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## Assessing bottom-trawling impacts based on the longevity of benthic invertebrates

**Running title:** Effects of bottom trawling on seabed biota

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39 **Abstract**

- 40 1. Bottom trawling is the most widespread human activity directly affecting seabed habitats.  
41 Assessment and effective management of the effects of bottom trawling at the scale of  
42 fisheries requires an understanding of differences in sensitivity of biota to trawling.  
43 Responses to disturbance are expected to depend on the intrinsic rate of increase of  
44 populations ( $r$ ), which is expected to be linearly related to the reciprocal of longevity.
- 45 2. We examine the relationship between the longevity of benthic invertebrates and their  
46 response to bottom trawling; both in terms of the immediate mortality following a trawl pass  
47 and their subsequent rates of recovery. We collate all available data from experimental and  
48 comparative trawling studies, and test how longevity influences these aspects of sensitivity.
- 49 3. The shortest-lived organisms (<1yr) increased in abundance shortly after experimental  
50 trawling but showed no response to trawling in longer-term comparative studies. Conversely,  
51 the abundance of biota with a life-span >1yr decreased by ~9% immediately following a trawl  
52 pass. The effect of bottom trawling in comparative studies increased with longevity, with a 2-  
53 3× larger effect on biota living >10yr than on biota living 1-3yr. We attribute this difference  
54 to the slower recovery rates of the longer-lived biota.
- 55 4. The observed relationship between the intrinsic rate of population increase ( $r$ , our metric of  
56 recovery rate) and the reciprocal of longevity matches theoretical expectation and predicts  
57 that the sensitivity of habitats to bottom trawling is higher in habitats with higher  
58 proportions of long-lived organisms.
- 59 5. *Synthesis and applications.* Where the longevity of a species or the longevity distribution of a  
60 community is known or can be inferred, our estimates of depletion and intrinsic rate of  
61 increase can be combined with high-resolution maps of trawling intensity to assess trawling  
62 impacts at the scale of the fishery or other defined unit of assessment. Our estimates of  $r$   
63 may also be used to estimate recovery times following other forms of seabed disturbance.

64 **Alternative language abstract: Dutch**

65 **Beoordelen van bodemvisserijeffecten op basis van de levensduur van bentische ongewervelde**  
66 **dieren**

- 67 1. Trawlvisserij op de zeebodem is de meest voorkomende menselijke activiteit met rechtstreekse  
68 gevolgen voor habitats op die zeebodem. Om de effecten van bodemtrawls op de schaal van de  
69 visserij te kunnen beoordelen en effectief te beheren, is inzicht nodig in gevoeligheid voor

70 trawlvisserij van verschillende dieren. Reacties op verstoringen zullen naar verwachting afhangen van  
71 de intrinsieke groeisnelheid van populaties ( $r$ ). We verwachten dat die omgekeerd evenredig is aan de  
72 levensduur van soorten.

73 2. We onderzoeken de relatie tussen de levensduur van bentische ongewervelde dieren en hun  
74 reactie op bodemtrawls; zowel in termen van de onmiddellijke sterfte na een trawltrek als van de  
75 daaropvolgende herstelpercentages. We verzamelen alle beschikbare gegevens uit experimentele en  
76 vergelijkende trawlonderzoeken en testen hoe de levensduur van soorten deze aspecten van  
77 gevoeligheid beïnvloedt.

78 3. De aantallen kortlevende organismen (<1 jaar) namen toe kort na experimentele trawlvisserij, maar  
79 vertoonden in vergelijkende studies op langere termijn geen reactie op trawlvisserij. De talrijkheid  
80 van dieren met een levensduur >1 jaar nam af met ~ 9% onmiddellijk na een trawltrek. Het effect van  
81 bodemtrawls in vergelijkende studies nam op langere termijn toe met de levensduur. Er was een 2-3 x  
82 groter effect op dieren die >10 jaar leefden dan op dieren die 1-3 jaar oud worden. We schrijven dit  
83 verschil toe aan het tragere herstel van de langer levende dieren.

84 4. De waargenomen relatie tussen de intrinsieke snelheid van de toename ( $r$ ) en het omgekeerde van  
85 de levensduur van een soort komt overeen met de theoretische verwachting. Hiermee wordt  
86 voorspeld dat de gevoeligheid van habitats voor bodemtrawls hoger is als er sprake is van hogere  
87 percentages langlevende organismen.

88 5. Synthese en toepassingen. Als de levensduur van een soort of de levensduur van een gemeenschap  
89 bekend is of kan worden afgeleid, kunnen onze schattingen van sterfte en intrinsieke groeisnelheid  
90 worden gecombineerd met hoge resolutie kaarten van de intensiteit van de trawlvisserij om de  
91 impact van trawlvisserij op de schaal van de visserij te bepalen.

92 Keywords: benthic invertebrates, bottom trawl, fisheries management, impact assessment, life-  
93 history meta-analysis, systematic review, seabed disturbance

#### 94 **Introduction**

95 Development of ecosystem-based fisheries-management strategies requires assessments of the  
96 distribution and impact of bottom trawling and the status of sediment-dwelling biota and habitats  
97 (Rice & Rochet, 2005; Rice *et al.*, 2012; Rijnsdorp *et al.*, 2016). Since biota and habitats differ in their  
98 degree of exposure and sensitivity to bottom trawling, knowledge of their distribution and sensitivity  
99 is required to assess impact and to formulate options or identify priorities for conservation and  
100 management.

101

102 Bottom trawls, here defined as any towed bottom-fishing gear including otter trawls, beam trawls,  
103 scallop dredges and hydraulic dredges, are used to catch fish, crustaceans and bivalves living in, on or  
104 close to the seabed (Sainsbury, 1986). Bottom trawling is by far the most extensive anthropogenic  
105 physical disturbance in the marine environment (Eigaard *et al.*, 2016), but also makes an important  
106 contribution to global food supply and accounts for 35M tonnes of annual fish catches (combined  
107 weight of reported landings, estimated discards and estimated landings from illegal, unreported and

108 unregulated (IUU) fishing; Watson & Tidd, 2018). Bottom trawling reduces benthic invertebrate  
109 biomass, numbers and diversity (Sciberras *et al.*, 2018), and selects for communities that are  
110 dominated by fauna with fast life-histories (Tillin *et al.*, 2006; van Denderen *et al.*, 2015) and/or that  
111 experience low rates of mortality per trawl pass (Pitcher *et al.*, 2000). These effects can lead to  
112 changes in community production, trophic structure and function (Hiddink *et al.*, 2006). While  
113 frequency of disturbance is governed by exposure to fishing activity, mortality rates ( $d$ ) depend on the  
114 penetration depth of the gear into the sediment (Hiddink *et al.*, 2017), the morphology and size of  
115 biota and their position relative to the sediment surface (Duplisea *et al.*, 2002; Sciberras *et al.*, 2018).  
116 Capacity to withstand disturbance depends on life-history characteristics, such that species with  
117 higher intrinsic rates of increase ( $r$ ), and therefore higher recovery rates, are expected to display  
118 smaller decreases in abundance in response to a given rate of mortality (Duplisea *et al.*, 2002).

