

1 **A model-based approach to standardizing American lobster (*Homarus americanus*)**
2 **ventless trap abundance indices**

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16 **Highlights**

- 17 • Model-based approaches were used to remove variability in U.S. American lobster abundance
18 indices.
- 19 • Accounting for site, day of year, and soak time in index construction provided insight on these
20 factors' relation to catch.
- 21 • Model- and design-based indices provided similar trends in lobster abundance, with reduced
22 variability in the indices and accounting for missing data.
- 23 • This advancement provides updated relative abundance indices for evaluation in future stock
24 assessments.

25

26 **Abstract**

27 Fishery-independent ventless trap surveys are an integral component to assessing American
28 lobster (*Homarus americanus*) population trends, as they can sample complex, heavily fished
29 habitats where most survey gear has difficulty accessing. U.S. American lobster stocks have been
30 assessed within state waters using a standardized ventless trap survey since 2006. However,
31 confounding survey attributes that may contribute to catch variability have not been investigated
32 and some discontinuity in sampling has resulted in missing estimates of abundance. We
33 constructed sex- and stock-specific generalized linear mixed models to discern the dynamics
34 between lobster catch and individual survey factors and removed these sources of variability
35 when producing continuous abundance indices. Soak time, day of year, and unique site had
36 measurable contributions to the variability in lobster catch per ventless trap. Generally, sex- and
37 stock-specific abundance indices from this model-based approach and a traditional design-based
38 approach exhibited similar trends. The two approaches' magnitudes and trends for the Gulf of
39 Maine were nearly identical. For Southern New England, model-based index trends were
40 smoother and lower in magnitude than the design-based estimates. The greatest difference in the
41 two approaches' trends for the Southern New England indices were in the early and terminal
42 years. This work serves as an example of how variability associated with fixed and random
43 effects of a survey can be accounted for when producing abundance indices used in stock
44 assessments.

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46 **Key Words:** ventless trap survey, American lobster, random effects, abundance indices

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48 **1. Introduction**

49 American lobster (*Homarus americanus*) supports the most valuable single-species
50 fishery in North America, with an ex-vessel value of \$US 1.7 billion in 2017 (DFO, 2020;
51 NOAA, 2020). Despite this economic significance, the species' dynamic life history can make it
52 difficult to infer population trends across life stages. Canadian lobster stock assessments
53 historically have relied on fishery-dependent data to derive catch per unit effort (CPUE) indices
54 for inferring population trajectories (DFO, 2013). However, using industry effort data may
55 provide misleading population trajectory signals given such CPUE data are derived primarily
56 from densely populated lobster areas where effort is concentrated, and do not necessarily include
57 the entirety of the population or stock bounds. Such dynamics have often led to hyperstability
58 when using CPUE to infer population abundance (Harley et al., 2001). Further, improvements in
59 fishing strategies and gear are constantly evolving and can influence CPUE (Smith and
60 Tremblay, 2003) yet are usually undocumented. While commercial CPUE can be useful in
61 assessing the performance of the fishery, these uncertainties have led to their exclusion as
62 abundance indices in U.S. American lobster stock assessments.

63 Historically, bottom otter trawl survey data from multi-species surveys have been used in
64 U.S. American lobster stock assessments as indicators of population trends (ASMFC, 2009;
65 ASMFC, 2015). Two of the greatest drawbacks for using otter trawl data to represent lobster
66 stock abundance are that (1) they typically do not sample in the preferable cobble or boulder
67 habitat of lobsters (Wahle and Steneck, 1991) in fear of entangling the trawl, and that (2) they
68 cannot fish in areas or paths where fixed fishing gear exists (Smith and Tremblay, 2003). Trap
69 surveys have often been considered a better method of assessing the structure-oriented lobster
70 population (Tremblay et al., 2009). To address this concern for the U.S. lobster stocks, states

71 within the northeast U.S. initiated complementary lobster ventless trap surveys designed to target
72 sub-legal lobsters (or recruits) across various habitat types in state waters (0-3 nautical miles
73 from shore). The Coastwide Ventless Trap Survey (VTS) employs a random stratified sampling
74 design and gear modeled after the commercial fishing configuration of traps strung together over
75 trawl lines, but with the inclusion of ventless traps to retain sub-legal and legal sized lobsters.
76 The catch of lobsters per ventless trap (CPVT) is used to provide an estimate of the relative
77 abundance of lobsters through space and time. The VTS was designed with depth and NMFS
78 statistical areas defining the stratification based on pilot surveys and analyses (Pugh and Glenn
79 2020), which found that a stratification scheme based on depth provided the most accurate
80 estimates of relative abundance, as represented by reduced variability in CPVT. The NMFS
81 statistical area inclusion in survey stratification is intended to account for the latitudinal gradient
82 of environmental drivers on lobsters, such as temperature. Mean CPVT estimates in concert with
83 strata-specific weights based on the survey area domain are then used to derive annual
84 abundance indices.

85 Male and female CPVT indices were first used in 2015 as fisheries-independent
86 assessment model inputs for the two U.S. stocks, Southern New England and the Gulf of
87 Maine/Georges Bank (ASMFC, 2015). This inclusion served as a major advancement for better
88 understanding the sub-legal lobster population and population trends in the summer, a season
89 that previously lacked relative abundance information. While the VTS abides by a standard set of
90 methods, unforeseen changes in sampling design can occur due to inclement weather and gear
91 loss. Many factors have been found to influence trap effectiveness and catchability for crabs and
92 lobsters (Miller, 1990; Tremblay and Smith, 2001; Smith and Tremblay, 2003; Geraldi et al.,
93 2009), but have been unaccounted for when constructing relative abundance estimates. Within

94 stock assessment models, survey indices are often scaled to the stock or population size using
95 non-linear, time-invariant relationships (Hilborn and Walters, 1992). While the estimated
96 parameters of these functions (i.e., catchability coefficients) capture the average magnitude of
97 survey catchability, process error in terms of time-varying catchability is often not considered.

98 Model-based approaches to standardizing fisheries-independent catch data have been
99 widely viewed as a useful tool in providing more accurate estimates of abundance indices
100 (Maunder and Punt, 2004). Utilizing these approaches provide several benefits, including an
101 understanding of the effects variables have on species observed abundance, reducing the
102 variance on the abundance predictions, corrected-abundance estimates for when these variables
103 deviate, a tool to help account for years with missed sampling, and a method to account for
104 random effects on observed abundances. This work aimed to understand the impact of several
105 covariates on lobster abundance trends for the inshore components of the Gulf of Maine (GOM)
106 and Southern New England (SNE) lobster stocks, and account for these sources of variability
107 when deriving abundance indices. Model-based approaches were utilized to derive sex- and
108 stock-specific annual abundance estimates from the lobster ventless trap surveys. Model-derived
109 male and female GOM and SNE CPVT indices were constructed and compared to those based
110 on the design-based approach to ascertain their differences and advantages, and ultimately the
111 significance of including factors documented to influence trap catchability in deriving abundance
112 estimates for the VTS.

113 **2. Methods**

114 *2.1 Survey design*

115 Beginning in 2006, the VTS has employed a random stratified survey design, using
116 NMFS Statistical Area (SA) and depth as the primary strata classifications. The SAs included in

117 the survey are 511, 512, 513, and 514 in the GOM region (no sampling is conducted in the
118 Georges Bank region of the stock), and 538, 539, and 611 in SNE. The survey is a cooperative
119 effort between state fisheries agencies and commercial lobstermen, in which lobstermen are
120 contracted to deploy and retrieve survey gear from their vessels with agency biologists aboard to
121 collect data (Pugh and Glenn, 2020). States that have or currently participate in the survey
122 include Maine, New Hampshire, Massachusetts, Rhode Island, and New York. The survey
123 design uses three depth strata that span the range of depths that lobsters are typically fished in
124 inshore waters: 1-20 m, 21-40 m, and 41-60 m.

