

1 **Linking Statistical and Hydrodynamic Modeling for Compound** 2 **Flood Hazard Assessment in Tidal Channels and Estuaries**

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11 **Keywords:**

12 Coastal Flood Hazards, Bivariate Statistical Analysis, Hydrodynamic Modeling,
13 Compound Flooding

14 **Highlights**

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16
17 - Compound flood hazard levels computed for tidal channels or estuaries.
18 - Bivariate statistical and hydrodynamic modeling is linked.
19 - Spatially distributed flood hazard levels computed for return period T .
20 - Method accounts for statistical and physical compounding effects.

21 **Abstract**

22 A method to link bivariate statistical analysis and hydrodynamic modeling for
23 flood hazard estimation in tidal channels and estuaries is presented and discussed
24 for the general case where flood hazards are linked to upstream riverine discharge
25 Q and downstream ocean level, H . Using a bivariate approach, there are many
26 possible combinations of Q and H that jointly reflect a specific return period, T ,
27 raising questions about the best choice as boundary forcing in a hydrodynamic
28 model. We show, first of all, how possible Q and H values depend on whether the
29 definition of T corresponds to the probability of exceedance of “ H OR Q ” or “ H AND
30 Q ”. We also show that flood hazards defined by “OR” return periods are more
31 conservative than “AND” return periods. Finally, we introduce a new composite
32 water surface profile to represent the spatially distributed hazard for return period
33 T . The composite profile synthesizes hydrodynamic model results from the “AND”
34 hazard scenario and two scenarios based on traditional univariate analysis, a
35 “Marginal Q ” scenario and a “Marginal H ” scenario.

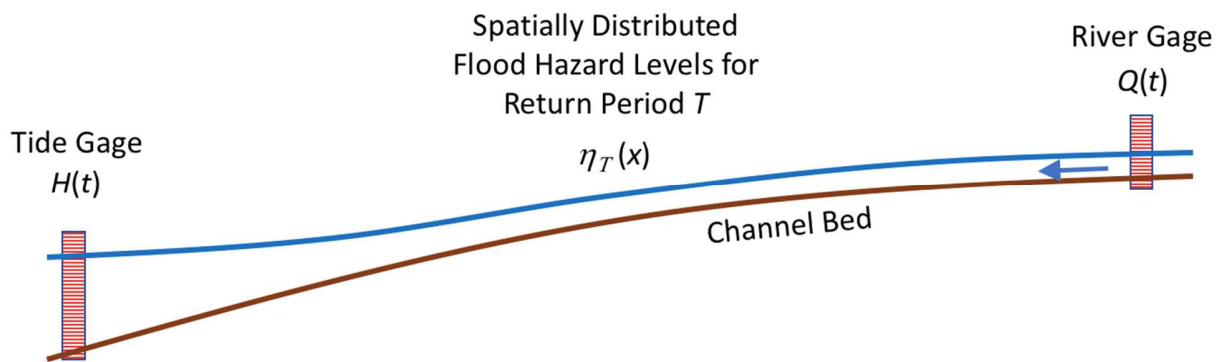
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37 **1. Introduction:**

38 Flood risk is increasing in coastal cities around the world due to several
39 factors including population growth, economic development, sea level rise,
40 subsidence, land use changes and intensification of rainfall (Hallegatte et al., 2013;
41 Hanson et al., 2011). By the year 2100, between 0.2–4.6% of global population and
42 0.3–9.3% of global gross domestic product may be exposed to coastal flooding if no
43 adaptation occurs (Hinkel et al., 2014).

44 Management of flood risk relies on statistical and hydrodynamic modeling
45 to delineate populations and assets exposed to flooding, anticipate and monetize the
46 consequences of flooding, and develop cost effective and socially robust
47 interventions including infrastructure projects, insurance programs, land use and
48 building code policy changes and emergency preparedness and response measures
49 (Sayers et al., 2013, Luke et al. 2018). To address risks, statistical and hydrodynamic
50 modeling is linked to delineate spatial fields of the intensity of flooding (e.g., depth
51 and velocity) for a set of exceedance probabilities, information which is
52 subsequently used to estimate average annual losses based on exposed assets and
53 their vulnerability to damage (Scawthorn et al., 2006). The linking of statistical and
54 hydrodynamic modeling is straightforward when addressing a single hazard such as
55 river discharge, Q . Flood risk is modeled by first performing univariate frequency
56 analysis of annual maximum discharge to estimate extreme values \hat{Q}_T for yearly
57 return periods T (e.g., \hat{Q}_{100} for 100 year return period discharge). Here, the hat
58 notation indicates annual maximum discharge and the subscript refers to the return
59 period. Second, hydrodynamic modeling is performed with \hat{Q}_T as a boundary
60 condition to characterize spatial fields of water surface elevation at each return
61 period, $\eta_T(x)$, where x represents distance along the river (FEMA, 2018). However,
62 coastal hazard assessment must account for interaction of river flooding, intense
63 rainfall, storm surge and waves and the likelihood of a coincidence in extreme and
64 non-extreme levels of these hazards which is also known as compounding effects

65 (Gallien et al., 2018; Moftakhari et al., 2017). One of the most important
 66 compounding effects is the interaction of river discharge and the downstream ocean
 67 level, H , in tidal channels and estuaries. Of the world's 32 largest cities, 22 are
 68 located on estuaries (Ross, 1995), at which the interactions between Q and H play a
 69 major role in flood risk estimation (Ward et al., 2018). In the U.S. alone, 140 million
 70 people (~50% of total population) live on the coast in close proximity to an estuary
 71 (Kennish, 2004).



72
 73 **Figure 1:** The spatial distribution of the T -year return period extreme water level, $\eta_T(x)$, in an estuary or tidal
 74 channel (blue line) depends on upstream river discharge $Q(t)$, downstream ocean water levels $H(t)$, and the
 75 system geometry and resistance to flow. This paper shows how bivariate statistical analysis and
 76 hydrodynamic modeling can be linked to compute $\eta_T(x)$.

77
 78 Figure 1 illustrates the estuarine flood hazard problem: the objective is to
 79 estimate spatially distributed extreme water levels, $\eta_T(x)$, for return period T in a
 80 tidal channel or estuary. We assume knowledge of the system geometry (e.g., bed
 81 elevation, channel width and shape) and resistance to flow (e.g., Manning resistance
 82 coefficient). We also assume that gauges provide time series records of boundary
 83 conditions: river discharge measurements representative of what enters the reach,
 84 $Q(t)$, and water level measurements, $H(t)$, representative of the downstream end of
 85 the reach. Hence, the key question becomes: how can statistical and hydrodynamic
 86 modeling be linked for the estuarine setting involving two gage records
 87 characterizing two different aspects of hydrodynamic extremes? Put another way,
 88 can the existing paradigm of univariate flood hazard modeling described above for

89 rivers be extended to account for a second gage in a relatively simple and
90 straightforward way? We write this mathematically as follows,

$$91 \quad \eta_T(x) = f(Q(t), H(t), p(x)) \quad [1]$$

92 where $p(x)$ represents the parameters describing the channel geometry and
93 resistance properties. Note that Eq. [1] can be extended for two-dimensional flood
94 hazard levels by interpreting x as two-dimensional vector representing geographical
95 coordinates.

96 In the absence of robust models for extreme water levels in estuaries, overly
97 simplistic bathtub models have received widespread use for estimating coastal
98 flooding hazards and the resulting human exposure at regional (Torresan et al.,
99 2012) and national/international levels (Dasgupta et al., 2011; Hinkel et al., 2010).
100 Bathtub models simply take an estimated extreme water level and extrapolate it
101 inland to estimate population and assets exposed to flooding, which neglects the
102 potential for flood stage to change with distance inland as a consequence of riverine
103 forcing and/or tidal damping/amplification (Lanzoni and Seminara, 1998). This
104 points to the potential for underestimation of flood consequences. On the other
105 hand, bathtub models may also overestimate flood consequences by failing to
106 account for flood defenses and the role of friction, inertia and storage in flooding
107 dynamics (Gallien et al., 2014; Sanders, 2017). With the integration of statistical and
108 hydrodynamic models to estimate extreme water levels, more robust estimates of
109 extreme water levels become possible as well as mechanistic routing of flood water
110 into adjacent urban areas to estimate flood impacts (Gallien et al., 2014, 2011).