119

120 Intrinsic rates of population increase tend to be higher in species with higher metabolic rates (e.g.  
121 Brown *et al.*, 2004; Savage *et al.*, 2004). Owing to compensatory adjustments in their life-histories,  
122 these species tend to have earlier maturity, higher annual reproductive output and higher natural  
123 mortality, the latter leading to lower average life span (Hoenig, 1983; Charnov, 1993). Conversely,  
124 species with a lower metabolic rate have later maturity, lower annual reproductive output and lower  
125 natural mortality, leading to greater average longevity. Compensatory trade-offs between these life-  
126 history parameters lead to “life-history invariants” allowing one life-history parameter to be used as a  
127 reasonable proxy for several others (Beverton & Holt, 1959; Charnov, 1993). We therefore infer that  
128 the effect of any given rate of trawl mortality on a population will depend on its life-history, whereby  
129 populations with low  $r$ , low natural mortality rates ( $M$ ) and greater longevity ( $T_{\max}$ ) have an increased  
130 sensitivity to trawling disturbance (Duplisea *et al.*, 2002). For example, Tillin *et al.* (2006)  
131 demonstrated that benthic epifauna with  $T_{\max} > 10\text{yr}$  decreased in abundance with trawling, but that  
132 no such reduction occurred for fauna in the same areas with  $T_{\max} < 2\text{yr}$ .

133

134 The widespread adoption of Vessel Monitoring Systems for fisheries has revolutionised our capacity  
135 to map the exposure of seabed habitats to trawling disturbance at high resolution (Eigaard *et al.*,  
136 2016). However, information on the benthic biota in the trawled areas may be sparse, and the  
137 distribution, variety or abundance of biota is often inferred by collating existing data. These data  
138 usually come from surveys and monitoring programmes that may not focus on the components and  
139 attributes of the community that provide effective indicators of trawling impacts. For example, many  
140 data are composed of counts by species but provide no information on body size or biomass, both of  
141 which are more informative indicators of trawling impacts (Hiddink *et al.*, 2006). Given that the  
142 dedicated collection of habitat data for the large-scale analysis of trawling impacts will require  
143 significant investment and long-term changes to existing surveys, current assessments of trawling  
144 impacts should seek to make best use of available data to describe the distribution and sensitivity of  
145 benthic biota. Approaches have included making better use of species count data by assigning traits

146 to species to describe aspects of the functional or life-history characteristics of a community (e.g.  
147 Tillin *et al.*, 2006). If the species present and their numeric abundances are known then traits such as  
148 longevity, that determine responses to elevated mortality, can be assigned to provide metrics of  
149 community sensitivity.

150  
151 Our primary objective is to examine the relationship between the longevity of benthic invertebrates  
152 ( $T_{max}$ ) and their response to trawling, both in terms of the mortality induced by the passage of a trawl  
153 and their recovery following trawling. The resultant relationships can be used to infer sensitivity to  
154 trawling for species of known longevity or communities with a known longevity distribution, for the  
155 purposes of assessing trawling impacts and informing habitat conservation and ecosystem-based  
156 fishery management. A secondary objective of this study is to demonstrate the use of the estimated  
157 fraction mortality per trawl pass (depletion  $d$ ) and  $r$  for calculating the relative benthic status (*RBS*) of  
158 seabed fauna after trawling. The *RBS* method was developed and described by Pitcher *et al.* (2017)  
159 and is a spatial and quantitative environmental risk assessment method for assessing trawling  
160 impacts.

161

## 162 **Materials and Methods**

163

164 We collated all available data from experimental and comparative trawling studies using a systematic  
165 review methodology. Three different types of studies were included: 1) *experimental* studies where a  
166 defined area was experimentally trawled, while another area was not trawled, 2) comparative  
167 “*trawled vs untrawled*” contrast (TU) studies where a commercially trawled area was compared with  
168 a reference that was not trawled or only lightly trawled, and 3) comparative *trawl gradient* studies  
169 (TG) which included studies of benthic communities over gradients of commercial trawling frequency.  
170 Experimental trawling studies were used to estimate depletion  $d$ . The TU studies were used to  
171 compare the effect of a typical bottom trawl fishery on the abundance of organisms with different  
172 longevity; these studies do not provide parameter estimates, but the outcomes from this analysis  
173 are used to examine the validity of the relationship between longevity and trawling response. The TG  
174 studies were used to estimate the intrinsic rate of increase  $r$  for these groups. The validity of  
175 estimates of  $d$  and  $r$  were tested using field observations from the North Sea.

176

### 177 Data sources and study inclusion criteria

178 Data were collated from published experimental and comparative studies of the effects of bottom  
179 trawling on seabed habitat and biota following a systematic review protocol (Hughes *et al.*, 2014),  
180 thereby avoiding selection bias. Studies were included if the abundance  $B$  (as numbers or biomass) of  
181 benthic species, genera and families, of either infauna and/or epifauna, was reported. Our analysis of  
182 comparative studies assumed that other environmental variables were not systematically related to  
183 trawling intensity, and we did not include studies in our analysis where this assumption was violated.

184 Gradients in trawling effort may be driven by fishing regulation and seabed obstructions, but are also  
185 observed in areas of homogenous habitat (e.g. Hiddink *et al.*, 2006). Data for number of individuals  
186 and biomass were combined in a single analysis, justified based on the analysis of Hiddink *et al.*  
187 (2017).

188  
189 Overall there were 32 articles reporting 53 experimental studies, and 31 articles providing data on 30  
190 comparative TU studies and 9 comparative TG studies. Four of the comparative TU studies provided a  
191 quantitative estimate of trawling intensity in both locations and were therefore included in both the  
192 TG and the TU analyses. The studies included in the analyses covered a wide range of trawl gears  
193 (otter trawl, towed dredge and hydraulic dredge), habitats (mud, sand, gravel and biogenic habitats),  
194 and faunal composition (infauna and epifauna). Studies were concentrated in temperate areas of  
195 Europe and North America, whereas very few studies were from tropical areas (see Tables S1 and S2  
196 and Figures S1 and S2 in Supporting Information).