125 Full description on the VTS can be found in ASMFC (2020). Briefly, all states have
126 sampled since 2006 but for three exceptions: New York sampled only from 2006-2009, New
127 Hampshire began sampling in 2009, and Massachusetts did not sample in 2013. All states except
128 Maine began sampling sites with one six-trap trawl lines, in which vented and ventless lobster
129 traps were alternated for three of each per trawl and spaced 60 feet apart (Table 1). Maine
130 deployed gear either as two three-trap trawls or as one six-trap trawl. Since 2015, Maine and
131 New Hampshire have exclusively fished ventless traps and abandoned sampling with vented
132 traps. Across states, sites are sampled twice per month with a targeted three-night soak time
133 (soak times have exceeded three nights when inclement weather delayed sampling). All traps are
134 baited when actively fishing, with bait type at the discretion of the contracted lobstermen. The
135 primary data stream from the survey is the number of lobsters caught in each trap, which is used
136 for estimating CPVT. However, for each lobster, several descriptors are also recorded: carapace
137 length to the nearest mm, sex, shell hardness, culls and other shell damage, external gross
138 pathology, mortality, the presence of extruded ova on females, and shell disease symptoms, with
139 bycatch similarly described where applicable.

140 *2.2 Data processing*

141 Samples considered for the model-based standardization included those whose sites
142 annual average position fell within the survey strata (Figure 1). All samples that fell out because
143 the sites have since been dropped from the survey domain were not included. This approach
144 resulted in a small portion of samples being dropped from Maine, New Hampshire,
145 Massachusetts, and Rhode Island, but also resulted in all samples from New York being
146 excluded.

147 Several data filters were also used to exclude samples from the modeling effort. Traps
148 that were not fishing effectively (e.g. torn netting, escape vent left open), and thus caught no or
149 few lobsters, were excluded from the analysis and considered as a missing trap given their
150 ineffective fishing. For ventless traps of this nature, their catch was not included in modeling
151 CPVT. Such vented and ventless traps were not incorporated into the total trap number in a
152 trawl. Further, only samples with soak times between one and six days were included. The
153 survey has traditionally targeted summer months (June, July, and August) for sampling;
154 however, due to occasional logistical or funding constraints, sampling timing sometimes varied.
155 Samples collected outside June-August were excluded from the analysis for consistency across
156 the survey domain. Lastly, only ventless trap catch was modeled; vented trap catch was excluded
157 from the analysis.

158 *2.3 Modeling approach*

159 Generalized linear mixed models were used to predict sex-specific lobster CPVT for both
160 stocks. This statistical framework expands upon that of generalized linear models by allowing
161 covariates to be modeled as fixed or random effects (Vidal et al., 2018). Random effects allow
162 for assessing variability among factors of repeated measures, or when randomly selected

163 variables are part of a larger population of which the bounds are not completely sampled (Bolker
164 et al., 2009; Deroba, 2018). Four individual models were built to predict the desired CPVT
165 response variable: male CPVT in SNE, female CPVT in SNE, male CPVT in GOM, and female
166 CPVT in GOM. Catch data reflected lobsters 53mm and larger to match the data needs of the
167 lobster stock assessment model (ASMFC, 2020). Lobsters smaller than 53mm can be caught in
168 the survey; however, the proportion of the catch less than 53mm is often small, suggesting that
169 these smaller sizes are not selected by the traps (Appendix A). Lobster CPVT was modeled at the
170 trap level, and a negative binomial error distribution was used to model CPVT in each of the four
171 models given the overdispersion in the catch data. Further, the models included zero-inflation to
172 account for the high frequency of zero catch. The models were constructed using R package
173 ‘glmmTMB’ (Brooks et al., 2017).

174 While the VTS collects many of the same data fields across states to derive lobster
175 abundance estimates, not all relevant data are collected by each state. Continuous, fine scale
176 sediment or bottom type data do not exist across the survey bounds to be included in the
177 standardization. Bathymetric slope was estimated for samples using NOAA’s National
178 Geophysical Data Center (NGDC) depth data for the Northeast U.S. Shelf using R package
179 ‘raster’ (Burrough and McDonnell, 1998), but corroboration between the associated depths and
180 those observed from the VTS was poor. Thus, bathymetry and slope data were not included in
181 the modeling. While bait type is believed to influence lobster catch rates, not all states have
182 consistently collected such data through time. Further, bait used in the SNE and GOM have been
183 primarily skate and herring, respectively, and may not provide enough variation in bait types for
184 the models to confidently assess catch variance associated with bait specifically. Lastly, position

185 of the trap within a trawl was not included based on the consistency of the information being
186 collected through time by states.

187 Several covariates were tested for use in modelling lobster CPVT. Year was modeled as a
188 fixed effect, while unique site, day of year, and soak time were modeled as random effects.
189 Designating factors as random effects allowed for isolating their population-level effect
190 estimates and removed the effects that a given factor's variation in sampling data may have on
191 the fixed effect of annual CPVT. The natural log of the total number of traps used in each haul
192 was used as an offset term to account for varied effort when traps on a trawl were lost. In
193 instances where traps were lost or did not function properly (e.g. torn netting, escape vent
194 accidentally open), these traps were subtracted from the total number of traps in a trawl. Despite
195 modeling ventless catch only, all trap types (ventless or vented) fishing properly were counted as
196 a trap within the trawl. With CPVT modeled at the trap-level, ventless traps within a trawl had
197 the same covariate values, capturing the degree of CPVT variability associated within given
198 spatiotemporal conditions or soak time. Survey depth stratification was incorporated into the
199 model by weighting samples based in the model on the areal proportion that their strata
200 comprised of the SA-Depth-State stratification. The weight served as multipliers on the model
201 log-likelihood contributions. Variables' significance in predicting CPVT were evaluated for each
202 model using backward-stepwise comparison and removing covariates subsequently. The model
203 variants were constructed by removing them sequentially using backward stepwise selection, and
204 the model variant with lowest Akaike information criterion (AIC; Akaike, 1973) value was
205 selected to derive model-based indices.

206 Temperature can influence lobster life history (Fogarty et al., 2007) and has been
207 commonly used to describe lobster habitat, abundance, and distribution (Chang et al., 2010;

208 Tanaka and Chen, 2015; Tanaka and Chen, 2016). However, temperature and others habitat
209 variables were not incorporated into the models. The primary reason for their exclusion was that
210 for the current U.S. lobster stock assessment, environmental influence on fisheries-independent
211 indices are incorporated directly within the assessment model when estimating catchability
212 (ASMFC, 2020.) When relating the survey index to the total population within the assessment
213 model, catchability parameterization allows for directly estimating the influence of temporal
214 environmental changes within a survey area on the index's relation to the estimated population
215 size. This approach can account for changes in the availability of lobster to survey domains over
216 time driven by changes in the environment (ASMFC, 2020.) Thus, incorporating yearly
217 environmental effects in the model-based approach would distort the final impact of the
218 environment on the CPVT. Incorporating the environmental effects for seasonal or spatial
219 dynamics were considered, but consistently measured data at the needed resolution was not
220 available across the survey domain. Further, many of these environmental drivers that vary
221 spatiotemporally are inherently incorporated via the day of year, year, and site covariates.
222 Including other environmental variables in addition to these covariates could risk model
223 overfitting or double counting the effects of a given ecological factor. As such, this index work
224 focused on survey, temporal, and gear configuration concerns as opposed to interannual changes
225 in the environment.

