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2	MR. HAIKUN XU (Orcid ID : 0000-0002-1269-6012)
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8	Evaluating the utility of the Gulf Stream Index for predicting recruitment
9	of Southern New England-Mid Atlantic yellowtail flounder
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11	HAIKUN XU ^{1a*} , TIMOTHY J. MILLER ² , SULTAN HAMEED ¹ , LARRY A.
12	ALADE ² , JANET A. NYE ¹
13	
14	¹ School of Marine and Atmospheric Science, Stony Brook University, Stony Brook, New
15	York, USA
16	² NOAA Northeast Fisheries Science Center, Woods Hole, Massachusetts, USA
17	Č
18	^a Current address: School of Aquatic and Fishery Sciences, University of Washington,
19	Seattle, Washington, USA
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36 Abstract

The justification for incorporating environmental effects into fisheries stock assessment 37 models has been investigated and debated for a long time. Recently, a state-space age-38 39 structured assessment model which includes the stochastic change in the environmental covariate over time and its effect on recruitment was developed for Southern New England-40 Mid Atlantic yellowtail flounder (Limanda ferruginea). In this paper, we first investigated 41 the correlations of environmental covariates with Southern New England-Mid Atlantic 42 vellowtail flounder recruitment deviations. The covariate that was most strongly correlated 43 with the recruitment deviations was then incorporated into the state-space model and 44 alternative effects on the stock-recruit relationship were estimated and compared. For the 45 model that performed best as measured by Akaike information criterion, we also compared 46 the estimates and predictions of various population attributes and biological reference 47 points with those from an otherwise identical model without the environmental covariate in 48 the stock-recruit function. We found that the estimates of population parameters are similar 49 for the two models but the predictions differed substantially. To evaluate which model 50 provided more reliable predictions, we quantitatively compared the prediction skill of the 51 two models by generating two series of retrospective predictions. Comparison of the 52 53 retrospective prediction pattern suggested that from an average point of view, the 54 environmentally-explicit model can provide more accurate near-term recruitment predictions especially the one-year ahead recruitment prediction. While for a specific near-55 56 term recruitment prediction from the environmentally-explicit model, the accuracy of 57 which is largely determined by the accuracy of the corresponding environment prediction the model provides. 58

59 Key Words: state-space model; Gulf Stream Index; recruitment; stock assessment;
60 Southern New England-Mid Atlantic yellowtail flounder; retrospective prediction;
61 prediction skill

62 INTRODUCTION

Whether we should incorporate environmental drivers explicitly into fisheries stock 63 64 assessment models has been investigated and debated for a long time (Walters & Collie, 1988, Haltuch & Punt, 2011, Punt et al., 2014, Szuwalski & Hollowed, 2016). Several 65 66 recent studies have shown that the environment cannot be ignored to better understand 67 some of the factors influencing changes in stock productivity of many fish populations (Vert-pre et al., 2013, Essington et al., 2015, Szuwalski et al., 2015), but incorporating 68 environmental drivers into stock assessment models remains elusive. However, Miller et al. 69 70 (2016) recently developed a state-space age-structured assessment model that allows for environmental covariates in the stock-recruit function. According to model fit and 71 72 retrospective pattern, they concluded that incorporating the effect of Mid-Atlantic cold pool dynamics on Southern New England-Mid Atlantic (SNEMA) yellowtail flounder 73 recruitment can improve model performance. 74

75 The physical environment in the SNEMA region is highly dynamic owing to variability 76 in both atmospheric and oceanographic processes. The North Atlantic Oscillation or NAO 77 is an atmospheric process known to have a profound effect on water temperature, storm 78 tracks and northern North Atlantic ecosystems (Drinkwater et al., 2003). The NAO has a 79 lagged effect on surface and bottom water temperature in the Northeast US as the signal propagates from the Labrador Sea (Mountain, 2012, Xu et al., 2015). In the Northeast US, 80 81 two current systems collide where cold water emanating from the Labrador Current to the 82 north and warm water moving from the south and east in the Gulf Stream meet (Greene et al., 2013). The position of the north wall of the Gulf Stream is the best leading indicator of 83 the relative strength of cold Labrador slope water and warm subtropical water and is highly 84 correlated with temperature on the shelf (Nye et al. 2011). In the SNEMA region, the Mid-85 86 Atlantic cold pool is a distinct remnant cold winter water at depth occurring from late spring to early fall, formed as a result of the strong seasonal thermocline in the SNEMA 87 region (Houghton et al., 1982). 88

Determining the cause of the low recruitment since the 1990's was argued to be one of the main sources of uncertainty in the most recent SNEMA yellowtail flounder benchmark assessment (NEFSC, 2012). The persistent low recruitment since the 1990's resulted in the

92 exploration of two recruitment scenarios in the benchmark assessment to account for the notable drop in stock productivity. The first scenario assumed that unfavorable 93 94 environmental conditions reduced stock productivity significantly since the 1990's such that the stock was considered rebuilt (albeit at a low level) and not overfished. By contrast, 95 the second scenario also accounted for greater historical recruitments prior to the 1990's 96 such that the stock was considered overfished. Therefore, making clear what processes are 97 responsible for the recruitment drop since the 1990's will be invaluable to improving 98 99 current understanding of the population dynamics and determining the stock status of SNEMA yellowtail flounder. 100

101 Recruitment of SNEMA yellowtail flounder may be dependent on temperature condition during the early life stages. SNEMA yellowtail flounder usually spawn in spring 102 103 and early summer, with a peak in May (NEFSC, 2012). Their fertilized eggs float at the surface for about 2 months, then larvae metamorphosis occurs and juveniles settle to the 104 105 bottom of the continental shelf (Sullivan et al., 2000). Both field observations (Sullivan et al., 2005, Sullivan et al., 2000) and modeling studies (Miller et al., 2016) have shown that 106 107 recruitment of SNEMA vellowtail flounder is closely related to the dynamics of the Mid-Atlantic cold pool. In the field, Sullivan et al. (2000) found that the SNEMA stock heavily 108 109 relies on the cold bottom water in the Mid-Atlantic cold pool as a thermal refuge in summer when water temperature reaches the annual maximum. Later on, Sullivan et al. (2005) also 110 found that stronger young-of-the-year cohorts were observed with colder and longer-lasting 111 Mid-Atlantic cold pools. 112

113 Based on the survey evidence from the field, incorporating the Mid-Atlantic cold pool dynamics in SNEMA vellowtail flounder stock assessment model was investigated in the 114 115 last benchmark assessment, attempting to explain the low productivity level since the 1990's (NEFSC, 2012). The Cold Pool Index (CPI), defined as the first principle 116 component of the Mid-Atlantic cold pool temperature and area matrix, was chosen in the 117 study to represent the thermal condition in the Mid-Atlantic cold pool. A negative 118 119 correlation was found between the CPI and the recruitment deviations from the Beverton-120 Holt stock-recruit function. Also, the CPI-incorporated Beverton-Holt stock-recruit function was found to fit data better than the traditional Beverton-Holt stock-recruit function without 121

any environmental covariate. Although this preliminary analysis demonstrated the negative
effect of cold pool temperature on SNEMA yellowtail flounder recruitment, the CPI was
not accepted in the baseline run in the last benchmark assessment as the low productivity
level since 1990 could not be fully explained by the CPI alone (NEFSC, 2012).

