

1 A question of scale: Weak evidence for broad regional synchrony in fish year-class strength
2 within or among species in inland lakes

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14

15 **Abstract**

16 Spatially-correlated abiotic and biotic conditions can potentially induce synchrony in the
17 dynamics of disparate populations or species. However, such potential synchrony among species
18 or populations may be tempered by dynamics operating at finer temporal and spatial scales, as
19 well as species-specific responses to environmental conditions. We examined within- and
20 among-species synchrony in year-class strength across 130 lakes in northern Indiana over 30
21 years to evaluate the relative scale of potential synchrony and its possible ecological mechanisms
22 in five recreationally important fish species: black crappie (*Pomoxis nigromaculatus*), bluegill
23 (*Lepomis macrochirus*), largemouth bass (*Micropterus salmoides*), redear sunfish (*L.*
24 *microlophus*), and yellow perch (*Perca flavescens*). Bluegill and black crappie recruitment was
25 significantly positively correlated when the species coexisted, and relative year-class strength of
26 both species was positively related to mean annual wind speed. However, there were few other
27 instances of recruitment synchrony between or within species, regardless of whether synchrony
28 was assessed within or among lake systems. In addition, habitat similarity and regional weather
29 patterns also played a limited and inconsistent role in shaping recruitment strength or synchrony
30 in these small inland systems. These results suggest that fish recruitment dynamics in small,
31 inland systems are most often a function of system-specific biotic interactions that mask limited
32 input from broader climatological influences, and that understanding recruitment in small lakes
33 will require examinations on appropriately fine spatial and temporal scales.

34 **Key words:** *climate; abiotic variables; e-folding scale; year-class strength; catch curve;*
35 *habitat*

36 **1. Introduction**

37 Broad spatial correlations in environmental conditions can potentially induce synchrony in
38 the characteristics of animal populations across wide geographic areas, a phenomenon
39 collectively referred to as the Moran effect (Cattanéo et al., 2003; Moran, 1953; Ripa, 2003).
40 However, the strength of these broad-scale environmental effects on populations, and therefore
41 the level of synchrony among populations, is tempered by dynamics operating at finer temporal
42 and spatial resolutions (Engen and Sæther, 2005; Ranta et al., 1997). Factors such as population
43 density (Grenouillet et al., 2008), biotic interactions among coexisting species (Durant et al.,
44 2007; Kelly et al., 2009), or local variations in ecosystem productivity (Wheeler et al., 2016) can
45 have major effects on resulting population characteristics and dynamics, influencing potential
46 Moran effects and limiting potential synchrony in population dynamics (Weber et al., 2017). In
47 fact, strong local variation in environmental conditions could induce more similar population
48 dynamics among coexisting species than in geographically disparate populations of the same
49 species (Howeth and Leibold, 2013; Robinson et al., 2013). Understanding the relative
50 importance of broad- and fine-scale environmental and climatic factors in synchronizing
51 population characteristics could therefore inform relevant scales of management (Feiner et al.,
52 2016; Hansen et al., 2015; Tonkin et al., 2017) and offer an avenue to predict the impacts of
53 broad-scale, long-term disturbances such as climate change on populations and communities
54 (Collingsworth et al., 2017; Hansen et al., 2017).

55 Fish recruitment, i.e., the survival of offspring to contribute to the adult population, has
56 represented an enigmatic and consistent problem for biologists seeking to understand how abiotic
57 and biotic conditions drive fluctuations in fish populations (Myers, 1998). Large fluctuations in
58 recruitment from year to year, often seemingly disconnected to variation in adult abundance

59 (Feiner et al., 2015; Gilbert, 1997), give the appearance of stochastic processes strongly
60 influenced by external forces (Koslow, 1992). Studies in large, interconnected marine or
61 freshwater systems have documented considerable intraspecific synchrony in year-class strength
62 (i.e., the abundance of recruits produced in a given year relative to average production levels)
63 across large geographic distances. These observations suggest that, at least in these large
64 systems, dispersal combined with broad climatic factors have the potential to structure the
65 survival and production of large year classes and drive similar dynamics among populations
66 (Cloern et al., 2010; Kelly et al., 2009; Myers et al., 1997; Ward et al., 2016).

67 Observations of broad-scale recruitment synchrony in large lakes coupled with research
68 showing stronger abiotic influences on recruitment in marine systems compared to stronger
69 biotic influences in freshwater systems suggests a paradigm where the importance of climatic
70 factors is correlated with system size. Smaller systems (i.e., inland lakes) should exhibit limited
71 recruitment synchrony compared to their larger, interconnected counterparts due to a lack of
72 dispersal of individuals among systems and stronger influences of local biotic interactions
73 (Houde, 1994; Janssen et al., 2014; Myers et al., 1997). However, this paradigm is challenged
74 by evidence for consistent responses of freshwater fish populations to broad-scale climate
75 patterns or climatological events (Bunnell et al., 2016; Hansen et al., 2015; Schupp, 2011; Weber
76 et al., 2017). In addition, there is some evidence that fishes inhabiting small freshwater systems
77 can show considerable synchrony in recruitment patterns over much greater geographic distances
78 than expected, even in the absence of high dispersal potential (Dembkowski et al., 2016;
79 Grenouillet et al., 2008; Phelps et al., 2008; Weber et al., 2017). Therefore, under some
80 conditions, fish recruitment (and management) could be better informed by understanding the

81 relative importance of biotic and abiotic influences operating at multiple spatial and temporal
82 scales, even when populations are relatively small or isolated.

