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Spatial synchrony in the response of a long range migratory species (*Salmo salar*) to climate change in the North Atlantic Ocean

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31

32 **ABSTRACT**

33 A major challenge in understanding the response of populations to climate change is to
34 separate the effects of local drivers acting independently on specific populations, from the
35 effects of global drivers that impact multiple populations simultaneously and thereby
36 synchronize their dynamics. We investigated the environmental drivers and the demographic
37 mechanisms of the widespread decline in marine survival rates of Atlantic salmon (*Salmo*
38 *salar*) over the last four decades. We developed a hierarchical Bayesian life cycle model to
39 quantify the spatial synchrony in the marine survival of 13 large groups of populations (called
40 stock units, SU) from two continental stock-groupings (CSG) in North America (NA) and
41 Southern Europe (SE) over the period 1971-2014. We found strong coherence in the temporal
42 variation in post-smolt marine survival among the 13 SU of NA and SE. A common North
43 Atlantic trend explains 37% of the temporal variability of the survivals for the 13 SU and
44 declines by a factor 1.8 over the 1971-2014 time series. Synchrony in survival trends is
45 stronger between SU within each CSG. The common trends at the scale of NA and SE capture
46 60% and 42% of the total variance of temporal variations, respectively. Temporal variations
47 of the post-smolt survival are best explained by the temporal variations of sea surface
48 temperature (SST, negative correlation) and net primary production indices (PP, positive
49 correlation) encountered by salmon in common domains during their marine migration.
50 Specifically, in the Labrador Sea/Grand Banks for NA populations 26% and 24% of variance
51 is captured by SST and PP, respectively and in the Norwegian Sea for SE populations 21%
52 and 12% of variance is captured by SST and PP, respectively. The findings support the
53 hypothesis of a response of salmon populations to large climate induced changes in the North
54 Atlantic simultaneously impacting populations from distant continental habitats.

55

56 **Key words:** Spatial covariation, climate change, stage-based life cycle model, marine
57 survival, Atlantic salmon, environmentally driven changes, bottom-up, hierarchical Bayesian
58 model

59

60 **1. INTRODUCTION**

61 Understanding the response of populations to global changes, in terms of demography and
62 adaptive capacity, is critical to support ecosystem-based management (Brown et al., 2011;
63 Edwards, Beaugrand, Hays, Koslow, & Richardson, 2010; Harley et al., 2006; Stenseth,
64 2002). A major challenge to understanding the response of populations to environmental
65 variations is to partition the effects of global drivers that likely impact multiple populations
66 simultaneously and synchronize their dynamics from the effects of drivers acting locally on
67 specific populations (Moran, 1953; Post & Forchhammer, 2002). This is also critical for a
68 better understanding of the mechanisms affecting the resilience of populations to global
69 change (Heino, 1998; Palmqvist & Lundberg, 1998).

70 Life cycle models that consider and incorporate the spatial and temporal heterogeneity of
71 ecological mechanisms and demographic responses are useful for examining the effects of
72 multiple factors that interact in a hierarchy of scales (Cunningham, Westley, & Adkison,
73 2018; Rochette, Le Pape, Vigneau, & Rivot, 2013; Stelzenmüller, Schulze, Fock, &
74 Berkenhagen, 2011). When combined with the analysis of multiple populations, these models
75 provide a powerful approach to partition the effects of factors impacting each population
76 specifically from those affecting groups of populations simultaneously (Lahoz-Monfort et al.
77 2013; Walter et al. 2017). In addition, signals that arise from multiple population relationships
78 are more likely to represent true biological processes rather than statistical flukes (Myers,
79 Mertz, & Bridson, 1997; Soberon & Nakamura, 2009) and as a result can be more informative
80 than separate analyses of single populations (Britten, Dowd, & Worm, 2016; Szuwalski, Vert-
81 Pre, Punt, Branch, & Hilborn, 2015; Zimmermann, Claireaux, & Enberg, 2019).

82 Separating out the different scales of interactions of ecological processes driving population
83 dynamics is particularly challenging in the case of highly migratory species, which can
84 interact with a multitude of single and/or synergistic factors at different points in time and
85 space during their life cycle. For instance, the life cycle of anadromous fish, such as
86 salmonids, relies on population-specific freshwater habitats for reproduction and juvenile
87 growth and marine habitats shared by multiple populations for feeding and maturation. This
88 makes these species sensitive to multiple environmental and anthropic stressors acting at
89 different spatial scales, with factors operating at sea potentially having synchronizing effects
90 on the dynamics of large groups of populations. For such species, identifying the space and
91 time domains associated with specific life stages that are most susceptible to conditioning the
92 population dynamics is a prerequisite to better understand population responses to global
93 changes (Cunningham et al., 2018) and to support improved management decisions and
94 actions at global and local scales.

95 Atlantic salmon (*Salmo salar*) is one of the most emblematic fish in the Atlantic Ocean. The
96 species reproduces in a large number (~ 2000) of rivers distributed in the eastern (Europe) and
97 western (North America) regions of the North Atlantic. Due to its highly evolved homing
98 ability, the species is structured into individual river populations, with specific and variable
99 freshwater habitat environments. During the freshwater phase, the population dynamics are
100 conditioned by local habitat quality and trophic resources (Elliott, 2001; Jonsson, Jonsson, &
101 Hansen, 1998; Milner et al., 2003). During the marine phase, populations originating from
102 distant continental habitats migrate to common feeding grounds in the North Atlantic, with
103 major concentrations located off West Greenland, in the Labrador Sea, and the Faroes Islands
104 and Norwegian Sea (Aas, Einum, Klemetsen, & Skurdal, 2010; D. H. Mills, 1989). In these
105 aggregations at sea, they are exposed to common environmental marine conditions and
106 fisheries.

107 Atlantic salmon populations from North America and Europe have undergone a widespread
108 decline in abundance over the last four decades (Chaput, 2012; ICES, 2017; Olmos et al.,
109 2019), but the mechanisms responsible for these declines are still unclear. The broad scale
110 pattern of decline has led to the hypotheses that major ecosystem changes in the North
111 Atlantic Ocean are the main driver of these declines (Olmos et al., 2019). The mechanisms for
112 this may include an indirect effect associated with an increase in sea temperatures (Beaugrand
113 & Reid, 2012; Friedland, Moore, & Hogan, 2009; Jensen et al., 2012). A major trophic shift in
114 the North Atlantic Ocean was documented in the early 1990's with trophic level changes

115 observed in the plankton communities upward to seabird populations (Beaugrand, Edwards,
116 Brander, Luczak, & Ibanez, 2008; Durant, Anker-Nilssen, & Stenseth, 2003; A. J. Pershing,
117 Head, Greene, & Jossi, 2010) which was hypothesized to exert bottom up control via
118 reductions in the abundances and the energetic value of prey across higher trophic levels (K.
119 E. Mills, Pershing, Sheehan, & Mountain, 2013; Otero et al., 2012; Renkawitz, Sheehan,
120 Dixon, & Nygaard, 2015). These changes may have been responsible for altered Atlantic
121 salmon growth at sea and consequently survival through size-dependent mortality (Friedland
122 & Reddin, 2000; Gislason, Daan, Rice, & Pope, 2010; Peyronnet, Friedland, Maoileidigh,
123 Manning, & Poole, 2007).

124 Broader scale analyses to date, however, suggest that despite the overall spatial coherence of
125 the trends in abundances and survival rates observed throughout the North Atlantic, the annual
126 and region specific variations between continental stock groups (CSG) in North America and
127 Southern Europe and among populations within a CSG are large (Olmos et al., 2019). This
128 may in part be explained by the diversity and complexity of migration routes at sea
129 undertaken by populations originating from different areas of the North Atlantic. As such, it is
130 challenging yet necessary to identify the space and time domains along the migration routes at
131 sea where salmon are exposed to favorable and unfavorable ecosystem conditions that may
132 strongly affect their survival.

133 Although the early Atlantic salmon post-smolt marine phase is often suggested as a critical
134 stage for survival (Friedland et al. 2003a, 2005, 2000; Thorstad et al. 2012; Chaput et al.
135 2018), the environmental conditions encountered later in the first year at sea can also be
136 important (Friedland et al., 2009; Friedland & Reddin, 2000; K. E. Mills et al., 2013). In
137 addition, the factors involved in the declines in survival may differ between populations.
138 Growth variations during the first summer at sea have been hypothesized as critical for the
139 survival of SE populations (Friedland et al. 2008; Friedland et al. 2014; McCarthy, Friedland,
140 and Hansen 2008; Peyronnet et al. 2007; Haugland et al. (2006) and Jensen et al. (2012)). In
141 contrast, variations in predation pressure in early spring have been hypothesized to be the
142 main driver of early post-smolt survival in southern NA populations (Friedland et al., 2014).

143 The mechanisms involved at various spatial and temporal scales, and the degree to which
144 these mechanisms and hence the responses are shared between populations remain largely
145 unclear. A simultaneous and joint analysis of multiple populations throughout the Atlantic
146 Ocean within a unified framework is needed to improve our understanding of the response of
147 Atlantic salmon populations to changes in the marine ecosystem.

148 In this paper, we rely on the modelling framework developed by Olmos et al. (2019) to
149 explore how environmental conditions encountered by Atlantic salmon in different space and
150 time domains along the marine migration routes may contribute to the variations of marine
151 survival in Europe and North America. Olmos et al. (2019) developed an age and stage-based
152 model for the collective analysis of the dynamics of thirteen geographically proximate
153 Atlantic salmon stock units (SU) from the eastern NA and SE CSG, and applied this model to
154 data over the period 1971-2014. The model provides a framework to quantify the spatial
155 coherence in the temporal variation of the post-smolt marine survival rates and in the
156 proportion of fish maturing after one winter at sea (1SW) in a hierarchy of spatial scales
157 across the North Atlantic. Olmos et al. (2019) reported on the strong coherence in temporal
158 variation of marine survivals among the 13 stock units of Southern Europe and North
159 America, represented by a collective decline in the marine survival over the 1971-2014 time
160 series. The results also provided evidence of covariation among geographically proximate
161 stock units, with the strength of the covariation that increases when going down to spatial
162 scale, thus suggesting the intricate influence of drivers acting at a hierarchy of spatial scale.

