

Quota allocation for stocks that span multiple management zones: analysis with a vector autoregressive spatiotemporal model

Running title: VAST spatial allocation

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Abstract

Allocating quotas among stakeholders requires an agreed upon formula. If the stock unit spans multiple management jurisdictions, the formula may require updated biomass estimates of the stock's spatial distribution with respect to those jurisdictions. Data for calculating stock biomass often come from fishery-independent surveys. While stratified random sampling is a common design, strata boundaries may not always align with state or national jurisdictions, requiring post hoc stratification and imputation to calculate area-specific biomass. The vector autoregressive spatiotemporal (VAST) model was explored as a tool for calculating fish biomass within subareas of a defined stock unit for three different stocks jointly managed by the United States and Canada on Georges Bank in the Northwest Atlantic Ocean. VAST estimated proportions of stock biomass in each nation's waters were compared with an existing allocation algorithm that utilizes a loess smooth through the average design-based swept area biomass from three fishery-independent surveys. The ability of VAST to impute biomass when no tows occur in a subarea of a survey stratum was demonstrated, as well as the relatively smoother biomass trend compared with design-based estimates, which may be desirable if the intent is to avoid large inter-annual swings in allocated quota.

32 **1. Introduction**

33 Allocating quotas among stakeholders is a management challenge, requiring a transparent
34 management process and buy-in from stakeholders. An agreed-upon formula that considers
35 factors such as total landings, capital investment and years fished for a specified reference period
36 (Anderson and Holliday, 2007) can provide the basis for achieving that transparency. If the
37 managed stock spans multiple management jurisdictions (adjacent states or provinces within a
38 nation, or multiple nations with an international boundary that intersects the stock unit), then
39 recent estimates of spatial distribution of the stock with respect to those jurisdictions could be a
40 component of the allocation formula that is regularly updated.

41 Globally, it is estimated there are 344 marine fish species whose stock boundaries span
42 the territorial waters of multiple nations and the high seas in between (Teh and Sumaila, 2015).
43 These include transboundary stocks, defined as fish that cross from the boundary of one
44 exclusive economic zone (EEZ) into the EEZs of one or more coastal countries; and highly
45 migratory stocks, such as tunas (*Thunnus* spp.) and billfish (*Istiophoridae*, *Xiphiidae*), that are
46 managed through multinational regional fisheries management organisations, such as the
47 International Commission for the Conservation of Atlantic Tunas (ICCAT). There are also
48 examples of inland multijurisdictional cooperation: The Great Lakes Fishery Commission
49 between the United States of America (USA) and Canada; The Lake Victoria Fisheries
50 Organization, which facilitates between Kenya, Tanzania and Uganda; and the Mekong River
51 Commission, which includes Cambodia, Lao PDR, Thailand and Vietnam as full signatories
52 (Lynch et al., 2016).

53 In the Northwest Atlantic Ocean, species of groundfish range along the coasts of the USA
54 and Canada, with individual stock boundaries that bridge the jurisdiction of one or more states
55 (coastal waters), federal waters, extending beyond coastal out to the EEZ, and transboundary
56 waters where the EEZ of the USA and Canada meet (the Hague Line), officially delimited by the
57 International Court of Justice in the Hague, The Netherlands (International Court of Justice,
58 1984).

59 On Georges Bank (Figure 1), three species are jointly managed by the USA and Canada
60 through a scientific body that performs assessments, the Transboundary Resource Assessment
61 Committee (TRAC). The species are: Atlantic cod (*Gadus morhua* L.), haddock
62 (*Melanogrammus aeglefinus* (L.)) and yellowtail flounder (*Limanda ferruginea* (Storer)). The

