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#### **Summary**

- 1. Causes of non-intentional mortality may pose conservation challenges for long-lived, migratory species. Recovery attempts for Atlantic sturgeon *Acipenser oxyrinchus oxyrinchus* populations in the USA have mainly involved closures of targeted fishing, but bycatch mortality from fisheries targeting other species remains a significant obstacle. Natural and fishing mortality levels are highly uncertain and difficult to separate, but quantifying spatial and temporal patterns of movements and total mortality can directly inform management policies regarding fishing activity that affects sturgeon. 2. Sub-adult sturgeon were tagged with acoustic transmitters to track their movements with receivers deployed in active fishing areas within the New York Bight. Multi-state mark- recapture models were used to quantify seasonal patterns in survival and migration while accounting for detection probabilities of tagged fish. 3. Movement patterns of sturgeon were highly variable among seasons along the Long Island Coast, with frequent south-westward movements during the increase in sea surface temperature in spring. North-eastward movements were most pronounced during winter, when temperatures were lowest. Sturgeon were less common along coastal Long Island during summer. 4. Larger fish transitioned among strata more frequently, but also had slightly lower survival than smaller fish which may result from selectivity for larger individuals caught incidentally in bottom trawl or gillnet fisheries. Weekly total mortality rates, including both natural and fishing mortality, averaged 0.24%. Highest weekly survival rates were observed during periods of decreasing sea surface temperature in autumn and winter, while lowest survival was observed during periods of increasing temperature in spring and summer while sturgeon migrated through areas of known bycatch. 5. *Policy implications.* Movement and survival patterns of Atlantic sturgeon suggest that late spring, coinciding with periods of high ocean bycatch in fisheries along the coast of 1. Cances of non-intentional mortality may pose conservation challenges for long-lived,<br>
26 any angle to the CSA have mainly involved closures of targeted fishing,<br>
29 conservations in the USA have mainly involved closure
- Long Island, is a particularly sensitive period for Atlantic sturgeon. Conservation efforts

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- remote sensing technologies to implement in-season fishery closures, thereby reducing incidental mortality of Atlantic sturgeon. Such bycatch management measures would aid
- in recovery attempts of a long-lived, migratory population with endangered status.

# **Key-words**

 *Acipenser*; animal movement; biotelemetry; bycatch mortality; capture-recapture; fishery closure; migration; model selection; population recovery; survival rate

## **Introduction**

 Management of living resources generally seeks to balance the abundance and mortality of populations to achieve sustainable exploitation. If populations are driven to low abundance, mortality rates must usually be reduced to allow recovery. This is particularly challenging with long-lived species [\(Heppell](#page-17-0) *et al.* 2005; [Jarić, Gessner & Lenhardt 2015](#page-17-1)) because mortality from anthropogenic factors accumulates during pre-adult life stages. Migrations pose further conservation challenges for some species as individuals are exposed to various stressors along habitats encountered. If critical mortality periods can be identified for populations in need of rebuilding, then spatio-temporal management approaches implemented at appropriate ecological scales (Dunn *et al.* 2016) may reduce mortality, even for long-lived and migrating species.

 Conservation concerns have developed for Atlantic sturgeon *Acipenser oxyrinchus oxyrinchus* (Mitchill, 1815) populations along the east coast of North America that led to threatened or endangered listings in U.S. waters [\(USOFR 2012\)](#page-18-0). Depressed population abundances resulting from two pulses of directed fishing in the late 1800s and late 1900s, combined with other factors including habitat loss [\(Smith 1985\)](#page-18-1) and pollution [\(Chambers](#page-16-1) *et al.* [2012\)](#page-16-1), led to a 1998 moratorium on targeted fishing. Despite the moratorium, populations continued to decline. One possible explanation for the lack of rebuilding is that sturgeon are still taken as bycatch (i.e. incidental catch in non-target fisheries) in marine fisheries operating along the U.S. Atlantic Coast [\(Stein, Friedland & Sutherland 2004;](#page-18-2) [ASMFC 2007;](#page-15-0) Neat *et al.* [2014;](#page-18-3) [Dunton](#page-17-2) *et al.* 2015). The timing of coastal migration periods strongly influences fishing *Actyonards*<br>
83 *Actyonard* movement: biotelemetry: bycutch mortality; capture-recapture; fishery<br>
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 [Hightower](#page-17-3) *et al.* 2015), but fishing times and locations can be managed [\(Crowder & Murawski](#page-16-2)  [1998\)](#page-16-2).

 Marine bycatch of Atlantic sturgeon occurs in coastal bottom trawl and gillnet fisheries [\(Stein, Friedland & Sutherland 2004\)](#page-18-2), particularly along the Long Island Coast in the New York Bight (Dunton *et al.* 2015). Summer flounder *Paralichthys dentatus* (Linnaeus, 1766) otter trawl fisheries overlap with sturgeon aggregations, while migration routes within 4.63 km of the coast expose sturgeon to gillnet fisheries generally targeting goosefish *Lophius americanus*  (Valenciennes, 1837) [\(Dunton](#page-17-4) *et al.* 2010; [Dunton](#page-17-2) *et al.* 2015). Detections of tagged sturgeon at acoustic receivers deployed along coastal Long Island showed broad movements across the region and two primary peaks of occurrence, in late spring and fall during migrations [\(Dunton](#page-16-3)  [2014;](#page-16-3) Dunton *et al.* 2015). However, variation among receiver stations in the probability of detecting tagged fish, as well as focus on mass detections rather than individual transitions, limit interpretation of movement and survival patterns over time.