163 Here, by taking advantage of the flexibility of the hierarchical model structure, we first extend
164 the modelling framework developed by Olmos et al. (2019) by explicitly modeling temporal
165 variation in post-smolt survival as the sum of trends in a hierarchy of spatial scales across
166 global to local SU-specific areas. This allows the investigation of the degree of synchrony in
167 Atlantic salmon post-smolt survival and explicitly quantifies the amount of variance that is
168 captured by trends at various spatial scales. Second, we investigate whether the temporal
169 variation in the marine survival can be explained by environmental variation encountered by
170 salmon during the early post-smolt marine phase when salmon use specific transit habitat, or
171 during the later phase of the first year at sea when salmon of different areas aggregate at
172 common feeding areas. We conducted an extensive review of the literature on post-smolt
173 migration routes to define the space-time domains associated with the early marine phase
174 (spatially specific to each SU or to small groups of SU with proximate freshwater habitat) and
175 late phase of the first year at sea (feeding areas common to large groups of SU). We then
176 assessed the relationships between the temporal variations of marine survival and
177 environmental covariates defined in those space-time domains including sea surface
178 temperatures, primary production indices, and large scale climate indices. Our prediction was
179 that the environmental conditions encountered in the common feeding areas should explain
180 the greatest part of the synchronous signal observed between the SUs, while environmental

181 conditions encountered during the early marine phase in transit habitat would not explain the
182 broader scale responses of these salmon populations.

183 **2. MATERIALS AND METHODS**

184 **2.1 General Model Outline**

185 Below we provide the main outlines of the model. Further details can be found in Olmos et al.
186 (2019).

187 The model is an age- and stage-based life cycle model (Fig. 1) that formulates the dynamics
188 of all SU in a single hierarchical framework. The spatial structure of the model is unchanged
189 from Olmos et al. (2019). The model considers thirteen stock units that each define
190 assemblages of river-specific Atlantic salmon populations reproducing in the respective North
191 American (NA) and Southern European (SE) CSG. The NA CSG consists of 6 SU (indexed
192 by $r = 1, \dots, 6$). The SE CSG consists of 7 SU (indexed by $r = 7, \dots, 13$) (Fig. 2).

193 The Atlantic salmon from a SU are considered to form a single homogeneous group with
194 similar life history and migration routes at sea. Juvenile salmon produced in each SU migrate
195 to the sea as smolts after 1 to 6 years in freshwater, with the proportions at age varying among
196 SUs. The model draws on explicit hypotheses about the migration routes at sea that generate
197 spatial segregation in salmon populations (Fig. 1 and 2). All salmon from NA and SE migrate
198 from their specific coastal area to reach a common feeding ground in the Labrador Sea and the
199 Norwegian Sea, respectively. After one winter spent at sea, some salmon mature and return to
200 their natal river to spawn while non-maturing salmon migrate to West Greenland. The
201 different SU in the model present two levels of aggregation (Fig. 1). During the first months at
202 sea, post-smolts of different SU are assumed to occupy spatially different transit habitats. In
203 the later phase of the first year at sea, they migrate to a shared feeding area common to all SU
204 of the same CSG, and where they are exposed to high seas fisheries operating on mixed SU.

205 The model is formulated in a Bayesian hierarchical state-space framework (Parent and Rivot,
206 2012; Rivot et al., 2004) that incorporates stochasticity in population dynamics as well as
207 observation errors. It assimilates information from the time series of data (1971 to 2014)

208 collated by ICES WGNAS (Working Group on North Atlantic Salmon; ICES, 2015, 2017).
209 These consist of: (i) annual estimates of the number of mature anadromous Atlantic salmon
210 that return to each of the 13 SU, by 1SW and 2SW maiden sea-age classes; (ii) annual
211 estimates of homewater catches for each SU by sea-age class; (iii) annual estimates of
212 commercial catches for the mixed stock fisheries at sea operating sequentially on
213 combinations of SU, and data on the SU origin of the catches (but see Olmos et al. 2019 for
214 further details).

215 The model was primarily designed to estimate the abundance of salmon at various life stages
216 along the life cycle, the exploitation rates in the fisheries, and two key parameters of the
217 marine phase: the post-smolt marine survival rates (from out-migrating smolts to the 1
218 January of the first winter at sea, referred as the Pre Fishery Abundance stage, or PFA) and
219 the proportions of fish maturing as 1SW, for each year and each SU. It explicitly considers
220 temporal covariation in those two key demographic parameters. For the present analysis, we
221 keep the original covariation model for the proportion of fish maturing as 1SW as defined by
222 Olmos et al. (2019), with temporal variations of this parameters modeled as a multivariate
223 random walk in the logit scale. Random variations are drawn from multivariate Normal
224 distribution with a 13×13 variance-covariance matrix. The model for temporal variation in
225 post-smolt marine survival, which is the focus of this paper, is modified from Olmos et al.
226 (2019), and temporal variation is modelled through an explicit decomposition of terms
227 associated with the various spatial scales, as detailed hereafter.

228 **2.2 Investigating the spatial synchrony in marine survival**

229 **2.2.1 Hierarchical decomposition of the temporal variations of post-smolt survival**

230 Different models for the temporal variation of post-smolt marine survival are tested (Supp.
231 Mat S1). In the reference model M1 (eq. (1), (2) and Tables 1, S1.1), temporal variation in
232 post-smolt survival is explicitly written as the sum of three components to partition out the
233 survival signal at three scales: a term capturing the synchronous signal between all SU, a term
234 capturing the synchronous signal within each CSG, and a term for the remaining temporal
235 variability specific to each SU.

236 Following the methodology developed by Grosbois et al. (2009) and Lahoz-Monfort et al.
237 (2011, 2013), post-smolt survival $\theta_{t,r}$ (in the *logit* scale) at year t in SU r within the CSG g
238 ($g=NA$ or SE) is modelled as the sum of independent normally distributed random terms:

239 (1)
$$\logit(\theta_{t,r}) = \beta_r + \delta_t + \alpha_{g_t} + \varepsilon_{t,r},$$

240 with β_r an intercept that is constant for all years and $(\delta_t, \alpha_{g_t}, \varepsilon_{t,r})$ that are identically and
241 independently normally distributed random terms:

242 (2)
$$\text{for all } t, \begin{cases} \delta_t \sim N(0, \sigma_\delta^2) \\ \alpha_{g_t} \sim N(0, \sigma_{\alpha_g}^2) \\ \varepsilon_{t,r} \sim N(0, \sigma_{\varepsilon_r}^2) \end{cases}$$

243 Time series of δ_t and $(\delta_t + \alpha_{g_t})$ characterize the synchronous part of the signal at two spatial
244 scales. The δ_t 's capture the trend that is common to all SU over the North Atlantic Ocean. The
245 $\delta_t + \alpha_{g_t}$'s characterize the amount of between year variation synchronous to all SU within
246 each CSG, $g=NA$ and $g=SE$. $\varepsilon_{t,r}$ are remaining random variations specific to each SU that
247 characterize the asynchronous part of the signal. Priors on parameters are all weakly
248 informative (Table 1).

249 Two embedded models of lower complexity were then considered (Supp. Mat. S1). Since
250 Olmos et al. (2019) have shown that there are correlations between SU, models with no
251 correlations between SU were not examined further. Analyses that considered environmental
252 covariates were based on the most complete model M1.

253 2.2.2 Quantifying synchrony

254 The different random terms $(\delta_t, \alpha_{g_t}, \varepsilon_{t,r})$ in eq. (1) are independent, therefore the total between
255 year variance of the post-smolt survival time series of each SU ($\logit(\theta_{t,r})$), denoted $V_{tot,r}$, is
256 the sum of the variance of the random terms:

257 (3)
$$Var_{tot,r} = \sigma_\delta^2 + \sigma_{\alpha_g}^2 + \sigma_{\varepsilon_r}^2$$

258 where $g=NA$ for $r=1,\dots,6$ and $g=SE$ for $r = 7,\dots,13$.

259 Synchrony at different spatial scales was quantified by calculating the Inter-Class Correlation
260 (ICC) based on the ratio of inter-annual variances (Grosbois et al., 2009; Lahoz-Monfort et
261 al., 2013, 2011). For each SU r , we calculated

$$262 \quad (4) \quad ICC_{\delta_r} = \frac{\sigma_{\delta}^2}{Var_{tot_r}}$$

$$263 \quad (5) \quad ICC_{g_r} = \frac{\sigma_{\delta}^2 + \sigma_{\alpha_g}^2}{Var_{tot_r}}.$$

264 where $g=NA$ for $r=1,\dots,6$ and $g=SE$ for $r = 7,\dots,13$. ICC_{δ_r} quantifies the amount of variance
265 of the survival time series that is captured by the global trend component δ . A large ICC_{δ_r}
266 means that the variance of the shared component (σ_{δ}^2 , synchronous part of the signal) is large
267 relative to the total variance of the time series. ICC_{g_r} quantifies the amount of variance that is
268 captured by the continental trend g to which the SU r belongs.

269 We then calculated synchrony indices as the average of ICC values:

$$270 \quad (6) \quad \overline{ICC}_{\delta} = mean_{all\ r}(ICC_{\delta_r}),$$

$$271 \quad (7) \quad \overline{ICC}_{NA} = mean_{all\ r\ in\ NA}(ICC_{NA_r}),$$

$$272 \quad (8) \quad \overline{ICC}_{SE} = mean_{all\ r\ in\ SE}(ICC_{SE_r}).$$

273 \overline{ICC}_{δ} is the amount of temporal variance that is synchronous among all SU and provides a
274 global index of synchrony over the entire set of SU in both NA and SE CSG. \overline{ICC}_{NA} and
275 \overline{ICC}_{SE} are the fraction of the between year variance accounted for by the NA ($\delta + \alpha_{NA}$) and
276 SE ($\delta + \alpha_{SE}$) synchronous components, respectively. They provide an index of synchrony
277 within each CSG.

278 **2.2.3 Testing the influence of environmental covariates in different space-time domains**
279 **along the migration routes**

280 The model was used to investigate correlations between times trends in post-smolt survival
281 and environmental covariates integrated over different space-time domains occupied by
282 salmon along the migration routes (Fig. 2). Specially, we considered two types of domains: (i)
283 domains visited during the early post-smolt phase as transit habitat specific to each SU (or to
284 groups of geographic proximate SUs); (ii) domains visited during the late post-smolt phase
285 that are common to all SU of the same CSG (Fig. 2 and 3).

286 In this section, we first describe the methods used to define the space-time domains. Then, we
287 present the ecological hypotheses tested and the associated environmental variables integrated
288 over those different domains. Finally, we detail the statistical method used to quantify the
289 amount of variance of the temporal variations of post-smolts survival that is captured by the
290 covariates in the different space-time domains.