226 *2.4 Design-based approach and comparisons*

227 The design-based abundance indices were constructed as used in the 2020 American
228 Lobster Stock Assessment (ASMFC, 2020), with brief description provided herein. The same
229 data sets generated from the data processing steps described previously were used for design-
230 based calculations. Because Massachusetts did not run a survey in the GOM or SNE in 2013 and

231 the Massachusetts data can significantly influence the combined indices, design-based indices
232 were not constructed for 2013 (ASMFC, 2015).

233 Survey sites were intended to not move within a year, so each survey site was treated as
234 an effective replicate and samples within the year as repeated measures. Sites would only change
235 slightly within a year if repeated gear loss threatened data collection. To get the yearly average
236 CPVT for a survey site, catch in the ventless traps were first averaged across ventless traps
237 within a trawl, then across trawls within a month, then across months in a year. Calculating the
238 indices from the survey site averages then used a standard stratified-random sampling
239 calculation, with the SA-Depth-State strata used in the survey design:

$$240 \quad Catch_y = \frac{\sum_{str} \frac{\sum_s Catch_{s,str,y}}{N_{str,y}} \times A_{str}}{\sum_{str} A_{str}}$$

241 where $Catch_{y,s,str}$ is the mean catch at site s in stratum str and year y , $N_{y,str}$ is the number of sites
242 in stratum str and year y , and A_{str} the area of a given stratum. As such, both the model and
243 design-based index units were CPVT, with the former using all traps in a predictive model, and
244 the latter arithmetically solving for weighted average CPVT. Both model-based and design-based
245 approaches were compared in terms of their magnitudes and relative trends. Trend differences
246 were examined after normalizing the indices by dividing each index by its time-series mean. The
247 resulting sex ratio for each stock and index type was also compared by dividing the male index
248 by the female index.

249 **3. Results**

250 *3.1 Model results*

251 When testing model variants, the models with the lowest AIC scores across sexes and
252 stocks were those excluding the number of traps offset (Table 2), of which all converged

253 successfully. Excluding the site effect resulted in the greatest difference in AIC scores and
254 weakened model fitness, suggesting that including site effect made the greatest improvement to
255 the model. Including the unique site effect improved model performance more for SNE models
256 than for GOM models, and for SNE females more than SNE males. The final models used to
257 derive abundance indices for all stocks and sexes included site, day of year, year, and soak time.

258 Random effect estimates indicated that in SNE, male and female CPVT were greater in
259 July than June and August (Figure 2). In the GOM, male lobster CPVT was greater in August
260 than June and July, whereas female lobster CPVT was similar in June and August, and slightly
261 lower in July (Figure 2). Random effect intercepts for SNE lobster CPVT did not suggest
262 discernible patterns with changes in soak time. Male CPVT indicated a slight positive effect with
263 increased soak time but was variable, and even more so for females (Figure 3). While SNE male
264 and female models shared similar patterns, the magnitudes were larger for females than males.
265 Soak time effects in the GOM CPVT suggested that for both males and females, catch increased
266 with increasing soak times. This pattern was more pronounced for males than females (Figure 3).
267 Random effects of site indicated spatial variability in its influence on CPVT for male and female
268 lobsters, but for SNE more than GOM. In SNE, both sexes' random effects for site were lower
269 inside the shallower estuaries (e.g. Buzzards Bay, Narragansett Bay) than in the deeper oceanic
270 environments (e.g. Block Island Sound, Rhode Island Sound) (Figure 4). In contrast to SNE, the
271 site effects were of smaller magnitude in GOM, and was less heterogeneous in the GOM survey
272 domain (Figure 4).

273 3. 2 *Abundance indices*

274 SNE model-based and design-based indices showed similar trends. Both SNE female and
275 male abundance indices have declined since 2006, with the model-based declines smoother and

276 less variable over time than those from the design-based approach (Figure 5). Trends from the
277 model-based approach were more similar between sexes than those of the design-based
278 approach, albeit slightly different in magnitude. SNE indices using the design-based approach
279 were higher than those of the model-based approach (Figure 5). The greatest deviations between
280 the two approaches for SNE were in recent years; modeled-based indices declined from 2017 to
281 2018, whereas the design-based indices increased (Figure 5). Further, the 2019 terminal year
282 dropped from 2018 in the design-based approach, whereas it slightly increased in the model-
283 based approach.

284 Corroboration in magnitude between model-based and design-based abundance indices
285 for the GOM was stronger than in SNE. Design-based indices were modestly lower than the
286 model-based indices throughout the time series. Both male and female GOM abundance indices
287 increased from 2007 through 2012, with abundance variable but still high through 2019 (Figure
288 5). Standard errors were greater for GOM model-based indices than design-based, likely
289 attributed to the processing of data at the individual trap level (i.e. model-based approach)
290 opposed to the average site catch level (i.e. design-based approach).

291 Trends between model-based and design-based approaches were similar across sexes and
292 stocks, except in the early and terminal years of the SNE indices (Figure 6). Corroboration in the
293 sex ratios between the approaches varied by stock. For GOM sex ratios, both model-based and
294 design-based approaches were similar, highlighting a trend towards female skewed catch (Figure
295 7). In SNE, the design-based approach also indicated a female skewed population in state waters;
296 however, the model-based approach indicated a male-skewed community through time, with
297 males up to 1.5 times more abundant than females (Figure 7).

298 **4. Discussion**

299 *4.1 CPVT sources of variability*

300 We have quantified several factors that influence catch rates of lobsters within the
301 Coastwide Ventless Trap Survey. Such confounding factors have long been suspected to
302 influence trap catch rates for crustacean species (Miller, 1990), and this work begins to assess
303 and account for these factors when deriving abundance indices by removing the variability in
304 catch associated with them. The year of sampling, day of year sampling occurred, unique site
305 sampled, and the soak time of the traps all influenced lobster CPVT. The site effect appeared to
306 have the greatest effect (Table 2), and more so in SNE than in the GOM (Figure 4). In SNE, we
307 hypothesize that the differences in site across the survey domain reflects differences in benthic or
308 thermal habitat. Narragansett Bay bottom temperatures have warmed significantly over the last
309 several decades (Fulweiler et al., 2015) and are consistently exceeding 20°C (ASMFC, 2020), a
310 temperature threshold believed to cause physiological stress and unsuitable conditions for
311 lobsters (Steenbergen et al., 1978; Dove et al., 2005). The SNE gradient of negative to positive
312 site effects moving away from the head of estuaries to oceanic environments may reflect lobster
313 preference for cooler oceanic waters than warmer estuarine waters. Wahle et al. (2015) noted that
314 bottom temperatures during this summer period were cooler at southern portions of Narragansett
315 Bay near Rhode Island Sound than the upper region of Narragansett Bay. Thermal habitat
316 suitability modeling for this species over the entirety of this survey domain, seasonality, and time
317 series would better test this hypothesis. While spatial differences in lobster settlement are
318 apparent within the GOM purportedly from differences in oceanography (Goode et al., 2019),
319 such differences are not strongly captured by the site effect for lobsters 53mm or greater.

320 Days in July in SNE corresponded to higher lobster CPVT rates, particularly for males.
321 Lobster CPVT in the GOM was greatest toward the end of August, with this pattern more

322 accentuated for males than females. The stock differences in temporal peak CPVT rates is likely
323 reflective of seasonal water temperature differences with latitude. Much of American lobster life
324 history and biological rates are influenced by temperature (ASMFC, 2020), and southern regions
325 like SNE generally warm seasonally earlier than northern regions such as the GOM. Soak time
326 effects indicated that there was either no discernible relationship between soak time and lobster
327 CPVT (e.g. SNE) or potential increases in CPVT at longer soaks (e.g. GOM) (Figure 3). While
328 our results did not provide a definitive relationship between lobster CPVT and soak time, they
329 were somewhat contrary to previous work. In the GOM, previous work has suggested lobster
330 CPVT increases logarithmically over a three-day soak, increasing over the first 24 hours and
331 then plateauing (Clark et al., 2015; Watson et al., 2019), thus perhaps the soak time effect
332 operates at a smaller time scale than targeted 3-day soak period in this monitoring program. The
333 effect of soak time on CPVT is also likely dependent on other factors, such as bait (Watson et al.,
334 2019), stock size, and spatio-temporal factors. For example, these previously reported saturation
335 rates have been conducted in shallow waters compared to the range of depths that the VTS
336 samples. Any depth-specific relations between soak time and CPVT that may exist would be
337 aggregated in this analysis. Evaluating interaction effects between these and other covariates in
338 future studies or model developments would be worthwhile in elucidating these complexities.