 To better understand seasonal migration patterns along coastal Long Island and variation in weekly mortality, we applied multi-state mark-recapture models [\(Arnason 1973;](#page-15-1) [Nichols &](#page-18-4)  [Kendall 1995;](#page-18-4) [White, Kendall & Barker 2006\)](#page-19-0) to detection data of Atlantic sturgeon tagged with acoustic transmitters. Accounting for time-varying detection probabilities at receiver stations allows time-varying movement rates among stations and time-varying survival rates to be quantified, thereby identifying high mortality periods which should be the focus of conservation efforts. Areas of distribution of sturgeon populations often span hundreds of kilometres of coastline [\(Dovel & Berggren 1983\)](#page-16-4), but we focus on detailed movements along coastal Long Island at temporal scales directly relevant to the management of fishing activity in which sturgeon are caught incidentally. The research provides the application of acoustic telemetry to estimate ecological rates that can inform spatio-temporal management strategies for species of conservation concern. 85 (Stein, Friedland & 1<br>86 Bight (D[u](#page-17-2)nton *et al.*<br>87 Eisheries overla[p](#page-17-2) wit<br>87 expose sturgeon to g<br>89 (Valenciennes, 1837<br>80 acoustic receivers de<br>193 expose sturgeon to g<br>193 acoustic receivers de<br>193 exposition and two

## **Materials and methods**

 The study area is the southern coast of Long Island, New York State, along which Atlantic sturgeon migrate and trawl and gillnet fisheries operate (Fig. 1). The shoreline shifts from north- south to east-west orientation at the mouth of the Hudson River. There is a mean southerly flow of shelf waters down the Mid-Atlantic Bight, with low-salinity water emanating from the Hudson River and Chesapeake and Delaware Bays [\(Townsend](#page-18-5) *et al.* 2006). Seasonal changes are strong in the study area, with sea surface temperatures (*SST*) ranging from 2–25°C (see Fig. S1 in Supporting Information). Seas are typically calmer in summer, and wind speeds increase during fall and winter, deconstructing the summer stratification (Fig. S1).

### **Fish tagging**

 Atlantic sturgeon were captured via targeted bottom trawling [\(Dunton](#page-17-2) *et al.* 2015). Captured fish were placed into 757 L holding tanks with flowing seawater, underwent standard sampling procedures, and were surgically implanted with ultrasonic transmitters (V16–6H, VEMCO, Bedford, Canada; 69 kHz; high-power output 158 dB re 1uPa @1m; variable transmission delay of either 50–130 s with a 1952 d tag life, or 70–150 s with a 2331 d tag life). All fish were handled under ethical guidelines detailed in National Marine Fisheries Endangered Species Permit #16422 to Stony Brook University. A total of 429 sturgeon were tagged and released over 22 field days between May 2010 and April 2012, most in spring or fall (Table 1). Most fish (93%) were released in 2011 or 2012 and most (95%) were released within proximity of station B (Fig. 1; Table 1). 13 River and C[h](#page-17-5)eapteake and Delaware Bays (Townsend *et al.* 2006). Seasonal changes are strong<br>
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 All but two tagged fish were detected at some later date following release (Table 1). Tagged sturgeon ranged from 57–181 cm fork length at time of release, corresponding to ages 3– 24 (Dunton *et al.* 2016). Genetic analysis of fin clips revealed that 77% of tagged fish were of Hudson River origin and 87.7% were of New York Bight Distinct Population Segment origin (O'Leary *et al.* 2014).

### **Receiver deployments**

 Acoustic receivers (VR–2W VEMCO, Bedford, Canada; resonant at 69 kHz) were arranged in 'listening lines' perpendicular to shore, designed to detect tagged sturgeon. Receivers were

 to cover the range of depths at which sturgeon are most commonly found [\(<20 m; Dunton](#page-17-4) *et al.* [2010\)](#page-17-4). Typically, 8 receivers were deployed in each line beginning 0.5 km from shore and extending offshore approximately 9 km, with adjacent receivers spaced 1.0–1.2 km apart. Receivers were anchored to the seabed at depths ranging from 7–31 m.

 Movements were accounted for among seven stations (Fig. 1). For relevance to management, primary interest was in movements among stations B–F along coastal Long Island, in areas of active bottom trawling and gillnetting. Stations A and G represented conglomerates of individual receivers deployed southwest and northeast of the main study area, respectively, which provided additional detection data to inform movement and detection estimates at stations B–F. Station A included receivers deployed in the Hudson River and along coasts of New Jersey, Delaware and further south. Station G included receivers deployed in Long Island Sound. Most receivers in these 'terminal' stations were deployed by project partners under data sharing 149 agreements. The number of receivers, locations and timing of deployments at terminal stations A and G varied across years and were less consistent than at stations B–F, so a more flexible approach was used for modelling detection probabilities at terminal stations. 166 equilibrary is the model of the number of all free that is a movements among stations B-F along constal Long Island,<br>166 example in an axe of grative bound rawing and gillibrating, Bations A and G represented conglore