126 After the last benchmark assessment, the effect of CPI on SNEMA yellowtail flounder 127 recruitment was further investigated in a new state-space age-structured assessment model (Miller et al., 2016). State-space models have the advantage of separately modeling time-128 129 varying stochastic processes and observation errors, and have recently become increasingly 130 popular due to the developments of software packages that can efficiently handle such models (Nielsen & Berg, 2014). This state-space assessment model allows CPI effects on 131 recruitment, and assumes stochastic changes of the CPI over time and accounts for errors in 132 133 the annual CPI observations (Miller et al., 2016). Comparison of the state-space models with and without CPI effects on recruitment indicated that the former had lower AIC and 134 135 provided less retrospective patterns (Mohn, 1999) in terminal year estimates of population attributes. This study further emphasized the importance of the environment in modulating 136 137 SNEMA yellowtail flounder recruitment.

In addition to understanding stock productivity and determining stock status, another 138 goal in fisheries stock assessment is to predict stock biomass trajectories under various 139 140 harvest scenarios (Quinn & Deriso, 1999, Haddon, 2010). Prediction skill is a term 141 popularly used in climate science referring to the ability of a model in predicting climate variables (Boer et al., 2013). It is usually assessed by generating a series of historical 142 143 climate predictions and comparing them with the corresponding observations (Meehl et al., 144 2009). Although a good prediction skill in historical predictions does not necessarily 145 guarantee a good prediction skill for the future, the historical prediction skill can inform us 146 about the uncertainty in model predictions for the future. In fisheries stock assessment, retrospective analysis is often done to evaluate the systematic bias in population estimates 147 148 in the terminal year when additional years of data are added (Mohn, 1999). Borrowing the 149 idea of model prediction skill from climate science, a series of retrospective prediction can 150 also be generated in a similar way for fisheries stock assessment models to evaluate the skill of models in predicting population attributes. Indeed, Brooks and Legault (2015) 151

recently have used the idea of the retrospective prediction to evaluate the predictive performance of New England groundfish stock assessment models, although their retrospective prediction scheme is different from that typically used in climate science.

155 The first objective of this paper was to examine the correlation of various atmospheric 156 and oceanographic covariates with SNEMA yellowtail flounder recruitment deviations. Until now, the examples of incorporating environmental effects directly into fisheries stock 157 assessment and management are still very limited (but see Schirripa, 2007, Hill et al., 2011, 158 and Miller et al., 2016). Thus, our second objective was to comprehensively compare the 159 160 estimates and predictions from the state-space assessment models with and without the 161 most strongly correlated climate process in the stock-recruit function. This comparison provided suggestions for future fisheries studies that incorporate environmental effects into 162 163 stock assessment models.

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165 DATA AND METHODS

166 The correlation of various large-scale atmospheric and oceanographic climate indices with 167 annual deviations in recruitment for the SNEMA yellowtail flounder stock were examined. Hydrology and ecosystem dynamics on the Northeast US Continental Shelf have been 168 known to be affected by the NAO - the dominant and most influential atmospheric 169 170 oscillation mode in the North Atlantic (Drinkwater et al., 2003, Mountain, 2012). The NAO index represents the scaled pressure difference between the two pressure centers of the 171 NAO, namely the Azores high (AH) pressure center and the Icelandic low (IL) pressure 172 center (Hameed & Piontkovski, 2004). The large-scale atmospheric indices investigated in 173 this study include the winter NAO index 174 from (http://www.cpc.ncep.noaa.gov/products/precip/CWlink/pna/nao.shtml) as well as the 175 pressures, latitudes, and longitudes of the IL pressure and AH pressure centers from 176 Hameed and Piontkovski (2004). In this study, the reason for including various indices 177 related to the two NAO pressure centers instead of only considering the NAO index is that 178 treating the two pressure centers as two separate systems can potentially explain a larger 179 180 portion of the NAO-induced variance (Hameed & Piontkovski, 2004). The oceanographic indices investigated in this study include two Gulf Stream related indices (Joyce & Zhang, 181 182 2010, Taylor & Stephens, 1998) and the previously investigated CPI. The two Gulf Stream 183 related indices are 1) the Gulf Stream Index (GSI), calculated by using water temperature at 200 m depth (Joyce & Zhang, 2010); and 2) the Gulf Stream North Wall (GSNW), 184 calculated by using sea surface temperature observation (Taylor & Stephens, 1998). The 185 186 two indices differ in the data source but both quantify the latitudinal position of the Gulf Stream: one at the surface and one at 200 m depth. Particularly, the GSI was shown to be a 187 good indicator of bottom temperature condition in the SNEMA region (Nye et al., 2011). 188

Assuming that the recruitment deviations from fitting to the Beverton-Holt stock-recruit function are at least partially related to environmental processes, the Pearson crosscorrelations between the recruitment deviations in log-space and environmental indices that lead recruitment by zero to two years were calculated. The lead time was designed to account for the delayed effects of some large-scale climate processes on the local environment in the SNEMA region. The recruitment and spawning stock biomass (SSB) time series used to fit to the Beverton-Holt stock-recruit function were extracted from the

196 baseline run in the most recent benchmark assessment (NEFSC, 2012). Recognizing that a significant correlation does not necessarily indicate causation (Hilborn, 2016) and the time 197 198 series used in the correlation analysis are from stock assessment models which are subject to various sources of uncertainty and bias (Brooks & Deroba, 2015), we also incorporated 199 the most significantly correlated environmental indices internally in the state-space 200 assessment model to compare model performance with respect to AIC and retrospective 201 bias. Following the method in Burnham and Anderson (2002), the Akaike weight was also 202 calculated for each model using AIC. 203

As in Miller et al. (2016), the environmental covariate (x) at time t is modeled as a random walk for t > 2:

$$x_t | x_{t-1} \sim N(x_{t-1}, \sigma_x^2)$$
 (1)

and the observation of which is

$$y_t | x_t \sim N\left(x_t, \sigma_y^2\right)$$
 (2)