339 Other variables may attribute to lobster CPVT and trap catchability that were not
340 accounted for in this analysis. The widely reported preference of cobble or boulder habitat for
341 lobsters (Wahle and Steneck, 1991) over low-relief (e.g. mud, sand) substrate is not explicitly
342 captured in the modeling. Previous studies have found mixed results regarding this; Tremblay
343 and Smith (2001) found that low-relief sites have lower lobster densities than boulder sites,
344 whereas Geraldi et al. (2009) found traps had higher counts on unstructured habitats than rocky

345 areas. The site effect may in part represent benthic habitat effects, but also likely includes other
346 components (e.g. bottom temperature, depth). Differences in end and middle traps of ventless
347 trap surveys have been found as a function of their distance apart; 1994 ventless trap sampling
348 off Cape Breton indicated middle traps had reduced lobster catch than end traps, but when
349 increasing the distance between traps in the trawl for the 1995 survey, no statistical differences
350 were found (Smith and Tremblay, 2003). Instances of lower catch in middle traps likely reflect
351 competition between the traps, whereas lack of differences may be indicative of non-overlapping
352 effective trap areas or trap saturation preventing detection of trap position differences. Bottom
353 temperature is likely manifested within the annual factor, which has been further examined via
354 the lobster stock assessment catchability parameterization (ASMFC, 2020).

355 Interspecific species interactions and behaviors can also influence lobster catch by
356 deterring them from entering the traps or being unable to because the traps have become quickly
357 saturated. Stocking experiments have indicated that the presence of American lobster in traps can
358 reduce the catch of *Cancer* sp. crabs, but the crabs' presence in traps does not significantly
359 impact lobster catch (Richards et al., 1983). However, a negative relationship between catch
360 rates of European lobster (*Homarus gamarus*) and *Cancer pagurus* for individual traps (Addison,
361 1995) suggests perhaps an antagonistic relationship between crabs and lobsters or a difference in
362 their local availability. Skerrett et al. (2020) further identified that lobster CPUE effects from
363 lobster-crab trap interactions can vary by the crab species. In the Coastwide Ventless Trap
364 Survey, bycatch often includes *Cancer* spp., demersal fish, and other invertebrate species;
365 however, it is unclear the extent that high bycatch samples reflect areas of greater bycatch
366 abundance compared to lobster, or whether lobsters in the area are at higher number than
367 observed but are not reflected in the catch due to these species interactions. The true influence of

368 altered catch rates of a target species due to trap saturation from bycatch or species interactions
369 are variable across taxa and trap type (Robichaud et al. 2011, Kersey and Clark, 2011; Bacheler
370 et al., 2013). Future analyses should include disentangling these species interactions from
371 fisheries independent trap survey results by incorporating benthic habitat dependencies, as
372 evaluating the different species habitat needs in the context of observed habitat type would
373 provide better insight into these interactions.

374 Density-dependent factors and intraspecific interactions have also been speculated to
375 influence catch. Lobster catchability has been found to not necessarily be constant with density,
376 as pre-stocking traps with lobsters has been found to reduce the catch of lobsters (Richards et al.,
377 1983). Similar findings were reported by Watson et al. (2019), where stocking lobsters before
378 deployment reduced catch and removing lobsters after 24 hours led to an increase in catch. Such
379 a phenomenon would have the potential to misrepresent an aggregated or highly abundant
380 species with lower and more uniform catch (Addison, 1995). Further, the intraspecies
381 interactions with American lobsters can vary with size; similar model standardization work for
382 sublegal and recruit lobsters per trap in Canadian surveys use the number of legal lobsters
383 observed to capture the behavioral dynamics between lobsters of different sizes (Cook et al.,
384 2018). Work with rock lobsters (*Jasus edwardsii*) has shown that intraspecific interactions that
385 vary by sex and size effect trap catch and potentially inferences on abundance resulting from trap
386 CPUE (Frusher and Hoenig, 2001; Ihde et al., 2006). These complexities resulting from
387 intraspecies behavioral interactions and their influence on CPUE warrant further investigation
388 for ensuring non-biased indices used for stock assessment modeling.

389 *4.2 Comparison of index methods*

390 Both the model-based and design-based approaches capture the trends in American
391 lobster relative abundance for the two stocks over the sampled time period: increasing abundance
392 in the GOM and decreasing or stable, yet low, abundance in SNE (Figure 5). Except for
393 interannual differences, these trends are corroborated by several other survey trends over the
394 stock bounds and modeled population sizes (ASMFC, 2020). The differences in model-based and
395 design-based indices were greater for SNE than for GOM (Figures 5 and 6). The greater
396 similarity between the two approaches in the GOM may be attributed to the number of samples
397 for each model. GOM models had over 5.7 times more samples for model fitting than SNE,
398 perhaps providing more information to improve mean CPVT predictions. However, the greater
399 variability in catch in SNE also likely attributes to the reduced corroboration between the design
400 and model-based index approaches. The increased variability is highlighted in the standard error
401 estimates for both index types (Figure 5) and can partially be attributed to the factors included in
402 the model types. For example, the inclusion of the site in the SNE models appears to cause
403 substantial deviation in scale and trend differences between model-based and design-based
404 abundance indices (Appendix A).

405 The model-based approach of formulating abundance indices at the individual trap level
406 as opposed to aggregate means over a season in the design-based approach provided both
407 benefits and drawbacks. By constructing abundance indices at the trap-level, the inherent
408 variability within a trawl (as incorporated with multiple traps with varying CPVT having the
409 same covariate values) were better accounted for. Further, not averaging ventless trap catch over
410 a site in the model-based approach allowed for including more information on observed CPVT
411 for informing abundance index variance; averaging CPVT data within a trawl in the design-
412 based approach may not account for this level of variability. As such, the variance in the model-

413 based estimates also increased as reflected in the standard error estimates (Figure 5). Despite the
414 increase in variance, index bias is generally reduced by accounting for factors affecting
415 catchability (Maunder and Punt, 2004).

416 One of the greatest benefits of the model-based approach is that it allows for estimating
417 abundances for 2013, where missing data from Massachusetts prevented the design-based
418 approach from calculating such data. GOM index estimates for 2013 indicated that year was one
419 of the top three years in abundance over the time series, where conversely, such estimates for
420 SNE indicated that year was either the lowest (females) or second lowest (male) abundance over
421 the time series (Figure 5). The model-based approach using random effects provides a solution
422 for estimating abundances for the stock unit in the event select state surveys are unable to be
423 conducted. The degree of missing data for a given year and stock should be evaluated prior to
424 relying on the model-based approach to estimate abundance.