111 Existing methods for flood hazard assessment (solving Eq. 1) in tidal
112 channels are limited (Hoitink and Jay, 2016). In particular, for the case where \hat{Q}_T
113 and \hat{H}_T are statistically independent, FEMA (2015) recommends the following
114 procedure to estimate the hazard for return period T : (1) univariate analysis of
115 annual maximum river discharge to estimate \hat{Q}_T and univariate analysis of annual
116 maximum total water level to estimate \hat{H}_T , (2) a pair of hydrodynamic model

117 simulations with one forced by \hat{Q}_T and a non-extreme H value (usually chosen as
118 mean higher high water) and the other forced by \hat{H}_T and a non-extreme Q value, and
119 (3) synthesis of the two hydrodynamic model simulations based on the pointwise
120 maximum water level across the two simulations. An appealing aspect of the FEMA
121 approach is that only two hydrodynamic simulations are required for each return
122 period, which is important because resources for flood mapping are limited (Burby,
123 2001) and because hydrodynamic flood simulation is computationally demanding
124 (i.e., many hours for one simulation) especially for urban areas where fine
125 resolution grids are needed to accurately depict flooding, e.g., Gallien et al. (2014,
126 2011). Hence, the FEMA (2015) approach is aligned with needs for simple and
127 efficient assessment approaches. Nevertheless, there are significant limitations. For
128 example, univariate statistical analysis is not appropriate when \hat{Q}_T and \hat{H}_T exhibit
129 statistical dependence also known as compound risks (Leonard et al., 2014;
130 Moftakhari et al., 2017; Zscheischler et al., 2018). Additionally, even when \hat{Q}_T and
131 \hat{H}_T are independent, extreme water levels may occur over the length of the tidal
132 reach due to the interaction of non-extreme boundary forcing values. FEMA (2018,
133 2015) does not presently offer guidance to address this situation. Broadly, the FEMA
134 (2015) guidance is recommends multi-hazard assessment based on the
135 predominant hazards, yet limitations of this approach are increasingly being
136 recognized (Hillier et al. 2015).

137 The aforementioned challenges of linking statistical and hydrodynamic
138 modeling can only be partly overcome with improved access to, and reduced costs
139 of, high performance computing systems that map flood hazards through Monte
140 Carlo simulation. That is, Monte Carlo simulation can be applied to depict thousands
141 of scenarios based on different combinations of \hat{Q}_T and \hat{H}_T , and depict flood hazards
142 based on the frequency of the pointwise exceedance of a water level thresholds
143 (Purvis et al., 2008). However, bivariate statistical analysis is needed in place of
144 univariate analysis to properly describe the correlation structure of the hazard

145 drivers and for sampling representative combinations of hazard drivers in Monte
146 Carlo simulations.

147 The objective of this paper is to present a solution method for Eq. 1 that
148 accounts for statistical correlation structure and physical compounding effects (e.g.,
149 backwater) between boundary forcing values \hat{Q}_T and \hat{H}_T while using only a small
150 number of hydrodynamic model simulations. Building on the existing methodology
151 recommended by FEMA (2015), we present a four-step method as follows:

152 (1) Bivariate statistical analysis of Q and H records to yield possible (\hat{Q}_T, \hat{H}_T)
153 pairs for return period, T .

154 (2) Selection of N specific (\hat{Q}_T, \hat{H}_T) pairs for hydrodynamic modeling. Here, we
155 recommend $N=4$ (more detail will follow) although other options are
156 possible.

157 (3) Hydrodynamic modeling of N scenarios defined by (\hat{Q}_T, \hat{H}_T) pairs identified in
158 Step 1 to yield spatial distributions of extreme water levels, $\eta_T^i(x)$, $i=1, \dots, N$.
159 Note that the subscript on η references return period and the superscript
160 references the scenario.

161 (4) Synthesis of hydrodynamic modeling results, $\eta_T^i(x)$, $i = 1, \dots, N$, to yield
162 $\eta_T(x)$.

163 The remainder of the paper presents this method in detail along with
164 applications. Section 2 presents methods and materials including data used in this
165 study, the bivariate statistical analysis methods to determine all possible
166 (\hat{Q}_T, \hat{H}_T) pairs, identification of four specific pairs useful for hydrodynamic
167 modeling, and methods for one-dimensional steady flow analysis and two-
168 dimensional unsteady analysis of extreme water levels. Section 3 presents results of
169 one-dimensional steady-flow analysis from several sites showing differences in
170 water level profiles arising from the (\hat{Q}_T, \hat{H}_T) pairs, and results of two-dimensional
171 unsteady analysis at a single site where we further examine the limitations of the 1D

172 modeling and the added benefits of 2D modeling for assessment of coastal flood
173 hazards. Here, we also introduce an extension of the FEMA (2015) method that
174 offers potential to systematically improve the assessment of coastal flood hazards
175 by linking bivariate statistical analysis and hydrodynamic modeling. We close the
176 paper with discussion (Section 5) and conclusions (Section 6).

177 Broadly, this work shows that extending the univariate paradigm of river
178 flood hazard assessment to a bivariate paradigm of coastal flood hazard assessment
179 is not straightforward, as has been reported for many other types of compound
180 hazard (Kappes et al. 2012). Nevertheless, a systematic approach is possible and
181 shown herein. Furthermore, results point to the possibility that the existing FEMA
182 approach underestimates flood hazards where compounding effects are strong, and
183 we present a simple method to make a better estimate.

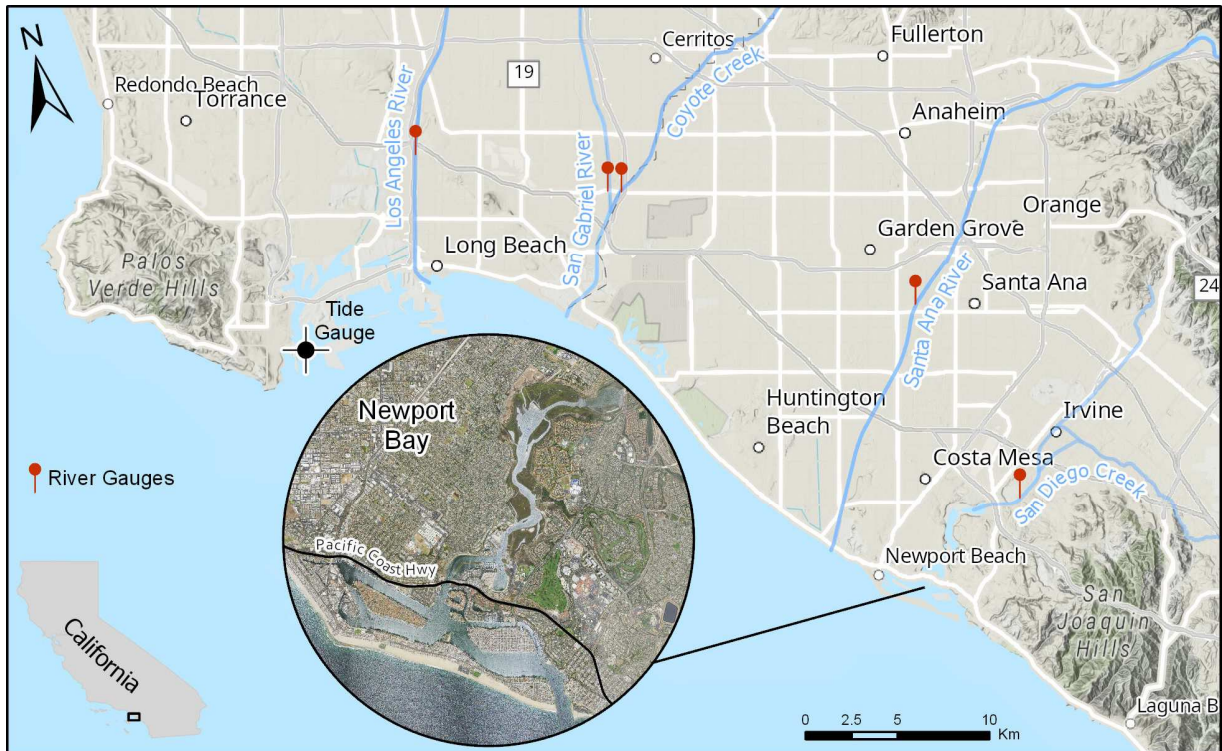
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2. Methods and Materials