291 **2.2.3.1 Defining specific and common CSG space-time domains**

292 We conducted an extensive review of the literature to define the key space-time domains
293 occupied by post-smolts over their marine phase across the North Atlantic Ocean (Fig. 2; Sup.
294 Mat. S2). Two types of space-time domains were defined. Specific domains are transit habitat
295 occupied by post-smolts during their first three months at sea of migration from the estuarine
296 and coastal areas to the common feeding area (Fig. 2, zones 1 to 9, Table S2.1, Fig S2.1).
297 These domains are specific to each SU or to small groups of geographic proximate SU (Fig. 3
298 and Fig. S2.1). Common CSG domains are space-time domains corresponding to the habitat
299 occupied by salmon in the later phase of the first year at sea (Fig. 2 domains A and B) and
300 associated to feeding areas common to all SU within the same CSG (Fig. 3).

301 The only exception is for the Southwest Iceland SU, which presents different migrations from
302 the other SU from SE CSG. Salmon from Iceland reach the sea later, in June, and do not
303 migrate to the Norwegian Sea (Guðjónsson, Einarsson, Jónsson, & Guðbrandsson, 2015).
304 Consequently, for the Southwest Iceland SU, the same spatial limits are defined for the transit
305 and common domains (Fig. 2 and Table S2.1).

306 2.2.3.2 Integrating environmental variables over space-time domains

307 The sea surface temperature (SST) and primary production (PP) were averaged over the
308 defined space-time domains (Fig. 3) and introduced as explanatory variables in the life cycle
309 model to assess the extent to which the temporal variations in the post-smolt survival could be
310 explained by environmental variations encountered in the specific or common domains
311 occupied by salmon during the first year at sea. We also examined the influence of two large
312 scale climate indices, the Atlantic Multidecadal Oscillation (AMO) and the North Atlantic
313 Oscillation Index (NAOI) to assess their influence on the temporal variations in the post-smolt
314 survival.

315 Sea Surface Temperature (SST)

316 Ocean warming is one of the major effects of climate change on marine ecosystems. An
317 increase in seawater temperature may affect survival differently (negative or positive) through
318 direct or indirect effects.

319 The direct physiological effect of an increase in temperature is difficult to predict as it can be
320 positive or negative, depending on the range of the temperature change relative to the species'
321 optima and tolerance. Atlantic salmon is an ectothermic species with a range of preferred
322 temperatures at sea between 2°C and 14°C (Holm, 2000; David G. Reddin & Schearer, 1987),
323 with the highest post-smolt captures being realized in temperatures between 4-10°C (D. G.
324 Reddin & Friedland, 1993). Then, by directly increasing metabolism, an increase in
325 temperature should increase growth potential of salmon, and in turn may have a positive
326 effect on marine survival provided that foraging resources are available in sufficient quantity
327 (Cunningham et al., 2018; Siegel, McPhee, & Adkison, 2017). By contrast, an increase of
328 temperature well above the optimum could have a negative effect on growth and marine
329 survival. However, based on the literature, we rather expect negative indirect effects of an
330 increase in seawater temperature on both growth and survival, through bottom-up control of
331 food resources available for salmon during the first year at sea (Beaugrand & Reid, 2012;
332 Friedland et al., 2009; Jensen et al., 2012).

333 SSTs were used to calculate the seawater temperature in each space-time domain and derived
334 from the HadISST1 datasets (See Sup. Mat. S3). Standardized anomalies of SST for each
335 space-time domain z (as defined in Table S2.1) and year t , denoted $SST_{z,t}^*$ were calculated as:

336 (9)
$$SST_{z,t}^* = \frac{\overline{SST}_{z,t} - \overline{SST}}{\sigma_{SST}}$$

337 where $\overline{SST}_{z,t}$ is the SST averaged over the space-time domain z (averaged over month and
338 space) for a particular year t , \overline{SST} is the SST averaged over spatial and temporal (months)
339 limits covered by all specific and common domains, and overall years t , and σ_{SST} is the
340 standard deviation calculated from the between year variability of the SST averaged over
341 spatial and temporal (months) limits covered by all specific and common domains. Note that
342 with this method, the anomalies are calculated relative to SST averaged over all space-time
343 domains (both common and specific) and covering both NA and SE post-smolt habitat.
344 Therefore, the contrast in absolute value and temporal (between year) variance between the
345 type of domains (specific versus common) and between the two CSG (NA and SE) is
346 conserved.

347 Primary Production (PP)

348 PP was considered as indicator of the ocean production which determines the prey availability
349 for salmon at sea and consequently expected to be positively correlated to post-smolt survival.

350 PP data are derived from the Earth System Model) developed by the Geophysical Fluid
351 Dynamic Laboratory (GFDL-ESM2M, Dunne et al., 2012) (see Sup. Mat. S3 for more
352 details).

353 Standardized anomalies of PP ($PP_{z,t}^*$) were calculated following the same approach as SST.
354 However, to match with the months of phytoplankton bloom, PP was integrated over the two
355 months April-May in both specific and common domains.

356 Atlantic Multidecadal Oscillation (AMO)

357 The AMO is a low-frequency and basin-wide climate index reflecting sea surface temperature
358 variability over the last century (Alheit, Drinkwater, & Nye, 2014; Enfield, Mestas-Nunez,
359 Trimble, & others, 2001).

360 Previous studies have reported on a negative correlation between temporal variations of
361 salmon abundance and the AMO in both North America (Crozon et al., 2005; K. E. Mills et
362 al., 2013) and in Southern Europe (Beaugrand et al., 2012). Friedland et al., (2014)

363 highlighted a differential response of salmon abundances from North America and Southern
364 Europe to the AMO. Based on these publications, we expect the positive AMO to negatively
365 impact post-smolt survival but with potentially different strength for NA and SE CSGs.

366 The effect of AMO on post-smolt survival is included using the average monthly value over
367 the entire post-smolt phase (May-December); data from 1975 to 2012 were considered (see
368 Sup. Mat. S3 for more details).

369 North Atlantic Oscillation Index (NAOI)

370 We used the winter NAOI (mean from December to March) as the NAOI is strongly
371 associated with climatic conditions during the winter (Sup. Mat. S3). Previous studies have
372 shown weak correlations between NAOI and salmon abundance (K. E. Mills et al., 2013;
373 Beaugrand and Reid, 2003, 2012). Our prediction is that high winter NAOI should be
374 associated with good feeding conditions because of positive temperature anomalies, and thus
375 be positively correlated with post-smolt survival. However, because NAOI described different
376 conditions in North America and Europe, our expectation is that the temporal variation of
377 NAOI will affect the two CSGs differently.

378 **2.2.4 Quantifying the influence of environmental variables in the different space-time** 379 **domains**

380 We developed a variant of the variance analysis method from Grobois et al., (2009) and
381 Lahoz-Monfort et al., (2011, 2013) to quantify the contribution of each covariate in the
382 different space-time domains to the temporal and spatial variations of post-smolt survival.
383 This also allows quantifying the contribution of covariates in generating synchrony at various
384 spatial scales.

385 Preliminary analysis showed that the time series of environmental variables exhibited an
386 important level of correlation, both between variables of different nature (i.e. variations of
387 SST, PP, AMO and NAOI are not independent) and between the different space-time domains
388 for the same covariate. Hence, the influence of each type of covariate was considered
389 separately, and for the same covariate, the influence in the different space-time domains at
390 different scales was also considered separately.

391 Time series of environmental covariates (X_{k_t}), defined at different spatial scales k (specific or
392 common CSG domains) were considered as an additional factor in model M1:

393 (11)
$$\text{logit}(\theta_{t,r}) = \beta_r + \delta_t + \alpha_{g_t} + \varepsilon_{t,r} + \gamma_k \times X_{k_t}$$

394 where k refers to the specific or common domains, and γ_k is the coefficient describing the
395 influence of covariate X_k on post-smolt survival (two separate models were built for the two
396 spatial scales). The γ_k were drawn a priori in a non-informative Uniform prior distribution.

397 For covariates SST and PP at the scale of specific domains, different coefficients γ_r for each
398 SU were considered. When covariates are considered at the scale of a common CSG domain,
399 two coefficients γ_{NA} and γ_{SE} were considered for the influence on NA and SE, respectively.
400 Because our expectation is that the effects of AMO and NAOI could be different between the
401 two CSGs, two coefficients γ_{NA} and γ_{SE} were considered for the influence on NA and SE,
402 respectively. Table 2 sums up the hypotheses tested and associated model configurations that
403 included environmental covariates.

404 Contribution of the covariates to the temporal variability of post-smolt survival

405 For each covariate considered independently, models were run with (*Cov*) and without the
406 effects of covariates (*NoCov*, all γ fixed to 0, equivalent to model M1). For each time series of
407 $\text{logit}(\theta_{t,r})$, the percentage of between year variance captured by covariate can therefore be
408 estimated by the ratio C_r :

409 (12)
$$C_r = 1 - \frac{\text{Var}_{tot_r}(\text{Cov})}{\text{Var}_{tot_r}(\text{NoCov})}$$

410 where $\text{Var}_{tot_r}(\text{noCov})$ corresponds to the total inter-annual variance for the model without
411 covariates as defined in eq. (3), and $\text{Var}_{tot_r}(\text{Cov})$ is the total inter-annual variance in the
412 model with covariates (i.e., the *residual* variance not captured by the covariate). The average
413 percentage of variance captured by a given covariate (denoted C_{TOT}) is then calculated over
414 all SU or over SU within each CSG. A high value of C_{TOT} corresponds to a high contribution
415 of the covariates in the trends of post-smolt survival.

416 Contribution of environmental covariates to generate synchrony or asynchrony in post-
417 smolt survival

418 To quantify the contribution of environmental covariates in generating synchrony or
419 asynchrony in survival, we also assessed the amount of variance captured by the covariates at
420 different levels of the spatial hierarchy:

$$421 \quad (13) \quad \Delta_{\delta} = 1 - \frac{\sigma_{\delta}^2(Cov)}{\sigma_{\delta}^2(NoCov)}$$

$$422 \quad (14) \quad \Delta_{\alpha_g} = 1 - \frac{\sigma_{\delta}^2(Cov) + \sigma_{\alpha_g}^2(Cov)}{\sigma_{\delta}^2(NoCov) + \sigma_{\alpha_g}^2(NoCov)}, \text{ for } g = NA \text{ or } SE$$

$$423 \quad (15) \quad \Delta_{\varepsilon_r} = 1 - \frac{\sigma_{\varepsilon_r}^2(Cov)}{\sigma_{\varepsilon_r}^2(NoCov)}, \text{ for } r = 1, \dots, 13$$

424 Δ_{δ} , Δ_{α_g} , and Δ_{ε_r} quantify the contribution of environmental covariates to the between year
425 variance at the global scale (general synchronous component), CSG-scale (synchronous
426 component within a CSG) or local scale (asynchronous component), respectively. Δ_{δ} and Δ_{α_g}
427 are positive if the covariate acts as a synchronizing factor. Indeed, if the covariate captures
428 part of the synchronous signal in components δ or α , the variance of the synchronous random
429 terms in the model should be lower when considering covariates (Table 2). Inversely, if Δ_{ε_r} is
430 positive, the variance of asynchronous terms is greater when considering covariates, meaning
431 that the covariate acts as an asynchronous agent (Table 2).