83 Beyond within-species recruitment synchrony, examinations of the synchrony of multiple
84 species either within or among systems are lacking (Edwards et al., 2007; Michaletz and Siepker,
85 2012; Rook et al., 2012), limiting potential inferences that could be made about the relative
86 importance of fine- and broad-scale environmental conditions to recruitment of entire
87 assemblages. Species-specific traits (e.g., spawning phenology, early life history, trophic niche)
88 could determine which environmental factors are most important to recruitment success in fishes,
89 on what scale such factors may act, and how environmental variation may structure recruitment
90 synchrony across populations and species. Therefore, examining recruitment synchrony within
91 and among several species coexisting in small lakes (thereby eliminating dispersal as a potential
92 mechanism for synchrony) offers an avenue toward distinguishing the mechanisms by which
93 environmental variables drive fish population dynamics on different spatial scales by addressing
94 several hypotheses simultaneously: 1) If fine-scale, within-lake interactions are most important
95 to fish recruitment and species respond similarly to environmental variation, there should be high
96 synchrony among species inhabiting the same lakes; 2) If broad-scale, climatological variation is
97 more important to recruitment success and species respond similarly to climatological influences,
98 there should be high geographic synchrony within and among species across different lakes; and
99 3) If interactions between species' traits and environment are most important to determining
100 recruitment success, synchrony may be low both within and among species at any geographic
101 scale.

102 We sought to address these predictions using a temporally (~ 30 years) and spatially (130
103 glacial lakes covering ~200 km in geographic distance) expansive set of observations of year-

104 class strength in five ecologically and recreationally important freshwater fish species that are
105 prevalent in North American lakes: black crappie (*Pomoxis nigromaculatus*), bluegill (*Lepomis*
106 *macrochirus*), largemouth bass (*Micropterus salmoides*), redear sunfish (*L. microlophus*), and
107 yellow perch (*Perca flavescens*). These species exhibit different traits and their recruitment has
108 been linked to both abiotic and biotic environmental conditions in inland lakes (Dembkowski et
109 al., 2017; DeVries et al., 2009; Garvey et al., 2002). Examining within- and among-species
110 variability in recruitment in a suite of species enabled us to infer how species traits and
111 environmental conditions interact to shape dynamics of lentic populations across geographic and
112 temporal scales.

113 **2. Methods**

114 *2.1. Fish data*

115 The five focal species of this study represent different reproductive life histories and
116 responses to environmental variation. The centrarchids (largemouth bass, black crappie, bluegill,
117 and redear sunfish) are nesting species exhibiting parental care of eggs until hatch, whereas
118 yellow perch exhibit no parental care. Yellow perch spawn earliest in spring at temperatures
119 between 4 to 15 °C (Feiner and Höök, 2015), followed by largemouth bass and black crappie at
120 temperatures between 15 and 20 °C, and bluegill and redear sunfish spawn latest at temperatures
121 between 20 to 25 °C (Cooke and Philipp, 2009). Bluegill and redear sunfish have the capacity to
122 produce multiple broods per year, whereas the other species typically spawn once per year. Each
123 species produces larvae that initially inhabit pelagic areas before transitioning to demersal
124 habitats later in summer and fall (Bryan and Scarneccchia, 1992; Faber, 1967), during which
125 conditions promoting sufficient young-of-year growth to allow individuals to survive overwinter
126 (e.g., warm temperatures, abundant invertebrate resources) are thought to represent an important

127 recruitment bottleneck (Bunnell et al., 2011; Miller and Storck, 1984; Santucci Jr. and Wahl,
128 2011). Largemouth bass, black crappie, and yellow perch all transition from invertivory to
129 piscivory to varying extents during their ontogeny (Mittelbach and Persson, 1998), whereas
130 bluegill and redear sunfish rely primarily on invertebrate prey throughout life (Huckins, 1997;
131 Werner and Hall, 1974).

132 Fish were collected from 130 lakes across northern Indiana during standardized annual
133 surveys performed by the Indiana Department of Natural Resources from 1987 to 2009 (Table 1,
134 Figure 1). During annual surveys, fish were collected during summer (May – October) using
135 three possible gear types: night electrofishing (DC, 5-6 amps, 530 volts, 60 pps with two netters,
136 in 15 minute transects in shallow nearshore habitats), overnight gill nets (76 x 1.8 m with five, 15
137 m panels of 1.3, 2.5, 3.8, 5.1, and 6.4 cm square mesh, fished on the bottom and set
138 perpendicular to shore beginning < 2 m deep), and overnight trap nets (13.7 m lead, two
139 rectangular 1.8 x 3 m frames, four 0.8 m diameter circular frames set 1.2 m apart with 1.5mm
140 stretched mesh, fished in waters < 1 m deep with the opening toward shore). Effort was
141 standardized to account for lake size (see Sullivan et al., 2015). Because surveys were
142 completed in multiple years for some lakes, we selected a single sampling year from each lake to
143 include in our analysis. Sampling years with the most available data (e.g., highest catch rates,
144 most fish aged) for each species were selected to represent each lake in the final dataset.

145 For each survey all fish were measured for total length, while a subset of fish also had
146 scales removed for age determination. Scale samples were removed from fish below the lateral
147 line, near the pectoral fin using a pocket knife and placed in a scale envelope. Scale impressions
148 were prepared on clear acetate slides using a hydraulic laboratory press or roller press as
149 described by (Nielsen and Johnson, 1983). Scale impressions were magnified and viewed under

150 microfiche readers by the IDNR district fisheries biologists that completed the survey. Scales
151 are known to underestimate the ages of older fish depending on species and reader experience
152 (Long and Fisher, 2001; Maceina and Sammons, 2006). However, precision and agreement of
153 age estimates using scales for fish younger than 6 to 7 years is similar to those of age estimates
154 using otoliths, spines, or opercular bones (e.g., Hoxmeier et al., 2001; Isermann et al., 2010;
155 Long and Fisher, 2001; Vandergoot et al., 2008). Less than 5% of our total dataset consisted of
156 fish older than seven years, suggesting that age estimates included in this study are likely
157 relatively precise. In addition, we used weighted regression to reduce the importance of old, rare
158 fish on the results of our analyses (see below). Therefore, our results should be robust to
159 potential biases arising from the use of scales in ageing fish.