2.1. Data

Analysis herein focuses on tidal channels/estuaries in southern California (See Figure 2) where flood hazards are affected by extreme ocean levels and flood discharges: the Los Angeles River (LAR), the Santa Ana River (SAR) and Newport Bay (NB). For ocean level analysis, hourly ocean water level measurements were obtained from the National Oceanic and Atmospheric Administration (NOAA) Los Angeles tide gauge (gauge ID: 9410660) at hourly intervals. Tide gauge measurements capture water level fluctuations from combined effects of tides, storm surge and other factors that affect sea levels on hourly and longer time scales, a reading that is sometimes called Total Water Level (TWL). Discharge measurements for the LAR were obtained from Los Angeles County Department of Public Works (LADPW) Station F319-R (LAR at Wardlow Road) and consisted of 92 years of annual maximum discharge data between 1928-2014. Discharge measurements for the SAR were obtained from USGS Gauge 11078000 (Santa Ana River at Santa Ana) and consisted of 94 years of annual maximum discharge data between 1923 and 2017. River discharge measurements for Newport Bay were obtained from the Orange County Department of Water Resources (OCDWR) Gauge 226 (San Diego Creek at Campus Drive) and consists of 39 years of instantaneous discharge (1978-2016). The San Gabriel River and Coyote Creek (see Figure 2) are not considered in this study since bivariate statistical analysis (see Section 3.1) showed no correlation between \hat{Q} and \hat{H} likely because of strong flow regulation from Whittier Narrows Dam located approximately 30 km from the coastline.

Topographic and bathymetric data for LAR, and SAR were taken from the 1 m resolution 2014 US Army Corps of Engineers National Coastal Mapping Program Topobathy Lidar DEM. Topographic and Bathymetric data for NB were based a DEM reported by Gallien et al. (2011) which merged several sources of topographic and bathymetric data.



213
 214 **Figure 2:** Compound flood hazards in southern California were examined for Los Angeles River, San Gabriel
 215 River, Santa Ana River, and Newport Bay using river discharge and ocean water level measurements at the
 216 gage locations shown.

217 **2.2. Bivariate Statistical Analysis**

218 Statistical analysis of extreme values of discharge and water level impacting
 219 a tidal reach or estuary are based on records of annual maximum values, although
 220 threshold-based approaches are also possible. Henceforth, we use a hat notation to
 221 indicate annual maxima data from the records of upstream discharge and
 222 downstream water level, \hat{Q} and \hat{H} , respectively. Record lengths of several decades or
 223 more are preferred to enable estimation of water levels at relatively low
 224 frequencies, to be used for the design of infrastructure (e.g., return periods of 50
 225 years or greater).

226 Bivariate statistical analysis begins with a test for correlation structure.
 227 While either linear or rank correlation coefficient measures can be used to assess
 228 the significance of the dependence between variables, tail dependence measures are

229 important for summarizing how extremes tend to occur simultaneously (Coles et al.,
 230 1999; Hao and Singh, 2016). Here, we employ joint density approaches over other
 231 alternatives (Hawkes et al., 2002; Heffernan and Tawn, 2004; Neal et al., 2013;
 232 Zheng et al., 2015) due to their flexibility, and computational/mathematical benefits
 233 (Hawkes, 2006; Salvadori et al., 2015).

234 According to Sklar's Theorem (Sklar, 1959), there exists a bivariate Copula
 235 function $C_{\hat{Q}\hat{H}}: [0,1] \times [0,1] \rightarrow [0,1]$ that formulates the joint distribution $F_{\hat{Q}\hat{H}}$ of the
 236 pair (\hat{Q}, \hat{H}) , with marginal distributions $F_{\hat{Q}}$ and $F_{\hat{H}}$, for all $(Q, H) \in R^2$, as:

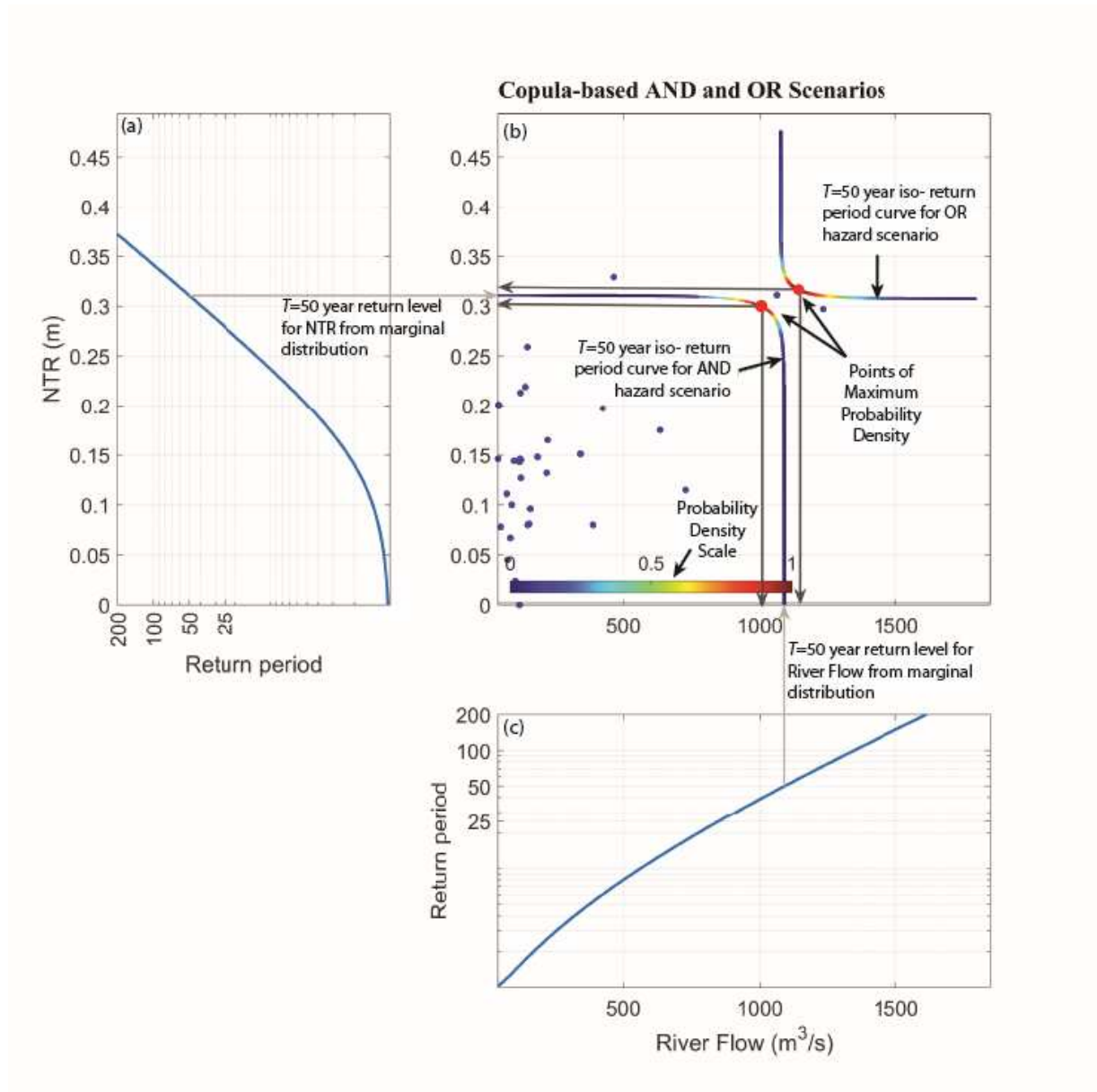
$$237 \quad F_{\hat{Q}\hat{H}}(Q, H) = C_{\hat{Q}\hat{H}}(F_{\hat{Q}}(Q), F_{\hat{H}}(H)) \quad [2]$$