197

#### 198 Longevity

199 Depletion and recovery parameters were estimated for benthic fauna with different longevity, which  
200 was defined as the maximum-recorded age of a species  $T_{\max}$  in years. Owing to scarce data and high  
201 uncertainty in  $T_{\max}$  estimates for individual species, longevities were assigned to taxa with a fuzzy-  
202 coding approach using the existing database of Bolam *et al.* (2017). This database assigns fractional  
203 scores to each of four  $T_{\max}$  categories (<1, 1-3, 3-10, >10yr, chosen to encompass the range of  
204 possible attributes of all the taxa), depending on the affinity of the species with these categories, and  
205 summing to one (Chevenet *et al.*, 1994; explained in more detail in Tillin *et al.*, 2006). Fuzzy coding  
206 allows taxa to exhibit multiple  $T_{\max}$  categories to different degrees, and helps to address the  
207 uncertainty in and absence of direct  $T_{\max}$  measurements for many benthic invertebrate species and  
208 expected differences in  $T_{\max}$  within species linked to latitude and environment. After coding, the  
209 abundance of each taxon was multiplied with its proportional affinity. Consequently, the abundance  
210 of each taxon within a study was divided across  $T_{\max}$  categories within that study in proportion to its  
211 fuzzy membership. Subsequently, the affinity-adjusted abundance was summed over all species in a  
212 study for each  $T_{\max}$  category, yielding a total abundance per  $T_{\max}$  category, which was used for the  
213 analyses below.

214

#### 215 Analysis of experimental studies

216 Depletion  $d$  by  $T_{\max}$  category was estimated using data collated from experimental studies, which  
217 quantified the abundance of biota following one or more trawling events. We used the log response-  
218 ratio (lnRR) for abundance  $B$  as the response variable, calculated as  $\ln(B_{\text{trawled}}/B_{\text{control}})$ . Depletion  $d$   
219 was estimated using weighted Meta-Analysis via Linear Mixed-Effects Models (using the function  
220 *rma.uni* in package *metafor* in R, Viechtbauer, 2010) according to the model:  $\ln RR_{ij} = \alpha_j + \beta_j \log_2(t_{ij} +$   
221  $1) + \alpha_i + \varepsilon_{ij}$  (eqn 1) for study  $i$  and  $T_{\max}$  category  $j$ ;  $\alpha$  is the random study effect;  $\varepsilon$  is the Gaussian error

222 term.  $t$  is the time in days since trawling. Depletion  $d_j$  was estimated as  $d_j = -(1-e^{-aj})$  (eqn 2). Time  $t$   
223 was transformed using a base-2 logarithm to provide a linear approximation to a logistic recovery as  
224 non-linear curves proved difficult to fit directly to the available data given the relatively low number  
225 of replicate studies. Studies were weighted by the inverse of their variance estimate, as is normal  
226 practice in meta-analyses. The combined variance for multiple species with the same  $T_{max}$  within a  
227 study and  $T_{max}$  category was calculated following Borenstein *et al.* (2009). Many experimental studies  
228 were excluded from the analysis as they did not include a measure of variance. The carrion generated  
229 by trawling as a result of the direct mortality of organisms on the seabed, and by discarding of by-  
230 catch, produces food subsidies for scavenging species and can lead to a short-term influx of  
231 scavengers in recently trawled areas (Ramsay *et al.*, 1997). Data for scavenging species collected  
232 within two days of experimental fishing disturbance were therefore removed from the data-set prior  
233 to analyses as in Sciberras *et al.* (2018).

234

#### 235 Analysis of comparative trawled vs untrawled (TU) contrast studies

236 Trawled vs untrawled (TU) contrast studies were used to compare the effect of a typical bottom trawl  
237 fishery on the abundance of groups of organisms with different longevities. Most TU studies did not  
238 give trawling frequency, but where trawling frequency was quantified, the mean trawling frequency  
239 was  $3.36 \text{ y}^{-1}$  in the trawled area (range: 0.2 to 12.9) and  $0.1 \text{ y}^{-1}$  in the reference area (range: 0.0 to  
240 0.4). TU studies were analysed using weighted Meta-Analysis via Linear Mixed-Effects Models with  
241 the log response-ratio (lnRR) for  $B$  as the response variable, by fitting the model  $\ln\text{RR}_{ij} = \ln(B_{\text{trawled}}$   
242  $_{ij}/B_{\text{reference } ij}) = k_j + \alpha_i + \epsilon_{ij}$  (eqn 3) with intercept  $k_j$  for  $T_{max}$  category  $j$ ;  $\alpha$  is the random study effect;  $\epsilon$  is  
243 the Gaussian error term. Studies were weighted by the inverse their variance estimate as described  
244 above. SI Text S1 describes how comparative TU studies were treated when trawling was not  
245 continuous but stopped or started in one of the areas during the study.

246

#### 247 Analysis of comparative trawl gradient (TG) studies

248  $RBS$  of seabed fauna (species or communities) can be estimated based on the assumptions that (1)  
249 the recovery of biomass or numbers (abundance  $B$ ) of biota after trawling is described by the logistic  
250 growth equation and the rate of recovery is therefore determined by the intrinsic rate of increase and  
251 (2) that the equilibrium solution of the logistic growth equation can be used to estimate abundance  $B$   
252 as a fraction of carrying capacity  $K$  for an environment subject to fishing disturbance ( $F$ ). This  
253 equilibrium solution of the logistic growth equation can be used to estimate  $RBS$ , which is abundance  
254  $B$  as a fraction of  $K$ . Gradient studies were used to estimate the intrinsic rate of increase  $r$  for species  
255 with different  $T_{max}$  following eqn 4, as described in detail in Hiddink *et al.* (2017).

256

257

$$258 \quad RBS = B/K = 1 - F d/r \quad (\text{eqn 4})$$

259

260

261 In this case, trawling effort ( $F$ ) is defined as the total area swept by trawl gear within a given area of  
262 seabed in one year divided by that area of seabed (units  $\text{y}^{-1}$ ). Depletion  $d$  is the fraction mortality per  
263 trawl pass estimated from experimental trawling studies. The intrinsic rate of population increase  $r$  is  
264 estimated from the slope of eqn 1. Eqn 1 suggests that  $r$  is constant, but in communities composed of  
265 species with a range of  $r$  values, trawling selects for species with faster life-histories that are more  
266 resilient, and therefore,  $r$  can be expected to increase with  $F$ . The  $r$  value that we want to estimate is  
267 for recovery of the community to its original unfished state, and therefore is the  $r$  at  $F = 0$ . We found  
268 that the relationship between community  $B/K$  and  $F$  for communities is well-approximated by a log-  
269 linear relationship. We, therefore, estimated  $r$  at  $F = 0$  and assuming a log-linear relationship between  
270  $RBS$  and  $F$  using a linear mixed effects model,  $\log_{10}(RBS_{ij}) = b1 F_i + b2 F_i \times T_{max\ ij} + \alpha_i + \epsilon_{ij}$  (eqn 5), where  
271  $b1$  is the general slope for effort  $F_i$  and  $b2$  is the slope for  $F_i \times T_{max\ ij}$  with  $T_{max}$  as continuous variable;  
272  $\alpha$  is the random study effect;  $\epsilon$  is the Gaussian error term. This model does not include the main  
273 effect of  $T_{max}$ , and although in most analyses this would be undesirable because it would lead to  
274 biased parameter estimates, it is appropriate here because the effect of trawling at  $F = 0$  has to be  
275 zero (Hector *et al.*, 2010). We modelled 'study' as a random effect, to account for the non-  
276 independence of observations within a study. Because the model would not converge using  $T_{max}$  as a  
277 categorical fixed effect, and because it is not possible to fit a continuous relationship between the  
278 intrinsic rate of increase and  $T_{max}$  with  $T_{max}$  as a categorical fixed effect,  $T_{max}$  categories were  
279 converted to a continuous variable following the method described in SI Text S2. Continuous  $T_{max}$  for  
280 each  $T_{max}$  category was 0.5 (<1yr), 2.1 (1-3yr), 6.5 (3-10yr) and 12.6 years (>10yr). The resulting  $r \sim T_{max}$   
281 relationship is not particularly sensitive to the exact values chosen here.