 Receiver station B was operational in years 2010–2013, but other stations C–F along coastal Long Island (Fig. 1) were only operational in 2012–2013. Once deployed, receivers were periodically retrieved to download detection data and then re-deployed for nearly continuous year-round coverage (Table S1). Occasionally individual receivers within a station were lost, yielding a gap along the listening line. We quantified the proportion of coverage at each station over time to incorporate this information as an index of sampling extent in analyses [\(Melnychuk](#page-18-7)  [2009b\)](#page-18-7). Receiver coverage at stations B–F varied over time, usually >75%, but occasionally 159 dropping  $\langle 50\% \text{ at some stations for short periods (Fig. S2)}.$ 

#### **Mark-recapture modelling**

 Multi-state mark-recapture models [\(Arnason 1973;](#page-15-1) [Nichols & Kendall 1995;](#page-18-4) [White, Kendall &](#page-19-0)  [Barker 2006\)](#page-19-0) were used to quantify detection (*p*), survival (*S*) and transition (*ψ*) probabilities throughout the study period. Transitions are movements among the seven receiver stations A–G, which are treated as mutually-exclusive 'states'. To provide the temporal resolution desired for addressing management concerns, weekly time steps were used. Parameters for each time step in

 require estimation over four years and seven strata would be prohibitively large for the sample size available [\(Burnham](#page-16-5) *et al.* 1987; [Lebreton](#page-17-6) *et al.* 1992). Instead, constraints were employed to efficiently model changes over time in *p*, *S* and *ψ* parameters while still allowing for weekly time steps. Constraints involved annual or seasonal pooling of parameters as well as pooling parameters among strata. In other words, we sought a desirable balance between model complexity for biological realism and model simplicity to avoid over-fitted models or poorly-estimated coefficients (Appendix S1).

 To quantify seasonal migration patterns, we associated weekly time steps with four seasons in each year: we considered 16 time steps for spring (1 March–20 June), 10 time steps for summer (21 June–29 Aug.), 10 time steps for fall (30 Aug.–7 Nov.) and 16 time steps for winter (8 Nov.–28/29 Feb.). For parsimony, pooling detection, survival and transition probability 178 parameters by season was necessary in some models.

 Detailed approaches and assumptions used for mark-recapture modelling are described in Appendix S1. Briefly, we constructed detection histories of individual tagged sturgeon at stations A–G and used these as input data to fit multinomial regression models containing parameters for survival, transition and detection probabilities over time. Environmental data were incorporated as covariates for estimating model parameters. Three survival sub-models, four movement sub- models and four detection sub-models were hypothesized; the resulting combinations of 48 candidate models were compared using formal model selection methods (Table S2). For the best two models, body size effects on survival and movement probabilities were evaluated (Table S3). Mark-recapture analyses sourced Program MARK [\(White & Burnham 1999\)](#page-19-1) through the R package RMark [\(Laake 2013\)](#page-17-7). 1911<br>
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### **Results**

 Model selection procedures showed strong support for four models within 1.0 AICc units of one another (Appendix S1; Table S3). These models include effects of season, *SST* and body size on transition probabilities (*ψ*), and effects of three environmental covariates (wind speed, wave height and *SST*) and receiver coverage on detection probabilities (*p*). Survival probabilities (*S*) in these four models are influenced by either *SST* or a categorical season effect, and influenced or

*[S.3FL, p.4, ψ.4FL]* (Table S4), and alternate results are shown for model *[S.2, p.4, ψ.4FL ]* (Table S5). General conclusions are consistent across the top four models.

#### **Survival patterns**

199 Survival probability estimates for weekly time steps  $t$ ,  $S_t$ , ranged from 97.06–99.98% over the duration of the study, with an annual geometric mean of 99.76% (Fig. 2). These imply average annual survival rates of 88.27% (SE 0.45%) under model *[S.3<sub>FL</sub>*, *p.4*,  $\psi$ .4<sub>FL</sub>]. Weekly  $S_t$  tended to be lowest in spring (99.44%, SE 0.15%), intermediate in summer (99.74%, SE 0.01%) and 203 highest in fall (99.89%, SE 0.01%) and winter (99.88%, SE 0.01%) (Fig. 2a). Weekly  $S_t$  were 204 not freely-varying, they were constrained as functions of  $SST<sub>t</sub>$  and change in sea surface temperature,  $\triangle SST_{t-1\rightarrow t}$ . The effect of  $\triangle SST_{t-1\rightarrow t}$  on  $S_t$  was greater than the effect of  $SST_t$  (Fig. 3; 206 Table S4). During weeks of rapidly increasing temperatures,  $\hat{S}_t$  were lowest, and were greatest during weeks of decreasing temperatures (Figs 2 & 3). Model *[S.2, p.4, ψ.4FL ]* showed similar A slight negative effect of fork length on *S* was observed. Smaller fish enjoyed a slight <sup>208</sup> results: weekly  $\hat{S}_t$  of 99.64% (spring), 99.54% (summer), 99.93% (fall), and 99.91% (winter) (Fig. 2b); annual geometric mean weekly survival, 99.76%; and average annual survival, 88.22%. 223 Seasonal movement<br>
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 survival advantage ( $\beta$ <sub>FL</sub> = -0.0094, SE 0.0053), but 95% confidence limits did not exclude zero (Table S4).