63 existing quota allocation formula incorporates historical utilization (1967–1994 total landings by
64 country) and proportion of biomass in each country's waters from a smoothed average of three
65 federal bottom trawl surveys. The latter are updated annually by incorporating data from the
66 latest survey (e.g. 2017) and dropping data from the earliest survey used in the previous year
67 (e.g. 1984) so that a 33-year window is maintained (Barrett & Brooks, 2018). While this
68 allocation formula is simple, transparent and well documented (Murawski & Gavaris, 2004), it
69 requires post-stratification of two of the surveys to accommodate the Hague line and
70 management unit borders for two of the stocks (Figure 2). These management boundaries were
71 established several decades after the USA survey strata were standardised. This post-
72 stratification occurs at a finer spatial scale and occasionally results in low or no sampling
73 occurring in the substrata due to the random allocation of tows at the spatial scale of the original
74 survey strata (although procedures have been put in place since the autumn 2018 survey to
75 ensure sampling in these substrata). Results of the allocation procedure may be sensitive to gaps
76 in survey sampling (substrata where no tows occurred) or very low sampling (substrata with one
77 to two tows), and the terminal year estimate of the loess smoothing algorithm (Cleveland, 1979)
78 applied to the survey average proportion.

79 Recent development of a vector autoregressive spatiotemporal (VAST) model (Thorson
80 & Barnett, 2017) provides a tool to estimate biomass in defined spatial areas, and could be useful
81 in quota allocation (Thorson, 2019). In the Northwest Atlantic Ocean, VAST has been used to
82 standardise indices for the northern shrimp (*Pandalus borealis* Krøyer) assessment (Cao et al.,
83 2017); to combine data from multiple fishery-independent surveys to predict density estimates
84 for cusk (*Brosme brosme* (Ascanius)) for use in habitat suitability indices (Runnebaum et al.,
85 2018), as well as to predict hotspots of cusk bycatch in the American lobster (*Homarus*
86 *americanus* Milne-Edwards) fishery (Runnebaum et al., 2020); to test the incorporation of
87 environmental covariates in a length-structured assessment of the American lobster (Hodgdon et
88 al., 2020); and to examine distribution shifts for cod (Guan et al., 2017) and summer flounder
89 (*Paralichthys dentatus* (L.)) in the region (Perretti & Thorson, 2019).

90 For this application, the ability of VAST to estimate spatial autocorrelation could
91 improve the estimates of biomass at unsampled locations near stock boundaries or in post hoc
92 substrata, and previous research has noted greater precision of indices compared with design-
93 based estimators (Thorson et al., 2015). Estimates of biomass were derived using VAST,

94 focusing on three stock units that are co-managed by the TRAC (Figure 2): eastern Georges
95 Bank (EGB) Atlantic cod (hereafter referred to as cod) and haddock; and GB yellowtail flounder
96 (hereafter referred to as yellowtail). The primary objective was to demonstrate the use of VAST
97 for estimating area-specific biomass and proportion of resource within each nation's jurisdiction.
98 The ability of VAST to overcome the identified limitations and sensitivities to the current
99 allocation method for these stocks is shown. Resource distribution and biomass results from the
100 current allocation method are compared with the results from VAST applied to the same data.

101

102 **2. Methods**

103 2.1. Data

104 Data from two fishery-independent survey programmes that sampled on both sides of the Hague
105 line were used in this analysis. The Northeast Fisheries Science Center (NEFSC) has conducted
106 spring and autumn bottom trawl surveys on the continental shelf of the Northeast United States
107 since 1968 and 1963, respectively. Data from 1985 to 2017 were used in this analysis (to align
108 with the 33-year moving window), with spring and autumn treated as two separate surveys
109 (Barrett & Brooks, 2018). The survey employs a stratified random design. Strata are defined
110 primarily by depth and latitude. Several gear and vessel changes have occurred over the course
111 of the survey (Miller et al., 2010; Johnston & Sosebee, 2014) and conversion factors to account
112 for these changes were applied as necessary.