160 To estimate the relative abundance of age classes for each species and lake, age-length
161 keys were developed using the aged and measured subsample from each species and lake by
162 dividing fish into 4 mm length bins and only including lakes with at least three year classes
163 present. Within datasets of aged fish, ages were linearly interpolated across length bins where no
164 aged fish were captured. Unaged individuals in length bins greater than the largest aged
165 individual were excluded from further analyses. Species- and lake-specific age-length keys were
166 then applied to the remainder of measured fish to estimate the abundance of each age-class,
167 which was then expressed in terms of catch per unit effort (fish hr^{-1} electrofishing, and fish net-
168 night $^{-1}$ for gillnetting and trapnetting). Because of inherent biases in how each gear samples the
169 fish community and different sizes of fish (Sullivan et al., 2015), catches from only one gear per
170 species were included in further analyses based on highest available sample size (and therefore
171 likely highest catchability of that species in that gear) and a previous study demonstrating high
172 catchability of specific species or guilds in each gear (Sullivan et al., 2015). Specifically,

173 bluegill and largemouth bass recruitment was indexed from night electrofishing catches, redear
174 sunfish from trapnet catches, and black crappie and yellow perch from gillnet catches.

175 *2.2. Determination of year-class strength and environmental conditions*

176 Once the relative abundance of age classes was determined for each lake and species, year-
177 class strength was determined using catch curve analysis (Maceina, 2004; Tetzlaff et al., 2011).
178 For each species and lake where at least three age classes were present, weighted linear
179 regression using the natural log CPUE of each age class as the response variable and age as the
180 predictor variable was used to develop a catch curve and predicted relative abundances for each
181 age class (Quinn and Deriso, 1999). To ensure that only age classes that were fully recruited to
182 the gear in each lake were used in the catch curve, we started each catch curve at the age of peak
183 abundance for each species and lake. The influence of rare, old age classes on the results of
184 each catch curve was limited by weighting observations by the natural log of the predicted
185 number of fish comprising each age class (Honsey et al., 2016). Quality of fit was further
186 ensured via visual inspection and assessment of catch curve explanatory power. The residuals of
187 each catch curve regression were then retained as relative indicators of year-class strength, where
188 positive residuals indicate that an age-class was more abundant than expected, suggesting a good
189 year class, and negative residuals indicate an age-class was less abundant than expected,
190 suggesting a poor year class. Catch curve residuals were then assigned to the year of birth for
191 that age class by subtracting the age of the fish from the year sampling occurred.

192 The abundance and age structure of species varied considerably across lakes, and catch
193 curve residuals were generally negatively skewed, complicating use of linear or parametric
194 models in assessing variability in recruitment. Furthermore, high recruitment variability

195 observed in many species may render the exact magnitudes of year-class strength less important
196 to managers than an assessment of whether recruitment was generally “good” or “poor” over
197 time or in response to environmental conditions (Hansen et al., 2015). Therefore, we
198 transformed catch curve residuals into a binary index of relative year-class strength by assigning
199 year classes with positive residuals (i.e., abundance was higher than expected – a “good” year
200 class) a value of one, and year classes with negative residuals (i.e., abundance was lower than
201 expected – a “poor” year class) a value of zero (following Honsey et al., 2016) to standardize
202 estimates of relative year-class strength across lakes.

203 In addition to understanding the extent of recruitment synchrony within and among
204 species in Indiana lakes, we were also interested in elucidating potential environmental drivers of
205 any observed synchrony. We obtained measures of lake habitat, including lake area (km^2),
206 maximum depth (m), the ratio of maximum to mean depth (a measure of littoral habitat
207 availability), and lake shoreline development index (a measure of lake circularity) from the
208 USGS National Hydrography dataset and Indiana Department of Natural Resources publications
209 (IDNR, 1993, 1966; Perry, 2011; USGS, 2004). As indicators of water quality, resource
210 availability, and oxythermal habitat quality (Downing et al., 1990; Magnuson et al., 1979;
211 Missaghi et al., 2017; Persson et al., 1991), Secchi depth (m), total phosphorus ($\mu\text{g L}^{-1}$),
212 dissolved oxygen at 1.5 m (mg L^{-1}), and water temperature at 1.5 m ($^{\circ}\text{C}$), were obtained from the
213 Indiana Clean Lakes Program (ICLP; <http://www.indiana.edu/~clp/indianalakeinfo.php>). These
214 metrics had been previously identified as potentially important and non-collinear variables
215 influence fish assemblage structure in Indiana lakes (Feiner et al., 2016). We characterized
216 regional weather patterns by obtaining daily wind speed (km h^{-1}) and air temperature ($^{\circ}\text{C}$) data
217 from four land-based NOAA weather stations and precipitation data (2.54-mm increments) from

218 six NOAA weather stations around northern Indiana from 1983 to 2010. From these sources, we
219 estimated annual mean daily wind speed, annual mean daily temperature, mean daily spring
220 (April – June) temperature, the number of winter days (mean daily temperature < 0 °C), and
221 cumulative annual and spring (April – June) precipitation for each year by averaging measures
222 across weather stations (see Supplement 1 and Feiner et al., 2016 for detailed information on
223 environmental and climatic variables). We scaled and centered each climatic variable to a mean
224 of 0 and standard deviation of 1 by subtracting the mean from each value and dividing by the
225 standard deviation (resulting in coefficients with units of change in response per one standard
226 deviation change in predictor) to standardize interpretations of environmental effects on
227 recruitment.