282

283 Because the slope in Eqn 4 equals  $d/r$ ,  $r$  can be estimated from the slope using the depletion  $d$   
284 estimate from the analysis of the experimental studies. We used the estimate for depletion rate  $d$  for  
285 biota with a lifespan of >1yr of  $d = 0.09$  as estimated in Figure 1A. However, it is necessary to take  
286 account of the log-linear relationship we used between  $RBS$  and  $F$ , and the uncertainty in the  
287 estimates (for a derivation see Hiddink *et al.*, 2017). The resulting uncertainty intervals capture both  
288 the probabilistic uncertainty resulting from the linear model and the vagueness inherent in the fuzzy  
289 coding of  $T_{max}$ .

290

291 Time to recovery was estimated as the time taken to recover from a given level of abundance to  
292 0.95K assuming logistic population recovery using the equations in Text S5 in Hiddink *et al.* (2017). We  
293 report the recovery time from 0.50 to 0.95K unless otherwise stated.

294

295 We checked the assumptions of all statistical analyses by visual inspection of the normalized residuals  
296 (Zuur *et al.*, 2009).

297



298 Assessment of benthic status

299 The performance of the estimates  $d$  and  $r$  for predicting benthic status was tested using field  
300 observations from the North Sea by applying the quantitative method of Pitcher *et al.* (2017). We  
301 modelled the predicted effect of trawling on  $B$  by  $T_{\max}$  category using Eqn 4. As an example, we  
302 describe how trawling affects the abundance of communities composed respectively of long and  
303 short-lived taxa in the Results. For this assessment we require  $F$ ,  $d$ ,  $K$  and  $r$  for each grid-cell (ideally  
304 around  $1 \text{ km}^2$ ) in the area of assessment, for each  $T_{\max}$  category. The values of  $r$  and  $d$  depend on the  
305 environmental conditions and fishing gears used in each cell respectively. Our analyses indicated that  
306 biota with a  $T_{\max} < 1\text{yr}$  are not consistently impacted by trawling, and therefore not included in our  
307 assessment.

308

309 Since we sought to compare sensitivity to trawling in different  $T_{\max}$  categories, we applied the  
310 estimates of  $d$ ,  $r$ ,  $B$  and  $K$  by  $T_{\max}$  category. For each  $T_{\max}$  category, a category-specific  $K$  was  
311 expressed as fraction of  $K$  for all  $T_{\max}$  categories ( $=1$ ). This fraction is habitat-specific and was derived  
312 by Rijnsdorp *et al.* (2018) from the cumulative biomass distribution of infaunal invertebrates by  $T_{\max}$   
313 category in the North Sea, as determined from grab and box core samples in untrawled locations (the  
314 effects of not including epifauna in these estimates is not known). Because we express category-  
315 specific  $K$  as a fraction,  $K = \sum K_j = 1$  (eqn 6) over  $T_{\max}$  categories  $j$  (although the approach could be  
316 based on absolute estimates of category-specific  $K$  if they were available).

317

318 The assessment method was tested by comparing the  $RBS$  predicted from trawling  
319 frequency with epifaunal biomass estimates at sites subject to different levels of trawling  
320 intensity in four North Sea areas with different habitat types (Dogger Bank - shallow sand,  
321 Hills - highly dynamic shallow sand, Northwest Rough - deeper sand, Silver Pit - deeper  
322 muddy sand). The sampling design and methods were described in Hiddink *et al.* (2006).  
323 Even at low trawling intensity, epifaunal biomass differs substantially among these areas of  
324 the North Sea (Hiddink *et al.* 2006). To address these differences in  $K$  a location effect was  
325 included in the linear model ( $B_{sk} = b1 RBS_{sk} + b2_k RBS_{sk} + \epsilon_{sk}$  (eqn 7), where  $b1$  is the general  
326 slope and  $b2$  the slope effect for location  $k$  with  $RBS_{sk}$  for replicate  $s$ ).

327

328 **Results**

329

330 Estimating depletion  $d$  from experimental studies

331 For  $T_{\max}$  categories 1-3, 3-10 and  $>10\text{yr}$  mean depletion was 9% (95% CI, 1.6-15.9%) per trawl pass  
332 (Figure 1a, Table S3) and the magnitude of depletion was not related to  $T_{\max}$ . For the  $T_{\max}$  category  
333  $<1\text{yr}$ , abundance increased immediately after trawling (Figure 1a, Table S3). This unexpected increase  
334 cannot be explained by an influx of scavengers as they were excluded.

335

#### 336 Comparative Trawled-Untrawled studies, TU

337 The abundance of benthic macroinvertebrates in trawled locations was lower than in reference  
338 locations for all  $T_{max}$  categories >1yr. There was no significant change in abundance for organisms  
339 <1yr (mean increase of 18%, 95% CI, -24% to +86%), but there was a significant decrease in  
340 abundance of 37% (95% CI, -57% to -7%) for organisms >10yr (Figure 1b). Whilst the effect of fishing  
341 was to reduce abundance of organisms, the response did not differ significantly among organisms  
342 with different  $T_{max}$  (test statistic  $Q_{M3} = 5.1$ ,  $P = 0.16$ ). Nevertheless, the estimate went from being not  
343 significantly different from zero for organisms <3yr (95% CI overlapping 0) to significant for organisms  
344 >3yr (Figure 1b).