113 Fisheries and Oceans Canada has conducted a spring bottom trawl survey on GB since
114 1986, with full coverage of strata beginning in 1987. Data from 1987 to 2017 were used in this
115 analysis (Barrett & Brooks, 2018). The survey employs a stratified random design. Strata are
116 defined primarily by depth, as well as the location of the international boundary and geographic
117 regions on the bank. Two vessels have been used over the course of the survey, but no
118 conversions are necessary as they are considered identical (Benoît, 2006; Stone & Gross, 2012).

119 2.2. Background

120 Details of the current TRAC quota allocation method can be found in Murawski and Gavaris
121 (2004). Briefly, design-based indices of swept area biomass (mt) are calculated for all three
122 surveys, for the country-specific area on each side of the Hague Line. Since the NEFSC survey
123 strata are split by both the Hague Line and the boundary of the EGB region (Figure 2), there are
124 years when no tows occurred within these modified strata due to the random allocation of tows

125 within the original strata. These cases have historically used either undocumented imputation to
126 fill strata with no tows, or they were assumed to be zero. For this application, no imputation was
127 performed, and strata biomass was assumed to be zero if no tows occurred there. The proportion
128 of biomass of each species within each country's waters was calculated from the estimated
129 biomass on each side of the Hague Line. A combined survey proportion for each country is
130 calculated as the simple average of the proportions from the three surveys for haddock and
131 yellowtail, while for cod the Canadian and NEFSC spring proportions were averaged first, and
132 this was averaged with the autumn NEFSC proportions. A loess smooth (span = 0.3) was then
133 applied to the country-specific proportion to remove unpredictable fluctuation and sampling
134 variation from the time series. To determine quota allocation estimates for the current year, the
135 result of the smoother (i.e. the current year estimate of proportion of a given stock in each
136 country's water) was weighted by 90% and the fraction of historic utilisation weighted by 10%
137 to determine the overall fraction of quota that was allocated to each country. This overall fraction
138 was then multiplied by the quota for the whole stock area (based on results from annual stock
139 assessments) to obtain country-specific quota amounts.

140 VAST is an open source (<https://github.com/James-Thorson-NOAA/VAST>) package in
141 the R statistical environment (R Core Team, 2019). To facilitate comparisons with the current
142 TRAC approach, all three survey data sets were combined to produce a single index, with an
143 uncertainty estimate, for each stock, with each of the three surveys treated as a vessel effect with
144 overdispersion (Grüss et al., 2017; Thorson, 2019, section 4.1). This differs from a recent VAST
145 study (Thorson et al., 2020) that modelled seasonal variation in spatial distribution of yellowtail.
146 It was noted that, while this previous work combined all three of the same surveys into one
147 model, the final output was separate indices of relative abundance or biomass for each season (or
148 survey), and thus was not suitable for the current objective of a single index, with an uncertainty
149 estimate, for each stock. Given the low number of tows east/west of the Hague Line in some
150 years, and to better inform points near the EGB boundary, species-specific spatial domains were
151 based on the larger stock unit defined by USA management for cod and haddock rather than the
152 smaller management unit agreed to for the TRAC. The USA and TRAC stock definition are
153 identical for yellowtail. From these stock-specific spatial domains, the predicted biomass was
154 estimated, and the proportion of biomass was calculated east/west of the Hague Line for either

155 EGB (cod, haddock) or GB (yellowtail). As VAST is by definition a smoother, a loess was not
156 applied to the calculated biomass proportion (as is done for the current allocation method).