Essentially, the incorporated environmental covariate is a random walk process with white noise. In this state-space assessment model, the environmental covariate can be incorporated into the stock-recruit function and therefore adjust the expected recruitment at time t:

$$log N_{t,1} | x_{t-1}, SSB_{t-1} \sim \mathcal{N}\left(g, \sigma_{N_1}^2\right)$$
(3)

where g is an environmentally-explicit Beverton-Holt stock-recruit function. Throughout 211 this paper, recruitment $(N_{t,1})$ is used to refer to the abundance of age-1 fish unless 212 otherwise noted. The environmental covariate (x) and abundance-at-age (N) are both 213 random-effect variables and estimated in ADMB based on empirical Bayes (Fournier et al., 214 215 2012). The state-space assessment model fitted to the data between 1973-2011, including 216 three abundance indices from bottom trawl surveys, two spawning stock indices from ichthyoplankton surveys, commercial catch, and annual age composition observations from 217 the three bottom trawl surveys and the commercial catch (see Miller et al. 2016). 218 219 Miller et al. (2016) found that performance of the state-space assessment model was improved by including CPI effect on recruitment. As the CPI was hypothesized to affect the 220

221 carrying capacity for pre-recruits, CPI was modeled as a "limiting factor" in the Beverton-222 Holt stock-recruit function (see Iles and Beverton 1998). However, Iles and Beverton (1998) 223 also considered effects of the environment on spawner density and(or) mortality 224 (controlling factor) and on pre-recruit mortality and(or) growth (masking factor) (Fry, 1971, 225 Neill et al., 1994). To evaluate the sensitivity of the state-space assessment model to the form of the environmentally-explicit Beverton-Holt stock-recruit function, we also 226 227 incorporated the most strongly correlated environmental covariate into the Beverton-Holt stock-recruit function as a controlling and masking factor. 228

229 After finding the best fitting environmentally-explicit stock-recruit function for SNEMA vellowtail flounder, we compared the estimates and predictions of three 230 population attributes (recruitment, SSB, and fully-selected fishing mortality (F)) and two 231 232 biological reference points (maximum sustainable yield (MSY) and SSB_{MSY}) provided by the two models with and without the environmental effect on recruitment. Both models 233 234 made five-year predictions for years 2012-2016 under the assumption that future F is at the level that produce the MSY (F_{MSY}). To evaluate which model can provide more reliable 235 236 population predictions for 2012-2016, the prediction skill of the two models were compared 237 by generating a series of retrospective predictions for each model.

238 In fisheries stock assessments, assessing model performance by generating a series of 239 retrospective assessments is not a new idea. Terminal year population estimates are of key 240 importance to stock status determination and harvest management, but are usually subject to retrospective bias (Mohn, 1999, Legault, 2009). Retrospective bias arises due to 241 misspecification in stock assessment models (Legault, 2009) and is usually evaluated in the 242 243 corresponding retrospective pattern generated through refitting the model to the data after 244 removing (peeling) its terminal year sequentially for several times (Mohn, 1999). Similar to 245 assessing the retrospective bias by generating a series of retrospective peels, the prediction skill of stock assessment models can also be assessed by generating a series of retrospective 246 predictions using the "true" F during the prediction years and then comparing the 247 248 retrospective predictions of population attributes with the corresponding "true" values. The 249 "true" values in this paper are defined as the estimates from the assessment using the full data from 1973 to 2011. The "true" F rather than "true" catch was specified in predictions 250

251 considering that the "true" catch can exceed the estimated population biomass and result in 252 negative population abundance (Brooks & Legault, 2015). As this paper is focused on 253 understanding environmental effects on SNEMA yellowtail flounder recruitment, we chose recruitment as the target population attribute in retrospective predictions. For each model, 254 13 retrospective three-year predictions were generated in a way similar to retrospective 255 peeling: first, the state-space model fitted to the data between 1973-2008 with three years 256 257 (2009-2012) recruitment predicted; then, the state-space model fitted to the data between 1973-2007 with three years (2008-2011) recruitment predicted; repeated this process in the 258 same manner until the state-space model fitted to the data between 1973-1996 with three 259 260 years (1997-1999) recruitment predicted. The mean relative difference (MRD) and mean absolute relative difference (MARD) of the 13 retrospective recruitment predictions from 261 262 the "true" recruitment were calculated for each prediction lead time (from one year to three years) to quantitatively compare the retrospective prediction skill among the candidate 263 models. The MRD and MARD for prediction lead year t were calculated as 264

MRD_t =
$$\frac{1}{13} \sum_{i=1996}^{2008} \frac{\theta_{i,t} - \Theta_{i+t}}{\Theta_{i+t}}$$
 (4)
MARD_t = $\frac{1}{13} \sum_{i=1996}^{2008} \frac{|\theta_{i,t} - \Theta_{i+t}|}{\Theta_{i+t}}$ (5)

respectively. $\theta_{i,t}$ is the *t*th recruitment prediction from the state-space model fitted to the data up to year *i*, and θ_{i+t} is the corresponding "true" recruitment the full data in year *i* + *t*.

267

268 **RESULTS**



While the correlation of recruitment deviations with the CPI was significant and stronger than with any atmospheric indices, the strongest correlations were observed with the two Gulf Stream related indices, especially the GSI (Table 1; Fig. 1). The GSI and recruitment deviations were negatively correlated with a lag of one year (Table 1). In other words, the 274 latitudinal position of the Gulf Stream negatively impacted the abundance of age 1 fish one275 year later.

The GSI and CPI were then separately incorporated in the state-space assessment model in the Beverton-Holt stock-recruit function either as a limiting, controlling, or masking factor. Based on AIC, the GSI-incorporated models performed better than the CPIincorporated models (Table 2), which is in agreement with the stronger correlation of GSI with recruitment deviations. Moreover, among the GSI-incorporated models, the one with GSI assumed to be a limiting factor performed best and more than two times more likely to be the best model than is the second-best one with GSI assumed to be a controlling factor.

