



RESEARCH ARTICLE

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Droughts in Amazonia: Spatiotemporal Variability, Teleconnections, and Seasonal Predictions

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Key Points:

- Spatiotemporal dynamics and teleconnections associated with Amazonia droughts are investigated based on the PDSI indices
- A drought forecast model for Amazonia is developed and tested based on the global SST field and sparse canonical correlation analysis
- This is the first study in the region that, based on cross-validation, leads to appreciable forecast skills for lead times beyond 4 months

Supporting Information:

- Supporting Information S1

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Abstract Most Amazonia drought studies have focused on rainfall deficits and their impact on river discharges, while the analysis of other important driver variables, such as temperature and soil moisture, has attracted less attention. Here we try to better understand the spatiotemporal dynamics of Amazonia droughts and associated climate teleconnections as characterized by the Palmer Drought Severity Index (PDSI), which integrates information from rainfall deficit, temperature anomalies, and soil moisture capacity. The results reveal that Amazonia droughts are most related to one dominant pattern across the entire region, followed by two seesaw kind of patterns: north-south and east-west. The main two modes are correlated with sea surface temperature (SST) anomalies in the tropical Pacific and Atlantic oceans. The teleconnections associated with global SST are then used to build a seasonal forecast model for PDSI over Amazonia based on predictors obtained from a sparse canonical correlation analysis approach. A unique feature of the presented drought prediction method is using only a few number of predictors to avoid excessive noise in the predictor space. Cross-validated results show correlations between observed and predicted spatial average PDSI up to 0.60 and 0.45 for lead times of 5 and 9 months, respectively. To the best of our knowledge, this is the first study in the region that, based on cross-validation results, leads to appreciable forecast skills for lead times beyond 4 months. This is a step forward in better understanding the dynamics of Amazonia droughts and improving risk assessment and management, through improved drought forecasting.

1. Introduction

The dynamics of the Amazonia ecosystem plays a significant role on the global biogeochemical cycles (McClain et al., 2001), wild fires (Chen et al., 2013), moisture transport to southeast South America (Drummond et al., 2008), regional and global climate (Nobre et al., 1991; Spracklen & Garcia-Carreras, 2015), and on local population (Marengo & Espinoza, 2016). In particular, extreme droughts in the Amazonia seem to have become more frequent in the last years (e.g., Marengo & Espinoza, 2016) and studies have shown their significant impacts on the regional ecosystem and biogeochemical cycles (Cochrane, 2003; Feldpausch et al., 2016; Laurance & Williamson, 2001; Maeda et al., 2015; Phillips et al., 2009), wild fires (Aragão et al., 2007), and on the local hydrological cycle (Lopes et al., 2016; Marengo et al., 2008a). Eventually, these changes in Amazonia propagate to distant regions because of large-scale atmospheric circulation dynamics. For instance, during austral summer, water vapor from Amazonia is transported to southeast South America (Arraut et al., 2012; Drummond et al., 2008; Silva & Ambrizzi, 2009). Changes in evapotranspiration in response to a drought will certainly change the water vapor transport—although the exact mechanisms and magnitude are still unknown—and affect water supply and hydropower generation in southeast South America.

Motivated by these ecosystem and socioeconomic impacts, a substantial number of studies (e.g., Arraut et al., 2012; Espinoza et al., 2011; Marengo, 1992; Marengo & Espinoza, 2016; Marengo et al., 2008b, 2011; Ropelewski & Halpert, 1987; Yoon & Zeng, 2010; and references therein) have tried to better understand the dynamics of droughts in Amazonia. The main cause of droughts in Amazonia has been attributed to the El Niño-Southern Oscillation (ENSO) and to a minor extent to the sea surface temperature (SST) variability in the Tropical North Atlantic. Warm SST anomalies in the eastern Tropical Pacific shift the descending branch of the Walker circulation over Amazonia and inhibit precipitation during the austral summer rainfall season. A warmer tropical north Atlantic will displace north the Inter-Tropical Convergence Zone from its

climatological position and therefore the ascending branch of the Hadley cell and reduce convection and precipitation over Amazonia.

Most studies on Amazonia droughts have primarily focused on rainfall variability (Aragão et al., 2007; Zou et al., 2016) and its teleconnections (Fernandes et al., 2015; Marengo, 1992; Yoon & Zeng, 2010), and the associated impacts on river flows across Amazonia (Espinoza et al., 2011; Lopes et al., 2016; Marengo et al., 2008b). We perceive that little is known about the compounding effects of rainfall and temperature over Amazonia, and how these two variables exacerbate drought impacts, particularly on the vegetation and natural ecosystem. On the other hand, the state of soil moisture could indicate water deficit caused by not only precipitation deficit but also excess evapotranspiration. Previous studies show that univariate drought risk assessment approaches based solely on rainfall can severely underestimate the underlying risk of extreme droughts (AghaKouchak et al., 2014; Shukla et al., 2015) and their impacts (Williams et al., 2013), and a multivariate approach can provide a more realistic assessment. In this sense, the use of drought indices that integrate different hydroclimate variables (e.g., rainfall, temperature, soil moisture, etc.), such as the Palmer Drought Severity Index (Palmer, 1965) and the Multivariate Standardized Drought Index (Hao & AghaKouchak, 2014) might reveal atmospheric and land conditions associated with water stress that are more suitable to investigate droughts and their impacts. In the case of Amazonia droughts, we perceive that just few, limited-scope studies (e.g., Dai, 2011; Dai et al., 2004; Jiménez-Muñoz et al., 2016; Joetzjer et al., 2013) have looked at drought indices and such field deserves further investigation, particularly in terms of understanding the variability of such indices, teleconnections associated, and predictive models for long-term forecasts.

Understanding and prediction of droughts across the world have been the subject of much research (e.g., Cook et al., 2010; Funk et al., 2014; Hoerling et al., 2012; Kwon et al., 2016; Lyon et al., 2012; Rajagopalan et al., 2000; Schubert et al., 2007; Seager et al., 2015; Thober et al., 2015) and some drought monitor systems (e.g., the U.S. Drought Monitor, <http://droughtmonitor.unl.edu/Home.aspx> and the Latin American Drought Monitor, <http://stream.princeton.edu/LAFDM/WEBPAGE/interface.php?locale=en>) plan or have already integrated seasonal drought forecasts. Here, we aim to advance the current Brazilian initiatives for drought monitoring (e.g., the INPE drought monitor, <http://clima1.cptec.inpe.br/spi/pt>) through a multivariate approach for drought analysis and prediction across Amazonia. This work is then carried out to better understand the spatiotemporal dynamics of droughts in Amazonia as informed by the widely used Palmer Drought Severity Index, PDSI (Alley, 1984; Dai et al., 2004; Palmer, 1965). Recently, it has been successfully used to study Amazonia droughts (Jiménez-Muñoz et al., 2016), mainly to investigate the relationship between the largest El Niño events in the last years (1982–1983, 1997–1998, and 2015–2016) and the drought severity and spatial distribution across Amazonia. We study sea surface temperature (SST) teleconnections and develop and test a long-term (up to 9 months lead time) forecast model for PDSI over Amazonia. We employ a linear model and exogenous predictors derived from the SST teleconnections and based on a sparse canonical correlation analysis (Zou et al., 2006), which essentially seeks to maximize the correlation of two projected sub spaces of two fields (Hotelling, 1936), while trying to minimize the noise effect in the projected variables by shrinking the associated canonical coefficients. After this introduction, this work is organized as follows. In the next section, we present the data set. In section 3, we analyze the spatiotemporal patterns of PDSI across Amazonia, and in section 4 we investigate the associated teleconnections with the global SST field. In section 5, we introduce the predictive model for PDSI and evaluate its forecast skill for different lead times based on a cross-validation procedure. Finally, in section 6, we offer a summary of the results found in the paper and present some conclusions.