432 **2.3 MCMC simulations and model checking**

433 Bayesian posterior distributions were approximated using Monte Carlo Markov Chain
434 (MCMC) methods using Nimble (<https://r-nimble.org>) (de Valpine et al., 2017). The Nimble
435 code for our model is available on GitHub: [https://github.com/MaxOlmos/SALMOGLOB-](https://github.com/MaxOlmos/SALMOGLOB-Life-Cycle-Model)
436 [Life-Cycle-Model](https://github.com/MaxOlmos/SALMOGLOB-Life-Cycle-Model). Two independent MCMC chains with dispersed initialization values were
437 used. The level of autocorrelation of MCMC chains is very high (still significant at lag 30).
438 The first 10^6 iterations were used as a burn-in period. To reduce the autocorrelation in the

439 MCMC sample used for final inferences, one out of 30 iterations post burn-in was kept and
440 the resulting sample of 30,000 iterations per chain was used to characterize the posterior
441 distribution. Convergence was assessed using the Gelman-Rubin statistic (Brooks & Gelman,
442 1998) as implemented in the R Coda package (gelman.diag()).

443 Following the methodology developed in Olmos et al. (2019), the model fit to each data
444 source was assessed by checking that the 90% credibility envelope of the posterior predictive
445 distribution of each variable contained the observation. In addition, Bayesian p-values
446 calculated from chi-square discrepancy tests (Gelman et al., 2014a) were calculated to check
447 the ability of the model to replicate a posteriori data similar to those observed. The likelihood
448 and the core structure of the population dynamic is the same as in Olmos et al. (2019), and
449 changes in the latent model structure do not affect the way the model fits the data. As in
450 Olmos et al. (2019), posterior predictive distributions show that the model fits well to all
451 observations, and posterior predictive checks do not indicate strong inconsistencies between
452 the model a posteriori and the data. Those results are not developed further in this paper (see
453 Olmos et al. (2019) for more details).

454 **2.4 Model comparisons**

455 We compared the parsimony of models using the W-AIC criterion. The WAIC is appropriate
456 to compare hierarchical models of any structure fitted to the same data sets (Gelman, 2014a;
457 Hooten & Hobbs, 2015; Watanabe, 2013). It can be considered as a generalization of the
458 Deviance Information Criterion (Gelman, 2014a; Vehtari, Gelman, & Gabry, 2017) and has
459 the advantage of being directly related to the posterior predictive ability of the model. Using
460 the common convention for information criteria on the deviance scale, differences of W-AIC
461 between models can be roughly interpreted according to the following rules of thumb: a
462 difference of 1-2 units offers little to no support in favor of a particular model; a difference of
463 between 4 and 7 units offers considerable support for the model with the lowest W-AIC; and a
464 difference of >10 units offers full support for the model with the lowest W-AIC (Burnham
465 and Anderson, 2002; Gelman et al. 2014a; Gelman et al. 2014b).

466 3. RESULTS

467 3.1 Quantifying spatial synchrony at a hierarchy of spatial scales

468 3.1.1 Model evaluation

469 Model M1, which explicitly partitions the signal into a common trend plus two separate trends
470 for each CSG, appears to be the best descriptor of the spatial coherence between SU and was
471 therefore retained in the subsequent analyses (Supp. Mat. S1).

472 3.1.2 Spatial synchrony in post-smolt survival

473 Results show a strong synchrony in the temporal variations of post-smolt marine survival
474 between all SU, but with a higher coherence within CSG (Fig. 4). The average \overline{ICC}_δ relative
475 to the global scale component across all SU is 37%, indicating a strong synchrony between all
476 time series of survival and the common trends. Time series of post-smolt survival show a
477 consistent decline across the 13 SU over the study period. The global scale component
478 exhibits a decrease in survival by a factor 1.8 (natural scale, not shown) with a strong drop in
479 1987, followed by a slight increase in the early 2000s, before slightly declining again until
480 2012 (Fig. 4a). The degree of synchrony with the common trend is variable depending on the
481 SU. ICC_{δ_r} are higher for SU within the SE CSG (Fig. 4b; average value of ICC_{δ_r} across all SU
482 in SE =45%) than within the NA CSG (average value of ICC_{δ_r} across all SU in NA = 24%),
483 indicating that the SU in the SE CSG are more strongly correlated with the global scale
484 component than the SU in the NA CSG.

485 Temporal variation in post-smolt marine survival within each CSG shows a stronger
486 coherence than among SU of the two CSGs, especially as the NA CSG presents a higher
487 synchrony ($\overline{ICC}_{NA} = 60\%$) (Fig. 4d) than the SE CSG ($\overline{ICC}_{SE} = 42\%$) (Fig. 4f). Common
488 CSG trends (calculated as $(\delta_t + \alpha_{g_t})$) revealed differences between NA and SE CSG (Fig. 4c
489 and 4e), however, both exhibit an overall declining trend, characterized by a sharp decline in
490 the 1990s. The survival in the NA component decreases over years with a strong decline by a
491 factor 3 (natural scale, not shown) during 1985 to 1995 (Fig. 4c) while SE shows a smaller

492 decline by a factor 1.9. The survival in the SE component also slightly increases between
493 2002 and 2007 although it remains relatively stable in the NA component after the decline of
494 the 1990s.

495 Within NA, Quebec ($ICC_{gr}=98\%$), Labrador ($ICC_{gr}=70\%$) and Newfoundland ($ICC_{gr}=65\%$)
496 are the SU that are the most strongly correlated with the global trend for NA (Fig. 4d). Within
497 SE, the strongest correlation between SU and the common trend is obtained for England &
498 Wales ($ICC_{gr}=98\%$), Eastern ($ICC_{gr}=58\%$) and Western Scotland ($ICC_{gr}=55\%$) (Fig. 4f).
499 Some SU like US in NA or N-Ireland in SE have specific trends that contrast with the average
500 CSG trend. Logically they show weaker ICC indices ($ICC_{gr}=30\%$ for US; $ICC_{gr}=18\%$ for N-
501 Ireland). N-Ireland exhibited a higher inter-annual variability compared to the SE common
502 component (Fig. 4e). US presents a stronger decline than the NA component (Fig. 4c).

503 **3.2 Influence of environmental covariates in different space-time** 504 **domains along the migration routes**

505 **3.2.1 Time series of covariates in the different space-time domains**

506 Time series of anomalies of covariates SST and PP exhibit some temporal variations over the
507 period considered (Fig. 5).

508 SST anomalies are in the same range between the specific and the common domains (Fig 5a
509 and 5b). Time series of SST exhibit an overall increase over the period. Inter-annual variance
510 of SST anomalies is generally higher in the specific than in the common CSG domains. In
511 NA, the increase in SST is higher in the common CSG domains than in the specific ones (Fig.
512 5a). In SE, the drop of SST observed in the common CSG domain at the beginning of the
513 1990's is stronger than in the specific domains (Fig. 5b). In NA, the time series of SST in the
514 different space-time domains show large temporal fluctuations, all marked by a strong
515 increase starting at the beginning of the 1990s, and again after 2003, following a decline
516 between 1998 and 2003 (Fig. 5a). The time series of SST for the SE CSG show different
517 signals than in NA, marked by a drop of SST between 1992 and 1995, and temperatures that
518 have decreased since 2009 (Fig. 5b).

519 Trends in time series of PP are weaker than for SST and slightly decreasing (Fig. 5a to 5d).
520 The decline of PP is slightly stronger in NA than in SE. By contrast with SST, PP anomalies
521 calculated in the specific domains are much higher than in the common CSG domains.

522 AMO shows a multidecadal variability and a marked increase over the time-series, especially
523 since the beginning of the 1990's (Fig. 5e). NAOI exhibits strong inter-annual variability,
524 with a general decline after 1990 and with considerable negative anomalies in 1996 and 2010
525 (Fig. 5f).

526

527 **3.2.2 Model comparisons**

528 Difference in W-AIC between model M1 and models with effects of covariates are weak and
529 do not allow us to select one particular model (Table S4.1). Model comparisons show that the
530 temporal variations of environmental covariates experienced by post-smolts in the common
531 CSG domains better explain the variance in post-smolt survivals than environmental
532 covariates in the specific space-time domains. The models considering a different effect of
533 AMO and NAOI for each CSG are supported by the data and present low W-AIC values,
534 similar to the ones of environmental covariates defined in the common CSG domains.

535 **3.2.3 Influence of PP and SST in the specific space-time domains**

536 Overall, temporal variations of SST and PP in the specific space-time domains occupied by
537 post-smolts during the first three months of marine migration only explain a low part of the
538 temporal variance of marine survival and the sign of their influence is not consistent across
539 SU.

540 The absolute values of regression coefficients of SST and PP anomalies strongly differ (Fig.
541 6a and 6b). However, no conclusions can be drawn from those differences as the method used
542 to calculate anomalies, anomalies of PP and SST in the specific time-space domains are not
543 centered on 0 and do not have the same variance. Still, the sign of regression coefficient and
544 the amount of variance explained by both covariates can be compared.