425 Inferences on lobster population sex ratios are challenging with trap surveys as males and
426 females are not spatially uniform due to differences in biological needs (Jury et al., 2019).
427 Additionally, differences in catchability between sexes can create biases with trap surveys;
428 whereas in Canada, Tremblay and Smith (2001) reported females were less catchable than males.
429 Within SNE's Buzzards Bay, shallower areas tend to have warmer waters and lead to a male-
430 skewed sex ratio, whereas deeper-cooler areas are reflective of a female-skewed ratio (ASMFC,
431 2010, Jury et al., 2019). The model-based approach for SNE captures this skewness, whereas the
432 design-based approach is reflective of a female-dominated catch (Figure 7). The difference does
433 not appear to be influenced by the sample weighting in the model-based approach (Appendix A);
434 however, the large number of shallow strata samples in the models may account for this
435 discrepancy in sex ratios between the model- and design-based approaches. The SNE VTS strata

436 are predominantly shallow (0-20m), where the deeper (21-40m) strata are primarily located in
437 the Rhode Island survey domain (Figure 1).

438 The model-based indices provided an alternative dataset for inclusion in stock assessment
439 model fitting (ASMFC, 2020). Understanding the influence of various factors on survey design
440 and catchability are paramount when trying to assess changes in abundance. Gear studies can
441 aim to test these factors through specific research and apply *post hoc* corrections (McManus et
442 al., 2020), but such efforts are often not possible due to funding and time constraints. This
443 modelling approach allows for testing these factor effects explicitly when deriving the abundance
444 estimates. Model-based approaches have become a favorable practice for several purposes,
445 including to account for survey variables introducing bias in observed abundance estimates
446 (Maunder and Punt, 2004; Venables and Dichmont, 2004), to incorporate environmental drivers
447 in abundance and distribution (Friedland et al., 2020), and to predict abundances through space
448 and time not sampled for a more complete understanding of population structure (Thorson et al.,
449 2015). This work serves as an example of standardizing abundance indices for structure-oriented
450 species that require non-traditional sampling techniques. In the context of lobster, the
451 standardized approach also provides new insights on the catchability and life history for an
452 iconic and valuable species. With the additional year of data now available (i.e. 2013), and
453 smoother trends from removing variability associated with catchability and not population
454 changes, the intent is for these indices to improve the ability of U.S. American lobster stock
455 assessment models to accurately predict population trajectories.

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467 **7. Author Contributions**

468 MCM designed the research study, and MCM and JK and executed the research. MCM
469 led the writing of the manuscript. All authors provided guidance throughout the study and
470 revised drafts of the manuscript.

471 **8. Compliance with Ethical Standards**

472 The authors declare they have no conflict of interest in these studies.

473 **9. Literature Cited**

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475 abundance as inferred from pot-caught samples. ICES Marine Science Symposium 199:294-300.

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680 Table 1. Description of Ventless Trap Survey by participating states.

State (Statistical Areas)	Operating Years	Trawl Configuration
Maine (511, 512, 513)	2006-2014	Single 6-trap trawls alternating vented and ventless traps or two 3-trap trawls with 3 vented traps on one trawl and 3 ventless traps on the other trawl
	2015-2019	One 3-trap trawl, all ventless traps
New Hampshire (513)	2009-2014	Alternating vented and ventless traps in single 6-trap trawls
	2015-2019	One 3-trap trawl, all ventless traps
Massachusetts (513, 538)	2006-2012, 2014-2019	Alternating vented and ventless traps in single 6-trap trawls
	2006-2019	Alternating vented and ventless traps in single 6-trap trawls
Rhode Island (539)	2006-2019	Alternating vented and ventless traps in single 6-trap trawls
New York (611)	2006-2009	Alternating vented and ventless traps in single 6-trap trawls

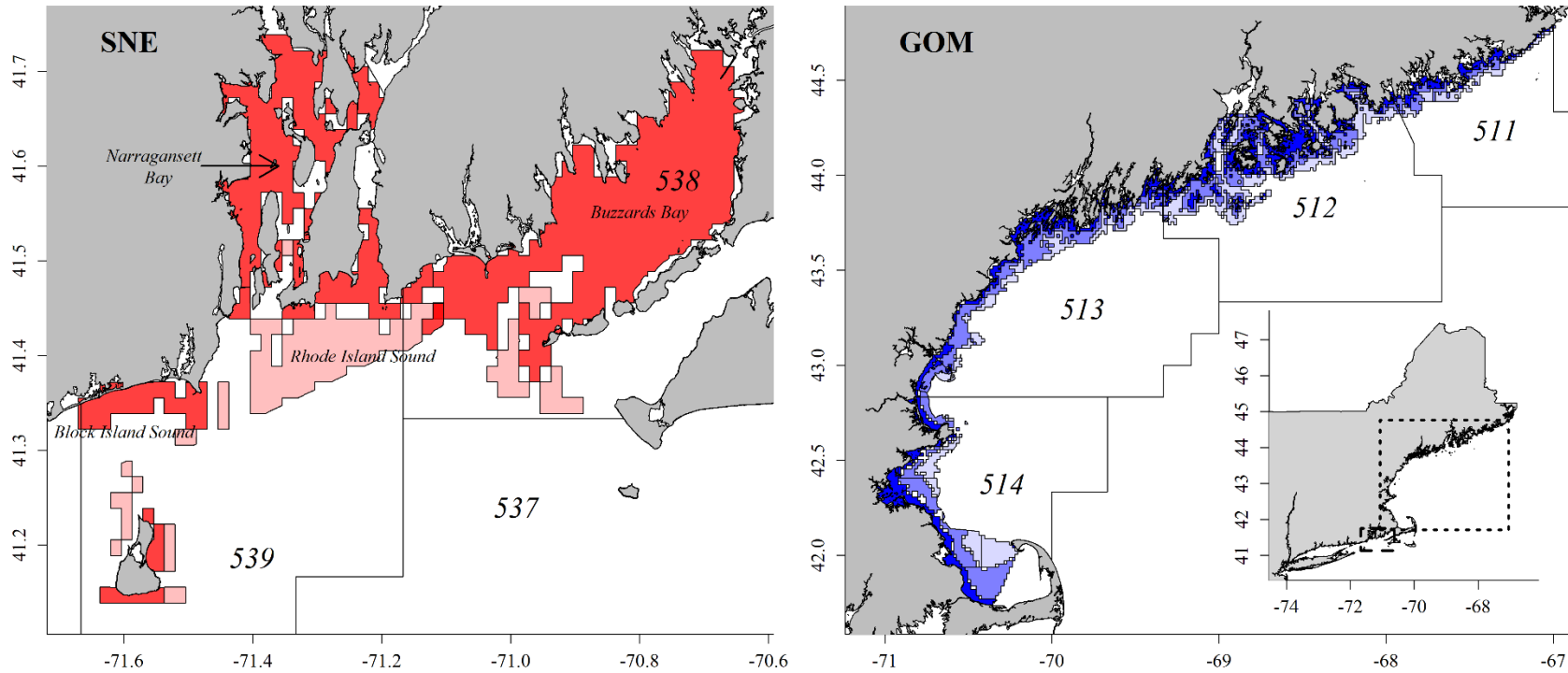
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693 Table 2. Stepwise comparison of model fits using varying model covariates. Akaike information
 694 criterion (AIC) are provided for each model variant. Smaller values within a model type indicate
 695 better fit. Bold values indicate the model variant with the lowest AIC score.