238 The multivariate model is constructed by fitting suitable univariate laws on
 239 the marginals, and an appropriate copula on the observed pairs (Genest and Favre,
 240 2007; Salvadori et al., 2007). Here, we use the method of Sadegh et al. (2018) which
 241 comprehensively analyzes the dependence structure of multiple drivers of flooding,
 242 and models them using copula functions to estimate return design values and their
 243 underlying uncertainties (Sadegh et al., 2018). This approach first selects a marginal
 244 distribution from 17 univariate distributions based on measures of goodness-of-fit
 245 including Akaike Information Criterion (AIC) and Bayesian Information Criterion
 246 (BIC), and then chooses a copula model from 26 copula functions. Copula model
 247 parameters are inferred through a Bayesian inference approach with Markov Chain
 248 Monte Carlo (Sadegh et al., 2018, 2017). The joint probability can refer to the
 249 exceedance of \hat{Q} AND \hat{H} or the exceedance of \hat{Q} OR \hat{H} (Salvadori et al. 2016), and a
 250 case can be made for the relevance of both to coastal flood hazard assessment. First
 251 of all, risk assessment should reflect the possibility that flooding is caused by either
 252 extreme river discharge or extreme ocean levels, which is consistent with the OR
 253 scenario. On the other hand, hydrodynamic modeling involves the simultaneous
 254 occurrence of an upstream discharge and downstream water level, consistent with
 255 the AND scenario.

256 Newport Bay data are used to illustrate this process. Figure 3 presents the
257 outcome of bivariate statistical analysis using both the OR and the AND hazard
258 scenarios using the method of Sadegh et al. (2018). There is no statistically
259 significant correlation between river flow and ocean water level at Newport Bay, but
260 correlation was found between river flow and non-tidal residual (NTR) defined as
261 the difference between TWL and the astronomical tide level. Hence, bivariate
262 statistical analysis that takes correlation structure into account is presented here
263 using river flow and NTR. Figure 3 illustrates the similarities and differences
264 between univariate and bivariate statistical analysis as well as the relative
265 complexity introduced by the copula based AND and OR hazard scenarios. In
266 particular, Figure 3 shows plots of the marginal distributions of NTR (Fig. 3a) and
267 river flow (Fig. 3c) representative of what has traditionally been used for univariate
268 flood hazard assessment, while Fig. 3b shows the copula-based AND and OR hazard
269 scenarios. The AND and OR hazard scenarios are shown as iso- return period curve
270 for $T=50$ year within a two-dimensional space whereby every point corresponds to
271 a possible (\hat{Q}_T, \hat{H}_T) pair for use in hydrodynamic modeling. Additionally, along each
272 iso- return period curve, there is a point of maximum probability density which
273 represents the most likely (\hat{Q}_T, \hat{H}_T) pairs given the correlation structure. Note that
274 the most likely (\hat{Q}_T, \hat{H}_T) pair for the OR hazard scenario exceeds extreme values
275 given by the marginal distributions, while the most likely (\hat{Q}_T, \hat{H}_T) pair for the AND
276 hazard scenario falls below the values given by the marginal distributions. This
277 shows that the OR hazard scenario will lead to boundary forcing that is more
278 conservative (meaning a more cautious approach to risk management perspective)
279 than the AND hazard scenario.

280 Given theoretical characteristics of the OR iso-return period curves (Salvadori
281 et al., 2016) boundary forcing associated with (\hat{Q}_T, \hat{H}_T) pairs at the ends of the OR
282 curve may far exceed values given by the marginal distributions for the same return

283 period. This shows that the OR scenario creates seemingly unrealistic (highly
 284 conservative) hazard scenarios in areas of low probability density.
 285



286
 287 **Figure 3:** Marginal probability distribution for (a) non-tidal residual (NTR) which is a surrogate for \hat{H} and (c)
 288 River Flow \hat{Q} ; (b) bivariate statistical analysis of (\hat{Q}_T, \hat{H}_T) pairs based on copula-based AND and OR hazard
 289 scenarios. Two iso- return period curves corresponding to $T=50$ year are shown. Note that ends of the copula-
 290 based scenario curves for return period T are aligned with the return level of the marginal distribution for
 291 return period T . Note also that along each $T=50$ year curve, there is a point of maximum probability density
 292 which (due to correlation structure) represents the most likely (\hat{Q}_T, \hat{H}_T) pairs. The most likely (\hat{Q}_T, \hat{H}_T) pair

293 for the OR hazard scenario exceeds extreme values given by the marginal distributions, while the most likely
294 (\hat{Q}_T, \hat{H}_T) pair for the AND hazard scenario falls below the values given by the marginal distributions.

295

296 Hydrodynamic modeling of compound flood hazards for return period T is
297 proposed based on four specific (\hat{Q}_T, \hat{H}_T) pairs taken from bivariate statistical
298 analysis as shown in Figure 3:

299 (S1) A “Marginal Q ” scenario defined by the T -year return period river
300 discharge and a non-extreme water level downstream (typically taken
301 as mean higher high water).

302 (S2) A “Marginal H ” scenario defined by the T -year return period ocean
303 water level and a non-extreme river flow (typically taken as the daily
304 average flow).

305 (S3) An “AND” scenario based on the (\hat{Q}_T, \hat{H}_T) pair with the highest
306 probability density along the AND iso-return period curve.

307 (S4) An “OR” scenario based on the (\hat{Q}_T, \hat{H}_T) pair with the highest
308 probability density along the OR iso-return period curve.

309

310 FEMA (2015) presently recommends hydrodynamic modeling of S1 and S2 and
311 estimation of $\eta_T(x)$ based on the maximum of the two. Hence, two additional
312 scenarios (S3 and S4) are considered here as a way of leveraging bivariate statistical
313 analysis.

314

315 2.3. Hydrodynamic Modeling

316 Dynamic changes in water surface elevation within estuaries can be
317 modeled with reasonable accuracy using shallow-water hydrodynamic models that
318 assume a constant fluid density and a depth-averaged horizontal velocity (FEMA,
319 2018, 2015; Sanders et al., 2010). Estuaries involve the mixing of riverine and ocean
320 water with different densities, and may be characterized by strong vertical density

321 stratification that acts as a major control on the velocity distribution and transport
322 (Geyer, 2010; Monismith, 2010). Consequently, three-dimensional models that
323 account for variable density from salinity and temperature are often needed to
324 estimate velocity distributions (Jay, 2010). However, the expression of density
325 effects on surface water elevation is weak and can be neglected when predicting
326 extreme water levels for the purpose of flood hazard modeling (Friedrichs, 2010). In
327 this study, water levels are modeled by solving one-dimensional, constant density,
328 steady-state shallow-water models (Chow, 2009) and two-dimensional, constant
329 density, depth-averaged, shallow-water equations (Kim et al., 2015). When a model
330 is set up for estuaries, tidal embayments, or tidal channels, the required boundary
331 conditions correspond to a time series of river discharge at the upstream boundary,
332 $Q(t)$, and a time series of water level, $H(t)$, at the downstream boundary. The
333 upstream and downstream boundaries of the modeled spatial domain are generally
334 placed adequately apart that compounding effects are avoided. This is not, however,
335 always possible in practice because tide gauges may be located within estuaries. As
336 an aside, we note that two-dimensional models require a spatial distribution of
337 boundary forcing and in practice, hydrodynamic models include methods to
338 distribute the total volumetric flow rate, $Q(t)$, across the inflow boundary while the
339 water level, $H(t)$, is typically assumed to be uniform across the outflow boundary.
340 Estuaries may also experience water level variability from internal forcing by winds
341 and waves, and accounting for these effects is outside the scope of this study.
342 However, the role of regional winds, waves, atmospheric pressure on water levels is
343 captured by this approach based on the measurement of water levels at the tide
344 gauge which are the basis as input for bivariate flood hazard assessment.