2. Climate Data Set

2.1. PDSI Drought Indices

We use a variant of the Palmer Drought Severity Index (PDSI), originally proposed by Palmer (1965) to investigate drought conditions. It is one of the most widely used indices for drought characterization, although it has some limitations (Alley, 1984; Hayes et al., 1999; Heim, 2002; Keyantash & Dracup, 2002; Mishra & Singh, 2010), particularly its underestimation of runoff, delayed response in detecting some droughts, and more suitability for agricultural droughts. PDSI is obtained by relating temperature, rainfall, and soil-water holding capacity to estimate a local water balance and define local moisture stress conditions. We refer the reader

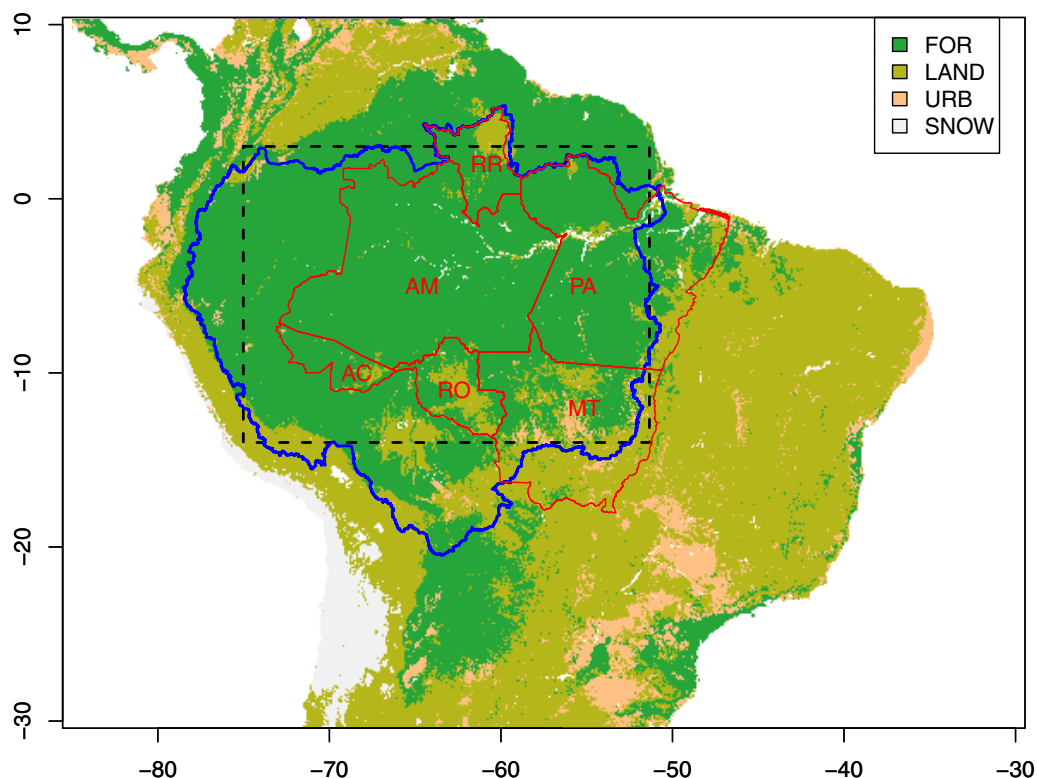


Figure 1. Grid domain (dashed line) defining the Amazonia region in this study to evaluate the PDSI gridded data. The Amazonia watershed is delimited by the blue line. The different land cover types follow the NASA land cover maps available at <http://eoimages.gsfc.nasa.gov/images/news/NasaNews/ReleaseImages/LCC/> (FOR = forests, LAN = lands, URB = Urban, and SNOW). The red line shows the political boundaries of the Brazilian federated states: Acre (AC), Amazonas (AM), Mato Grosso (MT), Pará (PA), Rondônia (RO), and Roraima (RR). The rainfall and temperature data are delimited by the grid domain (dashed line) and are available only within such states.

to Palmer (1965) and Wells et al. (2004) for the mathematical details on how to estimate the PDSI. Here we use the self-calibrated PDSI (Wells et al., 2004), which intends to replace empirical constants as used in the traditional PDSI calibration by dynamically estimated characteristics of the given location, which in turn makes the PDSI more comparable across space (Wells et al., 2004). The self-calibrated PDSI is usually referred in the literature as scPDSI. In this paper, for simplicity, we refer to it as PDSI. PDSI information is estimated by Dai et al. (2004) and consist of gridded, monthly values over Amazonia as defined in Figure 1 for the period 1980–2013. The data are available at <http://www.esrl.noaa.gov/psd/>.

2.2. Global Sea Surface Temperature

Here, we use monthly SST data from the ERA Interim global sea surface temperature archive for the period 1980–2013. The data set is interpolated onto a $2.5^\circ \times 2.5^\circ$ grid and is available at <http://apps.ecmwf.int/datasets/data/interim-full-moda/levtype=sfc/>.

2.3. Rainfall and Temperature

Monthly gridded ($0.25^\circ \times 0.25^\circ$) temperature and rainfall data for the period 1980–2013 are provided by Xavier et al. (2016). These data consist of interpolated daily rainfall and temperature observations from 3,625 rainfall gauges and 735 weather stations across Brazil available from different institutions (INMET, ANA, and DAEE). The interpolation schemes and validation procedures are described in Xavier et al. (2016). Note that these data are delimited by the Brazilian Amazonia boundary as defined by the states of Amazonas, Roraima, Pará, Rondônia, Mato Grosso, and Acre and shown in Figure 1. For each grid point, monthly anomalies of rainfall and temperature are obtained by removing from the observed value the long-term monthly mean for that grid point based on the 1980–2013 period. The data are also spatially constrained to the PDSI grid domain as shown in Figure 1.

3. Spatiotemporal Variability of Drought

The spatiotemporal dynamics of PDSI is first investigated using the principal component analysis (PCA, see Jolliffe, 2002 for more details) based on the centered and scaled PDSI data. The first three principal components (PCs) respond to 27%, 16%, and 11% of the data variability, respectively. Collectively, they explain 54% of the PDSI variance and, given that further components explain, individually, less than 6% of the data variance, we restrict our analysis to these first three PCs. Figure 2 shows the time series of these first three leading modes. In order to highlight the most extreme events, the PC values above the 90% and below the 10% empirical percentiles are colored in blue and red, respectively.

Based on the PC loadings (eigenvectors) associated with each mode (Figure 3), we can shed some light on the spatial patterns of droughts in Amazonia. The first mode (top plot of Figure 3) has positive loadings across the entire Amazonia, with slightly higher values in the central eastern region. This mode thus represents large droughts in Amazonia and if we consider the extreme events below the 10% percentile, these droughts took place in 1983, 1992, 1997–1998, and 2005 (top plot of Figure 2 and Figure 4). The second mode has a north-south (or meridional) kind of seesaw structure, with dry (wet) conditions south of about 6°S and wet (dry) conditions north of that latitude. Droughts with such pattern happened in several moments in the past (middle plot of Figure 2 and Figure 4), but it is worth mentioning the drought events in southern Amazonia that took place in 1988–1989, 2000, 2007, and 2010–2011. Some of these droughts have been identified in other studies (e.g., Marengo & Espinoza, 2016). In particular, referring to these two first modes, the 2005 and 2010 droughts are discussed in the next section. The third mode has a zonal type of dipole structure (bottom plot of Figure 3). In the following sections, we show that this third mode is

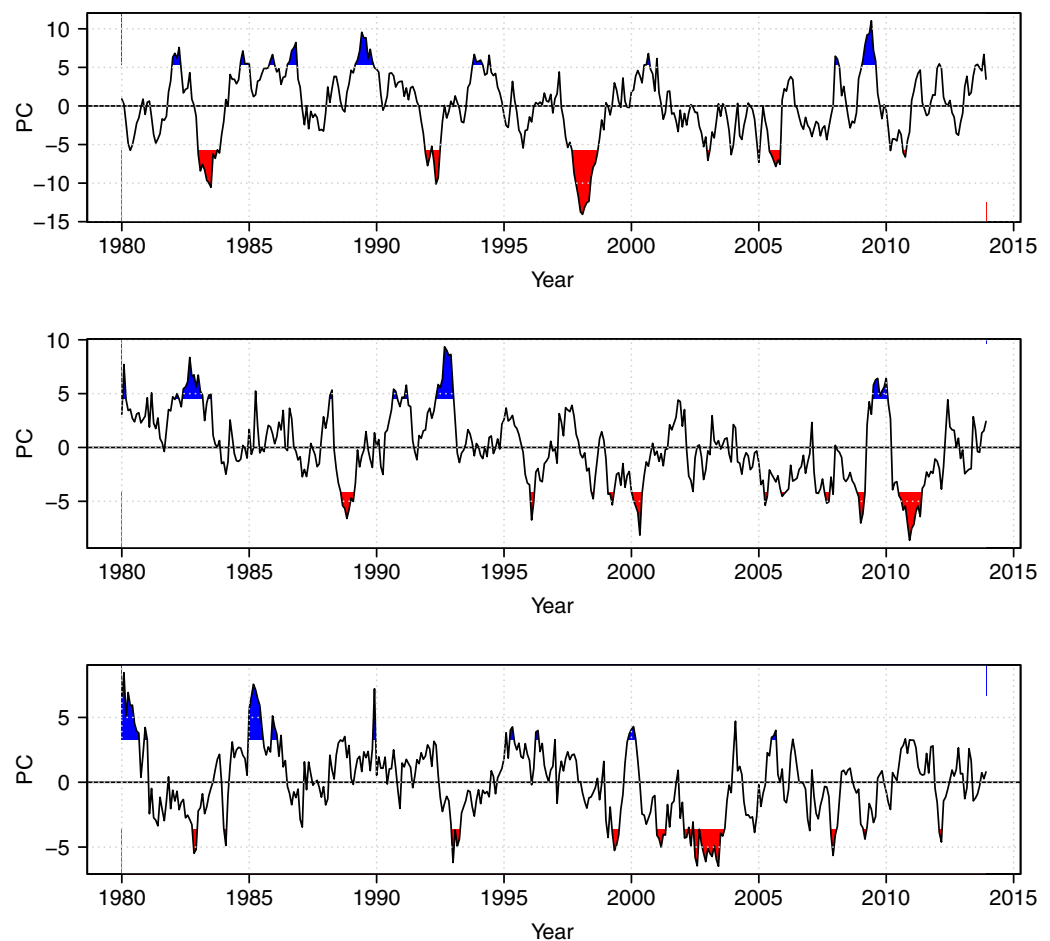


Figure 2. Time series of PDSI principal components (top three, from top to bottom plots). The blue and red colors identify points above and below the 90% and 10% percentiles, respectively.