Myers (1998) noticed that only 1 out of 47 environmental-recruitment correlations was 283 284 used in routine stock assessments at the time of his study, and moreover, most 285 environmental-recruitment correlations broke down when more years of data were added. As another measure of model performance, we compared the retrospective AIC values from 286 287 the best model above and from the best model Miller et al. (2016) found (with the CPI 288 assumed to be a limiting factor). Consistently in all seven peels, the GSI-incorporated 289 model had smaller AIC values than the CPI-incorporated model (Table 3), indicating that the GSI-incorporated model consistently outperformed the CPI-incorporated model over 290 291 time. The retrospective estimates of the environmental link parameter showed that the sign and degree of the GSI effect on recruitment were also consistent as addition years of data 292 293 were included (Table 3). The two models were also compared with respect to the Mohn's ρ , which was defined in this study as the mean of the seven relative differences in each 294 295 terminal year. Compared to the CPI-incorporated model, the GSI-incorporated model had larger Mohn's ρ for all three population attributes while the differences in Mohn's ρ are 296 negligible for SB and F (Table 4). 297

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299 *Effects of the GSI on predicting recruitment*

The estimated stock-recruit function in the two models with (R(SSB&GSI)) and without (R(SSB)) the GSI effect on recruitment was first compared. When recruitment is solely a function of SSB, the recruitment expected from a given SSB is always constant. However, when recruitment is also a function of environment and the environmental effect is strong, recruitment can vary dramatically with the environment (Fig. 2). Given that R(SSB) and R(SSB&GSI) differ in the stock-recruit function, the estimates and especially predictions of population attributes and biological reference points provided by the two models are expected to be different. It is important to note that R(SSB) was treated as a base model in this study to evaluate the consequences of incorporating an environmental covariate into a stock assessment model, it however was not considered in the last benchmark assessment because the stock-recruit relationship was not used (NEFSC, 2012).

311 As for recruitment, R(SSB) and R(SSB&GSI) provided similar estimates before 2011 (except in some individual years such as 1975-1980) but notably different five-year 312 313 predictions for 2012-2016 (Fig. 3a). Although under the same harvest scenario, R(SSB) predicted that future recruitment will be increasingly higher while R(SSB&GSI) predicted 314 315 that future recruitment will be persistently lower than that estimated in the terminal year. The SSB estimates provided by the two models were also similar and the SSB predictions 316 317 provided by the two models were also notably different (Fig. 3b). Specifically. R(SSB) provided higher SSB predictions than R(SSB&GSI), primarily due to higher recruitment 318 319 predictions from R(SSB). The recent unfavorable environmental conditions negatively affected recruitment and resulted in a decreasing stock size is predicted by R(SSB&GSI) 320 321 for the next five years. By contrast, the higher SSB predicted by R(SSB) provides an optimistic view that the stock size will slowly rebuild over the next five years. Same as 322 323 recruitment and SSB, F was also estimated to be similar in R(SSB) and R(SSB&GSI) (Fig.

324 3c).

While R(SSB) and R(SSB&GSI) provided similar F estimates, the estimated F_{MSY} from 325 the two models were notably different (Fig. 3c). Specifically, the F_{MSY} estimate from 326 327 R(SSB) is very close to the reference point from the most recent benchmark assessment $(F_{40\%})$, but notably smaller than that from R(SSB&GSI). Both MSY and SSB_{MSY} in this 328 state-space assessment model are functions of the incorporated environmental covariate, so 329 their estimates from R(SSB&GSI) varied annually with the GSI. The MSY estimates from 330 331 R(SSB&GSI) were relatively low since the 1990's (Fig. 3d) as unfavorable environmental 332 conditions (indicated by high GSI values) were more frequent during that time (Fig. 1). Note that the MSY estimates from R(SSB) is time-invariant also due to the fact that the 333

334 terminal year weight-at-age and selectivity were used in the calculation of MSY for all 335 previous years. In this case, only using a constant set of values for these allows us to look at 336 how MSY varies annually just due to the annual fluctuation in the environment. As for 337 $\log(SSB/SSB_{MSY})$, both the estimates and predictions from the two models differ substantially (Fig. 3e). When the GSI values were low (i.e., environmental conditions were 338 favorable) before 1990, the stock was estimated to be more productive and therefore 339 log(SSB/SSB_{MSY}) was estimated to be lower from R(SSB&GSI) than from R(SSB). 340 Conversely, when the GSI values were high (i.e., environmental conditions were 341 unfavorable) after 1990, the stock was estimated to be less productive and therefore 342 log(SSB/SSB_{MSY}) was estimated to be higher from R(SSB&GSI) than from R(SSB). 343

Overall, both models over-predicted recruitment after 1996, except in a few years 344 between 2003-2006 (Fig. 4). The recruitment predictions from R(SSB&GSI) were 345 346 generally higher than those from R(SSB) when the terminal year GSI values were lower than the long-term average and vice versa, as a result of the negative correlation between 347 the GSI and recruitment deviations. As expected, the recruitment predictions from either 348 349 model become more biased (larger MRD) and less accurate (larger MARD) as prediction lead time increases (Table 5). Generally speaking, incorporating the GSI into the stock-350 351 recruit function improved the accuracy of recruitment predictions as suggested by a smaller 352 MARD for R(SSB&GSI). Also, it reduced the bias in recruitment predictions as suggested by a smaller MRD for R(SSB&GSI). The importance of the incorporation to recruitment 353 354 predictions is most pronounced in the first prediction year and finally becomes negligible in 355 the third prediction year.

356 We compared each retrospective prediction pair from the two models and found that the 357 relative performance of the two models in predicting recruitment had dramatic year-to-year fluctuations and neither model consistently outperformed the other in predicting 358 recruitment (Fig. 4). Although both MRD and MARD are smaller for R(SSB&GSI), the 359 360 comparison indicates that R(SSB&GSI) fails to provide better recruitment prediction in all 361 13 retrospective prediction cases. Because the GSI is modeled as a random walk, the best 362 future prediction is the same as the estimate in the last observed year of data. However, the GSI had very large interannual fluctuation relative to the long-term average, so the annual 363

364 GSI prediction could be very different from the "true" GSI. As a consequence, the 365 recruitment predictions from R(SSB&GSI) could also be very different from the "true" 366 recruitment. For example, if the GSI did not change notably in the three prediction years (e.g., 2000-2002), the GSI and its effect on recruitment were found to be more accurately 367 predicted by R(SSB&GSI). In this case, the recruitment predictions from R(SSB&GSI) (the 368 bold solid line starting in 2000) were more similar to the "true" recruitments than the 369 370 predictions from R(SSB) which did not account for the unfavorable environmental conditions (the bold dashed line starting in 2000). In contrast, if the GSI changed 371 dramatically in the three prediction years (e.g., 2006-2008), the GSI and its effect on 372 373 recruitment were poorly predicted by R(SSB&GSI). In this case, the recruitment 374 predictions from R(SSB&GSI) (the bolded solid line starting in 2006) were further from the 375 "true" recruitments than the predictions from R(SSB) (the bold dashed line starting in 2006). 376