345

346 2.3.1 1D Steady State Modeling

347 One-dimensional (1D) steady state modeling of coastal flood hazards is
348 useful as a first approximation of flood hazard levels along tidal channels and

349 estuaries and can be done quickly with low computational effort. Application of 1D
350 analysis at several sites is performed to study how differences in the selection
351 process for (\hat{Q}_T, \hat{H}_T) pairs can affect the estimation of flood hazard levels. Flood
352 hazard levels are computed by solving the gradually varied flow equation under the
353 assumption of a rectangular channel with spatially variable Manning n_m , width w ,
354 and depth h as follows (Chow, 2009)

$$355 \quad \frac{d\eta}{dx} = -\frac{S_f}{1-Fr^2} \quad [3]$$

356 where x represents distance measured inland from the mouth of the estuary,
357 $Fr=Q/(gh^3w^2)^{1/2}$ is the Froude Number and $S_f=(n_m Q/w)^2/h^{10/3}$ represents the
358 friction slope. Eq. [3] is integrated with geometrical data for each site, $n_m=0.032$ m
359 $^{1/3}$ s, the downstream water level boundary given by \hat{H} , and a river discharge given
360 by \hat{Q} . Numerical integration is performed with the 4th/5th order Runge Kutta Scheme
361 ode45 supported by Matlab (Mathworks, Natick, MA). We note that the relatively
362 simple channel geometry and resistance approximation is used herein to examine
363 the relative differences between profiles from Scenarios S1-S4, and not to estimate
364 flood hazard along these rivers in an absolute sense. More detailed geometry and
365 resistance modeling will change the absolute value of flood hazard heights, but have
366 little impact on the relative difference between scenarios.

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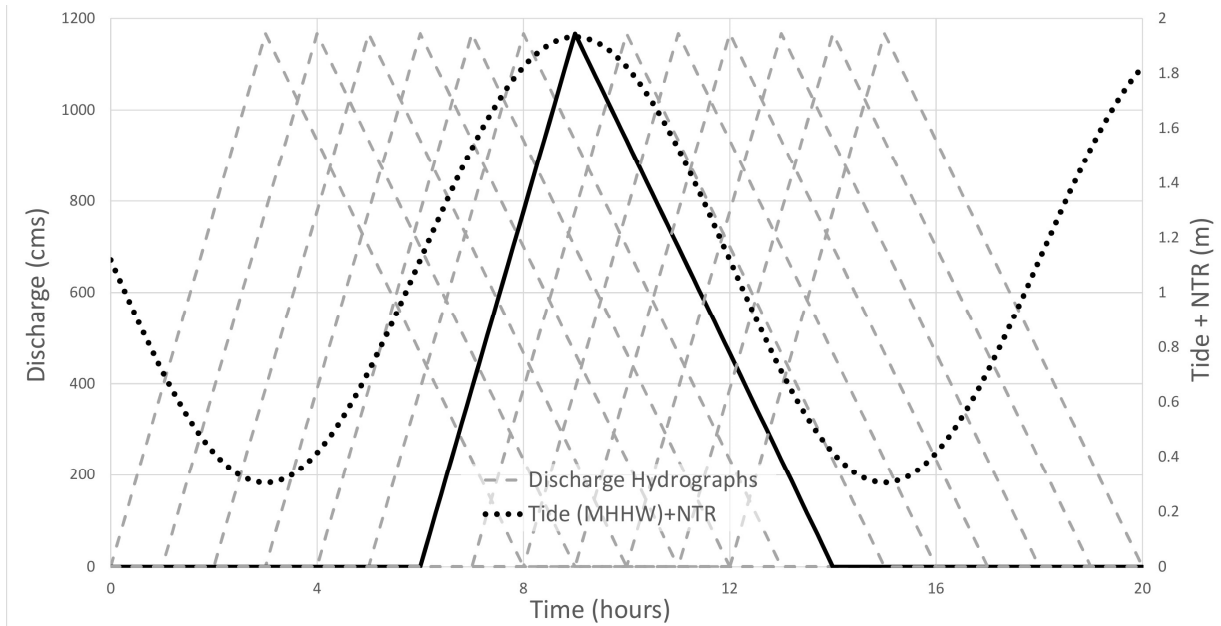
368 2.3.2 2D Unsteady Modeling

369 Two-dimensional (2D) unsteady modeling is performed at one of the four
370 sites, Newport Bay, to characterize limitations of the 1D steady state approximation
371 and to study how the relative timing of the flood peak and high tide level can affect
372 the flood hazard characterization. The 2D model BreZo (Begnudelli et al., 2008; Kim
373 et al., 2015) is applied based on a previous validation at Newport Bay (Gallien et al.,
374 2014, 2011). BreZo relies on an unstructured mesh of triangular elements with
375 varying size to capture the Bay's topography and bathymetry. The model was

376 originally setup to predict flood impacts in the urbanized portions of the Newport
377 Bay (City of Newport Beach, California, 2008). It features a fine resolution mesh
378 with 3 m average linear resolution cells across streets and land parcels, while across
379 the upper Newport Bay the average linear cell resolution is approximately 15 m.
380 The boundary conditions in BreZo are setup to specify the riverine discharge \hat{Q}
381 entering the upper Newport Bay at the outlet of the San Diego Creek, while water
382 level \hat{H} is specified along a boundary placed a short distance offshore of the
383 embayment.

384 For unsteady analysis, riverine discharge entering Newport Bay was
385 modeled with a triangular hydrograph with a peak value given by \hat{Q} and a time of
386 rise and total flood duration set to 3 and 8 hours, respectively, based on analysis of
387 instantaneous discharge measurements. Additionally, ocean water level changes
388 were modeled using a sinusoidal function with a 12 hour period (based on
389 semidiurnal tides) such that the maximum ocean water level equals \hat{H} . These
390 approximations followed preliminary modeling, which demonstrated that flood
391 heights in the upper bay were much more sensitive to the magnitude of the peak
392 flow than the duration of the event, within the range of observed values, and that
393 high-water levels were not sensitive to the precise shape of the tidal forcing. We
394 note that this may not be true in all systems, and thus these approximations are not
395 presented as a generalization but rather as a reasonable simplification given specific
396 site conditions.

397 Figure 4 presents the sinusoidal ocean forcing and the triangular inflow
398 hydrograph. To report the sensitivity of maximum water levels to the relative timing
399 of the peak inflow and peak high tide, the “OR” Hazard Scenario was repeated using
400 an inflow hydrograph that was shifted forward and backwater by as much as 6
401 hours, as shown in Figure 4.



402
 403 **Figure 4:** Inflow into Newport Bay was modeled with a triangular hydrograph with a duration and time to
 404 rise specified based on historical average values, and the ocean water level was modeled with a sinusoidal
 405 function with a period of 12 hours based on semidiurnal tides. The sensitivity of water level predictions to the
 406 time lag between peak river flow and high tide was examined with additional simulations involving forward
 407 and backwards time lags up to 6 hours.

408
 409 All modeling results are expressed in metric units and referenced to the NAD83
 410 State Plane horizontal coordinate system and NAVD88 vertical datum.

411 3. Results

412 3.1 Bivariate Statistical Analysis

413 Correlation analysis between \hat{Q} and \hat{H} defined by TWL revealed no
 414 statistical significance at these southern California sites due to relatively small
 415 storm surges compared to variability in high tide levels attributed to astronomical
 416 factors. However, correlation was found using NTR as a surrogate for \hat{H} at three of
 417 the four sites considered: LAR, SAR and NB.

418 Kendall tau and Spearman rho correlation coefficients between variables \hat{Q}
 419 and \hat{H} (defined using NTR) are presented in Table 1 along with p-values. A p-value
 420 of less than 0.05 suggests a correlation at 5% significance level. We apply the

421 method of Sadegh et al. (2018) to find the distribution functions that best describe
 422 the univariate distribution of river flow and NTR. Additionally, the JOE bivariate
 423 Copula function was then found to best describe the correlation structure between
 424 \hat{Q} and \hat{H} in all sites, given the marginal probabilities as shown in Fig. 3.