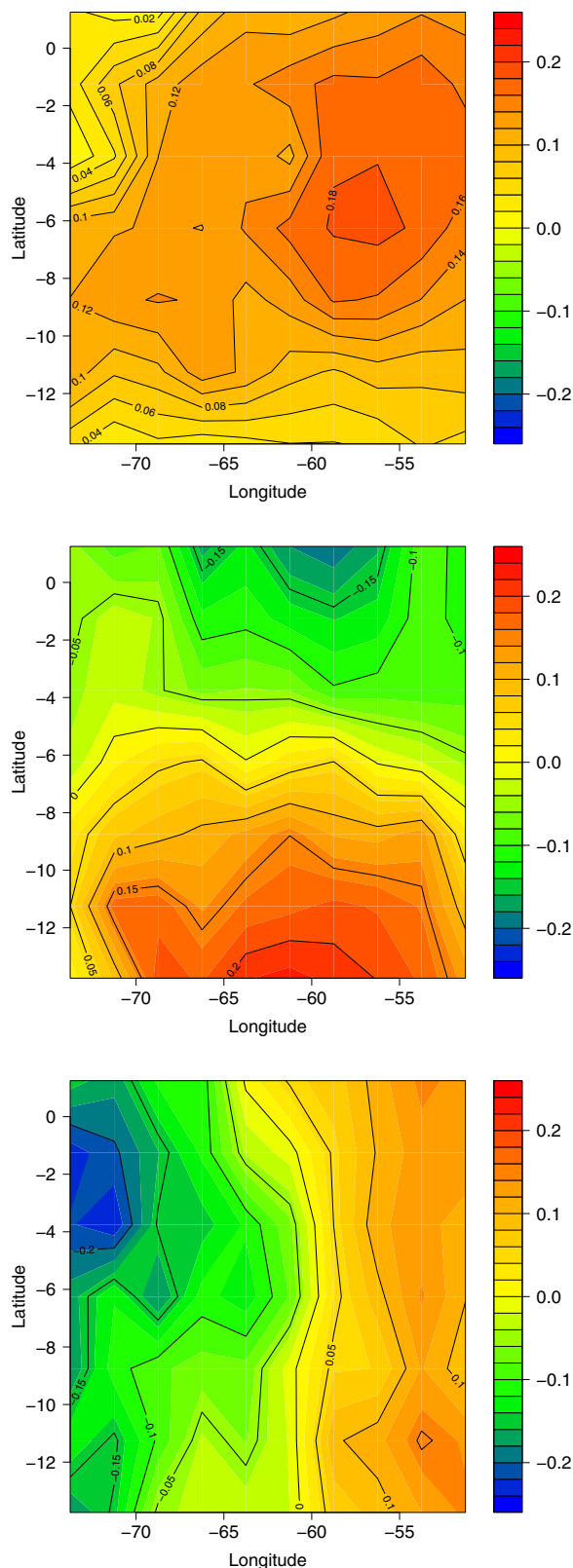


Figure 3. Top three (from top to bottom) loadings of PCA applied to the PDSI indices.

possibly associated with internal and local dynamics rather than large-scale climate processes. Hence, in order to limit the scope of this work, we do not analyze droughts associated with this mode as well as other drought patterns related to the remaining PCs. The specific years associated with wet events in the first and second PDSI modes (extreme events above the 90% percentile) are available in the supporting information.

3.1. Relationship Between PDSI, Rainfall, and Temperature

The average rainfall and temperature anomalies over the Brazilian Amazonia across the extreme events of the first three leading PDSI modes (Figures 2 and 4) are displayed in Figure 5. The first mode of PDSI is associated with low rainfall and high temperatures over the entire Amazonia (top plots in Figure 5). Negative anomalies in the second mode are associated with rainfall below average south of about 5°S, and slightly positive anomalies in the temperature over western and southeastern Amazonia and negative anomalies occurring in a small portion north of 5°S and south of about 10°S (middle plots in Figure 5). The third mode is associated with low rainfall in eastern Amazon and high temperatures in all but a small part in southern Amazonia. This information sheds light on the key drivers of drought, and highlights the role of temperature in Amazonia droughts. We provide in the supporting information the correlation maps between the first three leading PDSI modes and rainfall and temperature, and the conclusions are similar to those obtained for the composites (Figure 5). As shown, temperature plays a major role in droughts and deserves more attention, especially in light of the projected increase in future temperatures.

4. Drought Teleconnections With Global Sea Surface Temperature

The concurrent correlations of global sea surface temperature (SST) with the first three principal modes of PDSI are shown in Figure 6. The first mode has strong negative correlations in the ENSO region and in the extratropical South Pacific. This means that the positive anomalies in these regions are associated with negative anomalies in the first PDSI mode (i.e., drought conditions across the entire Amazonia). This pattern along with the negative correlations found in the tropical North Atlantic are related to major droughts in Amazonia (Jiménez-Muñoz et al., 2016; Marengo, 1992; Marengo & Espinoza, 2016; Yoon & Zeng, 2010; Zou et al., 2016). Other regions in the North Atlantic and over the entire Indian Ocean present statistically significant correlations but potential teleconnection mechanisms with Amazonia droughts are poorly understood and need further investigations.

The second mode (middle plot in Figure 6) displays a dipole kind of structure with the SST field across the Pacific basin, with positive correlations in the eastern part and negative correlations in the far western region. The tropical Atlantic shows negative correlations. These two patterns of correlations suggest that, based on the loadings of the second mode (Figure 3), droughts in southern (northern) Amazonia as shown in the middle plots of Figure 5 tend to be associated with negative (positive) or neutral anomalies in the SST across the ENSO region and positive (negative) anomalies in the tropical Atlantic.

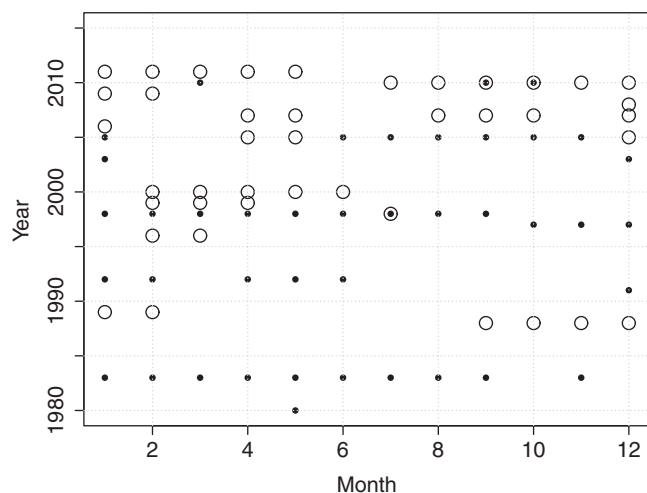


Figure 4. Periods of dry events (points below the 10% percentile) as identified for the first (black dots) and second (open circles) PDSI principal components.

Positive correlations are also found in the extratropical North Atlantic, but it is unclear the physical mechanisms responsible for such teleconnection.

A somehow similar conclusion was described by Zou et al. (2016), which suggested that droughts in northern Amazonia are influenced by El Niño events while dry periods in southern Amazonia are related to positive SST anomalies in the North Atlantic. They identified the 2005 drought as an event that hit southern Amazonia, while the 2010 drought was initially started in northern Amazonia and progressively hit southern Amazonia due to positive SST anomalies in the North Atlantic. Using the results obtained for the first two leading PDSI modes, we offer a slightly different but also complementary view of the 2005 and 2010 drought events. The 2005 drought started in southern Amazonia as identified by the negative anomalies in the second PDSI mode (Figure 4), and progressively hit the entire region (positive anomalies in the first PDSI mode, Figures 2 and 4). The drought event ended with dry conditions in southern Amazonia in the late 2005 and early 2006 (Figure 4). The 2010 drought event started in northern Amazonia in late 2009 as identified by positive

anomalies in the second PDSI mode (middle plot of Figure 2), then the dry conditions moved to southern Amazonia (Figures 2 and 4) and eventually hit the entire region in September–October 2010. These dry conditions had persisted in southern Amazonia until May 2011.

Finally, the third model displays (bottom plot of Figure 6) weaker correlations in several areas across the globe, with a more remarkable negative correlation region in the tropical Pacific centered around 150°W, suggesting that positive SST anomalies in this region tend to be associated with droughts in the eastern Amazonia, as shown in the bottom plots of Figure 5. However, based on the forecast results described in the next section, we believe that such large-scale teleconnections might play just a minor role on this third mode of variability.

5. A Drought Prediction Model for Amazonia

5.1. Technical Approach

Based on the field correlations (see Figure 6), our main goal here is to obtain SST predictors for the PDSI modes in order to provide forecasts at different lead times. Clearly, the maps show several locations that could be of just spurious correlations, which, although statistically significant, could be obtained just by chance or be associated with cross correlations within the SST field. The main objective is then to select just a few regions and grid points to obtain more robust SST predictors for the PDSI field.