228 Moran effects are assumed to operate via synchronous climatic conditions across space
229 driving synchrony in population dynamics (Engen and Sæther, 2005). We tested for spatial
230 correlations in our climatic variables by correlating annual measures of climatic (wind,
231 temperature, and precipitation) variables between all pairwise combinations of weather stations.
232 Correlations were highly significant across all wind (mean $\rho = 0.83$, $P \leq 0.001$) and temperature
233 variables (mean $\rho = 0.67$, $P \leq 0.016$; Table S2) and most precipitation correlations (mean $\rho =$
234 0.48, $P < 0.05$ in all but 4 out of 30 correlations; Table S3). Thus, climatic variables were
235 consistently spatially synchronous in northern Indiana during the study, supporting this key
236 assumption for the Moran effect.

237 *2.3 Data analysis*

238 *2.3.1. Within-lake synchrony between species*

239 As an initial comparison of the variability in relative year-class strength among species,
240 we calculated the coefficient of variation (CV; standard deviation/mean $\times 100$) of the absolute
241 value of catch curve residuals for each lake and species, then averaged CV across lakes for each
242 species. This index of recruitment variation was then compared across species using ANOVA
243 with type III sums of squares to account for unbalanced number of observations among species.

244 To evaluate within-lake recruitment synchrony between species, we identified all
245 instances where pairs of species had overlapping year classes in the same lake. Using this subset
246 of paired observations, we determined proportion agreement of relative year-class strength,
247 where each instance where both species produced either good year classes (i.e., both species
248 binary residuals equaled 1) or poor year classes (both species residuals equaled zero) were
249 counted as agreements (1) and cases where one species recorded a good year class and the other
250 a poor year class were recorded as disagreements (0) (Honsey et al., 2016; Zischke et al., 2017).
251 From this, the proportion agreement within each lake was determined as the number of year
252 classes with agreement divided by the total number of overlapping year classes in that lake. For
253 each species pair, we then calculated the weighted mean proportion agreement across all lakes
254 where both species were observed, weighting each observation by the number of years the
255 species pair overlapped in each lake; i.e., the proportion of “successes” (agreements) was
256 weighted by the number of trials (overlapping year classes).

257 We used two methods to statistically evaluate pairwise species recruitment synchrony.
258 First, we used a one-way weighted t-test (each observation was weighted by the number of years
259 of overlap) to determine whether mean weighted proportion agreement was greater than 0.5
260 (expected if recruitment in each species was random). Second, we used the raw residuals (i.e.,
261 before transformation to binary values) to perform repeated-measures Spearman-rank

262 correlations on paired recruitment observations ranked and nested within lakes (R package
263 ‘rmcorr’; Bakdash and Marusich, 2017). Each analysis was repeated for each species pair.

264 There was considerable variation among lakes in recruitment agreement. We further
265 sought to evaluate how in-lake habitat and water quality may have influenced variability in
266 recruitment synchrony between species using the suite of within-lake habitat and water quality
267 variables developed above. We used multiple logistic regression with mean percent agreement
268 in relative year class strength as the response and each environmental measure (Secchi depth,
269 total phosphorus, dissolved oxygen at 1.5 m, water temperature at 1.5 m, lake area, maximum
270 depth, maximum:mean depth, and shoreline development index) as additive explanatory
271 variables. To limit model overfitting and identify a most-parsimonious model, we then used
272 forward and backward stepwise model selection using AIC as a selection criterion. Model fit
273 was assessed using McFadden’s $R^2 = 1 - (loglikelihood(null)/loglikelihood(model))$, and
274 the Hosmer-Lemeshow test, which tests whether observations with similar predicted
275 probabilities have similar observed probabilities with a χ^2 test (where small χ^2 and large p-values
276 indicate good model fit).

277 2.3.2. Among-lake synchrony within species

278 We sought to evaluate spatial and environmental synchrony in recruitment within each
279 focal species using two methods. First, we examined the spatial scale of synchrony in
280 recruitment using the *e*-folding scale method (Honsey et al., 2016; Myers et al., 1997), which
281 determines the distance between populations required to decrease the agreement in recruitment
282 by a factor of e^1 . This analysis is normally performed on the correlation of year-class strengths
283 between pairs of lakes; however, due to the random sampling of lakes in our study, few lake

284 pairs had sufficient data overlaps to perform meaningful correlations. Instead, we identified all
285 instances where pairs of lakes had overlapping observations of year-class strength for a single
286 species and scored the agreement of recruitment strength as either 1 (both lakes had positive or
287 negative residuals) or -1 (one lake had a positive residual while the other had a negative
288 residual). We then averaged the agreement of each lake pair across all years where overlapping
289 year classes were observed to develop an index of recruitment agreement ranging from 1 (all
290 year classes agreed) to -1 (no year classes agreed). To determine the spatial extent of
291 recruitment synchrony, we then fit this index of agreement to the exponential decay equation:

292
$$\rho(d) = \rho_0 e^{-\left(\frac{d}{v}\right)},$$

293 where $\rho(d)$ is the agreement between a lake pair, ρ_0 is the agreement between lakes with no
294 geographic separation, d is the great circle distance between lakes (km), and v is the e-folding
295 scale parameter. We constrained ρ_0 to have values less than 1 (Myers et al., 1997) and weighted
296 each observation by the number of overlapping year classes observed (Honsey et al., 2016).