## **Movement patterns**

 Seasonal movements among detection strata were not unidirectional along coastal Long Island, they involved complex patterns. In spring, southwestward movements were more common than northeastward movements (Figs 4, S4 & S5). In summer, Atlantic sturgeon were uncommon along coastal Long Island (Fig. S3) and little movement occurred among strata as shown by the predominant intra-stratum loops in Fig. 4 (indicating a lack of transition to different strata in successive time steps; Fig. S6). In fall, weekly movements were slightly more frequent in both directions than in summer. In winter, sturgeon showed strong directional movement

223 Movement probabilities among strata,  $\psi^{stratum \to stratum}$ , were modelled with constraints of separation distance between pairs of receiver stations (Fig. S7). Separation distance had similar effects for southwestward and northeastward movements, with transitions being more frequent among strata that were closer together (Fig. 5; Table S4).

Movements over time,  $\psi_{t\to t+1}$ , were modelled with constraints of season,  $SST_t$  and  $\Delta SST_t$ 228 *1→t*. Northeastward  $\psi$  were greater when *SST<sub>t</sub>* was lower, while southwestward  $\psi$  showed little effect of  $SST_t$  *per se* on movements (Fig. 5). Instead, southwestward  $\psi$  were positively related to <sup>230</sup> *ΔSSTt–1→t*, with southwestward movements more common during periods of increasing temperature (Fig. 5). The interaction between  $SST_t$  and  $\Delta SST_{t-1\to t}$  on  $\psi_{t\to t+1}$  was also positive for <sup>232</sup> southwestward movements (Table S4), suggesting a greater tendency for southwestward <sup>233</sup> movement when temperatures were both high and increasing.

234 There was a strong positive effect of body size on overall  $\psi$  ( $\beta_{FL}$  = 0.0065, SE 0.0010). <sup>235</sup> Atlantic sturgeon with longer fork lengths at time of tagging were more likely to move among 236 strata in successive time steps (Tables S3  $\&$  S4).

#### <sup>237</sup> **Detection probabilities**

238 Estimates  $\hat{p}$  varied greatly among strata and seasons. For the aggregate stations A and G, <sup>239</sup> inconsistencies in the number and location of receivers deployed across years and seasons 240 required flexibility in modelling p, and indeed there was considerable variation in  $\hat{p}$  at these <sup>241</sup> terminal stations, especially in the first two years of the study (Fig. S8). In these years, fewer <sup>242</sup> tagged fish and fewer deployed receivers led to large seasonal changes and large uncertainties in  $\hat{p}$  at station A, but in the last two years  $\hat{p}$  were typically higher and estimated with greater <sup>244</sup> precision. Detections were less frequent at station G, so the relationship between *p* and *SST* was <sup>245</sup> constrained to be common across years, whereas relationships between *p* and *SST* were allowed <sup>246</sup> to vary among years at station A (Fig. S8). 252 Internal speed, when the speed of the speed of wind speed wind speed of wind speed and *SST*,  $m$  and *SST*,  $m$ ,  $m$  wind speed in the  $w$ ,  $w_1$ , the were nodeled with the southwestward  $\psi$  showed little<br>
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247 Along coastal Long Island,  $\hat{p}$  were typically <40% for stations C–E, <10% for station F 248 and ranged widely over time from 20–90% at station B (Fig. S8). The variability in  $\hat{p}$  over time <sup>249</sup> at stations B–F was captured by variability over time in receiver coverage and environmental <sup>250</sup> covariates. As expected, greater *p* was associated with greater receiver coverage (Fig. S9).

all three 2-way interactions on *p*, with negative  $\beta_{wind;wave}^{B-F}$  $\beta_{wind \text{.} wave}^{B-F}$ , positive  $\beta_{wind \text{.} SST}^{B-F}$  $\beta_{wind: SST}^{B-F}$  and positive  $\beta_{wave: SST}^{B-F}$  $\beta_{wave}^{B-F}$  effects observed. This implied that the overall positive effect of wave height on *p* was mediated by the interactions such that it lessened if both wave height and wind speed increased, and 255 increased if both wave height and *SST* increased. This resulted in  $\hat{p}$  tending to be greatest at intermediate levels of wind speed, wave height and *SST* (Fig. S9).

### **Discussion**

 Atlantic sturgeon use the Long Island Coast as a migration corridor, displaying strong seasonal patterns with peak occurrences in spring and fall. The few weeks during late spring of lowest survival may be a particularly sensitive time for sturgeon with respect to fishing activity, as they coincide with otter bottom trawl and gillnet fishing effort [\(Dunton](#page-17-2) *et al.* 2015). Although sturgeon are not targeted directly, incidental catch likely contributes to the high total mortality observed in late spring (Fig. 6). Causes of mortality were not directly observed with acoustic telemetry, but understanding the spatial extent and timing of mortality can aid in developing management strategies to increase sturgeon survival. Such mortality patterns can be assessed through real-time monitoring programs.