Here, we propose using the canonical correlation analysis (CCA, see Hotelling, 1936 for the original idea) for linking SST to PDSI fields in order to obtain highly correlated response and predictor variables in a subspace spanned by such fields (see, e.g., Barnett & Preisendorfer, 1987; Barnston & Ropelewski, 1992). However, the associated standard CCA transformation is performed through full matrices (with nonzero entries) multiplication, which tends to carry a significant amount of noise and less robust estimators in the prediction space. This drawback is usually seen when cross-validated predictions (with out-of-sample data) have limited skills as compared with high correlations found during the estimation phase. Our idea is to obtain SST predictors based on just a few number of grid points in order to avoid excessive noise in the predictor space. In other words, we want to find sparse SST grid points in Figure 6 that will lead to skillfully predictors for the PDSI modes. Our approach will follow then the ideas of sparse loadings originally developed for principal component analysis (e.g., Zou et al., 2006) and further extended to canonical correlation analysis (e.g., Witten et al., 2009). These approaches are also referred as robust (or regularized) methods for principal/canonical correlation analysis (e.g., Candés et al., 2011; Dehon et al., 2000; Hardoon & Shawe-Taylor, 2011; Jolliffe et al., 2003; Shen & Huang, 2008; Wang & Huang, 2016) and have been developed focused mainly on machine learning problems and applications. Only recently such specific methods have been explored in hydroclimate applications (Ho et al., 2016).

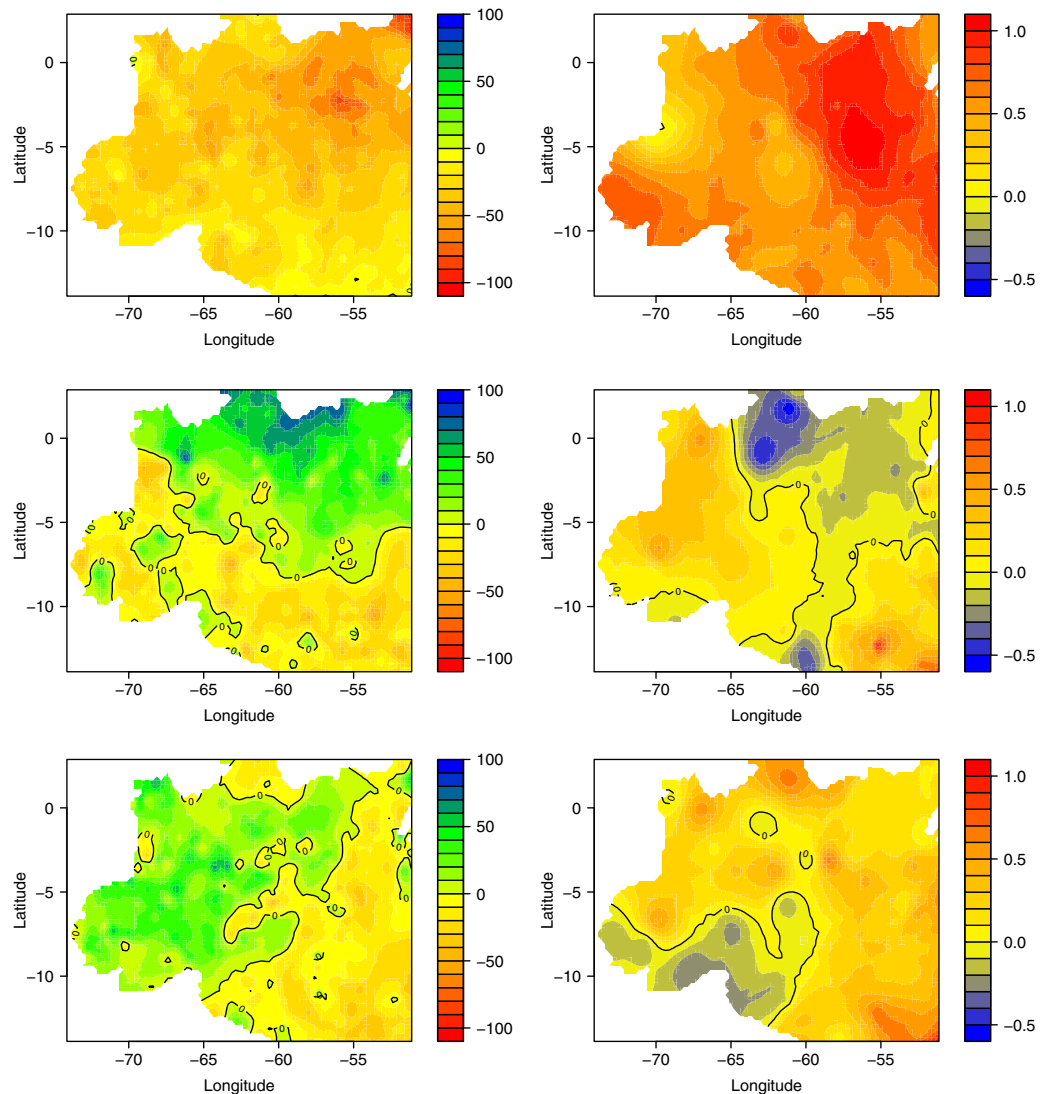


Figure 5. Composite analysis of anomalies in rainfall (left plots, scale in mm) and temperature (right plots, scale in °C) for the dry events of the top three (from top to bottom) PDSI principal components.

Let then \mathbf{Y} be the n by p matrix of the observed PDSI field over Amazonia (Figure 1), where $n = 408$ months (period 1980–2013) and $p = 70$ grid points. Let \mathbf{X} be the n by q matrix of the global SST field, where q is equal to 6,936 grids points. For simplicity, let us assume that \mathbf{Y} and \mathbf{X} are centered and scaled matrices. The basic idea from canonical correlation analysis originally proposed by Hotelling (1936) is to find two sets of basis vectors \mathbf{V} and \mathbf{U} such that the correlations of projections of \mathbf{Y} and \mathbf{X} onto these basis vectors ($\mathbf{Y}\mathbf{V}$ and $\mathbf{X}\mathbf{U}$) are maximized. $\mathbf{V} = [\mathbf{v}_1 \dots \mathbf{v}_l]$ and $\mathbf{U} = [\mathbf{u}_1 \dots \mathbf{u}_l]$ are the so-called canonical vectors (or weights, coefficients), where $\mathbf{v}_i, i = 1, \dots, l$, is a column vector with dimension p , \mathbf{u}_i is a column vector with dimension q and $l = \min(p, q)$. The projections $\mathbf{Y}\mathbf{v}_i$ and $\mathbf{X}\mathbf{u}_i$ onto the coefficients are called canonical variates and, while the (canonical) correlation between each pair $\mathbf{Y}\mathbf{v}_i$ and $\mathbf{X}\mathbf{u}_i$ is maximized and decreases from $i = 1$ to $i = l$, the cross canonical variates are uncorrelated, i.e., the correlation of $\mathbf{Y}\mathbf{v}_i$ and $\mathbf{X}\mathbf{u}_j$ is equal to zero for $i \neq j$.

The canonical coefficients \mathbf{V} and \mathbf{U} will likely have nonzero coefficients on all l variables, and this issue makes the method less robust to noise in the data and consequently with reduced cross-validated skill when one intends to predict \mathbf{Y} from \mathbf{X} . In order to obtain robust estimates for \mathbf{V} and \mathbf{U} that are less sensitive to noise, we will search for sparse vectors \mathbf{V} and \mathbf{U} such that a certain number of coefficients will be exactly equal to zero. This can be done by penalizing the canonical vectors in a procedure similar to the LASSO and RIDGE regressions (Hastie et al., 2001), in which a bound is introduced in the sum of the absolute or squared

norm value of the coefficients. Here we will adopt the methodology described in Witten et al. (2009) and will make use of the numerical codes available in the R language (Witten et al., 2013). The basic idea to obtain the sparse canonical coefficient is to solve the following optimization problem:

$$\max_{\mathbf{V}, \mathbf{U}} \mathbf{V}^T \mathbf{Y}^T \mathbf{X} \mathbf{U}, \text{ s.t. } \mathbf{V}^T \mathbf{Y}^T \mathbf{Y} \mathbf{V} \leq 1, \mathbf{U}^T \mathbf{X}^T \mathbf{X} \mathbf{U} \leq 1, P_1(\mathbf{V}) \leq c_1, P_2(\mathbf{U}) \leq c_2, \quad (1)$$

where P_1 and P_2 are convex penalty functions and here they will be the Lasso penalty:

$$P_1(\mathbf{V}) = \sum_{i=1}^l \sum_{j=1}^p |v_{ji}|$$

$$P_2(\mathbf{U}) = \sum_{i=1}^l \sum_{j=1}^q |u_{ji}|.$$

Witten et al. (2009) shows that feasible solutions are only obtained when the bounds c_1 and c_2 are in the range:

$$1 \leq c_1 \leq \sqrt{p} \text{ and } 1 \leq c_2 \leq \sqrt{q}.$$

Note that large values of c_1 and c_2 correspond to less penalization, i.e., the number of nonzero coefficients (or sparsity) is reduced. We refer the reader to Witten et al. (2009) for a thoughtful description and discussion of the method, including mathematical properties, algorithm implementation, selection of c_1 and c_2 , connection to other methods, and applications to genomic data.