345

#### 346 Estimating the intrinsic rate of increase $r$ from comparative gradient studies, TG

347 The effect of trawling on relative abundance was significant with a negative  $F \times T_{max}$  interaction ( $t_{226} =$   
348  $-3.24$ ,  $P = 0.0013$ ) indicating that organisms with different  $T_{max}$  respond differently to trawling. Each  
349 unit increase in  $F$  ( $\text{yr}^{-1}$ ) resulted in a mean decrease of 3.8% in abundance for organisms with  $T_{max} = 1-$   
350 3yr, 11.1% decrease for organisms with  $T_{max} = 3-10\text{yrs}$  and a 20.4% decrease in abundance for biota  
351 with a  $T_{max} >10$  yr (Figure 2a). A re-analysis excluding trawling intensities of  $>10 \text{ yr}^{-1}$  shows that the  
352 slopes of the  $\log_{10}(RBS)$  vs trawling frequency relationship for biota with a  $T_{max} = 1-3$  years were  
353 particularly influenced by data at  $F > 10\text{yr}^{-1}$  (Figure S4). Estimates of  $r$ , derived from the slopes in  
354 Figure 2 and the depletion rates  $d$  in Figure 1a, ranged from  $2.41 \text{ yr}^{-1}$  for biota with a  $T_{max} = 1-3\text{yr}$ ,  
355 through  $1.24 \text{ yr}^{-1}$  for biota with a  $T_{max} = 3-10\text{yr}$ , to  $r = 0.34 \text{ yr}^{-1}$  for biota with a  $T_{max} >10\text{yr}$ . These  $r$   
356 estimates translate to estimates of median time to recovery ( $T$ ) from  $0.5K$  to  $0.95K$  of  $1.2$  ( $0.55 - 2.78$   
357 uncertainty interval), of  $2.4$  ( $1.37 - 4.16$ ) and  $8.7$  ( $3.51 - 22.50$ ) years respectively (Figure 2b, Table  
358 S3). The relationship between  $1/T_{max}$  and  $r$  (Figure 3) is consistent with the linear relationship  
359 predicted by theory derived in Supplementary material Text S3 (or more precisely, based on these  
360 data we cannot reject theory that suggests there should be a linear relationship) and may be used to  
361 approximate  $r$  for species for which  $T_{max}$  is known .

362

#### 363 Assessment method and validation

364 Relative benthic status ( $RBS$ ) within a  $T_{max}$  category is predicted to decline linearly with trawling  
365 frequency (Figure 4a-c). The  $RBS$  of longer-lived biota declines faster because of a lower  $r$  (Figure 4c).  
366 The total biomass or  $RBS$  of a community with a high fraction of long-lived species therefore  
367 decreases further for a given increase in  $F$  because the long-lived taxa are much more sensitive to  
368 trawling (Figure 4d), while the  $RBS$  of a community with a high fraction of short-lived species shows a  
369 smaller response to trawling (Figure 4e).

370

371 The validation exercise demonstrated a positive correlation between the observed total epifaunal  
372 community biomass and predicted  $RBS$  over gradients of trawling intensity in the North Sea (Figure 5,

373 effect of *RBS*  $F_{1,29}=77.0$ ,  $P<0.001$ , interaction between *RBS* and location representing differences in *K*  
374 between areas,  $F_{3,29}=11.76$ ,  $P<0.001$ ,  $R^2=0.77$ ).

375

## 376 Discussion

377 This meta-analysis of experimental and comparative trawling impact studies demonstrates that  
378 decreases in the relative abundance of longer-lived fauna in trawled areas are greater than those of  
379 fauna with shorter lifespans. These differences are attributed to the lower intrinsic rates of increase  
380 of longer-lived fauna and, consequently, to their lower capacity to sustain trawling-induced mortality.  
381 The intrinsic rate of increase was estimated to be seven times lower for longer-lived (>10yr) than for  
382 shorter-lived fauna (1-3yr) and the corresponding time to recovery from 0.5*K* to 0.95*K* increased from  
383 1.22 yr (0.55 - 2.78 uncertainty interval) to 8.7 yr (3.51 – 22.50). The reported relationships can be  
384 used to estimate sensitivity to trawling for species or communities with a known longevity or  
385 longevity distribution and to estimate recovery times following trawling. While the community may  
386 not recover to the size and age composition associated with no trawling within this time, the recovery  
387 of abundance is likely to correlate to the recovery of ecosystem processes, such as bioturbation.  
388 Given that some of the longest-lived benthic biota in the longer-lived (>10yr) category can live for  
389 100+ yr (e.g. the ocean quahog *Arctica islandica*, 507 years, Butler *et al.*, 2013), full recovery to the  
390 pre-trawling age structure is unlikely to occur for such biota on any time-frame if any trawling  
391 persists. For taxa that form biogenic structures (e.g., the Ross worm *Sabellaria spinulosa* or the horse  
392 mussel *Modiolus modiolus*), the recovery rate of the biogenic structures will be slower than the  
393 recovery rates of the individual organisms that make up the structures (Rijnsdorp *et al.*, 2016).  
394 Therefore, where biogenic species are prevalent and the focus of assessment is on habitat status, the  
395 time required to rebuild the structures should be used to assess the sensitivity to trawling rather than  
396 the longevity of the individual organisms that build the structure. Consequently, the results of this  
397 study are most relevant for communities of free-living animals in and on soft and mobile sediments.

398

399 The relationship showing that the impact of trawling was greater for greater longevity in gradient  
400 studies was characterised by high uncertainty (eqn 5), in part because some relationships were driven  
401 by a few data points. Our systematic review methodology means that we deliberately included all  
402 available datasets worldwide that passed the screening criteria, so much of the uncertainty reflects  
403 real-world variation, caused, for example, by differences in the physical and ecological characteristics  
404 of different ecosystems. We consider our conclusions robust because the patterns observed in both  
405 the TU and TG studies are consistent, and match theoretical expectations. The large uncertainty  
406 nevertheless means that diverse responses are feasible because factors other than longevity  
407 contribute to faunal responses to bottom trawling.

408

409 Efforts to link longevity  $T_{max}$  and the intrinsic rate of increase  $r$  could be more rigorous and less vague  
410 (vagueness as expressed by the fuzzy coding process) if reliable point estimates of  $T_{max}$  were

411 available. However, in contrast with groups such as fish and birds, life-history data for benthic  
412 invertebrates are rather limited and ageing methods are uncertain or not developed for most groups,  
413 except bivalves. The fuzzy approach enabled assimilation of the patchy and inconsistent knowledge  
414 that does exist. However, unlike for probabilistic uncertainty, it is not clear how to assimilate and  
415 report the resulting vagueness in the subsequent meta-analyses or to carry forward the effects of the  
416 absolute loss of information that results from defuzzification into the next steps of the analysis.  
417 Consequently, the results of statistical tests may be ambiguous. For other steps in our method the  
418 treatment of uncertainty is not ambiguous as estimated uncertainty in  $T_{\max}$  (input) for any given  
419 species can be expressed as estimated uncertainty in  $r$  (output). In addition to working with  $T_{\max}$ ,  
420 there would, in the future, also be options to work with other life-history parameters that may be  
421 estimated more easily. These include using natural mortality, estimated directly or from size-based  
422 analyses of populations or the size or age at maturity. However, in practice, we are a long way from  
423 doing this for most benthic invertebrate species despite progress in some areas and for some groups  
424 (e.g. Palomares & Pauly, 2010). In general, for assessing fishing impacts and other effects of elevated  
425 mortality, a shift away from conventional sampling of species identity and numbers for multivariate  
426 community analyses towards a focus on life-history traits of populations would advance our capacity  
427 to assess fishing impacts and the impacts of other forms of seabed disturbance that result in mortality  
428 of benthic biota.