Model	Variables	AIC
GOM Females	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})+offset(\ln[\text{Traps No.}])$	29263
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})$	28686
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})$	28697
	$fe(\text{Year})+re(\text{Day of Year})$	28695
GOM Males	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})+offset(\ln[\text{Traps No.}])$	26976
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})$	26463
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})$	26517
	$fe(\text{Year})+re(\text{Day of Year})$	26519
SNE Females	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})+offset(\ln[\text{Traps No.}])$	8276
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})$	8245
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})$	9619
	$fe(\text{Year})+re(\text{Day of Year})$	9637
SNE Males	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})+offset(\ln[\text{Traps No.}])$	8720
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})+re(\text{Site})$	8692
	$fe(\text{Year})+re(\text{Day of Year})+re(\text{Soak Time})$	9696
	$fe(\text{Year})+re(\text{Day of Year})$	9695

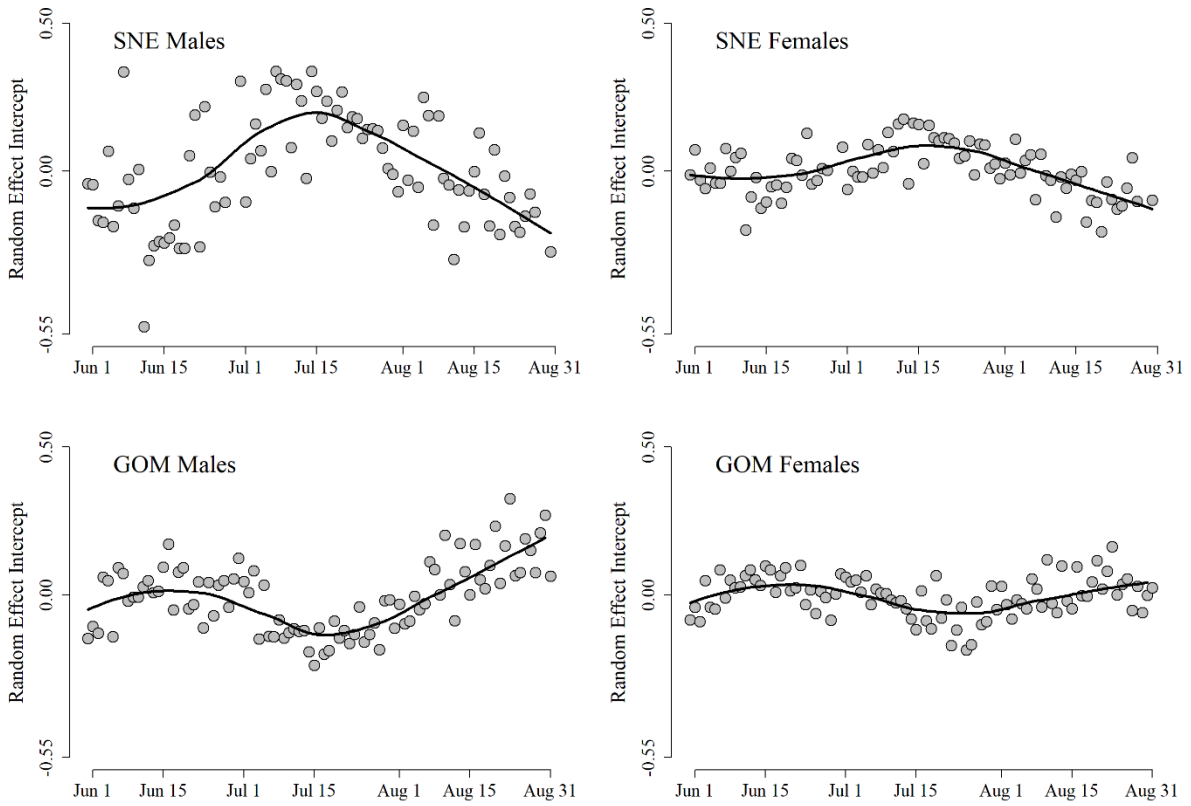
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699 Figure 1. Survey domain and stratification for the lobster VTS within the Southern New England (SNE, left) and
 700 Gulf of Maine (GOM, right) stock areas. Southern New England strata are 0-20m (dark red) and 21-40m (light red), whereas Gulf of Maine has 0-
 701 20m (dark blue), 21-40 (blue) and 41-60m (light blue) strata. NOAA Statistical Areas are presented within each region. The insert
 702 map of the Northeast United States in the right panel presents the geographical locations of SNE (dashed) and GOM (dotted) survey domains.
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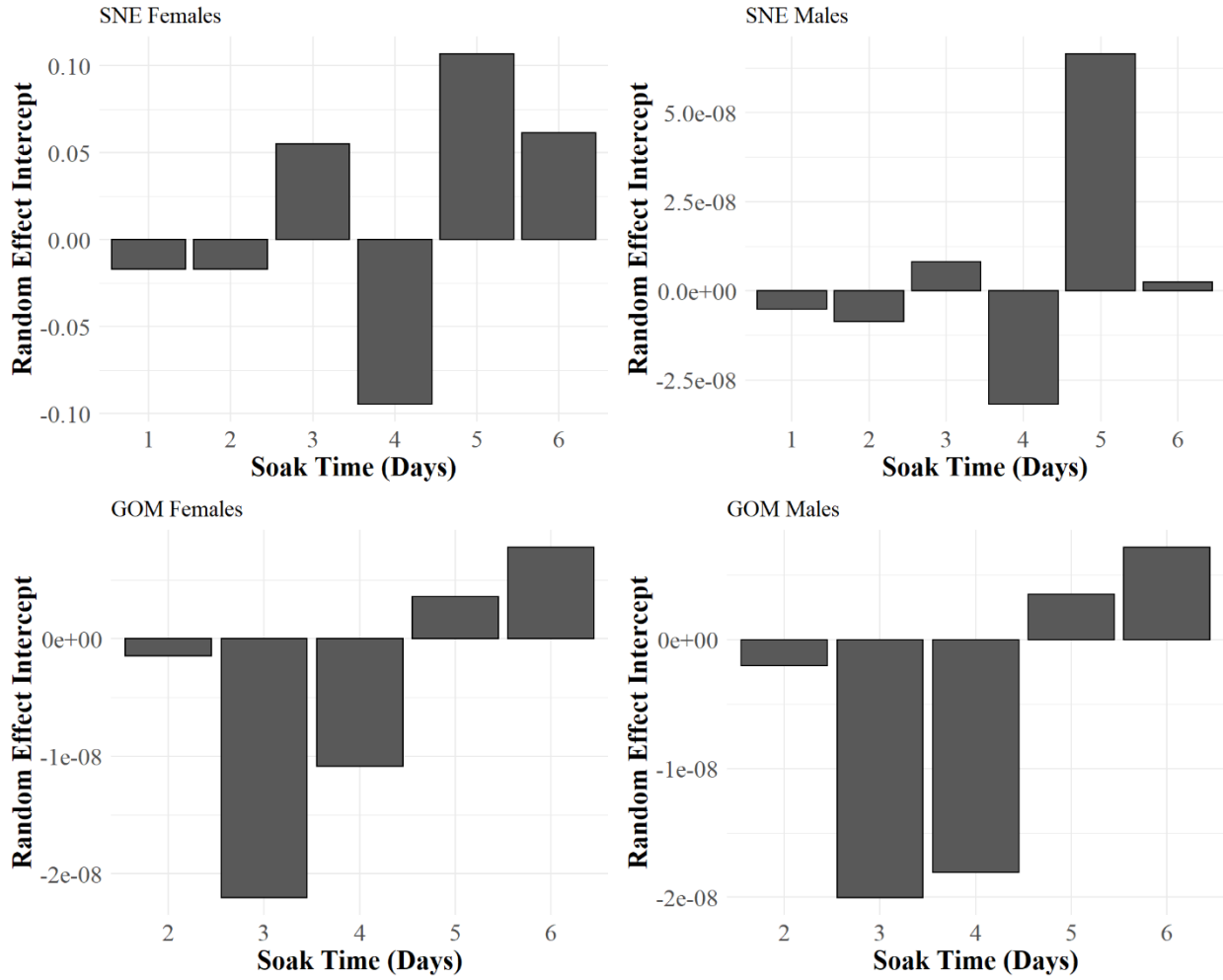


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707 Figure 2. Random effects intercepts for the day of year variable estimated for the male and
708 female SNE and GOM catch per ventless trap (CPVT) models. Lines represent loess fits through
709 the day of year intercept values.