5.2. Cross-Validation

One possibility to test the sparse CCA model as presented here would be to let \mathbf{Y} be the entire PDSI field and \mathbf{X} the global SST field. However, the resulting canonical variates could lead to difficult interpretation of the canonical coefficients and the physical reasons associated with such teleconnections. Hence, with the primary goal to illustrate the proposed method, we pick \mathbf{Y} as the first three PCs from PDSI (see Figure 2) and, for an initial evaluation of the sparse CCA method, we choose the value of c_1 so that the canonical coefficients \mathbf{v}_i are orthogonal unit vectors and \mathbf{V} is an orthogonal matrix of permutation of coordinate axes, so that the transformation $\mathbf{Y}\mathbf{V}$ leads to exactly \mathbf{Y} , but possibly with a permutation of the main modes. The different values of c_2 will be chosen following the formula:

$$c_2 = \min(1, \text{penalty}_x \cdot \sqrt{q}), \quad (2)$$

where penalty_x is between 0 and 1 and q is equal to 6,936 (number of grids points) in our case.

The results of the canonical coefficients \mathbf{u}_1 and \mathbf{u}_2 associated, respectively, with the first and second PDSI leading modes are shown in Figure 7 for three different values of penalty_x . The coefficients \mathbf{u}_1 (left plots of Figure 7) have a spatial pattern that somewhat resembles the correlation of the first PDSI mode with the SST field (top plot of Figure 6) but with several grid points with values equal to zero. As penalty_x increases and less penalization occurs, the number of nonzero \mathbf{u}_1 coefficients increases and the spatial pattern becomes closer to that of the correlation field displayed in the top plot of Figure 6. But interestingly, there is no significant changes in the correlation coefficient of the first PDSI mode with the canonical variate $\mathbf{X}\mathbf{u}_1$, being all equal to 0.75. Similar conclusions arise for the canonical coefficients

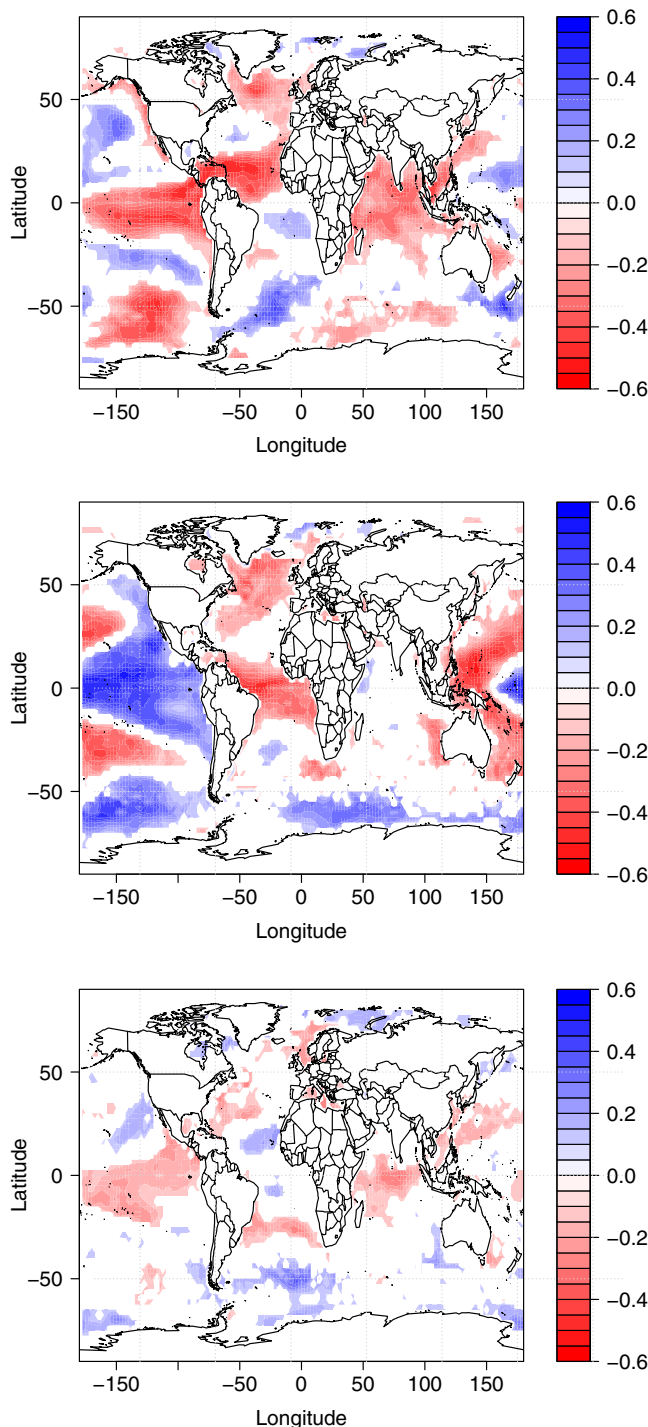


Figure 6. Correlation of monthly sea surface temperature (SST) with the first top three (from top to bottom) principal components of the PDSI indices. Only statistically significant correlations at the 5% significance level are shown.

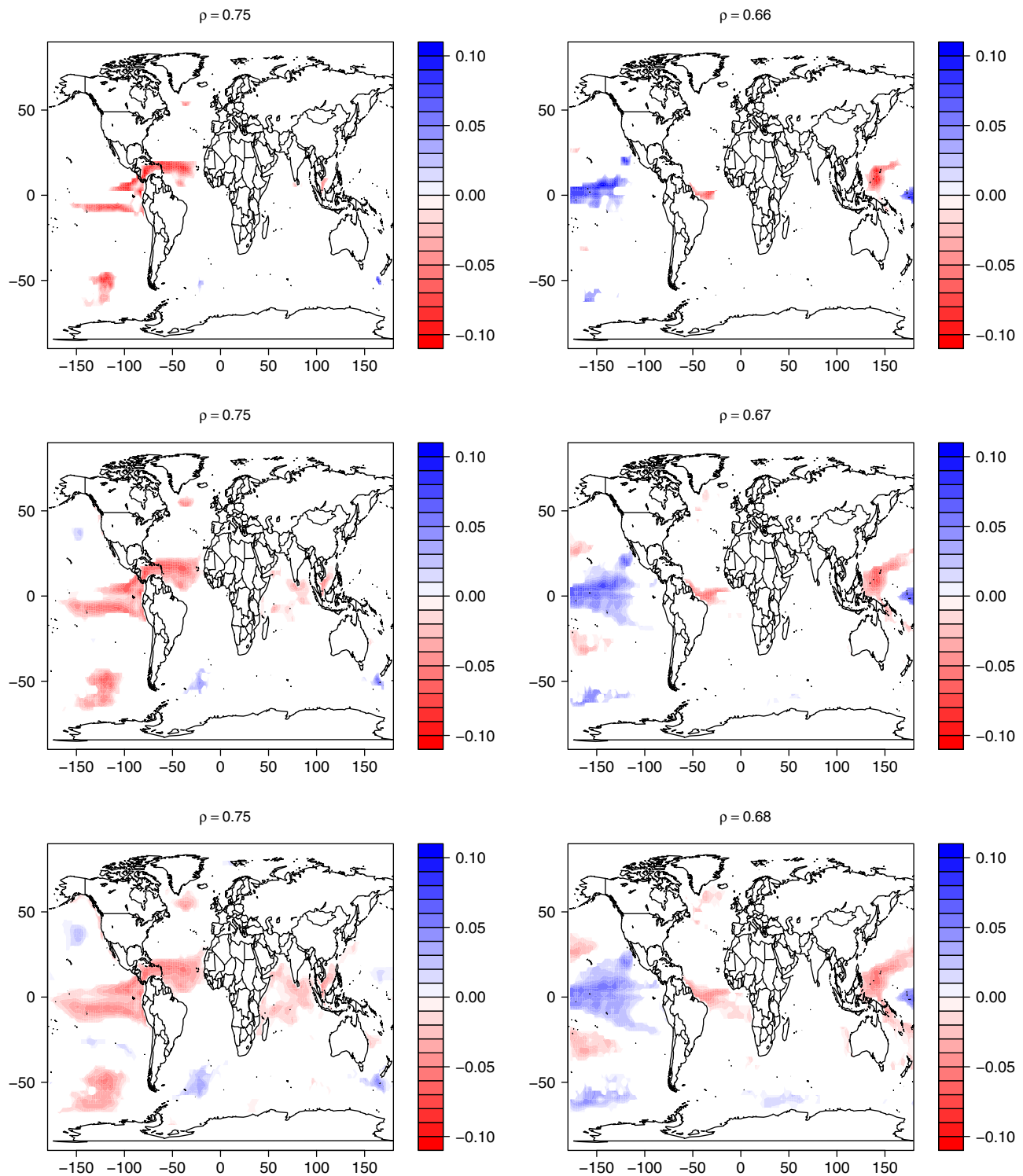


Figure 7. Canonical coefficients \mathbf{u}_1 (left plots) and \mathbf{u}_2 (right plots) for values of penalty λ equal to 0.2 (top), 0.3 (middle), and 0.4 (bottom). The Pearson correlation coefficient ρ of the canonical variates $\mathbf{u}_1 X$ and $\mathbf{u}_2 X$ with the first and second PDSI modes, respectively, is indicated in the top of each plot.