429

430 The linear relationship between reciprocal longevity and  $r$  is consistent with theoretical expectations,  
431 but this assessment was based on a comparison with just three data points given limitations of the  
432 longevity database and the absence of species-specific longevity estimates for most benthic species.  
433 Tests using more categories of longevity estimates or individual longevity estimates may in future give  
434 us more confidence that the data matches the theoretical expectation. Regardless of the shape of the  
435 relationship, communities with a greater mean longevity are shown to be more sensitive to a given  
436 rate of mortality due to trawling disturbance, and maintain a relatively lower  $RBS$ , consistent with  
437 observations. Assessments of impacts based on knowledge of the longevity distribution of the  
438 community are appealing because they account for the interaction of other forms of disturbance  
439 (primarily, but not exclusively, natural disturbance) and trawling disturbance. In the absence of  
440 trawling, the longevity distribution of benthic communities will vary among habitats, reflecting the  
441 influence the environment, including processes such natural disturbance of sediments by waves and  
442 tides, which affect the life-histories and mortality of species present (Rijnsdorp *et al.*, 2018). For  
443 example, large-scale sediment movements during winter storms in shallow areas, or summer hypoxia  
444 in the Baltic Sea, are associated with high natural mortality rates and with communities dominated by  
445 shorter-lived fauna with a higher  $r$ . Our results suggest these types of communities will be relatively  
446 more resilient to a given frequency of trawling disturbance than communities with low natural  
447 disturbance levels.

448

449 Synthesis and Applications

450 Quantitative descriptions of trawling impacts contribute to state of the environment reports,  
451 environmental risk assessments and advice on options for balancing the benefits of fish production  
452 and protection of the marine environment. At present, most descriptions are presented as overlays of  
453 the distribution of fishing activity on habitat distribution or sensitivity maps or spatial assessments of  
454 the relative reductions in total abundance resulting from trawling (e.g. Jennings *et al.*, 2012; Eigaard  
455 *et al.*, 2016). However, the new relationship between the longevity of species and  $r$  can be combined  
456 with high-resolution maps of trawling intensity to enable assessment of the risks and impacts of  
457 bottom trawling on the abundance of species or communities with defined longevity or recovery  
458 time. This would be achieved by generating estimates of  $r$  from longevity and using these within the  
459 *RBS* framework of Pitcher *et al.* (2017). It is important to take account of the substantial uncertainty  
460 around the  $r \sim$  longevity relationship by providing estimates of the mean as well as the upper and  
461 lower bounds of the effects of bottom trawling.

462

463 Quantification of the relationship between longevity and  $r$  is especially valuable; longevity is a species'  
464 characteristic that often matters for conservation and management agencies because it resonates  
465 with societal concerns about the loss of long-lived species (e.g. Garcia *et al.*, 2014). Our approach  
466 allows *RBS* and associated uncertainty to be estimated for any species with defined longevity deemed  
467 to be of 'concern' and for the presentation of options to show how alternative values of  $F$  would  
468 affect the *RBS* of species with different longevity. Outputs of such analyses will support demands from  
469 conservation and management agencies, the fishing industry and certification bodies for better  
470 quantitative information on the impacts of bottom trawling and the consequences of adopting  
471 alternate management options.

472

473 **Authors' Contributions**

474 JH, AR and SJ conceived the ideas and designed methodology. All authors contributed to data-  
475 collation. JH analysed the data and led the writing of the manuscript. All authors contributed critically  
476 to the drafts and gave final approval for publication.

477

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485 defuzzification.

486

487 **Data accessibility**

488 Comparative study and trait data available via the Dryad Digital Repository  
489 <https://doi.org/10.5061/dryad.426658h> (Hiddink *et al.*, 2018) and experimental study data available  
490 from (Sciberras *et al.*, 2018)

491

492

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593  
594  
595 **Figures**

596

597



598 Figure 1. The effect of bottom trawling on the abundance of taxa (as log response ratio - lnRR) with  
599 different longevities (mean, 95% confidence interval). a) Effect of a single trawl pass in experimental  
600 studies. The grey line and shaded bar indicate the mean effect ( $d = 0.09$ ) and confidence interval for  
601 biota with  $T_{max} > 1yr$ . b) The difference between control and impact locations in comparative TU  
602 studies of chronic trawling. If the 95% confidence interval overlaps  $\lnRR = 0$ , the effect was not  
603 significant. The right-hand axis indicates the responses in % changes. The number of studies included  
604 in each estimate is given below each bar.

605

606

607 Figure 2. The response of different longevity categories to trawling: (a) The relationship between  
608 trawling frequency and abundance (biomass or numbers) of taxa in gradient studies; (b) estimated  
609 recovery time to  $0.95K$  for depleted abundance as a function of estimated intrinsic rate of increase  $r$   
610 and fraction depletion ( $= -(B/K-1)$ ). In (b) the shaded areas indicate the 95% uncertainty intervals for  
611 estimates. See Figure S4 for an analysis with data points at trawling frequency  $> 10 \text{ yr}^{-1}$  removed.

612

613 Figure 3. Relationship between  $r$  and  $T_{max}$  estimated from gradient studies ( $r = 5.31 / \text{longevity}$ ,  $R^2 =$   
614  $0.96$ ,  $F_{1,1} = 73.9$ ,  $P = 0.013$ ). The points and error bars are  $r$  estimates and their 95% confidence  
615 intervals, while the solid line is the fitted regression line. The shaded areas indicate the regression fits  
616 through the upper and lower confidence intervals of the data (upper:  $r = 11.44 / \text{longevity}$ , lower:  $r =$   
617  $2.43 / \text{longevity}$ ).