377 **DISCUSSION**

378 This paper uses the state-space age-structured assessment model from Miller et al. (2016) to 379 explore other environmental covariates to explain SNEMA yellowtail flounder recruitment 380 variability and builds on it to evaluate alternative effects of environmental covariates within the Beverton-Holt stock-recruit relationship. Furthermore, we explore the ability of 381 382 environmental covariates to improve prediction of future recruitments. Specifically, we 383 incorporated indicators of climate variability directly into the stock-recruit function as a 384 limiting, controlling, and masking factor, respectively. Compared to the model without any environmental covariate, the model with GSI as a limiting factor performed better with 385 386 respect to AIC and provided recruitment predictions that were closer to the "true" recruitments estimated from the full data with respect to both MRD and MARD. However, 387 388 the recruitment predictions provided by the model with GSI were not closer to the "true" recruitments in every single retrospective prediction case. Indeed, we found that 389 390 recruitment predictions from the model with GSI can be further from the "true" values when GSI predictions from that model are far away from the "true" GSI. Therefore, we 391 392 suggest to treat the environmentally-explicit model as an alternative model, instead of the 393 best and only model, to be considered in population prediction and stock management. The 394 model with GSI as a limiting factor strongly suggested that the recent low productivity of

395 the SNEMA yellowtail flounder can be explained by the unfavorable environmental 396 conditions result from a northward shift of the Gulf Stream, and also, the stock has been 397 rebuilt relative to the current productivity level.

Many environmental indicators including the CPI were correlated with the recruitment 398 399 deviations taken from the stock assessment, but the correlation was strongest for the GSI. 400 The CPI is a local-scale environmental index representing bottom temperature condition in the Mid-Atlantic Cold Pool (NEFSC, 2012) while the GSI is a basin-scale environmental 401 402 index representing the latitudinal anomalies of the Gulf Stream path (Joyce & Zhang, 2010). 403 Comparison of model fits suggests that local bottom temperature in the Mid-Atlantic Cold 404 Pool should not be the only environmental factor affecting SNEMA yellowtail flounder recruitment. Indeed, the basin-scale GSI also indicates some other shelf physical/biological 405 406 conditions that potentially affect the recruitment: (1) shelf SST condition (Gawarkiewicz et al., 2012), which affects the physiology of SNEMA yellowtail flounder during the early 407 408 pelagic phase; (2) shelf current and eddy conditions, which affect larval transport and retention on the continental shelf (Hare & Cowen, 1996); and (3) shelf primary production 409 410 condition in spring (Saba et al., 2015), which affects food availability to the larvae. Hallett et al. (2004) argued that large-scale climate indices contain information on several local 411 412 processes, so potentially they could better predict ecological processes compared to local weather conditions when a mechanistic understanding of how local environment influences 413 414 a biological process is lacking. We hypothesize that the better performance of the GSI than the CPI in explaining SNEMA yellowtail flounder recruitment is due to the aggregation of 415 416 factors beyond the bottom temperature in the Mid-Atlantic cold pool that affects 417 recruitment.

Of the alternative type of environmental effects in the Beverton-Holt stock-recruit function, we found the "limiting factor" assumption where the carrying capacity of prerecruits is regulated by the GSI to perform best for SNEMA yellowtail flounder (Iles & Beverton, 1998, Neill et al., 1994). We noticed that the differences in AIC between the models with different forms of the stock-recruit functions (limiting, controlling, or masking factor) were smaller than those between different environmental covariates incorporated in the stock-recruit function (GSI or CPI). The retrospective AIC pattern showed that the GSI- incorporated model was consistently better than the CPI-incorporated model in all seven retrospective peels. According to Mohn's ρ , another important metric of model performance, the retrospective biases in the estimates of recruitment, SSB, and F from the GSI-incorporated model were minimally larger than those from the CPI-incorporated model.

429 Even though the GSI provided a lower AIC than the CPI, it is important to also assess 430 how the inclusion of GSI impacts model estimates and predictions. Comparison suggested that R(SSB) and R(SSB&GSI) provided similar estimates of recruitment, SSB, and F, but 431 distinct estimates of biological reference points and predictions of recruitment and SSB. 432 433 Indeed, the GSI affected the expected recruitment from the stock-recruit function. Although 434 recruitment estimates from the two models were similar, R(SSB&GSI) still had a much smaller AIC as the deviations between the estimated and expected recruitment were 435 436 generally closer to zero when the GSI effect on recruitment was included. When there are catch and survey data, recruitment estimates are informed by them and are also constrained 437 438 by a penalty term to not be far from the stock-recruit function expected. However, in the prediction period when no fisheries data are available, the best recruitment predictions are 439 440 from the stock-recruit function. Specifically, recruitment predictions from R(SSB) are based solely on SSB while those from R(SSB&GSI) are also profoundly affected by year-441 442 to-year fluctuations in the GSI. Since the only difference between R(SSB) and R(SSB&GSI) lies in the stock-recruit function, whether includes GSI effects on recruitment prediction is 443 the only possible source responsible for the large differences between the predictions of 444 recruitment and SSB from the two models. 445

In the first prediction year (i.e., 2012), the SSB predictions from R(SSB) and 446 R(SSB&GSI) are not differentiable, although the recruitment predictions from R(SSB) and 447 448 R(SSB&GSI) have been notably different. Since few SNEMA yellowtail flounder can be mature at age 1, recruitment minimally impacts SSB in the first prediction year. In the 449 second prediction year (i.e., 2013), the difference in recruitment prediction propagates to 450 age 2 at which maturity reaches 0.5, leading to divergent SSB predictions from the two 451 452 models. This divergence in SSB predictions then propagates back to recruitment one year 453 later (i.e., 2014) through the stock-recruit function, resulting in an even lower recruitment prediction from R(SSB&GSI), as the high predicted GSI has already led to a lower 454

recruitment prediction in 2014. The increasingly different predictions of recruitment and
SSB from the two models clearly show the importance of selecting the most appropriate
stock-recruit function to near-term population predictions.