157 2.3. VAST model

158 Details and equations of the VAST model have been published elsewhere (Thorson et al., 2015;
159 Thorson & Barnett, 2017; Thorson, 2019). Briefly, VAST is a spatiotemporal delta generalised
160 linear mixed model. The default delta model includes a logit-linked linear predictor for encounter
161 probability, and a log-linked linear predictor for expected catch rate, given a positive encounter.
162 There is also an option for a Poisson link model, which has a log-link for encounter probability.
163 The default error distribution for positive catch rates is the gamma. There is also an option for a
164 lognormal error distribution. The footnote in Table 1 shows the combinations of encounter
165 probability link functions and observed error distributions examined for each of the three stocks.
166 Altogether, there are 15 major decisions that must be made by users of VAST (Thorson, 2019).
167 The decisions made in the present analysis are summarised in Table 2. Another decision that
168 must be made is the number of prediction locations, i.e. knots. The recommendation of 1000
169 knots for index standardisation (Thorson, 2019) was used.

170 Model convergence was checked by ensuring that the Hessian of the likelihood function
171 was positive definite, and that the absolute value of the final gradient of parameters was less than
172 0.0001. The Akaike Information Criterion, or AIC (Burnham & Anderson, 2002), was used to
173 select the best model run for each stock. Additionally, standard diagnostic outputs from VAST
174 (e.g. Q-Q plot) were examined to ensure that there was no strong evidence of misspecification.

175

176 2.4. Comparison with the current TRAC method

177 The VAST estimates of biomass, with 95% confidence intervals, were compared with the
178 design-based swept area estimates used in the TRAC allocation. Comparisons were plotted for
179 east/west of the Hague Line for EGB cod and haddock, and GB yellowtail. Proportion of total
180 biomass east/west of the Hague Line was also calculated. Qualitative differences in biomass
181 trend and quantitative differences in biomass proportion were summarised; the TRAC allocation
182 is ad hoc and does not produce estimates of uncertainty so precision cannot be compared.

183 VAST performance was evaluated for cases where the current TRAC allocation method
184 has known limitations and sensitivities. As noted in above, there are a number of years in which
185 there has been zero, or only one to two, tows on either side of the Hague Line in the thin strata

186 along the edge of the bank, i.e. strata 17–18 and 21–22 (Figure 1). For the current TRAC
187 allocation method, biomass in these strata was assumed to be zero if no tows occurred. Thus, it
188 was of interest to compare VAST estimates of biomass for these strata. This was done using the
189 Georges Bank spring haddock data (1968–2015) that come with VAST to facilitate replication
190 by other investigators. Default delta-model settings were used with 1000 knots.

191 In addition to testing the performance of VAST when no tows occurred in a stratum,
192 VAST was evaluated to determine if estimates of annual biomass are sensitive to new years of
193 data as the 33-year window moves forward, as this is a known sensitivity for the current
194 allocation method. This was done by comparing the estimated biomass trends and proportion of
195 biomass trends east/west of the Hague Line for two consecutive time series: 1985–2017 and
196 1986–2018. This analysis was limited to one stock (cod). In particular, the behavior of terminal
197 year estimates from VAST were compared to determine if they are less sensitive than those from
198 the loess smoother used in the current approach. A Poisson link model was used with 1000 knots.

199

200 2.5. Software

201 NEFSC strata must be specified as a named list of area codes in VAST. Strata coordinates were
202 contained in the northwest_atlantic_grid that comes with VAST. As it does not contain the split
203 EGB strata (Figure 2), it was recalculated for this study using ArcGIS Desktop 10.7 (ESRI, Inc.,
204 Redlands, California, USA). The NEFSC survey strata were projected into a customised North
205 America Albers Equal Area Conic projection to maximise accuracy in later area calculations.
206 Another new feature class was created in the same projection consisting of 3.7×3.7 km (2×2
207 nmi) gridded polygons using the fishnet tool. A geometric intersection of these two feature
208 classes produced the final feature class. The centre longitude and latitude (in decimal degrees) as
209 well as the area (in km²) of each individual polygon were then calculated. The attribute table was
210 outputted as a text file for reading into R.

211 R version 3.6.0 and VAST version 2.1.0 were used in this analysis. VAST uses Template
212 Model Builder (Kristensen et al., 2016) to estimate fixed effects while integrating through
213 random effects, and the R-INLA package (Illian et al., 2012) to model spatial variation.
214 Versions of these packages used in this analysis were 1.17.15 and 18.07.12, respectively.