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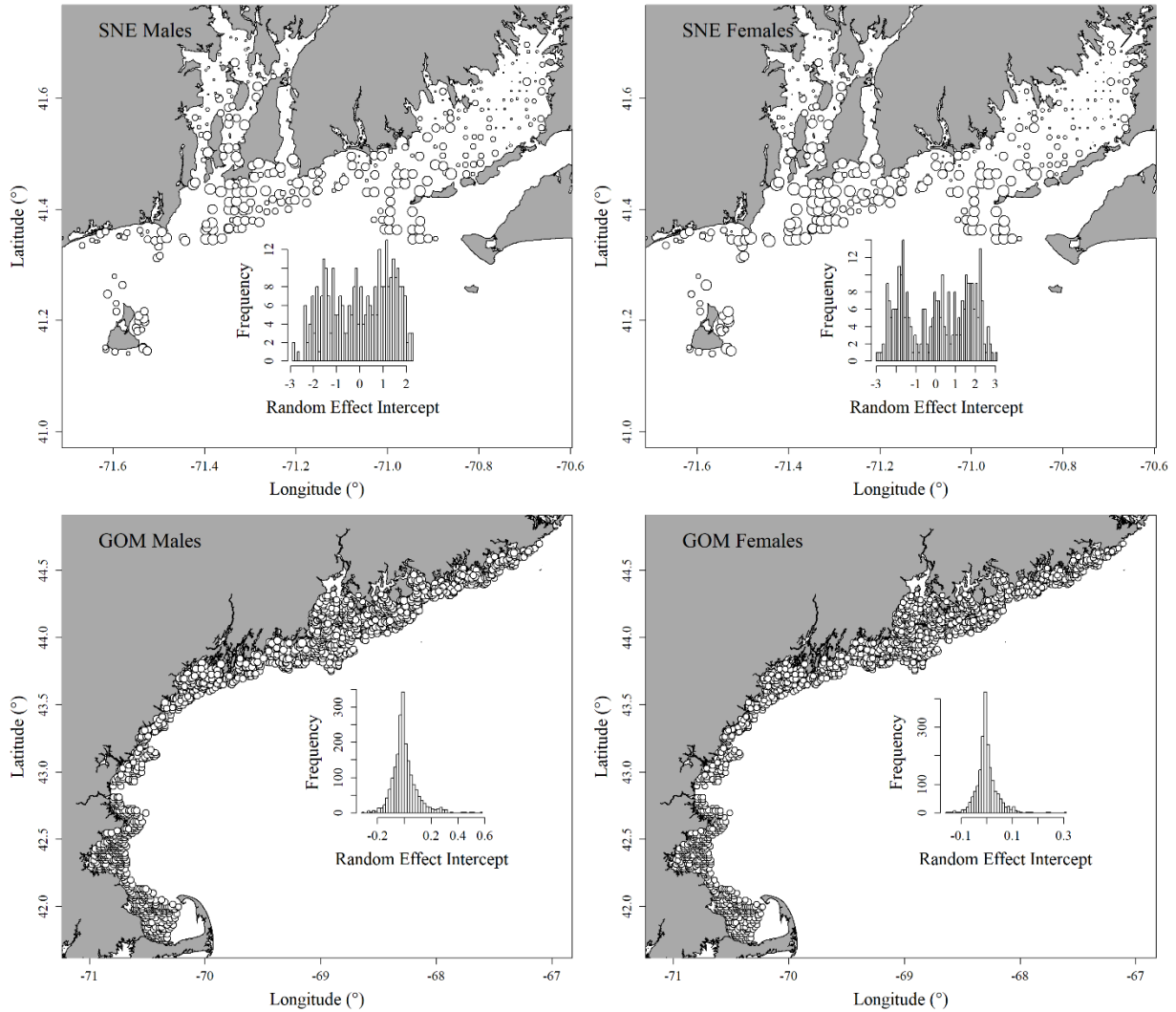


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712 Figure 3. Random effects intercepts for the soak time variable estimated for the male and female
 713 SNE and GOM catch per ventless trap (CPVT) models.

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717 Figure 4. Random effect intercepts for the site variable estimated for the male (left column) and
 718 female (right column) SNE (top row) and GOM (bottom row) catch per ventless trap (CPVT)
 719 models. Effects are plotted spatially representing their location, with the size of points relative to
 720 their value (more positive random effect intercepts correspond to larger points). Points within a
 721 sex-specific figure are relative to that sex only. Inset histograms present the random effect
 722 intercepts associated with the site variable for the respective catch per ventless trap (CPVT)
 723 model.

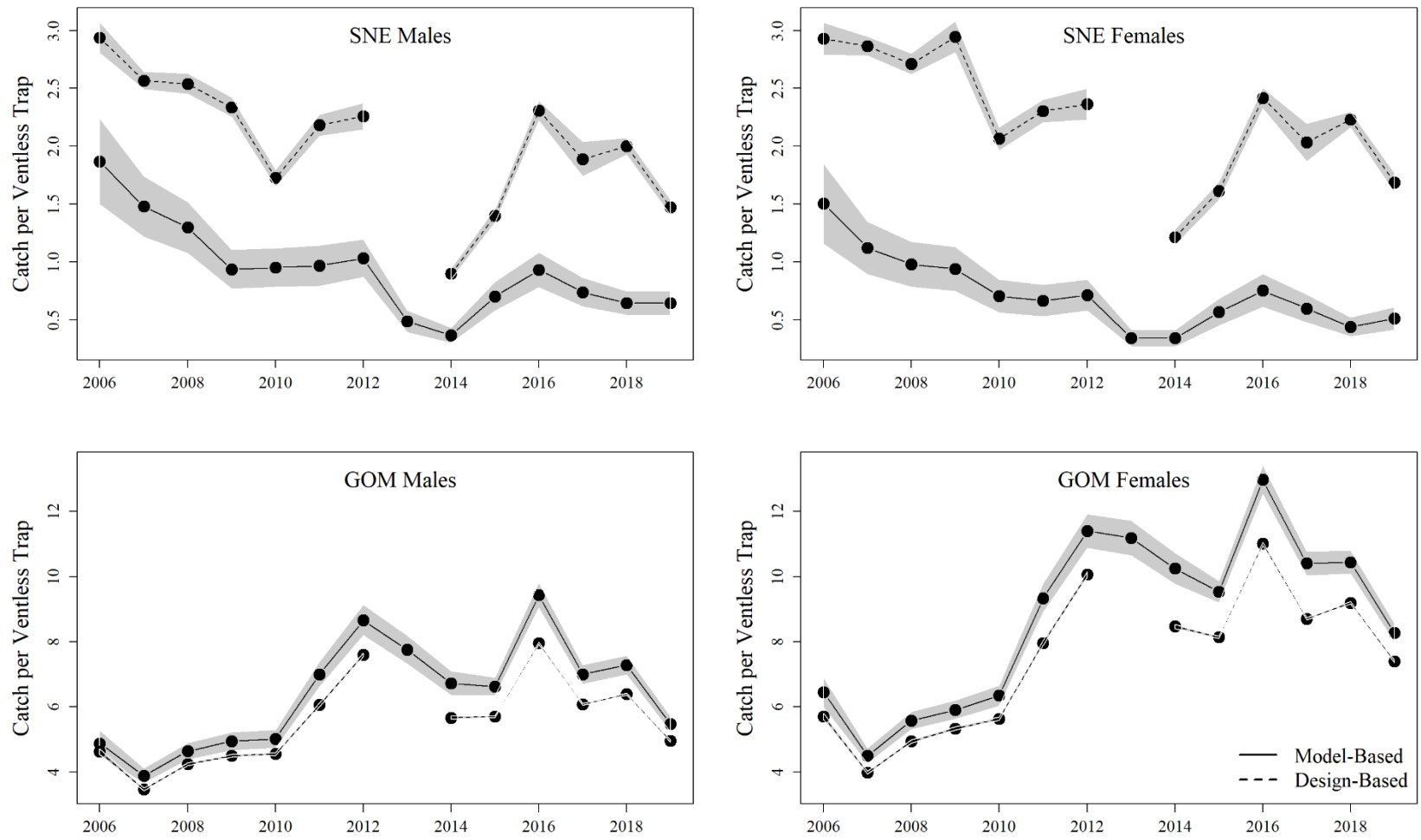


Figure 5. Annual lobster catch per ventless trap (CPVT) indices by sex and stock. Model-based approach mean indices (solid lines) for males and females and their associated standard error are presented with mean design-based indices (dashed line).