\mathbf{u}_2 , with a slightly increase in the correlation of the second PDSI mode and $X\mathbf{u}_2$ as penalty λ increases. The results for the canonical coefficients \mathbf{u}_3 lead to similar conclusions and are not shown here.

In order to evaluate the proposed methodology for PDSI forecasts at different lead times, we proposed a multivariate linear model for the PDSI data and perform a cross-validation procedure. For the sake of

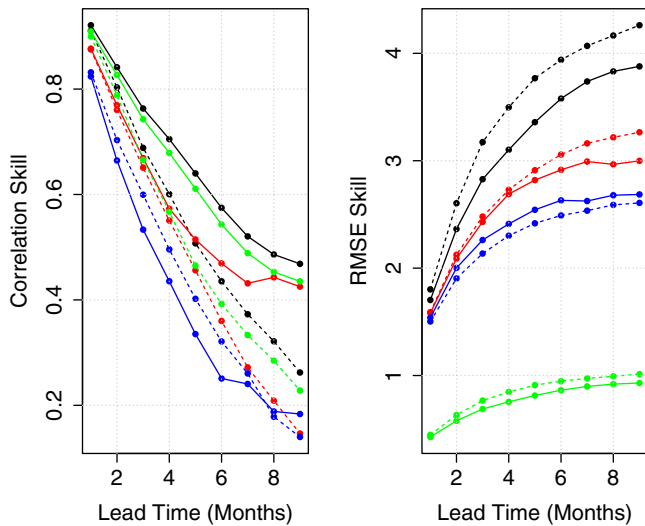


Figure 8. Cross-validated (left plot) correlation and (right plot) RMSE (skills for full model (solid line) and baseline (autoregressive) model (dashed line) for the first (black dots), second (red dots), and third (blue dots) principal components of PDSI. The green line shows the skill for the spatial average PDSI.

equation (3) are withdrawn, while the data for the remaining years are used to estimate the model parameters α , θ , β , and \mathbf{U} . The estimated model is then used to predict \mathbf{Y} for the first year. The entire procedure is repeated until forecasts are made for all years of the record. A baseline model based solely on the autoregressive and intercept terms (i.e., $\beta=0$ in equation (3)) is used to evaluate the real gain in model skill by including external information based on the SST data and the sparse CCA model. Predictions for the entire PDSI field and the spatial average PDSI are also evaluated by taking the predicted PDSI first three modes \mathbf{Y} and transforming them back to the original space using the estimated PCA loadings as described in section 3.

The cross-validated correlation and root mean square error (RMSE) skills obtained for \mathbf{Y} at different lead times are displayed in Figure 8. As expected, as the lead time increases the correlation skill drops while the RMSE increases for all three PDSI modes and for the two models tested.

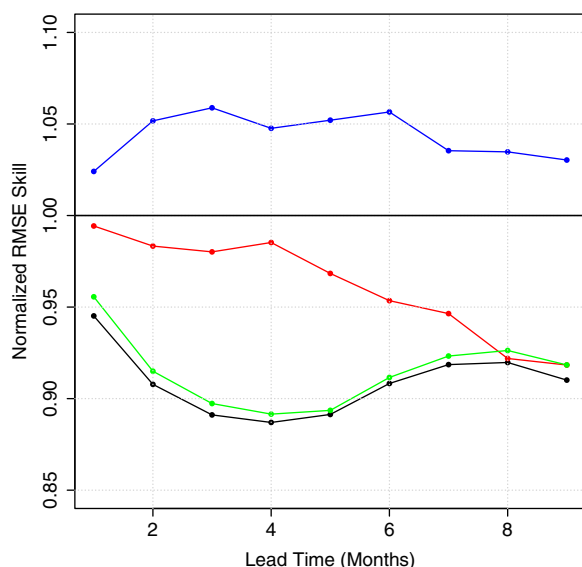


Figure 9. Cross-validated RMSE skill for the full model normalized with respect to the RMSE skill for the baseline model for the first (black dots), second (red dots), and third (blue dots) principal components of PDSI. The green line shows the skill for the spatial average PDSI.

simplicity, we will keep for all lead times tested a fixed value of 0.3 for penalty λ , which we believe, based on Figure 7, will yield to a good balance between model complexity and model skill. For c_1 and the correspondent penalty (similar to equation (2)), we observed that penalty = 0.3 yields to the desired orthogonal matrix \mathbf{V} for concurrent \mathbf{Y} and \mathbf{X} . Therefore, for the purpose of this work to illustrate the model skill, we keep this value also fixed for different lead times, which might not lead necessarily to an orthogonal matrix \mathbf{V} for all lead times.

Let us assume that \mathbf{Y}_t is the vector containing the observed values of the first three PDSI modes for a given month t . A multivariate linear model relating them to past values and to the canonical variates derived from the SST field is proposed here:

$$\mathbf{Y}_t = \alpha + \theta \cdot \mathbf{Y}_{t-\tau} + \beta \cdot (\mathbf{X}_{t-\tau} \mathbf{U}) + \epsilon, \quad (3)$$

where α is a 3-D vector of intercept terms, θ is a 3×3 matrix of autoregressive coefficients, β is a 3×3 matrix of exogenous coefficients, ϵ follows a zero-mean multivariate normal distribution with covariance Σ , and $\tau=1, \dots, 9$ is the lead time of the forecast in months.

The skill of the proposed model is evaluated through an 1 year leave-out cross-validation procedure, in which the observations of \mathbf{Y} for the first year of the record and the associated predictors as depicted in

equation (3) are withdrawn, while the data for the remaining years are used to estimate the model parameters α , θ , β , and \mathbf{U} . The estimated model is then used to predict \mathbf{Y} for the first year. The entire procedure is repeated until forecasts are made for all years of the record. A baseline model based solely on the autoregressive and intercept terms (i.e., $\beta=0$ in equation (3)) is used to evaluate the real gain in model skill by including external information based on the SST data and the sparse CCA model. Predictions for the entire PDSI field and the spatial average PDSI are also evaluated by taking the predicted PDSI first three modes \mathbf{Y} and transforming them back to the original space using the estimated PCA loadings as described in section 3.

The cross-validated correlation and root mean square error (RMSE) skills obtained for \mathbf{Y} at different lead times are displayed in Figure 8. As expected, as the lead time increases the correlation skill drops while the RMSE increases for all three PDSI modes and for the two models tested. For a given lead time, the model skills also reduced from the first to the third modes. For the first two PDSI modes, the full model presents higher skills when compared with the baseline one, which shows higher correlations and lower RMSEs only for the third PDSI mode. This suggests that this mode of variability could be related to other processes rather than the SST field. The correlation skill for the spatial average PDSI is close to the one obtained for the first PDSI mode, but the RMSE skill is relatively low compared with those obtained for all PDSI modes. And again, the full model has a better skill when compared with the baseline one. The relative gains of the full model are better seen in Figure 9, which shows its RMSE skill normalized with respect to the RMSE skill obtained from the baseline model. The biggest improvement is observed at 4 month lead time predictions, for the first PDSI mode and for the spatial average PDSI. The relative gain for the second PDSI mode increments as the lead time increases, while for the third mode the full model is not able to add any extra skill to the one already obtained by the baseline model.

Pointwise cross-validated correlations between the observed and predicted PDSI fields are shown in Figure 10 for 2 and 6 month lead times. For both full and autoregressive models, the highest correlations are found in the northeastern part of Amazonia and can be as