545 Overall, the signs of the coefficient of correlations between post-smolt marine survival and
546 SST and PP anomalies considered in specific domains do not indicate a consistent direction of

547 the effect across SUs. The 95% posterior credibility intervals for most coefficients include
548 zero, suggesting a limited influence of the variations in PP or SST in specific space-time
549 domains on the marine survival rate. However, some exceptions are observed. In SE CSG,
550 temporal variation of the marine survival in Ireland, Northern Ireland, Eastern Scotland, and
551 England and Wales are negatively correlated with those of SST. In average, the SST
552 coefficients associated with the northernmost SU of both CSGs (e.g. Southwest Iceland and
553 Labrador to a lesser extend) are positive, whereas the majority of those associated with the
554 southernmost SU are negative (e.g. Scotia-Fundy, US, Ireland, Northern Ireland, England and
555 Wales, Eastern Scotland). This result suggests that the effect of SST could be different
556 depending on the latitude, with a positive effect of increasing SST for SU in the northernmost
557 post-smolt habitats, and a negative effect on SU with the southernmost post-smolt habitats.
558 On average, PP and SST integrated over the specific space-time domains only captured a
559 small percentage of the total variance, with SST being slightly more influential than PP (C_{TOT}
560 = 0.2% and 11% for PP and SST, respectively) of the temporal variations of marine survival
561 in each SU (red bars Fig. 6g and 6h).

562 Both covariates integrated over the specific domains act mainly as asynchronizing agents,
563 meaning that the covariates capture part of the variance of the asynchronous component of the
564 survival rate but not of the common trends. PP contributes only to the asynchronous
565 component (Fig. 6g, orange bars). SST also essentially contributes to the asynchronous
566 components, although some part of the synchronous components is explained by temporal
567 variation of SST in the specific domains. SST in the specific space-time domains explains
568 33% of the variance of the global component (blue bars, Δ_{δ} ; Fig. 6h), 29% of the variance of
569 the SE CSG component (green bars, $\Delta_{\alpha_{SE}}$; Fig. 6h), but 0% of the variance of the NA CSG
570 component (green bars, $\Delta_{\alpha_{NA}}$; Fig. 6h). The local (asynchronous) influence of PP is highest for
571 Newfoundland ($\Delta_{\epsilon_{NFDL}} = 22\%$), France ($\Delta_{\epsilon_{FR}} = 12\%$), Scotia-Fundy ($\Delta_{\epsilon_{SF}} = 10\%$), and explains
572 less than 10% for the Gulf, US, Quebec, Ireland, Eastern Scotland, and Northern Ireland (Fig.
573 6g). The effect of SST is highest for South West Iceland ($\Delta_{\epsilon_{SW.IC}} = 86\%$), England and Wales ($\Delta_{\epsilon_{EW}} = 25\%$),
574 and France ($\Delta_{\epsilon_{FR}} = 19\%$), but explains only a low proportion of the specific
575 variance for the other SU (Fig. 6h).

576 3.2.4 Influence of environmental PP and SST in the common CSG domains

577 SST and PP integrated over space-domains shared by all SU within the same CSG later in the
578 first year at sea explain a larger proportion of the temporal variation of marine survival than
579 variables integrated in specific space-time domains. Also, the signs of the coefficients of
580 correlation between marine survival and variables indicate a consistent direction of the effect
581 across SU.

582 The regression coefficients associated with PP integrated over common CSG domains for NA
583 and SE are positive (Fig. 6c), whereas they are negative for SST (Fig. 6d). The coefficients
584 associated with the two covariates are stronger for NA than for SE; the 95% credible intervals
585 do not include zero for SST in NA and SE CSG and for PP in NA CSG, and 75% credible
586 interval do not include zero for PP in SE CSG. Temporal variations of SST capture a greater
587 proportion of the variance (C_{TOT} 16%) of the marine survival than does PP (C_{TOT} 10%) (Fig. 6j
588 and Fig. 6i, respectively).

589 SST and PP integrated over common CSG domains act as synchronizing agents for the
590 temporal variability of post-smolt marine survival. SST accounts for 42% of the variance of
591 the global trend (blue bar, Δ_{δ} ; Fig. 6j), and PP about 19% (blue bar, Δ_{δ} ; Fig. 6i). When
592 downscaling at the scale of CSG trends ($\delta + \alpha$), SST integrated over the common CSG
593 domains accounts for 26% and 21% of between year variance of the common trends for NA
594 and SE, respectively (green bars, $\Delta_{\alpha_{NA}}$ and $\Delta_{\alpha_{SE}}$; Fig. 6j). PP accounts for 24% and 12% of
595 between year variance of the common trends for NA and SE respectively (green bars, $\Delta_{\alpha_{NA}}$
596 and $\Delta_{\alpha_{SE}}$; Fig. 6i).

597 3.2.5 Influence of large scale environmental indices: AMO and NAOI

598 The AMO is negatively correlated with the trends in post-smolt survivals (Fig. 6e), but the
599 magnitude of the effect is higher for NA than for SE. The AMO index captures a high average
600 amount of variance ($C_{TOT} = 13\%$, red bar; Fig. 6k) and acts as a synchronizing agent of post-
601 smolt survival. The effect of AMO accounts for 29% (Δ_{δ}), 26% ($\Delta_{\alpha_{NA}}$) and 21% ($\Delta_{\alpha_{SE}}$) of the
602 global-scale, the NA CSG-scale, and the SE CSG scale, respectively and does not account for
603 a specific scale component, except for England and Wales ($\Delta_{\varepsilon_{EW}} = 14\%$) (Fig. 6k).

604 The NAO index is not correlated to post-smolt survival (Fig. 6f) and captures an insignificant
605 part of the variance at any spatial scale (Fig. 6l).

606 **4. DISCUSSION**

607 Understanding the demographic and ecological mechanisms shaping the response of
608 populations to climate change is a prerequisite for a science-based management and
609 conservation ecology (Koenig, 1999). A particularly challenging issue is to separate out the
610 effects of factors acting at various stages and spatial scales. This paper addressed this issue
611 with the Atlantic salmon as a case study. We elaborate on a hierarchical life cycle model
612 developed by Olmos et al. (2019) to analyze the dynamics of 13 large groups of populations
613 that sequentially occupy different habitats in the North Atlantic Ocean, with different
614 populations occupying distinct habitats during the first period of the marine phase and sharing
615 common habitats later on. The analyses provide a new quantification of the spatial synchrony
616 in post-smolt marine survival examined at a hierarchy of spatial scales, from a basin scale
617 (North Atlantic) to more local (national or regional) scales, and quantifies the amount of
618 temporal variation in the post smolt survival that is captured by environmental changes at
619 these spatial scales. To this end, we integrated explicit hypotheses on migration routes to test
620 how spatial and temporal variations in the marine environment shape the covariation in post-
621 smolt survival rate.

622 **4.1 Geography of covariation of post-smolt marine survival**

623 We partitioned the temporal variations of marine survival for 13 SU into three components
624 that capture (i) coherence of the signal between all SU (global scale), (ii) within each CSG
625 (NA or SE), and (iii) for each SU specifically (asynchronous components).

626 Consistent with results of Olmos et al. (2019), we found strong coherence in the temporal
627 variation of post-smolt marine survival among the 13 SU of NA and SE, characterized by a
628 decline in the common trend for the 13 SU over the 1971-2014 time series.

629 Results also highlight an increased coherence in temporal variation of post-smolt survival at
630 finer spatial scales. Synchronized dynamics are stronger among SU within the same CSG than
631 between SU of different CSGs. The common trends at the scale of NA and SE capture 60%
632 and 42% of the total variance of the temporal variations, respectively, with the remaining part
633 of the variability being explained by local SU variations. Within the same CSG, synchrony is
634 higher for geographically proximate SU, most likely explained by the similarity in post-smolt
635 habitat and migration routes at sea. Specifically, in the NA CSG, Labrador, Newfoundland
636 and Quebec are closer to the common feeding grounds in the Labrador Sea and Grand Banks.
637 Fish of those SU are likely to have similar migration routes during the first year at sea, which
638 would therefore explain the strong coherence in temporal variations of marine survival.
639 Similarly, in SE, post-smolt survival rates of SU with closed migration routes to the common
640 feeding grounds in the Norwegian Sea, such as Eastern Scotland, Western Scotland and
641 England and Wales (the most abundant salmon rivers, the Tyne, Dee and Lune, are in the
642 North of England) are correlated.

643 **4.2 Influence of environment variables in space-time domains** 644 **along the migration routes**

645 The geographic pattern of covariation in post-smolt survival suggests a response to spatially
646 correlated environmental drivers (Moran effect; Liebhold et al., 2004; Stenseth, 2002; Walter
647 et al., 2017). When arriving at sea, fish occupy different habitats sequentially along their
648 migration routes and at varying levels of population aggregation. We tested if the spatial
649 patterns of synchronicity in marine survival rate can be explained by temporal variations of
650 environmental conditions (SST and PP) encountered by the fish in those different habitats.

651 To support this, we developed an extensive review of the available information on migration
652 timing and migration routes from the mouth of the estuary in spring to the first over-wintering
653 stage at the end of the following autumn. Based on this review, we defined two types of
654 space-time domains: (i) associated with the early phase of the marine life (first two months
655 after the smolts migration); (ii) associated with the later phase of the first year at sea and
656 corresponding to common areas where salmon of different origins mixed to feed.

657 Results support the hypothesis of synchronous variations of post-smolt survival driven by
658 environmental factors affecting salmon in the feeding grounds where multiple populations

659 from a same CSG forage together in late summer/early autumn, in the Labrador Sea/Grand
660 Banks for NA CSG and the Norwegian Sea for SE CSG. Temporal variations of the post-
661 smolt marine survival are best explained by temporal variations of SST (negative correlations)
662 and PP (positive correlations) in those space-time domains than in the specific ones.

663 Our results are not fully consistent with Friedland et al. (2014) who found that warm
664 temperatures in early spring negatively affected the recruitment index in NA, whereas
665 recruitment in SE was negatively correlated with warmer SST in late summer. However, our
666 inferences are based on a model that considers the population dynamics of all 13 SU in the
667 North Atlantic Ocean in a single unified modelling framework. By contrast, data limitation
668 and insufficient spatial coverage in Friedland et al. (2014) may have hampered their
669 investigation of the spatial synchrony in a hierarchy of spatial scales. The authors compared
670 proxies of marine productivity for NA and SE based on heterogeneous data sources between
671 areas (global catch index for NA CSG and an index of marine survival based on tag returns
672 from the North Esk River (UK) for SE CSG) and their correlative approach was not based on
673 explicit hypotheses about salmon migration routes. Hence the lack of correlation between
674 spring SST and marine productivity for SE CSG may come from a mismatch between the
675 habitat occupied by salmon and the space-time domains where SST was considered.

