Data curation, Funding acquisition, Methodology, Project administration, Supervision, Writing – review & editing.

## Declaration of competing interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

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## Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.jglr.2024.102364>.

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