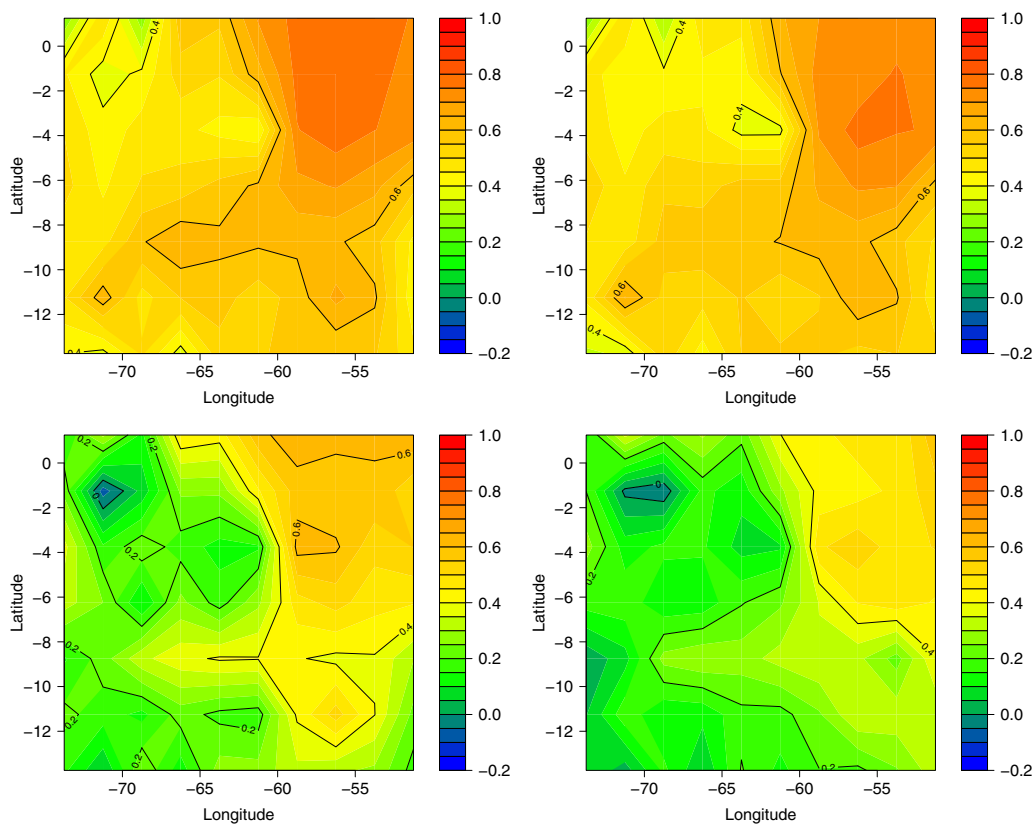


Figure 10. Correlation skill for cross-validated predictions at 2 month (top plots) and 6 month (bottom plots) lead times obtained from the full (left plots) and autoregressive (right plots) models.

high as 0.7 for 2 month lead time and 0.6 for 6 month lead time. The difference between the skill of the two models is more evident for 6 month lead time.

The ability of the full model to predict the spatial average PDSI is also shown in Figure 11 for 2 and 6 month lead times. Visually, the model is able to predict all the largest drought events with 2 months of antecedence and some of the events with 6 month lead time, although the magnitude of such events is underestimated in this case. Around 7% of the observed values are outside the 95% prediction intervals.

Finally, in Figure 12, we show the observed PDSI field averaged over the extreme events identified for the first and second modes (Figures 4 and 5) and the corresponding predicted PDSI field averaged over the same events for lead times of 1 and 6 months. In the cases of drought events associated with the first PDSI mode (left plots of Figure 12), which correspond to droughts across the entire Amazonia (Figure 5), the prediction model is able to highlight the driest region in northeastern Amazonia at 2 month lead time, while at 6 month lead time the predictions resemble the observed spatial pattern but underestimate the drought magnitude. Droughts related to the second PDSI mode, which features a dipole kind of pattern, with wet (dry) conditions in the north (south) and vice-versa, are also satisfactorily predicted by the proposed model at both lead times (right plots of Figure 12), particularly in terms of spatial variability. The drought magnitude is slightly underestimated in both cases.

The gain of the proposed model over climatology (i.e., benchmark model) is finally evaluated through the Brier skill score (Wilks, 2006). For each grid point of PDSI, we consider the 20% most extreme droughts, i.e., events with a cumulative probability of 20% (climatology). This probability is used to estimate the Brier score for the reference model. The Brier score for the proposed model is obtained through the cross-validated predictions by estimating, for each month and grid point, the cumulative probability associated with the respective 20% PDSI quantile obtained from the observed data. The results for 1, 2, 4, and 6 month lead times are shown in Figure 13. The skills are higher in the northeastern region and resembles the region of highest

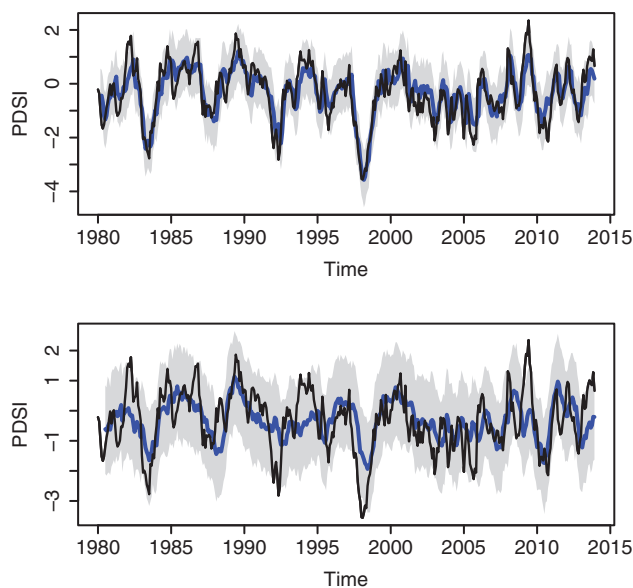


Figure 11. Observed PDSI (black line) and cross-validated predictions (blue line) at 2 month (top plot) and 6 month (bottom plot) lead times. The grey shaded region shows the 95% prediction interval.

correlations (Figure 10) and largest loadings of the first PC mode of PDSI (Figure 3), which has also the best predictive skill (Figure 8).

5.3. Comparison With SPI

In order to evaluate the drought patterns and predictions using a rainfall only based drought indicator, we performed the same analysis using the standard precipitation index (SPI, McKee et al., 1993) at 1 and 3 months derived from the observed precipitation. The results appear in supporting information. In general, we found that PDSI is well correlated with SPI at 3 months but not much with SPI at 1 month. This might be due to the persistence of temperature and its correlation with rainfall. The composite analysis of rainfall and temperature anomalies and the correlation with SST is similar to those obtained for PDSI (Figures 5 and 6). The prediction skill is also comparable to that obtained for PDSI (Figures 8 and 9) and therefore the entire results and conclusions are overall very much alike if we use SPI instead of PDSI.

6. Summary and Conclusions

In this paper, we use the self-calibrated Drought Palmer Severity Index (PDSI) to analyze droughts over Amazonia in terms of spatiotemporal variability and teleconnections. We then propose a seasonal forecast model for PDSI based on predictors obtained from the global SST field using the sparse canonical correlation analysis method.

The spatial dynamics of PDSI reveals three major distinguished patterns of droughts: the main one which covers the entire region with a center in the central east area, a second seesaw pattern with dry conditions in the south and wet conditions in the north and vice-versa, and finally a third east-west seesaw mode. These three modes together explain about 54% of the drought variability in Amazonia, and are related to the major droughts in the region identified in the literature based on rainfall and streamflow data. When taking only the extreme events (below the 10% percentile) associated with these three main modes, we found the major drought periods: 1983, 1988–1989, 1992, 1997–1998, 2000, 2005, 2006, and 2010–2011. This is a different perspective on droughts in Amazonia, since the PDSI takes into account not only rainfall but also anomalies in the temperature and soil moisture capacity. We also find that droughts in Amazonia as per the PDSI are characterized by compound events of low rainfall and high temperatures.

Concurrent correlations of the PDSI main modes with the global SST field show that the major droughts that cover the entire region are highly correlated with the SST in the ENSO region and in the tropical north Atlantic, agreeing with previous studies (Jiménez-Muñoz et al., 2016; Marengo, 1992; Marengo & Espinoza 2016). Statistically significant correlations are also found with the SST field over the Indian Ocean and in the midlatitudes of the North Atlantic and South Pacific. It is not clear whether these correlations are just a by-product of the intrinsic spatial variability of the SST global field or are associated with teleconnection mechanisms through the atmosphere. The second mode of PDSI related to the north-south seesaw drought pattern is found to be positively correlated with the SST almost entirely across the tropical Pacific basin, and negatively correlated with the SST over the tropical Atlantic. This finding suggests that droughts in southern Amazonia can be associated with La Niña or neutral conditions provided that there are positive anomalies in the tropical Atlantic SST. Hence, it helps explaining some major droughts in Amazonia not related to El-Niño conditions, such that in 2004–2005, and provides new insights into the association of SST with the spatial dynamics of droughts in Amazonia. The third PDSI mode associated with the east-west seesaw pattern has weak but yet statistically significant correlations with SST in several parts of the globe, including the ENSO region, but we believe, based on the forecast model results, that this third leading mode and its spatial pattern most likely result from local/regional processes and internal dynamics of Amazonia.