425

Site	Kendall		Spearman		Pearson		Distribution	
	Correlation Coefficient	p-value	Correlation Coefficient	p-value	Correlation Coefficient	p-value	Q	H
LAR	0.2616	0.0167	0.3880	0.0122	0.4666	0.0021	rayleigh	logistic
SAR	0.3357	0.0012	0.4907	0.0006	0.4303	0.0032	nakagami	logistic
NB	0.3229	0.0113	0.4478	0.0115	0.5659	0.0009	inverse gaussian	logistic

426 **Table 1:** Correlation coefficients between variables \hat{Q} and \hat{H} .

427

Site	S1: Marginal Q		S2: Marginal H		S3: AND		S4: OR	
	\hat{Q}	\hat{H}	\hat{Q}	\hat{H}	\hat{Q}	\hat{H}	\hat{Q}	\hat{H}
LAR	2435	1.611	10	1.901	2295	1.882	2531	1.914
SAR	693	1.611	10	1.983	640	1.965	730	1.997
NB	1090	1.611	10	1.922	1005	1.911	1167	1.930

428 **Table 2:** Scenario S1-S4 values of the 50-year return period river discharge (m^3/s) and ocean level (m above
 429 NAVD88) resulting from bivariate analysis as shown in Figure 3. Note that the AND and OR values correspond
 430 to maximum probability density (or likelihood).

431 Flood hazard levels \hat{Q} and \hat{H} for $T=50$ year resulting from the method of
 432 Sadegh et al. (2018) for Scenarios S1-S4 are presented in Table 2. Note that the
 433 “Marginal Q” scenarios all use a downstream water level corresponding to mean
 434 higher high water, and all of the “Marginal H” scenarios use a small (relative to the
 435 extreme flows) river discharge (\sim average daily flow) taken as $10 \text{ m}^3/\text{s}$. These
 436 results show that the LAR has larger river discharge values than SAR and NB, yet
 437 somewhat surprisingly, NB has larger river discharge values than SAR despite a
 438 much smaller watershed area. This is attributed to control of runoff by dams. Also,
 439 note that the water level of the “Marginal H” scenario differs between the three sites

440 despite all three relying on the Los Angeles Tide gauge. This is attributed to
441 differences in the record length arising from joint probability analysis of tide gauge
442 data and river gauge data.

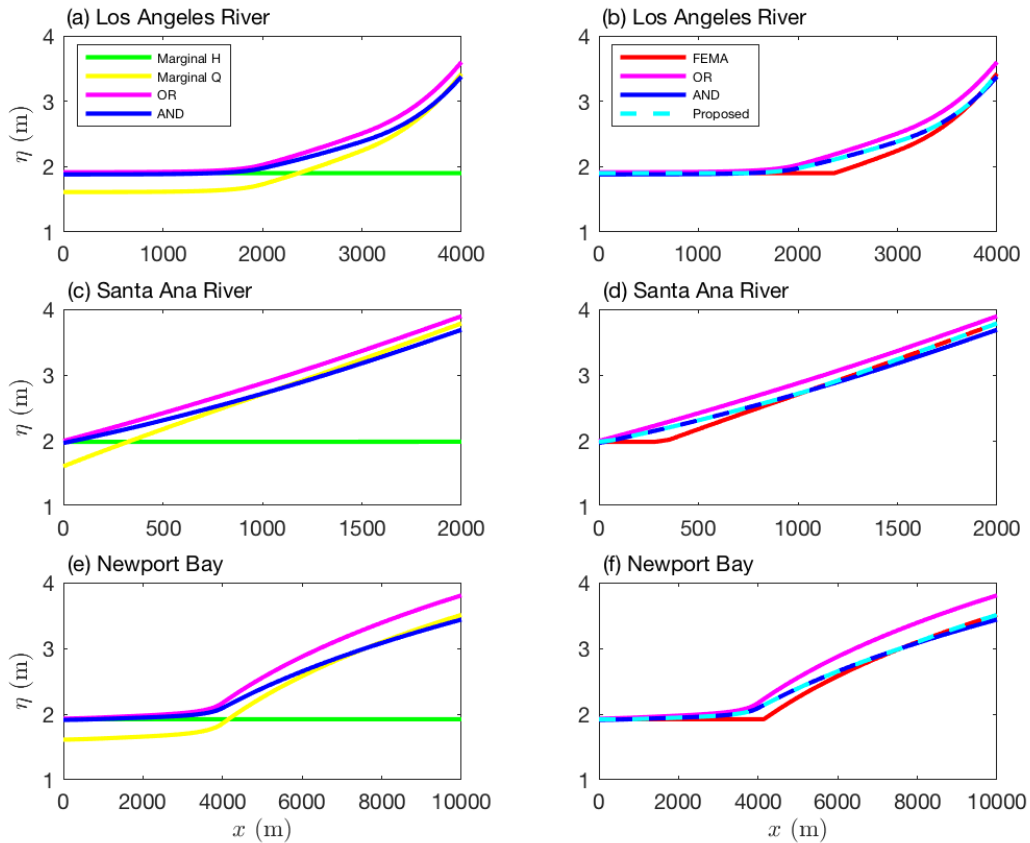
443

444 3.2 1D Flood Hazard Analysis

445 1D steady state water surface profiles were computed for the 12 sets of
446 (\hat{Q}_T, \hat{H}_T) pairs presented in Table 2 corresponding to S1-S4 at LAR, SAR and NB and
447 are presented in Figure 5a, c, and e, respectively. Focusing first on the “Marginal H”
448 and “Marginal Q” scenarios, these results shows that there is a transition in the
449 dominant factor controlling flood hazard levels along the length of the system with
450 H controlling flood hazards near the outlet and Q controlling flood hazards further
451 inland. The length of oceanic control is relatively long for NB (~4 km) and LAR (~2
452 km) and relatively short SAR (~ 300 m). Water levels from the “AND” scenario are
453 lower than the higher of the two marginal scenarios at inflow and outflow
454 boundaries, but, higher within an interior region where the marginal profiles
455 intersect. On the other hand, the “OR” scenario yields a water surface profile that is
456 always above both marginal scenario profiles. Conceptually, these results show that
457 the “OR” scenario represents a more conservative (i.e., cautious) representation of
458 the spatially variable water surface profile associated with return period T than the
459 “AND” scenario which is expected based on the magnitude of the boundary forcing
460 (see Table 2).

461 FEMA (2015) guidance recommends mapping of flood hazard levels in tidally
462 affected reaches based on the pointwise maximum of the two marginal scenarios.
463 This profile (labeled “FEMA”) is presented in Figure 5b, d, and f for LAR, SAR and
464 NB, respectively, alongside the “AND”, “OR”, and a proposed composite profile based
465 on the pointwise maximum of the “H Marginal”, “Q Marginal” and “AND” hazard
466 scenarios. At all three sites, the FEMA method underestimates the T year return
467 period water level, compared to the “AND” scenario, within a region where the

468 marginal water surface profiles intersect. Hence, the proposed composite profile is
 469 slightly higher than the FEMA method profile where there are physical
 470 compounding effects due to the interaction of river and oceanic influences on water
 471 levels; otherwise the proposed composite profile tracks the FEMA method profile.