676 Although previous papers reported some weak degree of association between NAOI and
677 Atlantic salmon dynamics (Beaugrand & Reid, 2012; K. E. Mills et al., 2013), no relationship
678 between winter NAOI and post-smolt survival was found in the present study. One reason
679 might be that the relationship between NAOI and salmon is not homogeneous between the
680 two sides of the North Atlantic Ocean, and even across a latitudinal gradient within a given
681 CSG. For instance, strong positive phases of the NAOI are associated with below-normal
682 temperatures in SE and in the North of the Labrador region, but with above-normal
683 temperature in Northern Europe and in the eastern coast of North America (Hurrell, Kushnir,
684 Ottersen, & Visbeck, 2003): the effect of NAOI on salmon may therefore not be uniform
685 across SUs or CSGs.

686 **4.3 Indirect growth-dependent mechanisms are likely involved**

687 Thermal conditions encountered by salmon in their open ocean feeding grounds are likely to
688 influence salmon survival. However, it is still unclear if salmon are influenced by direct

689 effects of warming (by increasing energy expenditure and metabolism), or by indirect effects
690 such as suboptimal food availability, predation, or migration timing.

691 The direct effect of SST warming appears less likely responsible for the observed patterns.
692 Ectothermic animals such as salmon have both their metabolic demand and growth potential
693 increasing with temperature (Siegel et al., 2017). Growth variations during the post-smolt
694 phase were hypothesized as being important for survival (Friedland et al., 2008, 2014;
695 McCarthy et al., 2008). As SST indices in the open ocean feeding grounds remained within
696 the optimal range (between 7°C-10,5°C) of Atlantic salmon (Holm, 2000; Reddin and
697 Schearer, 1987), the observed SST warming over the time period studied is not likely to have
698 a direct effect on marine survival, which is contrary to negative regression coefficients as well
699 as the overall declining pattern of marine survival.

700 The negative correlation between the temporal variation of SST, and the positive correlation
701 of PP with the common trends of marine survival rate rather suggest an indirect effect of SST
702 warming acting through bottom-up trophic mechanisms. The Norwegian Sea and the Labrador
703 Sea/Grand Banks, which are major feeding grounds for SE and NA CSG populations, are
704 sensitive areas to climate change in the North Atlantic (Beaugrand et al., 2008). Our findings
705 are consistent with a major trophic shift in the North Atlantic documented in the early 1990s,
706 with reported changes across trophic levels from phytoplankton communities to seabird
707 populations (Beaugrand et al., 2008; Beaugrand, Luczak, & Edwards, 2009; Durant et al.,
708 2003) . Subsequent reduction of the abundance and the energetic quality of prey may have
709 altered salmon growth at sea (K. E. Mills et al., 2013; Otero et al., 2012; Renkawitz et al.,
710 2015) and consequently survival through size-dependent mortality (Friedland & Reddin,
711 2000; Gislason et al., 2010; Peyronnet et al., 2007). Antagonistic effects of direct and indirect
712 mechanisms may also act in synergy. Indeed, although warmer temperature may imply a
713 higher and faster growth potential, sustaining higher metabolic rates also requires higher food
714 availability. Therefore, under limited resource conditions, warmer temperatures may well lead
715 to a decrease in growth (and then of survival) because energetic demand might outweigh
716 energy intake (Daly & Brodeur, 2015; Siegel et al., 2017).

717 Another possible indirect negative effect of warmer temperature could be increasing
718 metabolic costs and mortality through reduced concentrations of dissolved oxygen. Deutsch et
719 al. (2015) highlighted that the combined effect of dissolved oxygen loss and warming would
720 reduce the metabolic index (ratio of O₂ supply to an organism's resting O₂ demand) through
721 the upper water column (0 to 400m) by ~20% globally and by ~50% in mid-latitude Northern

722 Hemisphere oceans. Investigating the combined effect of temperature warming and oxygen
723 loss on Atlantic salmon habitat would be worth considering in the future.

724 **4.4 Local specificities in temporal variations**

725 Beyond the general patterns, our results reveal some differences in the temporal variations of
726 post-smolt survival and response to environmental variations between CSG and between SU.
727 Some of these can be explained by local specificities but are mostly the consequences of
728 complex yet unexplained mechanisms.

729 Results revealed a higher decline in post-smolt survival, a stronger coherence between SU and
730 a stronger effect of PP and SST in North America than in Southern Europe. This might result
731 from the particularly fast warming of the ocean in the Northwest Atlantic, especially in the
732 Labrador Sea/Grand Banks (Belkin, 2009; A. Pershing, Dayton, Franklin, & Kennedy, 2018;
733 Taboada & Anadón, 2012). Additionally, weaker synchrony observed in SE may be explained
734 by the diversity of marine environments and associated growth conditions encountered during
735 the post-smolt migration leading to the feeding areas in the Norwegian Sea. Post-smolt diet
736 reported by Haugland et al. (2006) shows high spatial and temporal variability in terms of
737 prey composition, with diet dominated by blue whiting in the Shelf Edge Current in the west
738 of the United Kingdom, and by sandeel and herring in the North Sea and the Norwegian Sea.
739 Such a high portfolio of potential prey depending on the area may reduce the synchronizing
740 effect of environmental fluctuations.

741 The sign of the correlations between post-smolt survival and SST integrated over the specific
742 time-space domains occupied by salmon during the first three months of marine migration are
743 not consistent across SUs. Specifically, a negative correlation between post-smolt survival of
744 Scotia-Fundy, US, Ireland, Northern Ireland and England and Wales and SST integrated over
745 specific domains may result from particularly warm temperatures in those space-time
746 domains. In particular, SST in the Gulf of Maine has increased faster than 99% of the global
747 ocean (A. J. Pershing et al., 2015), which could explain why US and Scotia-Fundy SU present
748 the strongest declines in post-smolt survivals.

749 A negative correlation between PP integrated over the specific space-time domains and
750 survival in the Labrador and Newfoundland and France SU was not expected. Salmon from
751 the Labrador and Newfoundland SU have a shorter migration to the common feeding grounds

752 (Bley & Moring, 1988; Friedland, 1994) and may directly migrate to the common domains to
753 feed. The spatial resolution of the specific space-time domain defined for the France SU, that
754 encompasses the Biscay Bay, the North Sea and the Western coast of UK, could be too large.
755 Populations from the North (Brittany and Normandy) and from South West of France may
756 have different migration routes. Mixing different ecosystems with different possible trophic
757 dynamics may have blurred the signal (Jensen et al., 2012; Haugland et al., 2006).

758 **4.5 Limits and future prospects**

759 In this study, we only considered a limited set of environmental covariates, namely SST, PP,
760 and large-scale climate proxies, AMO and winter NAOI. The limited set of tested variables
761 result from a trade-off between hypothesis testing about the mechanisms that drive post-smolt
762 survival and the availability of data over the required spatial and temporal scales.

763 Extending the approach to the Northern Europe (NE) SU would also allow extending the
764 gradient of environmental variation and may contribute to an even better understanding of the
765 response of Atlantic salmon populations to large scale ecosystem changes. Stock assessment
766 data for NE SU is only available for a shorter time series (starting in 1995 only; ICES, 2015).
767 Nevertheless, it is characterized by a general decline in productivity over the period,
768 suggesting a likely synchrony among the three CSG over the North Atlantic. Incorporating
769 this data into the model would allow for extending the modelling framework to all SU in the
770 North Atlantic.

771 Because of data limitations, the model structure forces all temporal variations in survival to
772 occur between the smolt migration and the Pre-Fishery stage. Indeed, as already discussed by
773 Massiot-Granier et al. (2015) and Olmos et al. (2019), the data currently available do not
774 allow partitioning out the temporal variations of the marine mortality that occur at different
775 periods of the marine phase. To better understand the effect of environmental variations on
776 marine survival, more data would be needed to partition out the natural mortality along the
777 migration routes.

778 We also assumed that the space-time domains sequentially occupied by salmon during the
779 first year at sea have not changed over the 1971 to 2014 period. However, both the timing of
780 smolt migration (Otero et al., 2012; Satterthwaite et al., 2014), and the boundaries of
781 favorable habitat at sea (Cheung et al., 2009; Poloczanska et al., 2013) have changed which

782 may have altered salmon migration routes (Guðjónsson et al. 2015). In addition, spring
783 plankton blooms and therefore the peak of higher trophic resources available for salmon may
784 be advanced in the season and may occur in different places (Edwards et al., 2010; Malick,
785 Cox, Mueter, Peterman, & Bradford, 2015; Parmesan & Yohe, 2003), thus potentially creating
786 a mismatch between salmon migration and available resources (Cushing, 1990).

787 Last, our findings have direct management implications. Indeed, post-smolt marine survival is
788 one of the main factors controlling Atlantic salmon stock productivity. Accurately accounting
789 for and forecasting temporal variation in the post-smolt marine survival will provide for a
790 more robust stock assessment and the provision of multi-year catch advice for the mixed-
791 stock fisheries occurring on these SUs within the North Atlantic (ICES, 2019; Vert-pre et al.
792 2013; Britten et al. 2016). Also, developing models that account for the effect of
793 environmental covariates on forecasting is critical to be able to integrate climate predictions
794 scenarios in those forecasts. In this perspective, building models that appropriately consider
795 how environmental changes can impact groups of populations simultaneously or differently is
796 therefore critical to develop appropriate management measures at various spatial scales.

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802

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1085 **Tables**

1086 Table 1: Prior distributions used for the parameters of the hierarchical structures. N refers to a normal distribution
 1087 and U refers to a uniform distribution.

SPATIAL COMPONENT	PARAMETERS	
	$r= 1:13$ $g= NA, SE$	Prior distribution
Specific Intercept	μ	$\sim N(0, \sigma = 10)$
	σ_{β}	$\sim U(0,5)$
Standard deviation Specific component	σ_{ε_r}	$\sim U(0,5)$
Standard deviation CSG-specific component	σ_{α_g}	$\sim U(0,5)$
Standard deviation Global component	σ_{δ}	$\sim U(0,5)$

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1090 Table 2: Summary of the hypotheses tested and the associated model configurations that included environmental covariates.