297 We further sought to evaluate whether environmental similarity of lakes, rather than their
298 spatial proximity, was a stronger driver of recruitment synchrony in these populations. To do so,
299 we calculated the univariate Euclidean distance of each environmental variable (Secchi depth,
300 total phosphorus, dissolved oxygen at 1.5 m, water temperature at 1.5 m, lake area, maximum
301 depth, maximum to mean depth ratio, and lake shoreline development index) between each
302 potential pair of lakes. We then used logistic regression with mean binomial (0 or 1, as in
303 section 2.3.1) proportion agreement in relative year-class strength as the response variable and
304 similarity in each environmental variable as the predictor variables. Each observation was
305 weighted by the number of year-class overlaps. In this analysis, we expected that lakes similar in

306 environmental variables (i.e., small Euclidean distances) should have higher agreement in
307 recruitment, leading to negative relationships between environmental dissimilarity and mean
308 proportion agreement in relative year-class strength. Forward and backward stepwise model
309 selection using AIC was used to select the most parsimonious model explaining recruitment
310 synchrony. Model fit was assessed using McFadden's R^2 and the Hosmer-Lemeshow test.

311 2.3.3. Among-lake synchrony between species

312 Our next question of interest was whether year-class strength exhibited synchrony
313 between species at a regional (among-lake) level, which allowed us to expand our scope of
314 inference to the entire time series of data, rather than being limited to specific cases where year
315 classes from two species were sampled in the same lake. To evaluate this question, we sought to
316 correlate mean relative year-class strengths between species across time. Specifically, we first
317 averaged binary relative year-class strengths across lakes for each species to develop an annual
318 measure of whether recruitment was generally good or poor each year for each species, thereby
319 creating a continuous, normally distributed variable ranging from 0 to 1 (where 0 indicates poor
320 recruitment across lakes that year, and 1 indicates good recruitment across lakes that year). We
321 then used Spearman rank correlation to evaluate the correlation of mean relative year-class
322 strengths between each potential species pair, weighting each observation by the fewest number
323 of lakes for which relative year-class strength was determined for either species that year.

324 2.3.4. Climatic drivers of year-class strength

325 We used generalized multiple logistic regression modeling to determine the effect of
326 climatic variables (wind, temperature, winter days, and precipitation) on mean annual
327 recruitment for each species. Binary relative year-class strength (as developed above) was the

328 response, the six climatic variables were additive predictors, and we included random intercepts
329 for lake identity. Model fit was assessed by the area under the ROC curve (AUC, a measure of
330 classification performance which ranges from 0.5 – 1.0, with values > 0.8 considered good fitting
331 models) and the Hosmer-Lemeshow test. Significance of climatic variables was determined
332 using t-statistics and p-values using Satterthwaite's type III degrees of freedom (R package
333 'lmerTest'; Kuznetsova et al., 2016).

334 **3. Results**

335 *3.1. Within-lake synchrony between species*

336 In total, recruitment from at least one species was indexed in 130 glacial lakes located across
337 northern Indiana. Bluegill and largemouth bass were ubiquitous across the region (~100 lakes
338 and ~500 year-class observations), whereas yellow perch, redear sunfish, and black crappie were
339 observed less often (< 25 lakes and < 100 total year classes; Table 1, Fig. 1). Coefficients of
340 variation in catch curve residuals were variable among lakes and species, ranging from 72 to
341 191%. After natural log-transforming CV to meet normality assumptions, there was evidence
342 that within-lake variation in relative year-class strength significantly differed among species
343 (ANOVA: $F_{4,260} = 5.50$, $p < 0.001$). Specifically, largemouth bass recruitment exhibited
344 significantly lower mean CV of catch curve residuals across lakes than bluegill (*post hoc*
345 Tukey's test, $p < 0.001$), indicating reduced recruitment variability in largemouth bass compared
346 to bluegill. However, this difference was small in magnitude (16%) and no other species
347 differed in CV (*post hoc* Tukey's test, $p > 0.15$). Generally, the standard deviation of catch
348 curve residuals was slightly less than their average magnitudes (range in mean CV: 82 – 98%).

349 We observed generally minimal recruitment synchrony between species when limiting the
350 data to paired observations of year classes within lakes. While mean percent agreement in year-
351 class strength was greater than 50% in 5 of 10 possible species pairs, only bluegill and black
352 crappie recruitment agreement agreed significantly more often than 50% of the time (Table 2).
353 Supporting this, there was only one species pair, bluegill – black crappie, that had a significant,
354 positive correlation between year-class strengths (Figure 2a). There were, however, three
355 species pairs (black crappie – largemouth bass, bluegill – redear sunfish, and redear sunfish –
356 yellow perch) that exhibited significant, negative correlations in year-class strength (Table 2;
357 Figure 2b-d). There were few significant effects of in-lake habitat or water quality
358 characteristics on recruitment agreement between species – bluegill and largemouth bass
359 synchrony was positively related to the ratio of maximum to mean depth, and bluegill and yellow
360 perch synchrony was negatively related to lake area (Table S4, Fig. S1).

361 *3.2. Among lake synchrony within species*

362 *e*-folding scale regression indicated very low spatial synchrony in year-class strength among
363 Indiana glacial lakes (Table 3), as indicated by small (or negative) and non-significant effects of
364 v , suggesting either no synchrony at all, or synchrony limited to less than 10 km among lakes
365 (Figure 3).

366 Due to the lack of spatial synchrony, we then sought to evaluate whether environmental
367 similarity was driving year-class strength agreement between lakes for each species. Model
368 performance was generally very poor (low McFadden's R^2 and low Hosmer-Lemeshow test p-
369 values; Table S5). Bluegill year-class strength agreement was significantly, positively related to
370 differences in lake depth, and redear sunfish year-class strength agreement was significantly,

371 negatively related to differences in temperature at 1.5m depth (Figure S2). No other species
372 yielded significant relationships between recruitment synchrony and lake environmental
373 similarity (Table S5).