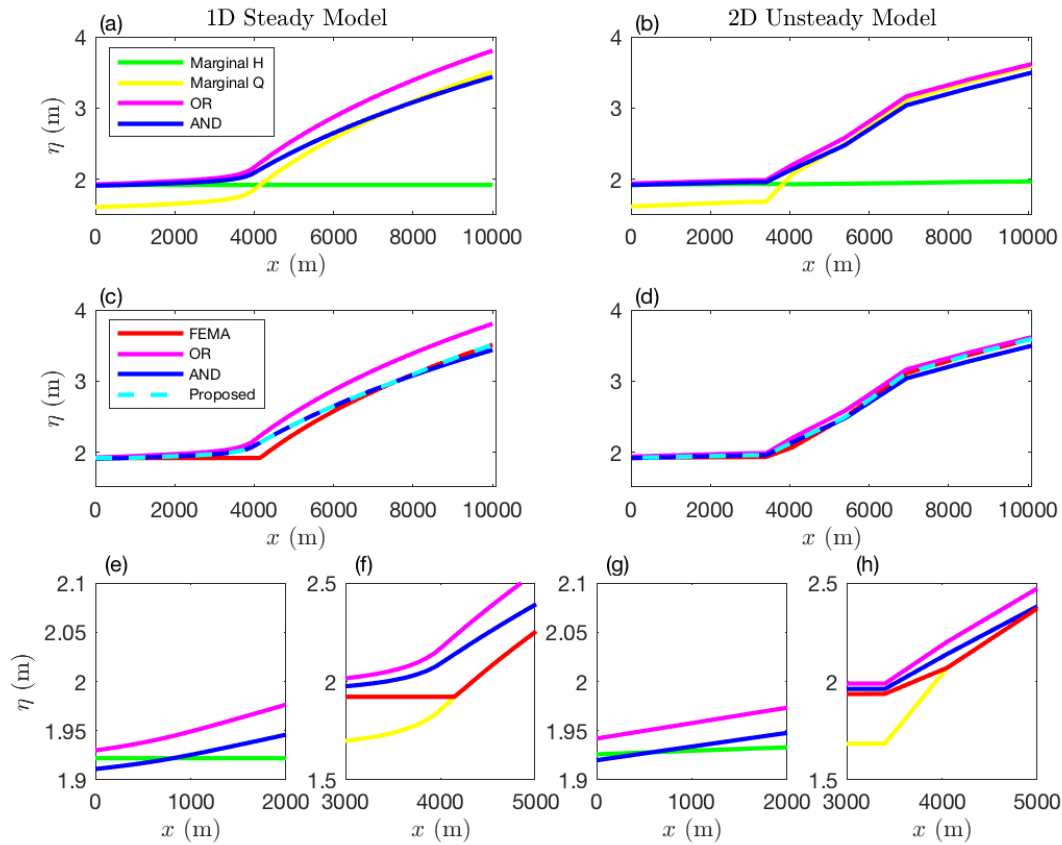
472
 473 **Figure 5:** Steady state water surface profiles versus distance from mouth, x , based on “Marginal H”, “Marginal
 474 Q”, “OR” and “AND” scenarios for (a) LAR, (c) SAR and (e) NB. Composite profiles based on the FEMA
 475 methodology and a proposed extension that considers the “Most Likely AND” scenario for (b) LAR, (d) SAR
 476 and (f) NB.

477

3.3 2D Flood Hazard Analysis

478
 479 Two-dimensional modeling of flood hazards in Newport Bay leads to spatially
 480 and temporally distributed water levels. Hence, flood hazards are mapped based on
 481 the point maximum water level attained over an unsteady simulation covering the

482 rise and fall of a flood peak with magnitude \hat{Q} and the rise and fall of an ocean tide of
 483 height \hat{H} . Results are presented first for the case of the temporal coincidence in the
 484 two peaks, and later the sensitivity of the results to the time lag between peaks is
 485 shown.



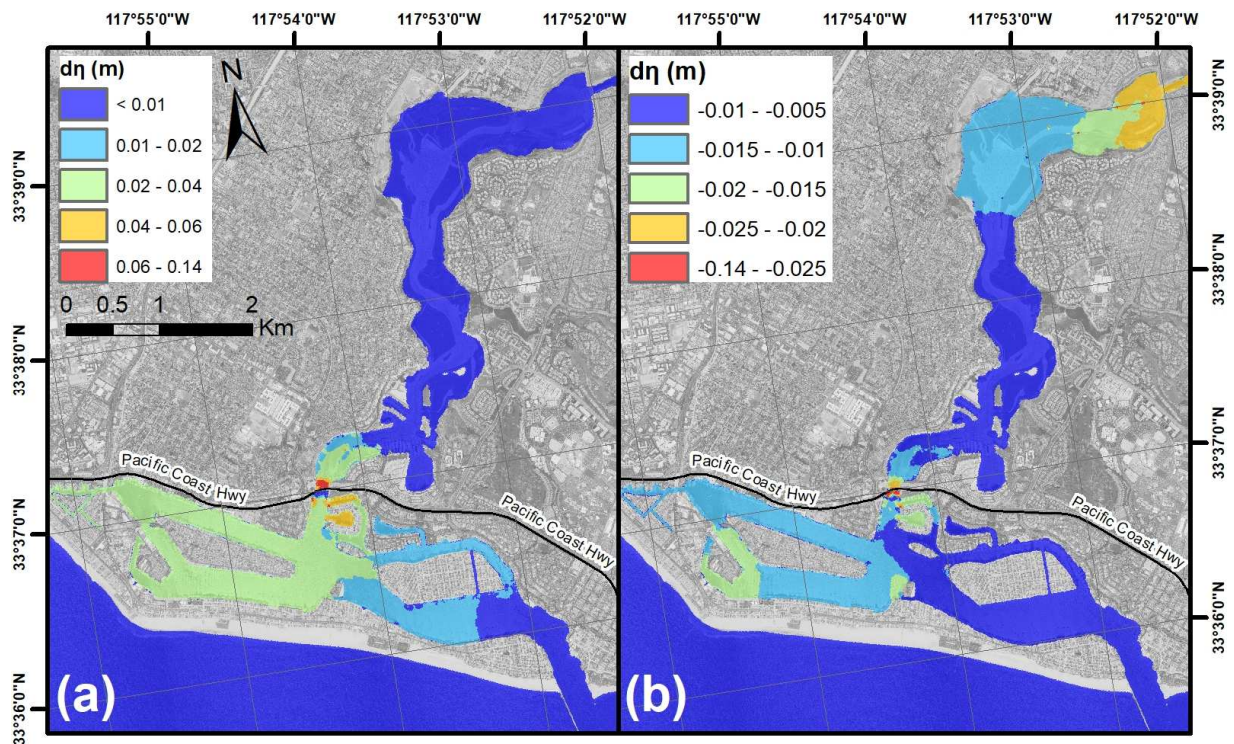
486
 487 **Figure 6:** NB flood hazard levels versus distance from mouth, x , based on 1D steady model (a,c,e,f) and 2D
 488 unsteady model (b,d,g,h). 2D results based on 5 points selected from along the main channel and linearly
 489 interpolated between points. Using both 1D and 2D approaches, the “AND” scenario predicts lower flood
 490 hazard levels compared to the maximum of the marginal profiles at the mouth ($x=0$) and head ($x=10000$ m),
 491 and higher flood hazard levels near where the marginal profiles intersect ($x=4000$ m); the “OR” scenario
 492 predicts the highest water levels everywhere. Differences between 1D and 2D models attributed to
 493 unsteadiness and treatment of complex geometry.

494
 495 A comparison of extreme water level scenarios (S1-S4) along the main
 496 channel of Newport bay using 1D and 2D methods is shown in Figure 6. For each

497 scenario, the modeling method has little impact on water profiles in the lower bay (x
498 < 4000 m) while differences are evident in upper bay ($x > 4000$ m) and attributed
499 mainly to differences in the treatment of complex system topography/bathymetry.
500 That is, the 1D model assumes a rectangular cross-sectional with a width and depth
501 based on the main channel, while the 2D model resolves both channel and
502 floodplain/marsh topography allowing for greater conveyance at higher flood
503 stages. Similar to the 1D model results presented earlier, the 2D results show that
504 the “AND” scenario predicts lower flood hazard levels compared to the maximum of
505 the marginal profiles at the mouth ($x=0$) and head ($x =10000$ m), and higher flood
506 hazard levels near where the marginal profiles intersect ($x =4000$ m). Additionally,
507 the “OR” scenario predicts the highest water levels everywhere. A magnified view of
508 water surface profiles at the mouth is presented in Fig. 6e (1D) and 6g (2D) and at x
509 $=4000$ m in Fig. 6f (1D) and 6h (2D).