SPATIAL SCALE k	TEMPORAL PERIOD OF THE POST-SMOLT PHASE	COVARIATE X_t	MODELS	Covariate contributes to the decline in post-smolt-survival if	Covariate acts as synchronizing agent at the global scale if	Covariate act as synchronizing agent at the CSG scale if	Covariate act as asynchronizing agent if
Influence of covariates in the specific space-time domains	Spring of the first year at sea	PP SST	$\text{logit}(\theta_{t,r}) = \beta_r + \delta_t + \alpha_{g_t} + \varepsilon_{t,r} + \gamma_r \times X_{r_t}$ <p style="text-align: center;">with $\gamma_r \sim U(-6,6)$ $r = 1:13$</p>	$C_{TOT} > 0$	$\Delta_\delta > 0$	$\Delta_\alpha > 0$	$\Delta_{\varepsilon_r} > 0$
Influence of the covariates in the common space-time domains	Late summer of the first year at sea <i>Large Scale Indices</i>	PP SST AMO NAO	$\text{logit}(\theta_{t,r}) = \beta_r + \delta_t + \alpha_{g_t} + \varepsilon_{t,r} + \gamma_g \times X_{g_t}$ <p style="text-align: center;">with $\gamma_g \sim U(-6,6)$ $g = NA \text{ or } SE$</p>				

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1093 Legends

1094 Figure 1: Structure of the age- and stage-based life cycle model and covariation structure among the 13 stock units
1095 (adapted from Olmos et al. 2019). Sources of covariation include: 1) covariation in the time series of post-smolt
1096 survival and proportion maturing as 1SW (depending on the model structure M1, M2 and M3; see Table 1); 2)
1097 Covariation through fisheries operating on mixtures of SU at sea. Red boxes refer to NA SU ($i = 1, \dots, 6$), blue boxes refer
1098 to SE SU ($j = 1, \dots, 7$) and purple boxes refer to both NA and SE SU.

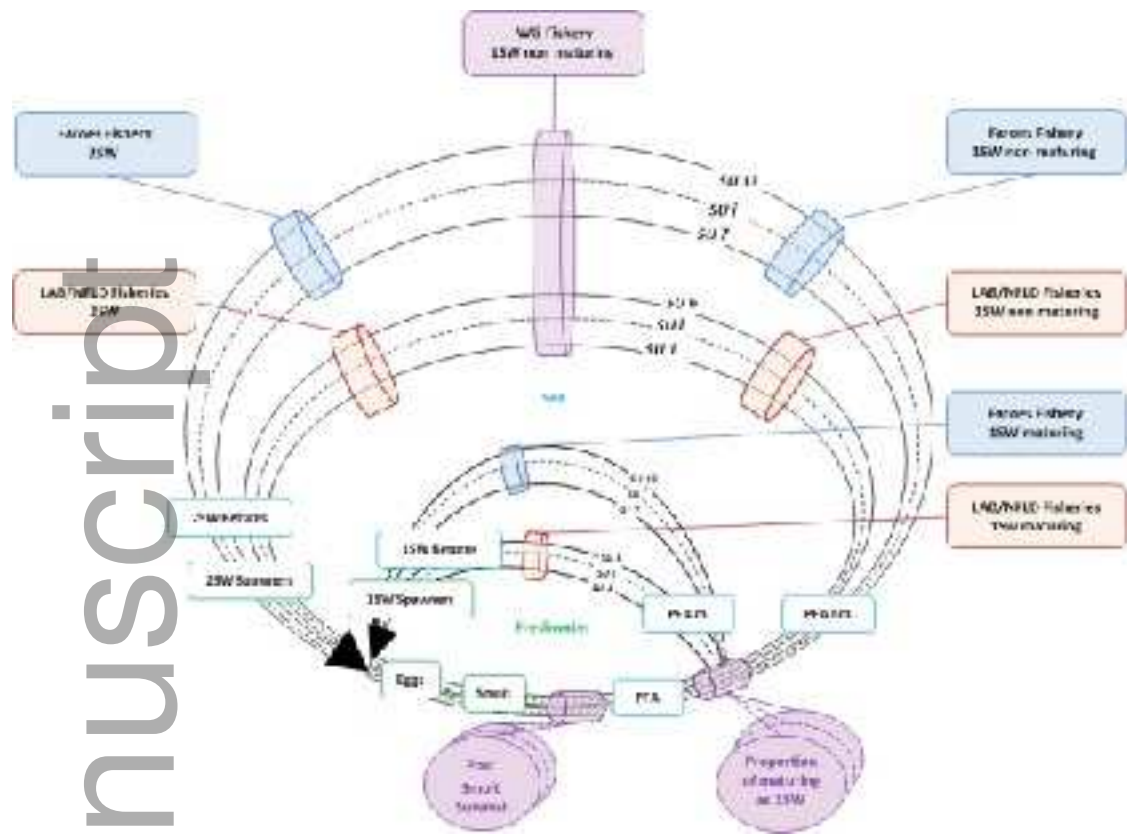
1099 Figure 2: Location of the 13 stock units and specific and common space-time domains considered in North Atlantic.
1100 Stock units of North America: NFLD = Newfoundland, GF = Gulf, SF = Scotia-Fundy, US = USA, QB = Quebec, and
1101 LB=Labrador. Stock units in Southern Europe: IR = Ireland, E&W = UK(England and Wales), FR = France, E.SC =
1102 UK(Eastern Scotland), W.SC = UK(Western Scotland), N.IR = UK(Northern Ireland), and SWIC= Southwest Iceland).
1103 Specific and common CSG space-time domains are in orange and green respectively. See Table S2.1 for
1104 correspondence between space-time domains and SU.

1105 Figure 3. Theoretical representation of the hierarchy of space-time domains. Orange: Space-time domains defined as
1106 transit habitat occupied by post-smolts during their first two months at sea (specific domains). Green: Space-time
1107 domains corresponding to the habitat occupied by salmon in the later phase of the first year at sea, associated with
1108 feeding areas, common to all SU within the same CSG (Labrador Sea and the Norwegian Sea for NA and SE CSG
1109 respectively) (common domain).

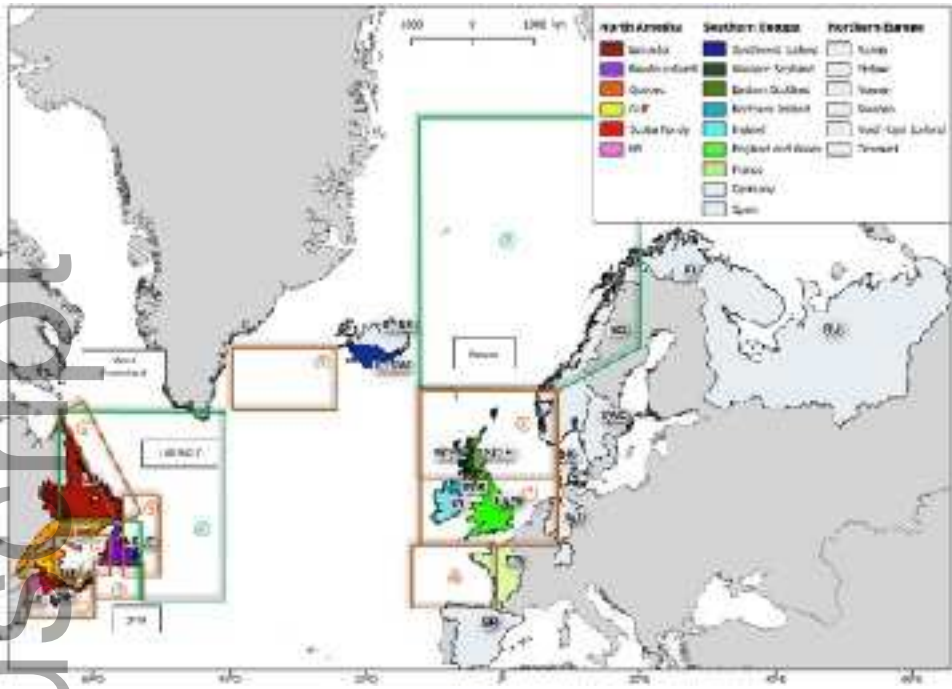
1110 Figure 4: Left panel: Large scale component trends (medians of marginal posterior distributions and 95% credibility
1111 interval (shaded area)) and individual time-series medians for each SU estimated. (a) Global component and the 13 SU
1112 (c) NA component and NA SU (e) SE component and SE SUs. Right panel: Synchronicity indices ICC quantifying the
1113 proportion of variance captured by the average trend for each SU within the global component (b), the NA CSG (d) and
1114 the SE CSG (f). The average ICC (ICCM) is indicated by the thick line. The thick line corresponds to the median of the
1115 ICCm and the shaded area represent the 95% credibility interval.

1116 Figure 5: Time series of environmental covariates: Sea Surface Temperature in NA (a) and SE (b); Primary Production
1117 in NA (c) and SE (d). For (a) to (d): time series defined in the common space-time domains (green color) and time
1118 series defined in the specific space-time domains (color range from pink to maroon); (e) standardized Atlantic
1119 Multidecadal Oscillation (AMO); and (f) standardized North Atlantic Oscillation Index (NAOI).