#### **Population recovery**

 Atlantic sturgeon are long-lived and not typically able to withstand high fishing mortality [\(Musick 1999\)](#page-18-8). Natural mortality has been estimated at  $M = 0.07$  based on a life-history- invariant relationship with assumed 60-year longevity [\(ASMFC 1998;](#page-15-2) [Kahnle, Hattala &](#page-17-8)  [McKown 2007\)](#page-17-8). An acceptable fishing rate, defined as half the virgin level of eggs-per-recruit, <sup>272</sup> was estimated at  $F_{50\%} = 0.03$  for fully-recruited individuals [\(ASMFC 1998;](#page-15-2) Kahnle, Hattala & [McKown 2007\)](#page-17-8). Additionally, population projections predict that 6% total annual mortality would put recovering populations at risk [\(ASMFC 2007\)](#page-15-0). Approximately 6% cumulative mortality was estimated for tagged sub-adult sturgeon during relatively short periods of 12–17 weeks in late spring (Fig. 6), and total annual mortality estimates were twice this level (11.7%). The body size range for which mortality thresholds of fully-recruited sturgeon were estimated does not completely overlap with that of sub-adult migrants telemetered in this study, but the increased. This resulted in the time of the studies and SST increased. This resulted in the nding to be greatest at intermediant stead of virtual speed, wave height and SST (Fig. S9).<br>
279 **Discussion**<br>
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 mortality have not been well-studied for this endangered species, but are likely low given the surgical implantation of tags and the relatively low tag:body mass ratio compared to other fish telemetry studies. If mortality along coastal Long Island is greater than typical levels along the greater Atlantic Coast and Hudson River, and if our study period is representative of longer-term trends, then annual total mortality of the Hudson River population is presently too high to allow for recovery. This is consistent with Atlantic sturgeon populations further south, for which annual mortality is considered sufficient to hinder recovery [\(Hightower](#page-17-3) *et al.* 2015).

 The weak relationship between body size and total survival was in the opposite direction than expected and is suggestive of size-selective bycatch mortality. Variability in natural mortality due to body size is common in marine fishes [\(Lorenzen 1996\)](#page-17-9), with larger fish enjoying a size advantage, but in this study survival was slightly greater for smaller fish. As total mortality is comprised of both natural and fishing mortality, the unexpected pattern observed for total mortality may result from either greater selectivity for larger individuals in incidental fisheries or a higher mortality rate of discarded bycatch for larger individuals, potentially outweighing any opposite pattern for natural mortality. Size distributions of sturgeon bycatch are dome-shaped (Dunton *et al.* 2015), and body sizes of sturgeon tagged in this study were mostly in the lower portion of these distributions, within the region of positive slope in which larger individuals account for more bycatch than smaller individuals. This provides support for size- selective fishing mortality (and thus considerable overall fishing mortality) in fisheries that catch sturgeon incidentally. gradia rangual total mortality of the<br>trends, then annual total mortality of the<br>for recovery. This is consistent with Atla<br>annual mortality is considered sufficient<br>The weak relationship between b<br>than expected and is sug

 Seasonal survival estimates can be used to approximate the reduction in annual mortality that could be achieved if seasonal closures were implemented in the state waters of New York during the spring peak in occurrence (April–June). Sturgeon bycatch in trawl and gillnet fisheries is relatively high in spring and summer months along coastal Long Island [\(Dunton](#page-17-2) *et al.* 2015). Assuming that the estimated average weekly mortality in fall and winter of 0.12% is comprised entirely of natural mortality, and applying this weekly mortality to a full year, yields annual natural mortality of 5.88%. The difference between annual total mortality, 11.73%, and annual natural mortality results in estimated annual fishing mortality of 5.86%, attributable to incidental catches. If the assumption that fall and winter mortality is comprised of only natural mortality is incorrect, and instead bycatch mortality occurs throughout the year, the estimate of annual

#### **Seasonal movements**

Atlantic sturgeon movement and survival patterns varied with *SST* in complex ways.

 Northeastward movements were associated with low temperature periods in winter, when weekly survival was high. The scarcity of detected sturgeon along coastal Long Island in summer did not result from lower *p* in summer (Fig. S8), but instead reflected lower occurrence (Fig. S3). In spring, during the greatest weekly increases in *SST*, weekly survival was low and southwestward movements along the coast (often into the Hudson River) were frequent. It is unknown whether sturgeon directly respond to low temperature (in winter) or rapidly increasing temperature (in spring) by increasing movement rates northeastward or southwestward, respectively, or whether temperature is simply correlated with other seasonally-changing factors. Whether the behavioural response is to temperature or to other temperature-related factors, it appears to have survival consequences. The relationship with temperature offers opportunities for implementing spatio-temporal fishery closures during spring migration periods but then relaxing restrictions in summer months, consistent with sturgeon movements [\(Breece](#page-16-6) *et al.* 2016; Dunn *et al.* [2016\)](#page-16-0).

 Individual and population characteristics also affected sturgeon migration behaviour. The observed influence of body size on movement rates is not surprising, as larger fish are typically stronger swimmers than smaller fish during migration periods. There is also some indication of differential migration behaviour among individuals within the Hudson River [\(Dunton 2014\)](#page-16-3). Notably, conclusions regarding direction of movement along coastal New York depend on the geographic scale considered. Limiting focus to only stations B–F along coastal Long Island revealed northeastward spring migrations and southwestward fall migrations [\(Dunton 2014\)](#page-16-3), whereas the larger area considered in the present study revealed the opposite pattern, driven largely by spring movements into the Hudson River (part of station A) and fall movements out of the river. Although tagged fish were primarily of Hudson River origin, fish from other populations also migrate along the New York coast [\(Dunton](#page-16-7) *et al.* 2012; [Wirgin](#page-19-2) *et al.* 2015). These findings together suggest a diversity of migration behaviours both among individuals and among populations. Future research is needed to relate migration behaviour to genetic origin and compare survival between alternate migration routes. survival was high. The scan<br>result from lower *p* in sum<br>spring, during the greatest<br>movements along the coast<br>sturgeon directly respond t<br>spring) by increasing move<br>temperature is simply corre<br>behavioural response is to<br>s