Using the sparse CCA method, we are able to find, among the SST grid points statistically correlated with the PDSI modes, only a small number of nodes to compose a predictor index for Amazonia droughts. When

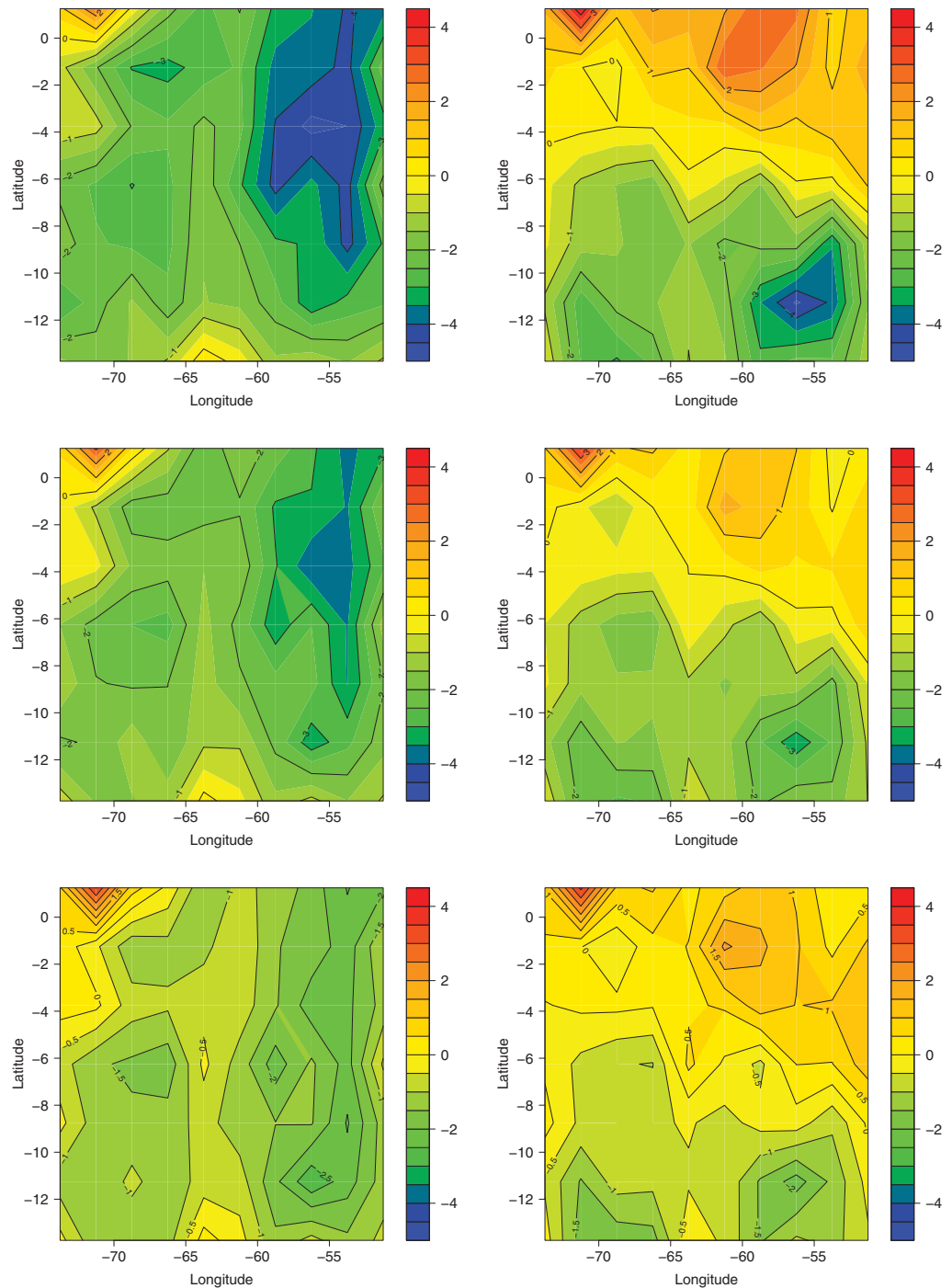


Figure 12. Composite analysis of extreme droughts in the first (left plots) and second (right plots) PC from PDSI. The top plots show observed values while the predictions at 2 month and 6 month lead times are shown in the middle and bottom plots, respectively.

compared to an autoregressive baseline model, the inclusion of these exogenous predictors in a forecast model for PDSI leads to an increase in the forecast skill. In a cross-validation procedure, we obtain correlations between the observed and predicted spatially average PDSI up to 0.60 and 0.45 for lead times of 5 and 9 months, respectively. These values are in the same order of magnitude or even above correlations obtained from more complex models developed for drought and El Niño predictions (e.g., Barnston et al., 2012; Thober et al., 2015). The proposed model is also able to satisfactorily predict the spatial variability of

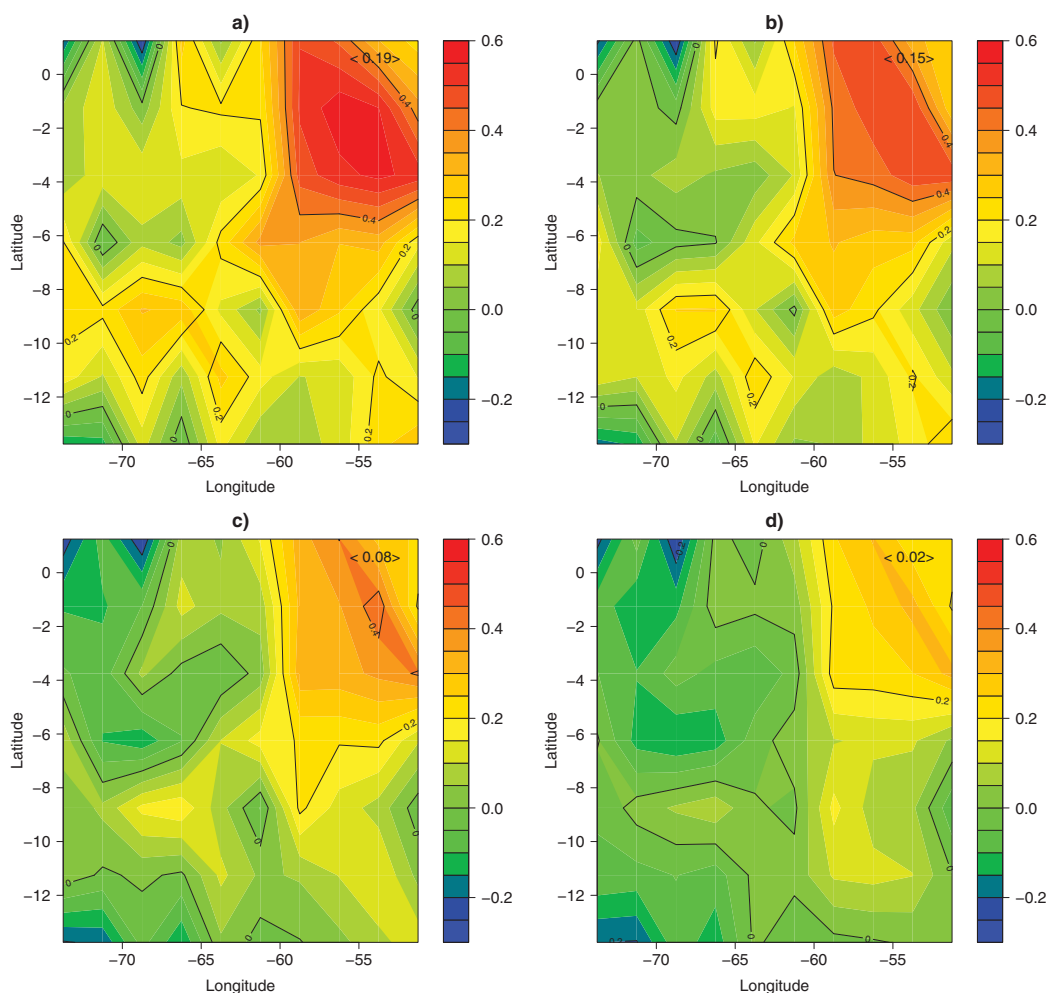


Figure 13. Brier skill score for (a) 1 month, (b) 2 month, (c) 4 month, and (d) 6 month lead times. The spatial average of the corresponding score is shown in the top right corner of the plots.

PDSI, although the magnitude is generally underestimated. The gain over climatology, as evaluated by the Brier skill score, is on average 20% for 1 month and 2% for 6 month lead times, although values up to 58% are obtained in the northeastern region of Amazonia.

To the best of our knowledge, the Amazonia drought forecast model proposed here is the first in the literature to provide, based on cross-validation results, satisfactory skills for lead times beyond 4 months. Furthermore, the approach offers the potential to capture the spatial variability of Amazonia droughts. The model is relatively easy to implement and can be used in operational mode as a tool for risk assessment in Amazonia, integrating Brazilian initiatives of drought monitor systems (e.g., the INPE drought monitor at <http://clima1.cptec.inpe.br/spi/pt> and the FUNCEME drought monitor at <http://msne.funceme.br/>). The inference process of the proposed model considered the penalization parameter of the sparse CCA constant across the different lead times, and this is one aspect that can be improved in future studies. Moreover, other climate variables such as sea level pressure and geopotential height fields can be added to the model to potentially improve the forecast skills, particularly at short lead times, and with time varying coefficients. The PDSI shows high correlations with rainfall across Amazonia but the possibility to use other drought indices (e.g., Hao & Singh 2015; Mishra & Singh, 2010; Mu et al., 2013; Rajsekhar et al., 2015; Shukla & Wood 2008) to better represent dry conditions in the region could be also explored in future studies. Finally, we believe the results obtained and the model proposed in this work are a step forward to improve the management of drought impacts in Amazonia, including the local and regional societal impacts, water resources, and the ecosystem response.

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Erratum

The originally-published version of this article interchanged Figure 11, Figure 12, and Figure 13. The error has been corrected, and this may be considered the official version of record.