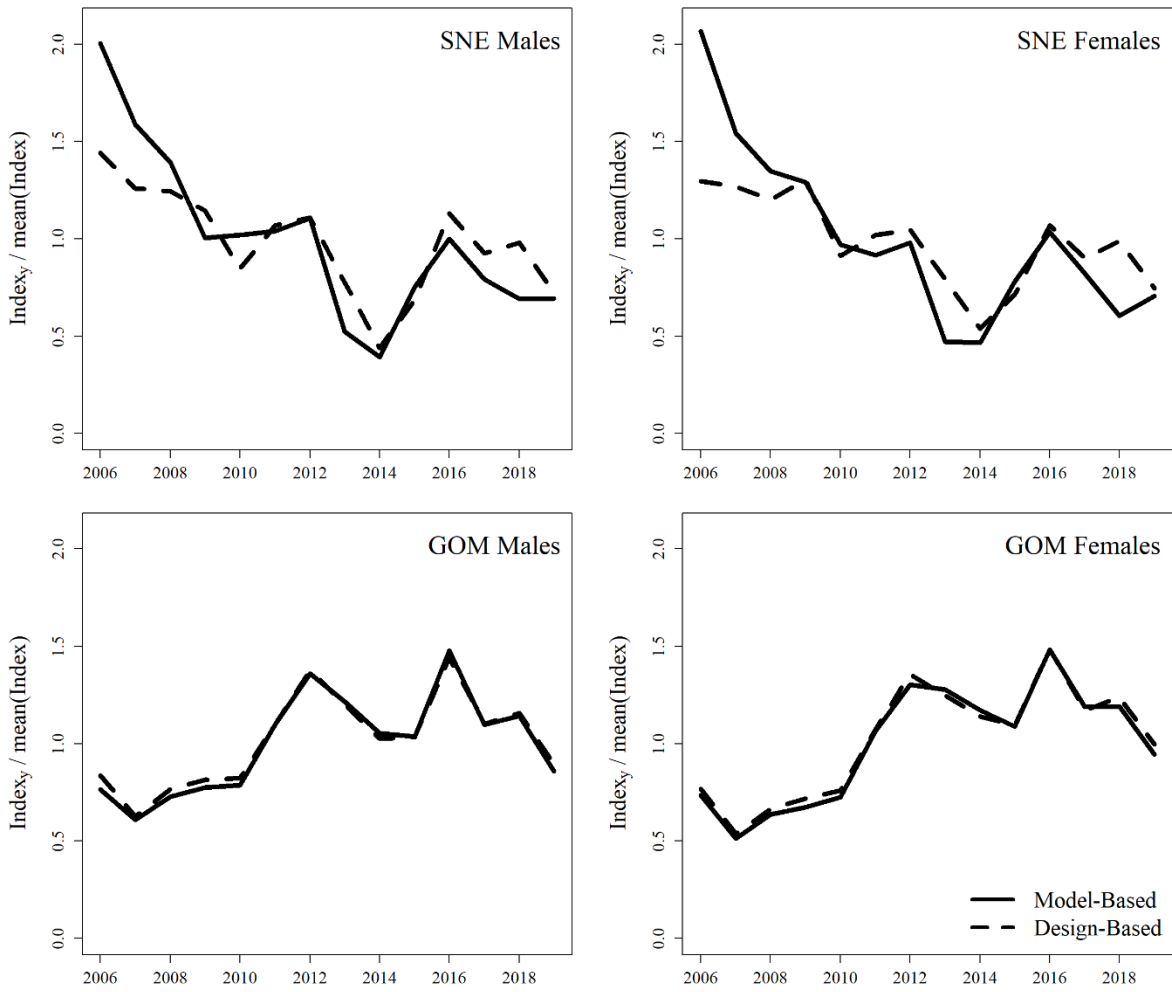


Figure 6. Normalized male and female SNE and GOM CPVT indices using the model-based and design-based approach.

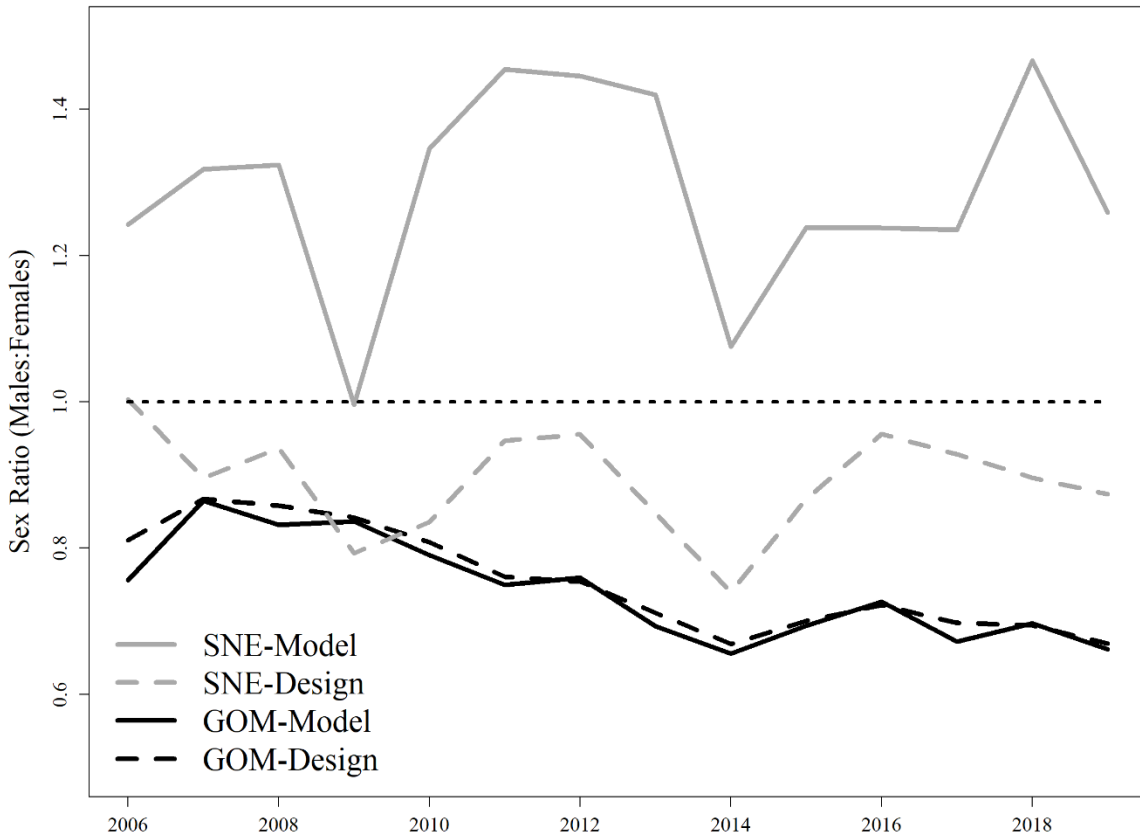


Figure 7. Annual sex ratios for indices using model-based and design-based approaches. The horizontal dotted line reflects where the sex ratio is 1:1.