 An appropriate balance between biological realism and statistical robustness (e.g. avoidance of over-fitting a model to sparse data) is often challenging in mark-recapture modelling. In multi- state models, complexity increases non-linearly and general models (in which separate parameters are estimated for each stratum and each time step) become prohibitive for modest sample sizes [\(Nichols & Kendall 1995;](#page-18-4) [White, Kendall & Barker 2006\)](#page-19-0). The candidate models we considered were highly constrained. Time steps were weekly, but in no sub-models were there unique parameters estimated for individual weeks. Parameters were instead constrained to be common across time steps within a season or to be functions of environmental covariates. Similarly, movements among strata were constrained to be functions of environmental covariates and the separation distance between pairs of strata. This approach permitted a high temporal resolution of movement and survival parameters without having models become too complex for the sample size of tagged fish [\(Melnychuk 2009a\)](#page-18-9). For survival, near-equivalent support was found for temperature-constrained and categorical seasonal effects, and only slight body size effects were observed, but annual survival rates as well as weekly detection and transition probabilities were similar across the top four candidate models. 324 sample sizes (slichols & Kendall 1995; White, Kendall & Barker 2006). The candidate move considered were highly constrained. Time steps were weekly, but in no sub-models were the unique parameters estimated for indivi

 The study lasted four years, but sample size of tagged fish and sampling effort by receivers were greater in 2012–2013. Survival and movement sub-models did not involve any year-specific effects, so estimates for 2010–2011 are heavily weighted towards detection data from 2012–2013. Detection probability sub-models, however, did involve year-specific effects, designed to capture the change in sampling effort between the first two and last two years. Most sturgeon were tagged and released near station B, but the study was long enough to ensure sufficient mixing among stations following release dates (97.6% of sturgeon released near station B were later detected at a different station). Releases were also spread out over several dates which further aided in mixing and ensuring that observed patterns of tagged fish are likely representative of the population and its seasonal movements.

 The spatial focus of this study was the Long Island Coast, involving strata B–F; receiver stations were purposefully placed in areas of active trawling and gillnetting to quantify movement and survival patterns. Terminal strata A and G were outside this area and are of less management interest for the present study, but they provided detection data which informed movement estimates within coastal Long Island. Because strata A and G were aggregations of

 stratum G in successive time steps (loops in Fig. 4) do not imply that tagged fish were not moving frequently, rapidly, or great distances within these terminal strata. Future analyses involving migration patterns along the entire Atlantic coast could identify whether mortality

patterns change when shifting focus from a regional to a coastwide scale.

#### **Planning of spatio-temporal fishery closures**

 Conservation efforts targeted at reducing sturgeon bycatch in coastal New York could increase survival of the Hudson River population in particular, and other populations to some extent, thereby helping to reverse abundance declines. Protection from marine bycatch during occupation of the Hudson River, and *de facto* protection provided in many states by trawling limitations, leaves New York and Connecticut coastal waters lacking protection from commercial fisheries during Atlantic sturgeon migrations [\(Dunton](#page-17-4) *et al.* 2010). The higher estimated mortality rates in late spring compared with other seasons likely result from bycatch of bottom trawl and gillnet fisheries along coastal Long Island [\(Dunton](#page-17-2) *et al.* 2015). Implementing fishery closures during this critical period in this region would seem promising for decreasing bycatch mortality of the Hudson River population [\(Crowder & Murawski 1998;](#page-16-2) Neat *et al.* [2014;](#page-18-3) Dunn *et al.* 2016).

 The period of late spring and the region of coastal Long Island are shorter and smaller than annual time scales and the full area of population distribution, respectively, but socioeconomic considerations still preclude fishery closures over this full period and region. More plausibly, closures could involve specific weeks and locations along coastal Long Island. To assist with in-season predictions of critical times, the observed non-linear influence of change in *SST* on sturgeon survival can be used. Lowest survival occurred as temperatures rose quickly; if particularly low survival can be avoided by reducing fishing activity during weeks of rapid temperature increase (e.g. >1.5 or 2  $^{\circ}$ C·week<sup>-1</sup>), this could increase overall annual survival more than actions targeted at other times of year. The specific, weekly-resolution timing of *SST* increase during late spring varies among years (Fig. S1), and temperature data collected from weather and oceanographic buoys along coastal New York can be used in real time to predict the movement and presence of sturgeon in potential fishing locations. Additionally, some acoustic receivers can remotely communicate detections of tagged sturgeon, which could also inform the From the excale spatio-temporal fishery closures<br>
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survey of  $\frac{1}{2}$  abundant River population in particular, and other

 Such real-time observations have the potential to restrict in-season fishery closures to the times and locations of highest bycatch risk while allowing continued local fishing activity in times and locations of lower bycatch risk. Atlantic sturgeon could thus become an example of using acoustic telemetry and remote sensing technologies to reduce non-intentional mortality of a long-lived migratory species with endangered status to promote population recovery.