374 *3.3. Among lake synchrony between species*

375 Examinations of correlations in year-class strength at the regional level revealed a similar
376 lack of synchrony as within-lake analyses. Although 6 of 10 possible species pairs exhibited
377 positive correlations of mean year-class strength across lakes, only one correlation was even
378 marginally significant: a positive correlation between redear sunfish and largemouth bass (Table
379 5; Figure S3).

380 *3.4. Climatic drivers of year-class strength*

381 There were relatively few significant relationships between binary year-class strength and
382 regional climatic variables across all species (Table S6). Bluegill and black crappie recruitment
383 exhibited a significant positive relationship with annual mean daily wind speed, and yellow
384 perch recruitment was positively related to spring precipitation. No other species exhibited
385 significant relationships between year-class strength and climatic variables. In addition, model
386 performance was generally poor, with AUC less than 0.8 and generally small Hosmer-Lemeshow
387 p-values, indicating the predictive ability of our selected climatic variables was limited for the
388 recruitment of these species.

389 **4. Discussion**

390 Examining recruitment patterns in five species across 130 lakes allowed us to evaluate
391 several potential predictions for how environmental variables at multiple spatial scales influence

392 and potentially synchronize fish population dynamics. We observed surprisingly few instances
393 of synchrony in recruitment either between species inhabiting the same lakes, within species
394 inhabiting different (even environmentally similar lakes), or among species at a regional scale.
395 In addition, there was inconsistent evidence of environmental or climatic variables influencing
396 recruitment of each species. These results give strongest support to our third prediction: that
397 interactions between species-specific traits and local-scale environmental variation, both abiotic
398 and biotic, are most important in driving recruitment variation in these types of small glacial
399 lakes. Our results support the paradigm that the scale of recruitment synchrony and importance
400 of climatological variables is correlated to system size (Myers et al., 1997).

401 In general, spatial synchrony among populations of the same species is thought to be limited
402 between small, disconnected lakes. The lack of dispersal between water bodies and variable
403 biotic, habitat, and environmental conditions among these types of systems would seem to
404 reduce the amount of potential environmental similarity fish populations would experience, even
405 if broad-scale climatic conditions were correlated across large geographic areas. The few
406 published studies examining recruitment synchrony in our target species, in concert with our
407 findings, would seem to support this hypothesis, especially in the sunfishes. Bluegill recruitment
408 has been shown in multiple studies to be largely determined by biotic conditions (Kaemingk et
409 al., 2013; Parkos et al., 2011; Santucci Jr. and Wahl, 2011), with limited influence of abiotic
410 variables and minimal among-population synchrony (Edwards et al., 2007; Tomcko and Pierce,
411 2011). Similarly, largemouth bass recruitment has been related to system productivity, prey and
412 habitat availability, and density-dependent effects (Michaletz and Siepker, 2012; Miller and
413 Storck, 1984; Paukert and Willis, 2004; Post et al., 1998), with few observed relationships to
414 climate patterns (but see Rypel, 2009). Similar dynamics appear to influence black crappie

415 recruitment (Bunnell et al., 2011; Guy and Willis, 1995; Maceina, 2003). The evidence for
416 synchrony in yellow perch populations is mixed. Honsey et al., (2016) and Dembkowski et al.,
417 (2016) both observed extensive recruitment synchrony in the Great Lakes and inland lake
418 systems, respectively, but yellow perch between Muskegon Lake and Lake Michigan, and in this
419 study, showed little concordance even at short geographic distances (Janetski et al., 2013). In
420 addition, yellow perch recruitment has been extensively related to biotic conditions like predator
421 abundance and density-dependent dynamics (Forney, 1971; Irwin et al., 2009). Thus, we
422 conclude that, within species, recruitment synchrony is likely to be minimal across inland lake
423 populations.

424 One species pair, black crappie and bluegill, exhibited synchronous dynamics in their
425 correlations of year-class strength when coexisting and in their regional positive response to
426 mean annual wind speed. Both bluegill and black crappie prefer nesting sites sheltered from
427 wind (Pope and Willis, 1997; Stahr et al., 2013); thus, altered spawning site selection in windy
428 years may affect recruitment success similarly in both species. Wind speeds may also alter lake
429 mixing, primary production, and zooplankton phenology and distribution, influencing potential
430 forage for larval and juvenile fishes (de Souza Cardoso and da Motta Marques, 2009; Gauthier et
431 al., 2014). However, a precise mechanism remains unclear, as wind speeds have not been
432 strongly linked to larval distribution or abundance in either species (Kaemingk et al., 2011; Post
433 et al., 1995), although Pope et al., (1996) also found a positive relationship between black
434 crappie recruitment and wind speed. Beyond wind, similar responses to biotic conditions may
435 structure synchrony between species. Black crappie and bluegill responded similarly to
436 largemouth bass removals in Alabama lakes (McHugh, 1990), and juveniles of both species
437 share similar habitat and trophic niches (Holland and Huston, 1985; Knights et al., 1995; Werner

438 et al., 1977). Thus, a combination of within-lake interactions and regional climatic patterns may
439 result in similar population dynamics between these two species, where conditions benefitting
440 black crappie recruitment, a goal of fishery biologists seeking to improve fisheries for this
441 recreationally popular species (Boxrucker and Irwin, 2002), may also benefit bluegill
442 populations.