215

216 3. Results

217 Diagnostics (Supporting Information) suggested no obvious misspecification for any model run.
218 AIC indicated that a Poisson link for encounter probability was the best for all three stocks
219 (Table 1). In terms of error distribution, gamma was best for cod and haddock, while the
220 lognormal was better for yellowtail. It is also worth noting that for cod, the Δ AIC between run 2
221 and run 1 was 0.65, indicating essentially no difference in the link function used for encounter
222 probability.

223 All four runs estimated similar trends for each of the stocks with only minor differences
224 in biomass, although there were slightly higher biomass estimates with the lognormal error
225 distribution (runs 3 and 4) for haddock (Figure 3). The coefficient of variation (CV) suggests
226 reasonably precise biomass estimates (Figure 3): CVs for cod were less than 0.35; while CVs for
227 haddock and yellowtail were, with the exception of 1986, less than 0.3. For cod and haddock, run
228 1, with a gamma error distribution and logit link had the lowest CV, while Run 4 (lognormal
229 error distribution with Poisson link) almost always had the highest CV.

230

231 3.1. Comparison with current TRAC method

232 VAST estimates of biomass, with 95% confidence intervals, were compared with the TRAC
233 allocation swept area estimates in Figure 4. VAST proportions of biomass in USA waters (i.e.,
234 west of the Hague Line) were also compared with TRAC estimates of resource proportion, as
235 well as the loess smooth for the latter. In general, the trend of population biomass on each side of
236 the Hague Line was similar for the two methods, but the estimates following the current TRAC
237 allocation method showed strong annual variability (Figure 4, top and middle rows). The annual
238 variability was similarly present in the TRAC estimates of average proportion of biomass per
239 country (Figure 4, bottom row). These large annual swings in proportion would translate to large
240 annual swings in quota allocation, if taken at face value. The loess smoothed proportion of
241 biomass, which is the final step currently used for allocation, showed a more stable trend through
242 time. Comparing the loess smoothed annual proportions with the VAST estimated proportions,
243 there were differences in the direction of the trend for the first points in the time series (1985) for
244 all three stocks, and for the last point in the time series (2017) for yellowtail. In any given year,
245 the VAST estimated proportion in USA waters differed from the loess smoothed proportion
246 between ± 0.0005 and ± 0.32 , depending on the stock.

247 In each stratum area where no tows occurred, VAST predicted non-zero biomass. As
248 these were narrow strata with very little fraction of the total stock area, the predicted biomasses
249 were not very large, and ranged from 0.2 t to 2195 t (Table 3), with a trend towards higher
250 biomass estimates in recent years due to the historic high biomass for haddock (Figure 4). VAST
251 estimates for strata with no tows ranged from 0 to 72% of the total biomass estimated in those
252 years. Similarly, in each stratum where only one to two tows occurred, VAST predicted biomass
253 ranging from 0.3 t to 7935 t, as compared with 0 t to 20,964 t from the TRAC estimates, with the
254 largest differences occurring in stratum 21 and 22 east of the Hague Line (Table 3).

255 The consecutive fits of VAST to 1985–2017 and 1986–2018 EGB cod data showed no
256 sensitivity to the biomass estimate in 2017, as opposed to the consecutive fits with the TRAC
257 method (Figure 5): in the loess for 1985–2017 the estimated trend “chases” the high biomass
258 estimate in 2017; but when the loess is applied for 1986–2018, the smoothed estimate for 2017 is
259 much lower. The VAST estimate of biomass in 2017 in both fits was stable because VAST uses
260 spatial correlation across the spatial domain to inform estimates, whereas the TRAC method
261 relies on stratum means with occasionally few or no tows in a given stratum. In addition, the
262 loess smoother predicted values at the terminal year were unstable because that point was only
263 informed by data from earlier years, and when additional data were added for later years these
264 can influence estimates that had previously been endpoints.