Interesting enough, a higher F_{MSY} was estimated by R(SSB&GSI) than R(SSB) and the 458 459 recruitment estimates (when GSI equal 0) from R(SSB&GSI) were also higher than from 460 R(SSB). It indicates that the stock was estimated to be more productive at low SSB levels when the GSI effect is included in the stock-recruit function. The reason for which, we 461 suspect, is that the recent low recruitments from R(SSB&GSI) were partly attributed to 462 463 unfavorable environmental conditions, instead of being solely attributed to reduced SSB as those from R(SSB). In other words, when environmental conditions are neutral, the 464 productivity and consequently F_{MSY} of SNEMA yellowtail flounder estimated from 465 R(SSB&GSI) should be higher than R(SSB) in which the environmental effect on 466 recruitment is not included. 467

468 As shown earlier, the five-year recruitment predictions from R(SSB) and R(SSB&GSI) 469 differ substantially. To evaluate which model can provide more reliable recruitment 470 predictions, we generated 13 retrospective predictions for each model and then compared those predictions with the "true" recruitments which were defined in this study as the 471 472 estimates from the full data. Generally speaking, the differences between the predicted and "true" recruitments are smaller in R(SSB&GSI), especially in the first prediction year in 473 474 which recruitment prediction is a function of the relatively reliable GSI observation in the last year of observations. The second and third recruitment predictions are functions of the 475 increasingly unreliable GSI predictions from R(SSB&GSI), so the inclusion of GSI effect 476 477 on recruitment leads to relatively small improvement in those recruitment predictions. As 478 expected, comparison of either MRD or MARD indicated that incorporating the GSI into the stock-recruit function cannot reduce the difference between the predicted and "true" 479 recruitment beyond a lead time of two years. In addition, we also made year-by-year 480 comparison of the retrospective recruitment predictions from the two models to evaluate 481 482 whether R(SSB&GSI) consistently outperformed R(SSB) over time in predicting 483 recruitment. The year-by-year comparison suggested that the accuracy of recruitment predictions from R(SSB&GSI) is largely dependent on the accuracy of those years' GSI 484

485 prediction from R(SSB&GSI). Because the GSI had large interannual fluctuations relative to the long-term average, its predictions from a random walk model in R(SSB&GSI) could 486 487 be considerably biased. If it happens, the recruitment predictions from R(SSB&GSI) are 488 also expected to be considerably biased, albeit in the opposite direction. However, if the environmental covariate incorporated into the stock-recruit function is a low-frequency 489 decadal oscillation such as the Pacific Decadal Oscillation, the random walk model is more 490 likely to provide reliable near-term predictions for the environmental covariate. However, 491 another problem can arise when incorporating a low-frequency decadal oscillation into the 492 stock-recruit function. Haltuch and Punt (2011) found that when fisheries data are relativity 493 494 short in time compared to the period of the incorporated decadal oscillation, stock 495 assessment models are not able to correctly tell whether the incorporated environmental 496 process has a significant effect on recruitment or not.

This study evaluated the skill of the state-space assessment model in predicting 497 498 recruitment via implementing the retrospective prediction scheme that is popularly used in 499 the climate science community. However, the skill of an assessment model in predicting 500 recruitment needs to be interpreted differently from that of a climate model in predicting climate variables. Generally speaking, the predicted climate variable such as sea surface 501 502 temperature can be observed through either in situ or remote sensing method. Therefore, the prediction skill can be evaluated by comparing model predictions with the 503 504 corresponding observations that are relatively credible. By contrast, some population attributes predicted (e.g., recruitment and SSB) by stock assessment models are inherently 505 506 unobservable in the field. This study evaluated the skill of the state-space model in 507 predicting recruitment by comparing model predictions with the corresponding estimates 508 from the full data, which are model output and less accurate than direct climate observations. As a result, a good skill in predicting recruitment does not necessarily equals 509 accurate recruitment predictions. For instance, a high prediction skill can possibly exist 510 511 when retrospective recruitment predictions and recruitment estimates from the full data are both systematically biased in the same direction to a large extent. Thus, like Mohn's ρ , the 512 retrospective prediction skill is only one metric of model performance and should be 513 evaluated together with other model diagnostics. 514

515 In brief, this paper provides two major suggestions for future fisheries studies that incorporate environmental effects into stock assessment models. First, care should be taken 516 even when the model with an environmental covariate fits data better. When the 517 518 environmental covariate is poorly predicted, the model with that environmental covariate can provide less accurate predictions than the model without any environmental covariate. 519 In future work, alternative time series models for an environmental covariate should also be 520 considered in the state-space model to potentially improve its predictive performance. For 521 instance, a stationary autoregressive process of order greater than 1 has been found to be 522 robust in predicting the GSI and its effect on silver hake distribution for the near-term 523 524 (Davis et al., 2017). Second, analyzing the retrospective prediction pattern to quantitatively evaluate model prediction skill is recommended before making management decisions 525 526 based on model predictions, in the same way that retrospective pattern is analyzed as a regular procedure in stock assessments to evaluate the biases in terminal year model 527 estimate. 528

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536 **REFERENCES**

- 537 Boer, G., Kharin, V. and Merryfield, W. (2013) Decadal predictability and forecast skill. *Climate*
- 538 *dynamics*, **41**, 1817-1833.
- Brooks, E.N. and Deroba, J.J. (2015) When "data" are not data: the pitfalls of post hoc analyses
 that use stock assessment model output. *Canadian Journal of Fisheries and Aquatic Sciences*, **72**, 634-641.

542	Brooks, E.N. and Legault, C.M. (2015) Retrospective forecasting-evaluating performance of stock
543	projections for New England groundfish stocks. Canadian Journal of Fisheries and Aquatic
544	Sciences, 73 , 935-950.
545	Davis, X.J., Joyce, T.M. and Kwon, YO. (2017) Prediction of silver hake distribution on the
546	Northeast US shelf based on the Gulf Stream path index. Continental Shelf Research, 138,
547	51-64.
548	Drinkwater, K.F., Belgrano, A., Borja, A., Conversi, A., Edwards, M., Greene, C.H., Ottersen, G.,
549	Pershing, A.J. and Walker, H. (2003) the response of marine ecosystems to climate
550	variability associated with the North Atlantic Oscillation. Geophysical Monograph Series,
551	134, 211-234.
552	Essington, T.E., Moriarty, P.E., Froehlich, H.E., Hodgson, E.E., Koehn, L.E., Oken, K.L., Siple, M.C.
553	and Stawitz, C.C. (2015) Fishing amplifies forage fish population collapses. Proceedings of
554	the National Academy of Sciences, 112, 6648-6652.
555	Fournier, D.A., Skaug, H.J., Ancheta, J., Ianelli, J., Magnusson, A., Maunder, M.N., Nielsen, A. and
556	Sibert, J. (2012) AD Model Builder: using automatic differentiation for statistical inference
557	of highly parameterized complex nonlinear models. Optimization Methods and Software,
558	27, 233-249.
559	Fry, F. (1971) 1 The effect of environmental factors on the physiology of fish. Fish physiology, 6, 1-
560	98.
561	Gawarkiewicz, G.G., Todd, R.E., Plueddemann, A.J., Andres, M. and Manning, J.P. (2012) Direct
562	interaction between the Gulf Stream and the shelfbreak south of New England. Scientific
563	reports, 2 .
564	Greene, C.H., Meyer-Gutbrod, E., Monger, B.C., McGarry, L.P., Pershing, A.J., Belkin, I.M.,
565	Fratantoni, P.S., Mountain, D.G., Pickart, R.S. and Proshutinsky, A. (2013) Remote climate
566	forcing of decadal-scale regime shifts in Northwest Atlantic shelf ecosystems. Limnol.
567	Oceanogr, 58, 803-816.
568	Haddon, M. (2010) Modelling and quantitative methods in fisheries. CRC press.
569	Hallett, T., Coulson, T., Pilkington, J., Clutton-Brock, T., Pemberton, J. and Grenfell, B. (2004) Why
570	large-scale climate indices seem to predict ecological processes better than local weather.
571	Nature, 430, 71-75.