#### **Application of acoustic telemetry to conservation**

 Acoustic telemetry was effective in meeting study objectives—to estimate seasonal movements and mortality of Atlantic sturgeon, and to identify critical periods and areas which conservation actions may target—and has advantages over traditional recapture methods. Acoustic tags allowed repeated detections in multiple locations, providing more information per tagged individual than a single location and time of recapture typically provided by traditional methods; increased information gain is essential for the limited sample sizes typically used in research of endangered or depleted species. Repeated detections across widespread locations allowed for weekly resolution of mortality estimates, and at the same time allowed for quantifying broader movement and survival patterns across years and throughout the region. Traditional mark- recapture methods would not achieve these simultaneous levels of resolution without unrealistic tag numbers and levels of tag recovery effort. Traditional methods also rely on tags being returned by fishers; reporting rates in some fisheries may be low, inconsistent, poorly estimated, and potentially biased against areas and times of high bycatch that could trigger fishery closures (Brownie *et al.* 1985). A well-designed network of local acoustic telemetry studies can provide large-scale monitoring opportunities to inform regional conservation strategies while simultaneously providing high-resolution movement and mortality patterns to identify localized threats. Future research is needed to evaluate alternative (or complementary) approaches for avoiding incidental catch—such as move-on rules (Dunn *et al.* [2014\)](#page-16-9), gear alteration, and tow- length reduction—in terms of research cost, lost fishery revenue, and efficacy of bycatch reduction. As acoustic telemetry technology becomes more widely adopted in marine and freshwater systems, it has the potential to address applied ecological questions at multiple spatial and temporal scales relevant to management and conservation concerns. **Applicati[o](#page-16-8)n of accustic**<br>1096 **Application of accustic**<br>1096 **Application of accustic**<br>1098 and mortality of Atlantia<br>1098 and mortality of Atlantia<br>1099 actions may target—and<br>110 allowed repeated detectical<br>110 ind

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 Savoy (CTDEEP) and other ACT Network participants for providing additional detection data of Atlantic Sturgeon outside of our study area.

## **Data accessibility**

 Sequences of detection histories of individual tags used in this study as input data for multi-state mark-recapture models are available in the Knowledge Network for Biocomplexity online data repository (Melnychuk, 2016). In the file, 'Melnychuk\_etal\_JApplEcol\_inputData.txt', each digit of 'ch' represents a weekly time step. The first letter A–G in the sequence represents release, and all remaining letters A–G represent detection at stations A–G within the time step. Zeros following release indicate time steps of no detections, and zeros preceding release are used to align the differing release dates across tagged fish. Tags, 'tag', are numbered T1–T426, Dates on which individual fish were tagged and released, 'relDate', are given. Fork lengths, 'FL', are given in cm. 256 Catalant Montain Properting 2<br>
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## <sup>556</sup> **Tables**

<sup>553</sup> *Biology,* **86,** 1251-1270.

<sup>557</sup> Table 1. Release times, numbers and body sizes of tagged Atlantic sturgeon throughout the study <sup>558</sup> period

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<i>Biology</i> , <b>86</b> , 1251-1270.					
<b>Tables</b>					
					Table 1. Release times, numbers and body sizes of tagged Atlantic sturgeon throu
period					
Release date	Number	Number	Nearest	FL	Range FL
	released	detected	station to	$mean \pm SD$	$(cm)$ $\dagger$
		thereafter	release*	$(cm)$ †	
$2010 - 05 - 18$	3	3	B	$76.0 \pm 0.5$	$75.5 - 76.5$
$2010 - 05 - 19$	9	9	$\bf{B}$	$91.9 \pm 14.3$	72.0-109.0
$2010 - 06 - 14$	8	8	B	$104.4 \pm 8.5$	$94.5 - 122.0$
$2010 - 06 - 15$	4	4	B	$117.2 \pm 12.0$	$105.0 - 130.0$
$2010 - 10 - 23$	$\mathbf{1}$	$\mathbf 1$	$\mathbf{A}$	104.0	
$2010 - 10 - 25$	3	3	$\mathbf{A}$	$137.2 \pm 38.4$	$103.5 - 179.0$
$2011 - 04 - 13$	$\overline{2}$	$\overline{2}$	$\mathbf{A}$	$65.9 \pm 12.2$	$57.2 - 74.5$
$2011 - 04 - 14$	3	3	$\mathbf{A}$	$66.7 \pm 3.4$	$64.2 - 70.6$
$2011 - 04 - 25$	$\mathbf{1}$	$\mathbf{1}$	A	59.6	
$2011 - 04 - 26$	$\overline{4}$	$\overline{\mathcal{A}}$	A	$73.6 \pm 8.7$	$65.1 - 83.7$
$2011 - 04 - 27$	$\overline{4}$	$\overline{4}$	$\mathbf{A}$	$72.8 + 7.3$	$64.5 - 82.0$
$2011 - 05 - 24$	68‡	$67 \ddagger$	$\, {\bf B}$	$114.6 \pm 14.8$	$76.7 - 155.0$ $\ddagger$
$2011 - 08 - 11$	47	47	B	$112.5 \pm 16.0$	75.6-138.3
$2011 - 09 - 11$	40	40	$\bf{B}$	$103.2 \pm 19.1$	71.9-164.0
$2011 - 10 - 11$	26	26	$\, {\bf B}$	$101.0 \pm 17.3$	77.3-127.0



\* See labels in Fig. 1.

 $\ddagger$  Fork length (FL) summary statistics for released fish; SD = standard deviation.