618

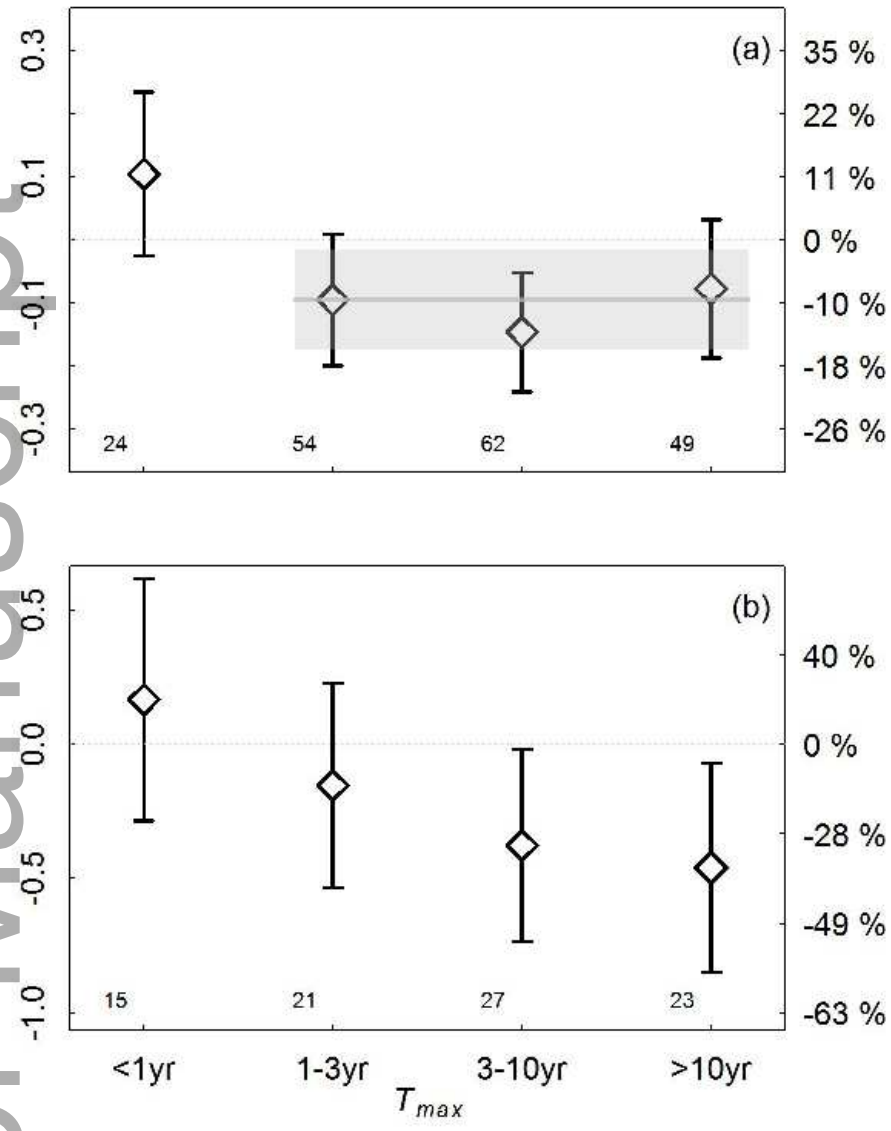
619

620 Figure 4. The response to trawling of biota of different longevity. (a-c): The predicted relative benthic  
621 status  $RBS$  of three longevity categories in response to trawling with their confidence intervals. (d-e):  
622 The response of predicted  $RBS$  to trawling for two hypothetical benthic communities with different  
623 longevity distributions: (d) scenario for a community with a large fraction of long-lived fauna, (e)  
624 scenario for a community with a high fraction of short-lived fauna.

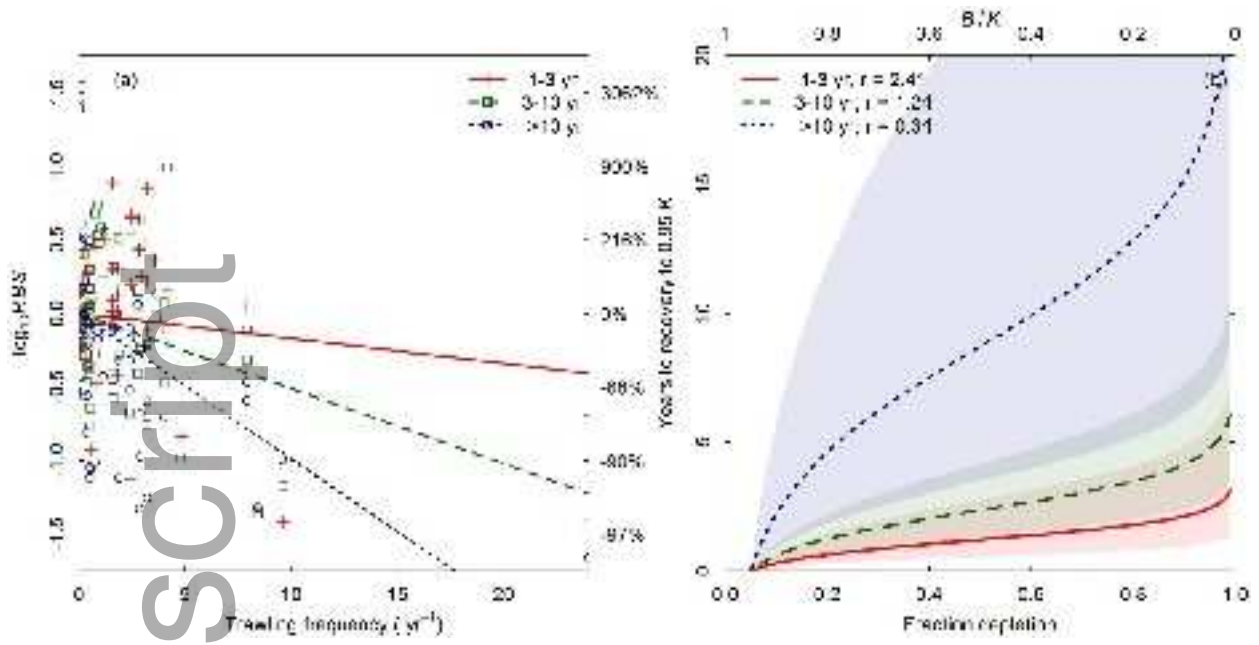
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626

627 Figure 5. Relationship between the observed total community biomass of epifauna and predicted  
628 relative benthic status  $RBS$  in four fishing grounds in the North Sea with gradients of trawling  
629 intensity. Black lines indicate the fitted regression, grey lines the 95% confidence intervals. Bubble  
630 sizes are proportional to the trawling intensity ( $\text{yr}^{-1}$ ) at each site.