443 Habitat availability and water quality can act as important regulators of fish population
444 dynamics, influencing everything from spawning site preferences to larval growth and survival
445 (Sass et al., 2017). Therefore, despite limited geographic synchrony within or among species, we
446 also attempted to discern whether environmental characteristics could potentially drive
447 recruitment synchrony. Evidence was limited, but there was an indication that agreement
448 between bluegill and largemouth bass was positively related to littoral habitat availability within
449 lakes (maximum:mean depth), and similarity in lake depth and water temperature were
450 significantly related to within-species agreement in bluegill and redear sunfish recruitment
451 among lakes, respectively. Bluegill generally spawn at depths less than 2 m (Gosch et al., 2006)
452 – steep sided (i.e., high ratio of maximum to mean depths), deep lakes may offer less littoral
453 habitat, limiting bluegill to smaller areas or making them more susceptible to environmental
454 variability by offering lower quality habitat to juvenile bluegill and largemouth bass, thereby
455 increasing synchrony compared to lakes with abundant and variable littoral habitats (Gaeta et al.,
456 2014; Sass et al., 2006; Werner and Hall, 1974). There have been few investigations of redear
457 sunfish recruitment, but studies on similar species (e.g., bluegill) have shown some importance
458 for spring temperatures to improve growth and promote overwinter survival (Santucci Jr. and
459 Wahl, 2011). Therefore, lakes with similar thermal conditions may generally produce similar
460 effects on redear sunfish recruitment success, while differences between thermally distinct lakes

461 may result from interactions with other abiotic or biotic variables. Similar nonlinear or
462 interacting effects in recruitment dynamics have been observed in other species. For instance,
463 walleye recruitment was found to be more resilient to warming temperatures when largemouth
464 bass abundance was low in Wisconsin lakes (Hansen et al., 2017), and water transparency
465 interacted with zooplankton abundance to structure bluegill recruitment in a Nebraska reservoir
466 (Kaemingk et al., 2013). These results suggest that the interplay of abiotic and biotic conditions
467 can complicate the assessment of recruitment patterns in lentic species under dissimilar
468 environmental conditions.

469 More broadly, our results suggest that biotic interactions between species likely represent
470 critical drivers of recruitment success. While there was only a single positive correlation in
471 relative year-class strength between species, there were significant negative correlations between
472 black crappie and largemouth bass, bluegill and redear sunfish, and yellow perch and redear
473 sunfish relative year-class strength. The observed negative correlations primarily occurred
474 between species that inhabit relatively similar trophic niches. Both black crappie and largemouth
475 bass transition from benthic invertebrate prey to fish prey during their ontogeny, and an inability
476 to switch to fish prey can significantly reduce juvenile performance (Ellison, 1984; Mittelbach
477 and Persson, 1998; Olson, 1996). In addition, negative relationships between black crappie
478 populations and largemouth bass abundance have been observed elsewhere, suggesting that
479 strong largemouth bass year classes may depress black crappie recruitment (McHugh, 1990;
480 Schultz et al., 2008). Similarly, bluegill, redear sunfish, and yellow perch all prey on a
481 combination of zooplankton and benthic invertebrates during their first year of life (Feiner and
482 Höök, 2015; Huckins, 1997; Werner and Hall, 1974). Potentially, competitive interactions
483 between these species may have led to these negative correlations, where strong year classes for

484 one species limited year-class success by others and vice versa. Additional work to uncover the
485 mechanisms by which these species potentially interact could further elucidate whether these
486 types of biotic interactions may influence their recruitment.

487 Our conclusions contrast some recent work suggesting that freshwater species can exhibit
488 recruitment synchrony on geographic scales approaching those observed in large, marine
489 systems, including percids (Beard et al., 2011; Dembkowski et al., 2016; Honsey et al., 2016),
490 cyprinids (Grenouillet et al., 2008; Marjomäki et al., 2004; Phelps et al., 2008), and coregonids
491 (Bunnell et al., 2010; Myers et al., 2015). The differences we observed may have resulted from
492 differences in life history, system size, or desynchronization during ontogeny. First, most
493 species previously shown to exhibit recruitment synchrony are broadcast spawning fishes with
494 no parental care, which may expose their offspring to stronger influences of spatially-correlated
495 climatic conditions (Pope et al., 1996; Beard et al., 2011; Honsey et al., 2016; Janetski et al.,
496 2013). This contrasts the centrarchids included in this study, which exhibit parental care and
497 show stronger effects of biotic variables on recruitment (Edwards et al., 2007; Post et al., 1998;
498 Santucci Jr. and Wahl, 2011). Secondly, the systems in our study were generally much smaller
499 than those within which other freshwater species have exhibited broad-scale synchrony, most
500 notably in the Great Lakes, where potential larval dispersal and connected habitats may boost
501 potential recruitment synchrony similarly to large marine systems (Bunnell et al., 2016; Honsey
502 et al., 2016). Finally, within-lake conditions may act to desynchronize populations later in life.
503 Our catch curve analysis was estimating year-class strength based on the abundances of adult
504 (age 2+) fish. Recruitment synchrony in inland lakes has previously been observed in analyses
505 indexing recruitment using the abundances of early life stages (larvae or young-of year; Beard et
506 al., 2011; Dembkowski et al., 2016; Marjomäki et al., 2004). Potentially, early life abundances

507 may be synchronized by climatic conditions, but become unlinked by within-lake processes
508 during ontogeny. For instance, Grenouillet et al., (2008) observed synchrony in the abundance
509 of age-0 roach (*Rutilus rutilus*) along the Rhône River, but little synchrony in age-1 abundances.
510 In small systems, therefore, biotic or other in-lake environmental variation is likely most
511 important to recruitment variability and may act to decouple adult abundances. Thus, our
512 observations of minimal spatial synchrony are potentially typical for these study species.