572 Haltuch, M.A. and Punt, A.E. (2011) The promises and pitfalls of including decadal-scale climate 573 forcing of recruitment in groundfish stock assessment. Canadian Journal of Fisheries and 574 Aquatic Sciences, 68, 912-926. 575 Hameed, S. and Piontkovski, S. (2004) The dominant influence of the Icelandic Low on the position of the Gulf Stream northwall. Geophysical research letters, 31. 576 577 Hare, J.A. and Cowen, R.K. (1996) Transport mechanisms of larval and pelagic juvenile bluefish 578 (Pomatomus saltatrix) from South Atlantic Bight spawning grounds to Middle Atlantic 579 Bight nursery habitats. *Limnology and Oceanography*, **41**, 1264-1280. 580 Hilborn, R. (2016) Correlation and causation in fisheries and watershed management. Fisheries, 41, 18-25. 581 582 Hill, K.T., Crone, P.R., Lo, N.C., Macewicz, B.J., Dorval, E., McDaniel, J.D. and Gu, Y. (2011) 583 Assessment of the Pacific sardine resource in 2011 for US management in 2012. NOAA 584 Technical Memorandum NMFS-SWFSC, 487. 585 Houghton, R.W., Schlitz, R., Beardsley, R.C., Butman, B. and Chamberlin, J.L. (1982) The Middle Atlantic Bight cold pool: Evolution of the temperature structure during summer 1979. 586 587 Journal of Physical Oceanography, **12**, 1019-1029. 588 Iles, T. and Beverton, R. (1998) Stock, recruitment and moderating processes in flatfish. Journal of Sea Research, 39, 41-55. 589 590 Joyce, T.M. and Zhang, R. (2010) On the Path of the Gulf Stream and the Atlantic Meridional 591 Overturning Circulation. Journal of Climate, 23. 592 Legault, C.M. (2009) Report of the retrospective working group. NOAA NMFS Northeast Fisheries 593 Science Center Reference Document, 09-01. 594 Meehl, G.A., Goddard, L., Murphy, J., Stouffer, R.J., Boer, G., Danabasoglu, G., Dixon, K., Giorgetta, 595 M.A., Greene, A.M. and Hawkins, E. (2009) Decadal prediction: can it be skillful? Bulletin of 596 the American Meteorological Society, **90**, 1467-1485. 597 Miller, T.J., Hare, J.A. and Alade, L.A. (2016) A state-space approach to incorporating 598 environmental effects on recruitment in an age-structured assessment model with an 599 application to Southern New England yellowtail flounder. Canadian Journal of Fisheries 600 and Aquatic Sciences, 73, 1261-1270. 601 Mohn, R. (1999) The retrospective problem in sequential population analysis: An investigation 602 using cod fishery and simulated data. ICES Journal of Marine Science: Journal du Conseil, 603 **56,** 473-488.

- Mountain, D.G. (2012) Labrador slope water entering the Gulf of Maine—response to the North
 Atlantic Oscillation. *Continental Shelf Research*, 47, 150-155.
- Myers, R.A. (1998) When do environment–recruitment correlations work? *Reviews in Fish Biology and Fisheries*, **8**, 285-305.
- NEFSC (2012) 54th Northeast Regional Stock Assessment Workshop (54th SAW) Assessment
 Report. US Dept Commer, Northeast Fish Sci Cent Ref Doc. 12-18; 600 p.
- Neill, W.H., Miller, J.M., Van Der Veer, H.W. and WINEMIllER, K.O. (1994) Ecophysiology of marine
 fish recruitment: a conceptual framework for understanding interannual variability.
 Netherlands Journal of Sea Research, **32**, 135-152.
- Nielsen, A. and Berg, C.W. (2014) Estimation of time-varying selectivity in stock assessments using
 state-space models. *Fisheries Research*, **158**, 96-101.
- Nye, J.A., Joyce, T.M., Kwon, Y.O. and Link, J.S. (2011) Silver hake tracks changes in Northwest
 Atlantic circulation. *Nature communications*, 2, 412.
- Punt, A.E., A'mar, T., Bond, N.A., Butterworth, D.S., de Moor, C.L., De Oliveira, J.A., Haltuch, M.A.,
 Hollowed, A.B. and Szuwalski, C. (2014) Fisheries management under climate and
- 619 environmental uncertainty: control rules and performance simulation. *ICES Journal of*620 *Marine Science*, **71**, 2208-2220.

621 Quinn, T.J. and Deriso, R.B. (1999) *Quantitative fish dynamics*. Oxford University Press.

622 Saba, V.S., Hyde, K.J., Rebuck, N.D., Friedland, K.D., Hare, J.A., Kahru, M. and Fogarty, M.J. (2015)

- 623 Physical associations to spring phytoplankton biomass interannual variability in the US
- Northeast Continental Shelf. *Journal of Geophysical Research: Biogeosciences*, **120**, 205220.
- Schirripa, M.J. (2007) Status of the sablefish resource off the continental US Pacific Coast in 2007.
 Pacific Fishery Management Council, Portland, OR.
- Sullivan, M.C., Cowen, R.K., Able, K.W. and Fahay, M.P. (2000) Spatial scaling of recruitment in four
 continental shelf fishes. *Marine Ecology Progress Series*, 207, 141-154.
- Sullivan, M.C., Cowen, R.K. and Steves, B.P. (2005) Evidence for atmosphere–ocean forcing of
 yellowtail flounder (Limanda ferruginea) recruitment in the Middle Atlantic Bight. *Fisheries Oceanography*, 14, 386-399.
- Szuwalski, C.S. and Hollowed, A.B. (2016) Climate change and non-stationary population processes
 in fisheries management. *ICES Journal of Marine Science*, **73**, 1297-1305.