265

266 **4. Discussion**

267 The primary objective of this study was to demonstrate the use of VAST to estimate the
268 proportion of stock biomass in country-specific waters for multijurisdictional allocation of
269 quotas. This was contrasted with the current approach that uses the stratified random sampling
270 design to estimate average proportion of biomass in each management zone, to which a loess
271 curve is then fit. An advantage of the VAST approach is that it provides an objective way to fill
272 strata with no tows and inform strata with few tows based on tows in neighboring strata, which is
273 important when strata are post-stratified to deal with allocation issues. Another advantage of
274 VAST is that it smooths large inter-annual fluctuations that are potentially due to outlier tows or
275 low sample size rather than real trends in the population. This was observed for USA cod
276 biomass in 2005 and 2010, and Canadian yellowtail biomass in 2008 and 2009, and led to
277 estimated biomass spikes in those years in the design-based method. Previous comparisons of

278 VAST versus design-based indices have also found that the latter exaggerated temporal
279 variability (Cao et al., 2017). However, Hodgdon et al. (2020) found that model-based
280 abundance indices were not intrinsically better than design-based indices and should be tested for
281 each species individually, as has been done in this study.

282 Across the four model configurations defined by the link function (logit or Poisson) and
283 assumed error distribution (gamma or lognormal), no observed error distribution was
284 consistently best. The gamma was better for cod and haddock; while the lognormal was better for
285 yellowtail. There were two extreme outlier tows for yellowtail, one in both 2008 and 2009
286 (Figure 4), and it was hypothesised that the lognormal error distribution provided a better fit due
287 to its heavier tail compared to the gamma. Regarding the link function, AIC indicated that the
288 Poisson link models were generally best, but in one case (cod) there was essentially no difference
289 from the conventional delta-model. These findings support the recommendation of Thorson
290 (2019) to compare the performance of conventional delta-models versus Poisson link models,
291 and to explore multiple distributions.

292 The general trend in proportion of biomass was similar for the VAST estimate and the
293 loess smoothed estimate. Although absolute differences of 0.0005–0.32 existed, there was no
294 consistent trend in the direction of which country had more or less proportion estimated in a
295 given year. The current TRAC method allocates quotas for the upcoming year based on the
296 current year's estimate of resource distribution, implicitly assuming it will be similar. The
297 application of a loess with a span of 0.3 provides some responsiveness, while removing
298 fluctuations that may be due to sampling variation (Murawski & Gavaris, 2004). Given the
299 relative similarity of trend between VAST and TRAC estimates of biomass proportion, the
300 question is whether one method is more appropriate than the other. One advantage of the VAST
301 approach was highlighted in model performance explorations, focusing on strata areas where no
302 tows, or only one to two tows occurred. In all instances where no tows occurred, VAST
303 estimated non-zero biomass; with the current TRAC approach, ad hoc imputation would be
304 necessary to infer biomass. Thus, VAST provides a better way to fill a missing stratum than
305 assuming it has zero biomass, and it is preferable to ad hoc imputation for transparency and
306 reproducibility considerations. The performance of VAST for models such as this has been tested
307 previously through simulation (Thorson et al., 2015; Grüss & Thorson, 2019; Grüss et al., 2019;
308 Johnson et al., 2019; Brodie et al., 2020), and future simulation experiments exploring

309 performance for the combined surveys method proposed here are recommended. Specifically,
310 future work could explore the accuracy of VAST imputed biomass in these cases.