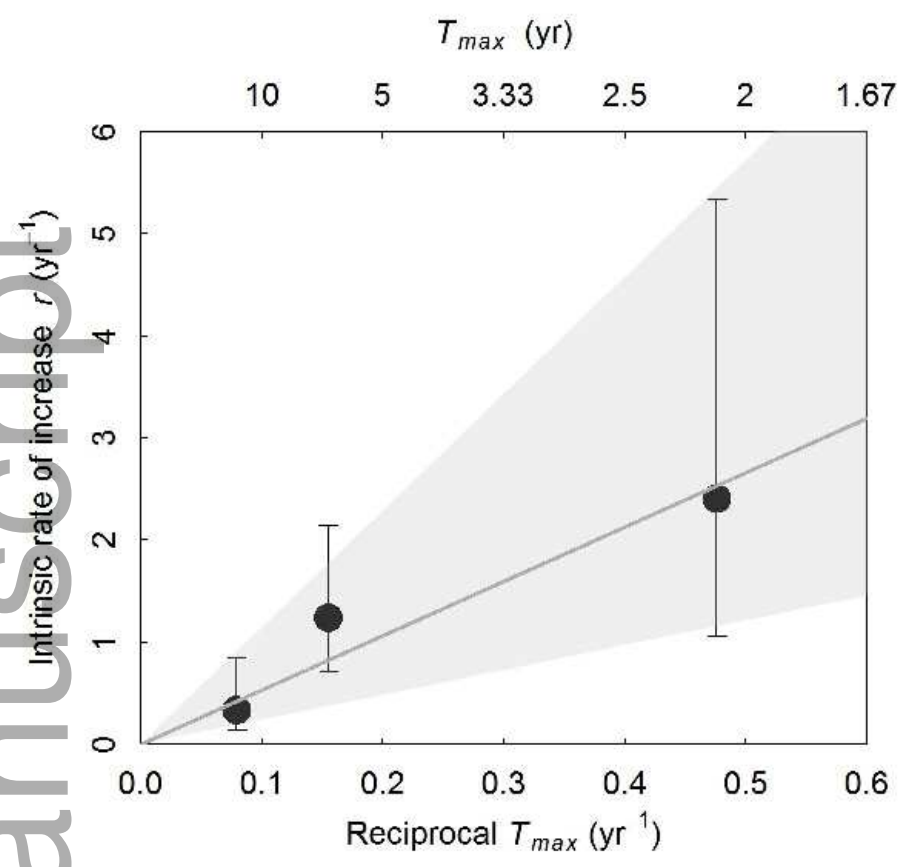


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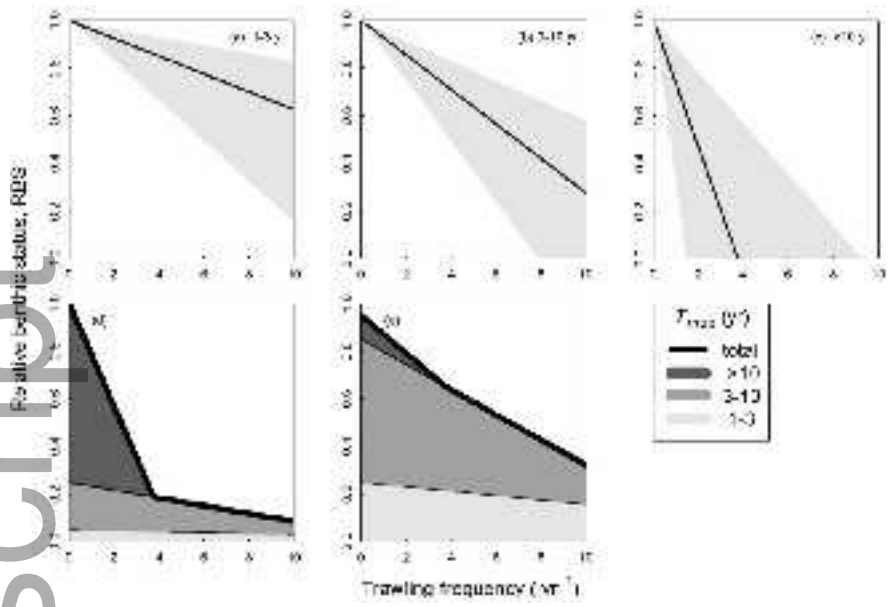


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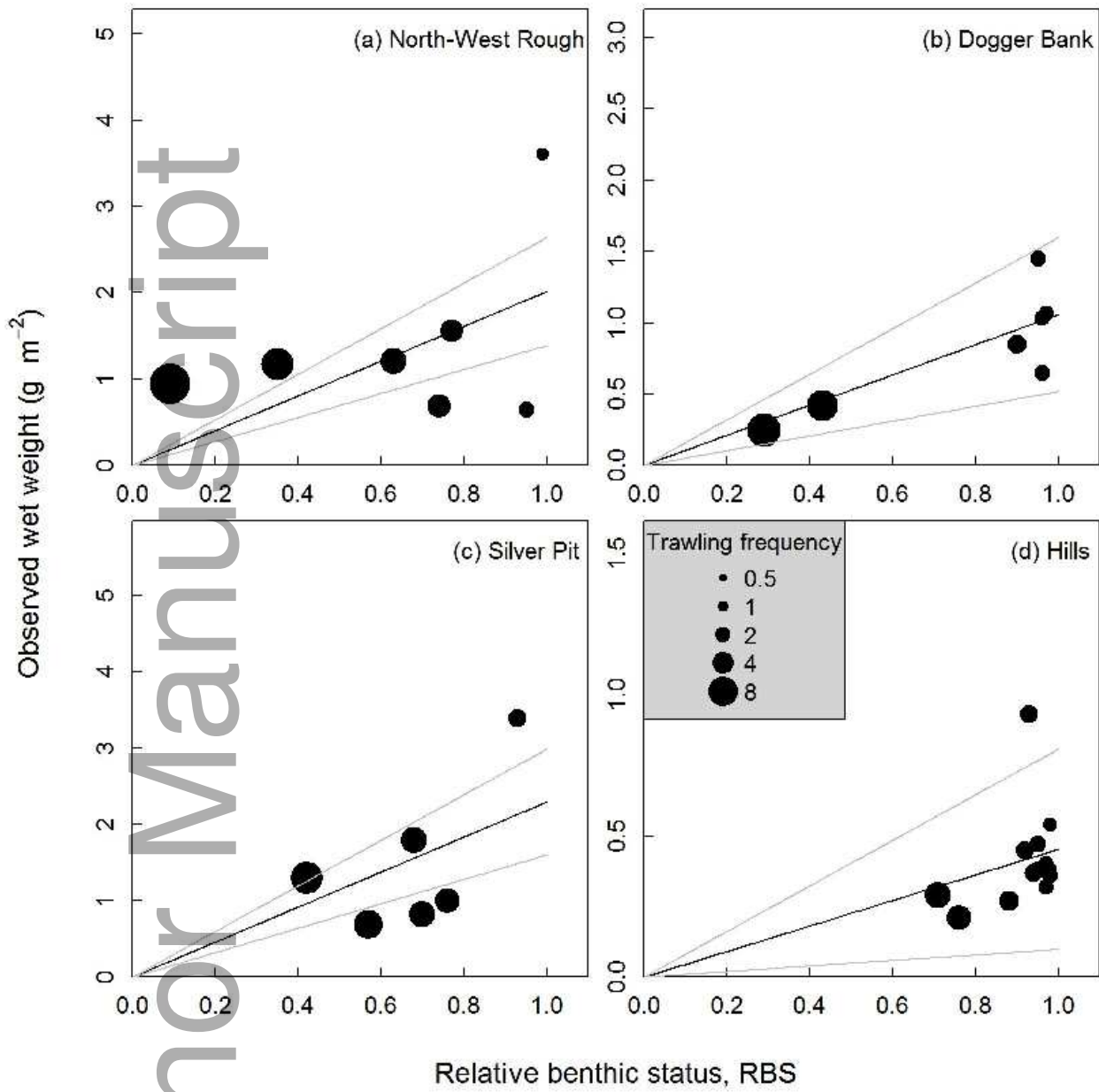
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