513 Findings of spatial synchrony in fish populations have been identified as a potential avenue
514 for determining the proper spatial scale for management actions or in predicting the responses of
515 species to climatological events or climate change at broader regional scales (Hansen et al.,
516 2015; Tonkin et al., 2017). While these are useful goals when recruitment synchrony is present,
517 our results underscore the importance of understanding how habitat and biotic interactions at the
518 lake level regulate recruitment success. Our observations of at least some within-lake
519 environmental influences on recruitment or recruitment synchrony and a near-complete lack of
520 spatial synchrony or regional climate effects would indicate that, in smaller lakes, spatial
521 variability in habitat may be more important than geographic proximity. Therefore, after
522 assessing recruitment synchrony within and among lakes and species across more than a hundred
523 systems, we suggest that recruitment dynamics for the species included in this study are more a
524 function of the interplay between species traits, biotic interactions, and habitat rather than any
525 large-scale climatic disturbances. Thus, inferences about recruitment may be limited in the
526 absence of data taken at sufficiently fine temporal and spatial scales.

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821

822 Table 1. Lake and species sample sizes used in analyses. Below the diagonal are the number of
823 lakes with species pairs, and above the diagonal are the total number of species pair observations
824 across lakes and years. Values on the diagonal represent the number of lakes for that species,
825 with total number of year classes observed across lakes listed in parentheses.

Species	Black crappie	Bluegill	Largemouth bass	Redear sunfish	Yellow perch
Black crappie	13 (46)	31	31	2	9
Bluegill	10	107 (499)	271	64	64
Largemouth bass	10	71	98 (502)	53	52
Redear sunfish	1	19	15	23 (95)	16
Yellow perch	3	19	16	5	24 (91)

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828 Table 2. Upper: weighted mean (SE) proportion agreement in year-class strength (below
 829 diagonal) and p-value from one-way weighted t-test testing whether proportion agreement was >
 830 0.5 (above diagonal). Lower: repeated-measures Spearman rank correlation coefficients between
 831 paired observations of year-class strength within lakes (below diagonal) and associated p-values
 832 (above diagonal). Dashes (--) indicate insufficient sample size for a comparison and bolded
 833 values are significant at $p < 0.05$. See Table 1 for sample sizes.

<i>Mean proportion agreement in year-class strength</i>					
Species	BLC	BLG	LMB	RES	YEP
BLC		0.006	0.829	--	0.409
BLG	0.74 (0.08)		0.355	0.970	0.500
LMB	0.39 (0.11)	0.51 (0.03)		0.279	0.934
RES	--	0.36 (0.07)	0.55 (0.08)		0.928
YEP	0.56 (0.21)	0.50 (0.07)	0.40 (0.06)	0.31 (0.10)	
<i>Correlation of year-class strength within lakes</i>					
	BLC	BLG	LMB	RES	YEP
BLC		0.006	0.042	--	0.690
BLG	0.567		0.651	0.008	0.928
LMB	-0.436	-0.032		0.187	0.094
RES	--	-0.388	0.216		0.041
YEP	0.186	0.014	-0.279	-0.595	

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837 Table 3. *e*-folding scale model results to determine extent of spatial synchrony in the recruitment
838 of five sportfish species in northern Indiana, including model coefficients (SE) ρ_0 (the expected
839 correlation when distance is zero) and v (the distance (km) required to decrease synchrony by e^1),
840 model degrees of freedom (df), and residual standard error (RSE). Minimal spatial synchrony
841 was observed across all species.

Species	ρ_0	v	df	RSE
Bluegill	-0.307 (1.644)	0.574 (2.245)	2261	1.050
Largemouth bass	0.375 (0.784)	1.012 (1.96)	2059	1.045
Redear sunfish	0.545 (0.323)	8.992 (9.835)	81	1.174
Black crappie	0.176 (0.178)	-90.867 (77.246)	20	1.146
Yellow perch	0.007 (0.031)	-31.572 (35.311)	90	1.103

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843

844 Table 4. Weighted Spearman rank correlation coefficients (r) for mean annual binary year-class
845 strength between species (below diagonal). Observations were weighted by the number of lakes
846 sampled each year. The number of years included in the analysis are shown in parentheses, and
847 p-values for correlations are listed above the diagonal.

	Black crappie	Bluegill	Largemouth bass	Redear sunfish	Yellow perch
Black crappie		0.551	0.624	0.787	0.414
Bluegill	0.148 (20)		0.497	0.545	0.961
Largemouth bass	0.122 (20)	0.145 (26)		0.049	0.902
Redear sunfish	0.072 (18)	-0.142 (22)	0.434 (22)		0.806
Yellow perch	-0.221 (17)	0.012 (22)	-0.029 (22)	0.059 (21)	

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850 Figure Captions

851 Figure 1. Location of northern Indiana, USA, lakes where black crappie, bluegill, largemouth
852 bass, redear sunfish, and yellow perch were sampled by the Indiana Department of Natural
853 Resources from 1982 to 2010. Size of dots correspond to number of year classes observed in
854 each lake.

855 Figure 2. Significant, repeated measures Spearman correlations of relative year-class strength
856 (YCS) were observed between a) bluegill and black crappie, b) largemouth bass and black
857 crappie, c) redear sunfish and bluegill, and d) yellow perch and black crappie. Only bluegill and
858 black crappie recruitment was positively correlated.

859 Figure 3. Relationships between mean scaled proportional agreement in relative year-class
860 strength and geographic distance (km) between lakes for a) black crappie, b) bluegill, c)
861 largemouth bass, d) redear sunfish, and e) yellow perch sampled in northern Indiana, USA,
862 glacial lakes. Circles (gray) represent observed agreement, with symbol size denoting the
863 number of overlapping year classes for each lake pair. Solid black line is the predicted *e*-folding
864 scale relationship for each species. Minimal spatial synchrony was observed across species.