‡ Three fish from 2011–05–24 release group either died or had tags extruded near station B

shortly after release and were detected continuously thereafter. These fish were excluded from

analyses resulting in n=426, of which 424 were detected thereafter.

## **Figure Captions**

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 Figure 1. Map of study area. Dark circles show acoustic receivers maintained by Stony Brook University; lighter circles show receivers maintained by Atlantic Cooperating Telemetry Network partners where Atlantic sturgeon were commonly detected. Receiver stations A–G were treated as distinct states in multi-state models: A, all areas south and west of B including Hudson River (area shaded); B, Rockaway Peninsula and Sandy Hook Spit (area shaded); C, Jones Beach; D, Fire Island; E, Shinnecock; F, Montauk Point; G, all areas north and east of F including Long Island Sound (area shaded). Major water bodies and two-letter state abbreviations are labelled. Inset shows U.S. mid-Atlantic coast, with dashed box showing main study area. Scale bar applies to main map. **2012-02-2 <b>c**  $\frac{1}{2}$  and  $\frac{1}{2}$  and  $\frac{1}{2}$  B  $\frac{1}{2}$  101.4+18.0 76.0-149.6<br>
2012-04-05  $\frac{22}{2}$  22 B  $104.4+19.7$   $57.2-181.0$ <br> **FLAB**<br> **FLAB**<br> **FLAB**<br> **FLAB**<br> **FLAB**<br> **FLAB**<br> **FLABE**<br> **FLABE**<br> **FLABE**<br> **FL** 

Figure 2. Weekly Atlantic sturgeon survival probability estimates under (a) model *[S.3FL , p.4 ,* 

 surface temperature (*SST*, °C) measured at the New York Harbor Entrance Buoy is overlaid in (a); one period of four time steps had no available *SST* data, during which *SST* values were linearly interpolated. Seasons are shown with grey shading: spring = lightest; winter = darkest. Primary vertical axis does not begin at 0.

 Figure 3. Effects of sea surface temperature (*SST*, °C), change in *SST* between successive time steps (Δ*SST*) and their interaction on survival probability estimates. Solid lines show partial survival predictions, with lower 95% confidence band (dashed). Vertical axes do not begin at 0. 

 Figure 4. Seasonal transition probabilities among strata. Estimated transition probabilities for each time step are averaged across each season, separately for each of 49 possible transitions. Season durations are: spring, 1 March–20 June; summer, 21 June–29 Aug.; fall, 30 Aug.–7 Nov.; 591 winter,  $8 \text{ Nov.} -28/29$  Feb. Northeastward transitions are shown by arches above station labels, southwestward transitions are shown by arches below station labels, and lack of movement to different strata in successive time steps is shown by loops adjacent to station labels. Average transition probabilities into each stratum and from each stratum sum to 1, with line thicknesses proportional to these 0–1 probabilities. Figure 3. Effects of sea surface<br>
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 Figure 5. Effects of separation distance between receiver stations, sea surface temperature (*SST*, °C), change in *SST* between successive time steps (Δ*SST*) and interaction of *SST* and Δ*SST* on transition probability estimates. Estimates are shown only for movements involving stratum D, for clarity. Partial survival predictions (solid line) and 95% confidence bands (dashed) are shown.

 Figure 6. Cumulative mortality during periods when tagged Atlantic sturgeon were frequently detected at one or more stations B–F in the final two years of the study. Selected intervals are arbitrary, chosen to correspond to periods of greatest numbers detected (2012-03-01 to 2012-06- 21; 2012-08-30 to 2012-12-20 and 2013-04-26 to 2012-07-12). Cumulative mortality is calculated as 100% minus the product of weekly survival estimates during the interval.

- Additional Supporting Information may be found in the online version of this article.
- **Appendix S1.** Detailed methods for mark-recapture modelling approaches.
- **Table S1.** Number and duration of acoustic receiver deployments at stations B–F throughout the study period.
- **Table S2.** Initial information-theoretic comparison of candidate mark-recapture models.
- **Table S3.** Final information-theoretic comparison of mark-recapture models involving possible
- effects of body size on survival and transition probabilities.
- **Table S4.** Estimated coefficients for AICc-best model *[S.3FL , p.4 , ψ.4FL ].*
- **Table S5.** Estimated coefficients for model  $[S.2, p.4, \psi.4_{FL}]$ .
- **Fig. S1.** Time series of wind speed, wave height and sea surface temperature throughout the study period.
- **Fig. S2.** Index of detection coverage by weekly time step for detection stations B–F.
- **Fig. S3.** Number of tagged fish detected per day at stations during the study period.
- **Fig. S4.** Estimated transition probabilities for northeastward movements along the Long Island Coast.
- **Fig. S5.** Estimated transition probabilities for southwestward movements along the Long Island Coast.
- **Fig. S6.** Implied transition probabilities for remaining in the same stratum between successive time steps. **Fig. S5.** Estimated transition-theoretic comparison of candidate mark-recepture models.<br> **Fable S3.** Final information-theoretic comparison of mark-recepture models involving p<br>
effects of body size on survival and trans
- **Fig. S7.** Matrix of separation distances between pairs of strata.
- **Fig. S8.** Estimated detection probabilities at stations A–G.
- 630 **Fig. S9.** Effects of wind speed (m·s<sup>-1</sup>), wave height (m), sea surface temperature (*SST*, °C), 2-
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