510 Figure 6 also shows the FEMA (2015) composite profile and the proposed
511 composite water surface profile, which takes the pointwise maximum of the
512 marginal scenarios and the “AND” scenario. In this case, the 2D modeling predicts a
513 smaller difference between composite profiles (~ 6 cm) than 1D modeling (~ 15 cm)
514 and this is attributed mainly to the treatment of complex topography. Nevertheless,
515 small height differences can be significant with respect to the delineation of flood
516 hazard zones where floodplain topography is relatively flat. For example, a vertical
517 height of 6 cm on a slope of 1/1000 implies a 60 m change in horizontal position
518 which is larger than many land parcels in developed areas.

519 Figure 7a shows the spatial distribution of the differences between water
520 surface levels based on the FEMA method and the proposed method. The largest
521 differences ($d\eta \geq 1$ cm) are found in the lower NB and are maximum at the
522 constriction between upper and lower bay located at Pacific Coast Highway. Figure
523 7a also shows differences in the western part of NB, off line from the main channel
524 connecting San Diego Creek to the mouth of NB ,



525
 526 **Figure 7:** Difference in water surface elevations between (a) the proposed composite profile and the FEMA
 527 composite profile, and (b) “OR Hazard” scenarios with coincident peaks at hour 9 vs scenarios with shifted
 528 peaks that achieve maximum water surface level.

529 The effect of the relative timing of peak inflow and high tide to achieve
 530 maximum water surface elevations using the “OR” Hazard scenario was found to be
 531 relatively small compared to differences between hazard scenarios (i.e., AND vs.
 532 OR). Figure 7b shows color contours of the difference in water surface elevations
 533 between the scenario where hydrograph peaks are matched in time (peak at time=9
 534 hours in Figure 4) and water surface elevations obtained by shifting the peaks in
 535 time to achieve maximum water level. In the uppermost section of the upper bay
 536 maximum water level is achieved by delaying peak river discharge by one hour
 537 (peak Q at hour 10 in Figure 4), while in the lower bay maximum water level is
 538 achieved by advancing peak Q by one hour (peak Q at hour 8 in Figure 4). The
 539 difference between water levels based on the timing of high tides vs peak flow was
 540 found to be less than 3 cm across the majority of the bay. This constitutes only about
 541 a third of the difference between the proposed composite profile and the FEMA

542 (2015) composite profile. This result provides *a posteriori* validation of selecting Q
543 and H as the basis for bivariate statistical analysis at this site as opposed to other
544 system attributes such as the time lag between river flow and high tide.

545

546 **4. Discussion**

547 Multivariate statistical analysis is limited here to two variables (bivariate
548 analysis) chosen to represent annual maximum river discharge and ocean level.
549 Hazard scenarios beyond two variables are possible using copula-based methods,
550 however expansion to higher dimensions can have drawbacks including uncertainty
551 bounds so large that no conclusion may be drawn from its results (see Bevacqua et
552 al. (2017) for example). For the sites considered, the importance of the randomness
553 of annual maximum river discharge and ocean level (over other variables) justifies
554 the formulation of the bivariate statistical analysis problem around these two
555 variables. In systems where flood hazards are controlled by randomness in other
556 factors such as waves or rainfall or even uncertain internal processes, a different
557 approach would be needed. Examples of internal processes include the frictional
558 interaction between streamflow and tidal levels (Kukulka and Jay, 2003a, 2003b;
559 Moftakhari et al., 2013, 2016). Applications at a broader set of sites are warranted to
560 better understand the broader applicability of the proposed method of bivariate
561 analysis for flood hazard assessment in tidal channels and estuaries.

562 Composite profiles derived from the pointwise maximum of water levels
563 predicted by two or more hydrodynamic modeling scenarios offer a practical
564 approach for delineating compound flood hazards in tidal channels and estuaries for
565 a return period, T . Only a limited number of (relatively expensive) hydrodynamic
566 model simulations need to be completed despite an infinite number of possible
567 forcing scenarios based on bivariate statistical analysis. The most likely “AND
568 Hazard” was identified as a promising candidate for extending FEMA (2015)
569 guidance on flood hazard mapping in tidal channels and estuaries to improve

570 assessment of compound flood hazards. That is, results here suggest that the FEMA
571 (2015) may underestimate the flood hazard level over an interior section of tidal
572 reaches and estuaries where high water levels are sensitive to both riverine
573 discharge and ocean levels. In NB, this section corresponds to the urbanized lower
574 bay where exposure and vulnerability to flooding is highest. Moreover, an important
575 implication is that extreme water levels may be higher at certain points within a
576 system from combinations of river discharge and ocean heights that both fall below
577 the return levels given by the marginal distribution (i.e., univariate analysis). On the
578 other hand, the most likely “OR Hazard” scenario results in boundary forcing that
579 exceeds the return levels given by univariate analysis, and it produces water levels
580 that are higher than the marginal scenarios and the “AND Hazard” scenario. The “OR
581 Hazard” scenario could be useful when there is interest in using a single
582 hydrodynamic modeling scenario to represent compound flood hazard levels and to
583 avoid the need to compute composite profiles from multiple hydrodynamic
584 modeling scenarios. Importantly, the “OR Hazard” represents a more conservative
585 interpretation of the T year return period hazard compared to traditional univariate
586 assessment methods as well the “AND Hazard” scenario.

587 This paper points to the possibility of more robust framework for mapping
588 coastal flood hazards in tidal channels in estuaries that takes advantage of recent
589 advances in multivariate statistical modeling (e.g., Sadegh et al. 2018) and
590 hydrodynamic coastal flood hazard mapping (e.g., Gallien et al. 2011, Luke et al.
591 2018) and is in line with the limited resources and past practices of flood hazard
592 mapping (Burby 2001, FEMA 2015). In short, the method involves: (1) bivariate
593 statistical analysis where correlations in extreme values exist, (2) selection of a
594 limited number of \hat{Q} / \hat{H} pairs for return period T , (3) hydrodynamic modeling of
595 the chosen pairs to produce extreme water levels, and (4) synthesis of the model
596 results to provide a spatial distribution of water level associated with return period
597 T . If this approach is taken to be more robust than existing methods (and more

598 research will be needed at a broader set of sites to make this assessment), then the
599 limited testing presented herein points to existence of compound flood hazards that
600 are presently underestimated by the existing FEMA method for tidal channels
601 (FEMA 2015). That is, at all three sites, there was a reach of the channel where the
602 proposed composite profile accounting for the “AND Hazard” scenario was higher
603 than the composite profile accounting only for marginal scenarios as a result of
604 physical compounding effects. Recent research has also shown that flood hazard
605 zones in the U.S. are underestimated due to poor representation of pluvial flood
606 hazards (e.g., Wing et al. 2017).

607 Finally, we note that the hydrodynamic modeling shown here for Newport
608 Bay is for only a single return period to demonstrate the hybrid statistical-
609 hydrodynamic framework. To consider other return periods, the corresponding iso-
610 probability curve (e.g., Figure 3) needs to be selected to define the preferred
611 scenario.

612

613 **5. Conclusions**

614 A method of linking statistical analysis with hydrodynamic modeling to map
615 compounds in tidal channels and estuaries is presented for cases where flood
616 hazards are associated with both high river discharge (upstream forcing) and high
617 ocean levels (downstream forcing). Bivariate statistical analysis is introduced to
618 create combinations of river discharge and ocean levels suited for hydrodynamic
619 modeling, and extreme water levels produced by hydrodynamic models are
620 synthesized to create a composite water surface profile representative of return
621 period, T . The method accounts for compound flood hazards in two ways. First, it
622 accounts for statistical correlation between upstream and downstream forcing
623 which represents one dimension of compound hazards. Secondly, hydrodynamic
624 modeling accounts for physical compounding effects. Importantly, this work shows
625 that water levels at interior points of a tidal channel or estuary resulting from the

626 bivariate “AND” hazard scenario can be higher than water levels from marginal
627 scenarios even though the boundary forcing is smaller than the corresponding
628 marginal scenarios. This is attributed to physical compounding effects, i.e., nonlinear
629 interactions between discharge and water level described by shallow-water wave
630 theory. This work also shows that if a single scenario is needed to depict spatially
631 distributed compound flood hazard levels, the bivariate “OR” hazard can be used
632 and results here show that it provides a conservative assessment of the T year
633 return period hazard.

634

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640

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