635	Szuwalski, C.S., Vert-Pre, K.A., Punt, A.E., Branch, T.A. and Hilborn, R. (2015) Examining common
636	assumptions about recruitment: a meta-analysis of recruitment dynamics for worldwide
637	marine fisheries. Fish and Fisheries, 16, 633-648.
638	Taylor, A.H. and Stephens, J.A. (1998) The North Atlantic oscillation and the latitude of the Gulf
639	Stream. <i>Tellus A</i> , 50 , 134-142.
640	Vert-pre, K.A., Amoroso, R.O., Jensen, O.P. and Hilborn, R. (2013) Frequency and intensity of
641	productivity regime shifts in marine fish stocks. Proceedings of the National Academy of
642	Sciences, 110, 1779-1784.
643	Walters, C.J. and Collie, J.S. (1988) Is research on environmental factors useful to fisheries
644	management? Canadian Journal of Fisheries and Aquatic Sciences, 45, 1848-1854.
645	Xu, H., Kim, HM., Nye, J.A. and Hameed, S. (2015) Impacts of the North Atlantic Oscillation on sea
646	surface temperature on the Northeast US Continental Shelf. Continental Shelf Research,
647	105, 60-66.
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650	TABLES
651	
652	Table 1. The Pearson correlation coefficients between SNEMA yellowtail flounder
653	recruitment deviations and various environmental indices with a lag of one to three years.

Positive lags mean the environment leads the recruitment. IL stands for Icelandic low and AH stands for Azores high. Only correlations that are significant at the 95% confidence level are shown and the coefficient marked in bold represents the correlation is significant at the 99% confidence level.

Environmental index	Lag 1	Lag 2	Lag 3
IL Pressure	0.38		
IL Longitude			
IL Latitude			
AH Pressure			
AH Longitude			-0.36

AH Latitude			-0.38
NAO		-0.37	-0.38
GSI	-0.52		
GSNW	-0.41		
CPI	-0.39		

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Table 2. The fits of the state-space assessment model with different environmentallyexplicit stock-recruit functions. These fits are compared according to AIC and Akaike weight, which are shown in column three and four, respectively.

Model	Stock-recruit function	ΔΑΙΟ	ω (AIC)
R(SSB)	$\frac{SSB}{b + aSSB}$	13.89	0.00
R (CPI _{limiting} , SSB)	$\frac{SSB}{b + aSSBe^{cCPI}}$	5.40	0.04
R(CPI _{masking} , SSB)	SSB be ^{cCPI} + aSSB	4.53	0.06
R(CPI _{controlling} , SSB)	$\frac{SSB}{b + aSSB}e^{cCPI}$	3.93	0.08
R (GSI _{limiting} , SSB)	$\frac{SSB}{b + aSSBe^{cGSI}}$	0.00	0.55
R (GSI _{masking} , SSB)	$\frac{SSB}{be^{cGSI} + aSSB}$	4.94	0.05
R(GSI _{controlling} , SSB)	$\frac{SSB}{b + aSSB}e^{cGSI}$	1.67	0.24

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Table 3. The retrospective difference in AIC between the model with the GSI as a limiting factor and that with the CPI as a limiting factor (column 2) as well as the retrospective estimates of the environmental link parameter and the associated standard deviation from the model with the GSI as a limiting factor. Positive difference in AIC corresponds to the GSI-incorporated model outperforms the CPI-incorporated model and vice versa.

	Peel	AIC(CPI) - AIC(GSI)	c (sd)
	0	5.40	1.53 (0.37)
	1	5.17	1.52 (0.37)
U	2	5.48	1.53 (0.37)
\mathbf{O}	3	5.55	1.51 (0.36)
	4	7.31	1.50 (0.34)
	5	12.08	1.46 (0.29)
	6	10.51	1.51 (0.32)

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Table 4. The Mohn's ρ of SSB, F, and recruitment from the state-space assessment model with and without environmental (GSI or CPI) effect on recruitment. A smaller Mohn's ρ (absolute value) corresponds to less retrospective bias and consequently better model performance.

Model	ρ (SSB)	ρ(F)	ρ(R)
R(SSB)	0.11	-0.14	0.24
R (CPI _{limiting} , SSB)	0.11	-0.14	0.22
R (GSI _{limiting} , SSB)	0.14	-0.16	0.36
	Model R(SSB) R (CPI _{limiting} , SSB) R (GSI _{limiting} , SSB)	Model ρ (SSB)R(SSB)0.11R (CPI_limiting, SSB)0.11R (GSI_limiting, SSB)0.14	Model ρ (SSB) ρ (F)R(SSB)0.11-0.14R (CPI_limiting, SSB)0.11-0.14R (GSI_limiting, SSB)0.14-0.16

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	Prediction lead time	1 year	2 years	3 years
	MRD - R(SSB)	1.23	1.53	1.68
	MRD - R(SSB&GSI)	0.89	1.26	1.59
	MARD - R(SSB)	1.45	1.73	1.76
	MARD - R(SSB&GSI)	1.04	1.50	1.77

Table 5. The mean relative difference and mean absolute relative difference in the near-term recruitment predictions from R(SSB) and R(SSB&GSI).

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680 FIGURE CAPTIONS

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Figure 1. Compare the natural log of SNEMA yellowtail flounder recruitment with the two most related environmental indices (GSI and CPI). The recruitment time series is from the baseline run in the most recent benchmark assessment (NEFSC, 2012) and the CPI is scaled to have the same variance as the GSI for easier comparison. Both environmental indices in this figure have been shifted one year backward to account for the one-year lag between them with recruitment.

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Figure 2. The estimated Beverton-Holt stock-recruit function from R(SSB) (dashed line) in
comparison to those from R(SSB&GSI) (solid lines) under various GSI values.

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Figure 3. The first row shows the estimated and predicted recruitment (a) and SSB from the two comparing models under the F_{MSY} harvest scenario. The second row shows the estimated F and F_{MSY} (c) and the estimated MSY (d) from the two comparing models. The third row shows the estimated and predicted log (SSB/SSB_{MSY}) from the two comparing models under the F_{MSY} harvest scenario (e). In this figure, the color dashed lines and vertical error bars represent the 95% confidence interval, and the black vertical dashed lines mark last year of data.

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Figure 4. The retrospective prediction patterns from R(SSB) (dashed lines) and R(SSB&GSI) (solid lines). The two black lines represent the "true" recruitment from each model, i.e., the recruitment estimates when performing the assessment on full data. The dots are the terminal year recruitment estimates and color of lines and dots represents the last year of data available for each retrospective prediction, from dark blue (1996) to dark red (2008). The two pairs of recruitment predictions mentioned in the discussion are highlighted as bold color lines.

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