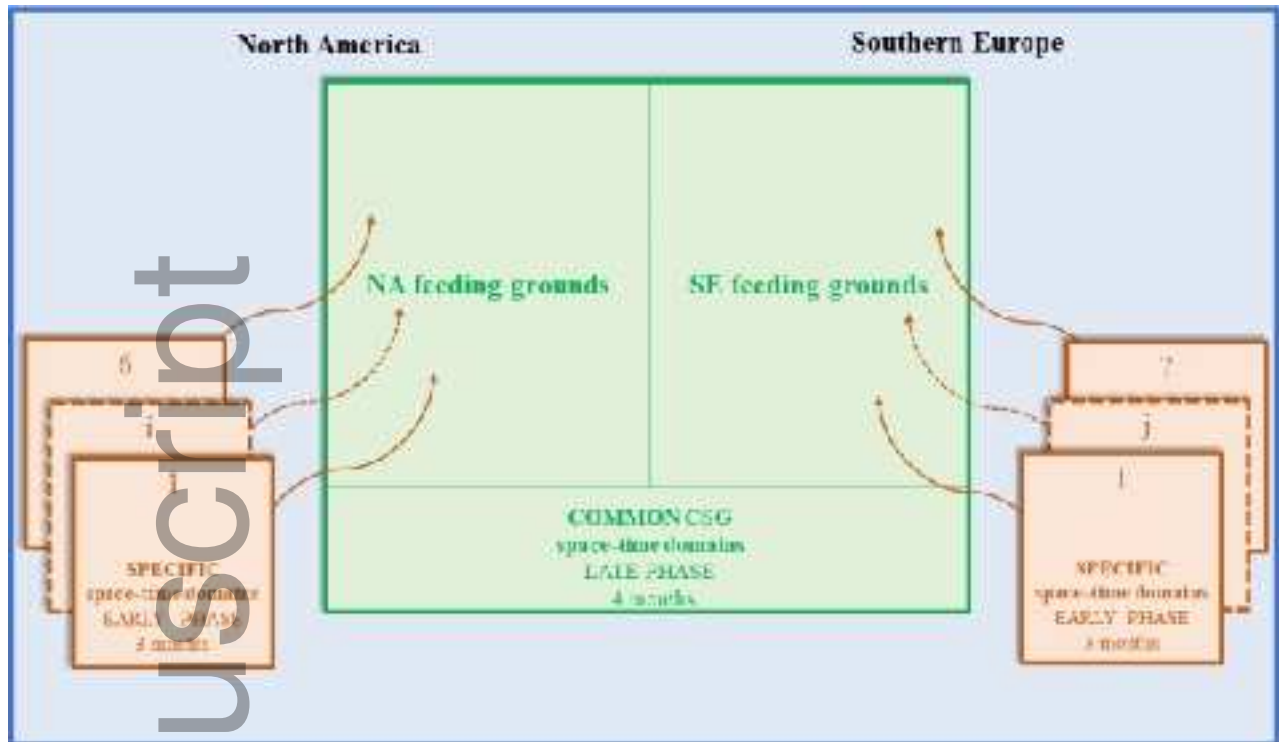
1120 Figure 6: Regression coefficients ((a)-(f)) and fraction of temporal variation of post-smolt survival accounted for by
1121 effect of environmental covariates ((g)-(l)) defined in the specific space-time domains ((a) & (g)) (PP), ((b) & (h))
1122 (SST), in the CSG space-time domains ((c) & (i)) (PP), ((d) & (j)) (SST) ((e) & (k)) (AMO), ((f) & (l)) (NAOI). Left panels:
1123 marginal posterior distributions for the regression coefficients. Thick point is the median, and the different
1124 thicknesses of lines represent the 50%, the 75% and the 95% posterior credibility intervals. Right panels: fraction of
1125 variation accounted for by the covariates. C_{TOT} is the average fraction captured by the covariates when considered
1126 over all stock units (red). Blue, green, and orange barplots represent the contribution of environmental covariates to
1127 the between year variance at the global-scale (Δ_δ), CSG-scale (Δ_{α_g}) and specific scale (Δ_{ϵ_r}) respectively.



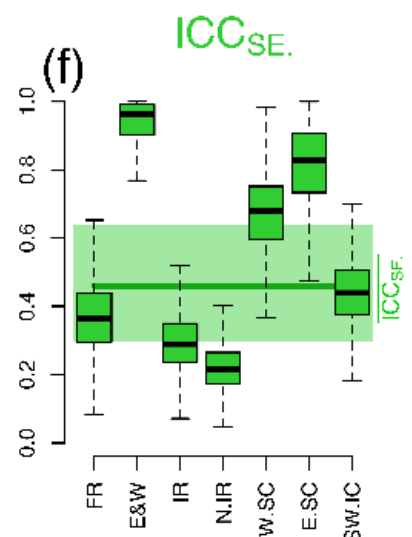
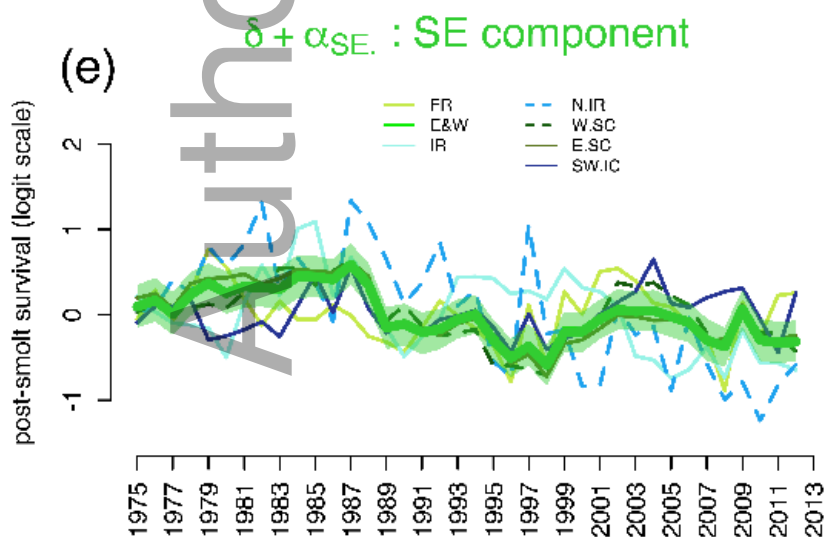
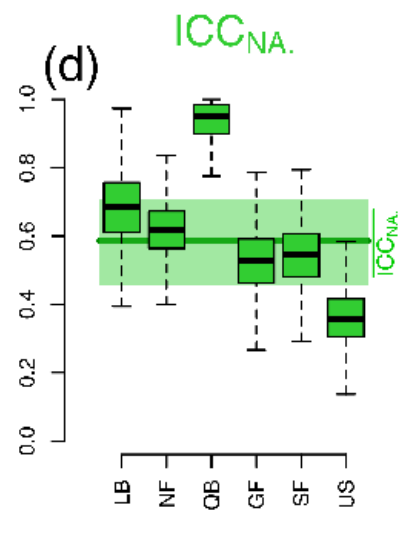
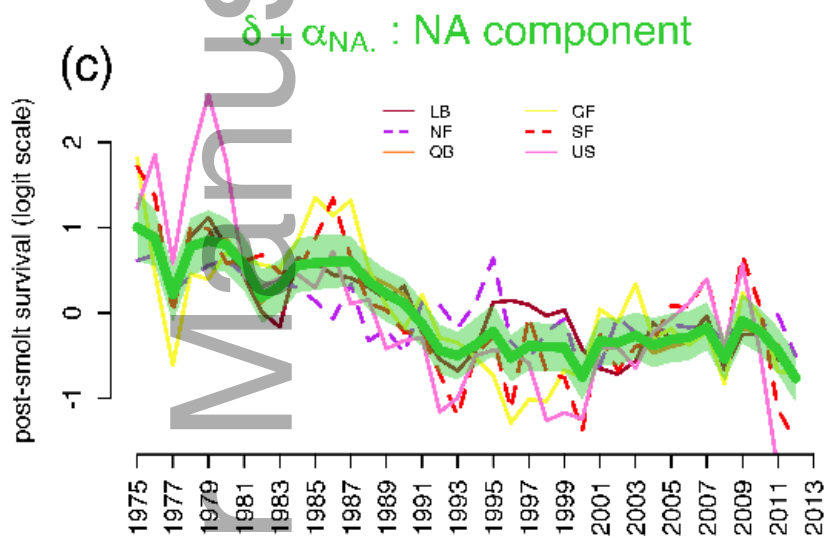
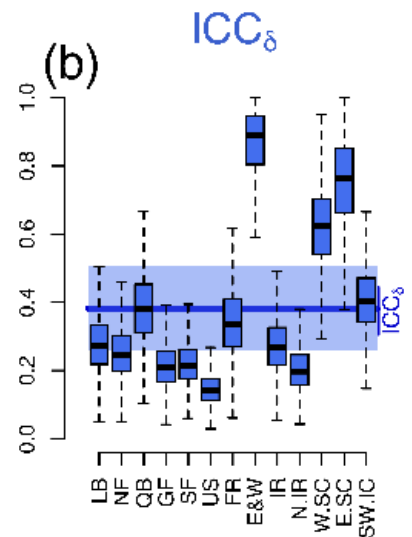
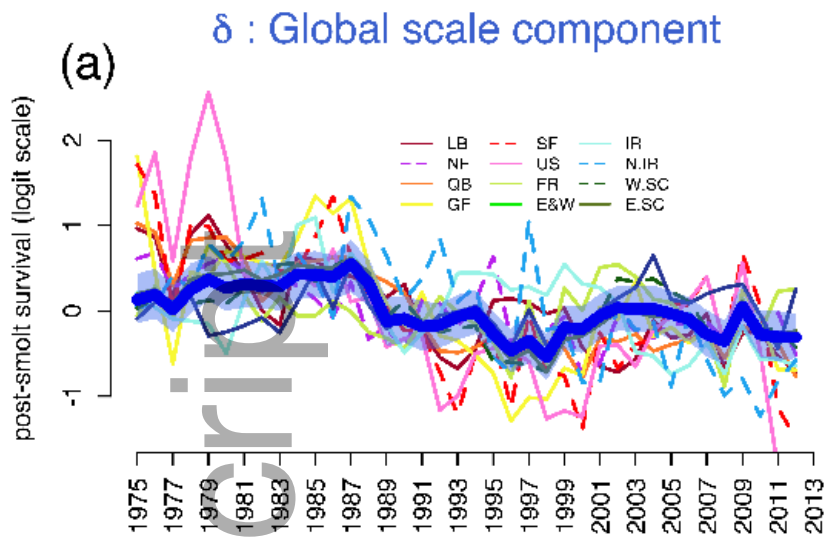
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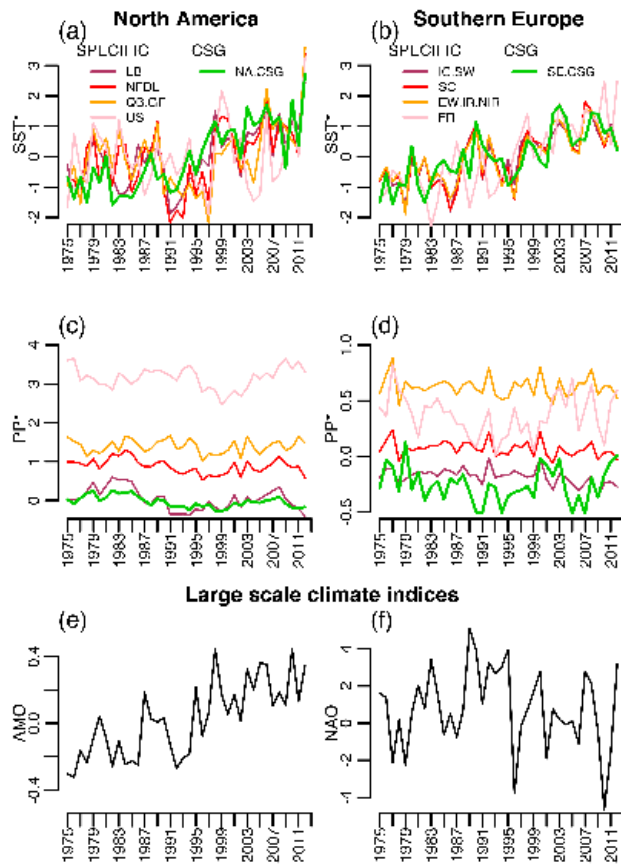


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