311 The USA proportion (Figure 4) suggests potentially significant differences in annual
312 allocations between those from the current TRAC method and the VAST estimates. Confidence
313 intervals are not available for the TRAC allocation method, so it is not possible to tell whether
314 intervals for the two methods overlap, but there are several instances for each stock where the
315 TRAC loess estimates are outside the 95% confidence interval for the VAST estimates: 1994 and
316 2015 for EGB cod; 1994, 1998, 2001, 2003, 2008, 2012 and 2014–2016 for EGB haddock; and
317 2003, 2006 and 2011 for GB yellowtail. Although these differences did not consistently favour
318 one country, the country-specific quotas are obtained by multiplying proportion of country-
319 specific biomasses by a total allowable quotas (based on results from a stock assessment of the
320 total stock area); thus, in a given year the scale of total allowable quota could make the
321 magnitudes of those differences in proportion financially significant. Moreover, it is not
322 uncommon that available quota for one species can impact fishers' decisions about targeting and
323 effort for other species, ultimately affecting individual vessel income. Thus, even seemingly
324 minor differences can be important.

325 In the present study, the ability of VAST to calculate biomass within subareas of a
326 defined stock unit for three stocks jointly managed by the USA and Canada on the Georges Bank
327 was explored. A similar approach could be taken for other transboundary species. For example,
328 quota for Pacific halibut (*Hippoglossus stenolepis* Schmidt) is shared between the USA and
329 Canada based on regional biomass estimates (Cox et al., 2013). A review of the allocation
330 suggested that a combined spatio-temporal smoothing applied to each year could help to retain
331 spatial consistency in biomass across regulatory areas. Other examples include stocks in the
332 Northeast Atlantic, such as Barents Sea cod, which has experienced a poleward displacement
333 since the late 1980s (Gullestad et al., 2020).

334 VAST offers some advantages over currently applied methods to allocate quota in
335 situations where there are multijurisdictional considerations. It overcomes sampling gaps,
336 provides estimates of biomass with uncertainty bounds, and is less sensitive to outlier tows and
337 sharp trend changes at the end of the time series. The cost of this more sophisticated method is
338 potentially revisiting model selection as data are updated. Formal restratification, to address
339 management boundaries that are defined after survey strata definitions and ensure proper tow

340 allocation, would be another way to avoid sampling gaps. However, inevitable challenges such
341 as weather delays and mechanical problems may be insurmountable in some years, leaving gaps
342 in sampling despite best intentions to avoid such occurrences. Thus, a method such as VAST
343 provides reliability against the unforeseen and unavoidable realities of field sampling.

344 Spatial boundaries help make making environmental issues more manageable (Lidskog et
345 al., 2011). On the other hand, it is known that such boundaries are a political and management
346 construct, which may not be aligned with, nor respected by, ecological and human components
347 essential to commercial fishing (Song et al., 2017). Ecologically, boundary mismatches can
348 create added pressure on fish stocks, resulting in overfishing (Song et al., 2017). Given changes
349 in fish distribution worldwide, whether due to climate change or stock expansion, survey data
350 should be used not only to set catch limits, but to re-examine catch shares (Fernandes and Fallon,
351 2020). This VAST analysis illustrates how temporal changes in spatial distributions can be
352 modeled and incorporated into the allocation of stocks that span multiple management
353 jurisdictions, thereby aiding in the conservation of fishery resources.

354

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358

359 **References**

360 Anderson, L. G., & Holliday, M. C. (2007). The design and use of limited access privilege
361 programs. NOAA Technical Memo NMFS-F/SPO-86.

362 <https://www.fisheries.noaa.gov/webdam/download/92512602>

363 Barrett, M. A., & Brooks, E. N. (2018). Update of allocation shares for Canada and the USA of
364 the transboundary resources of Atlantic cod, haddock, and yellowtail flounder on
365 Georges Bank through fishing year 2019. Transboundary Resources Assessment
366 Committee Reference Document 2018/04.

367 <https://repository.library.noaa.gov/view/noaa/24817>

368 Benoît, H. P. (2006). Standardizing the southern Gulf of St. Lawrence bottom trawl survey time
369 series: Results of the 2004–2005 comparative fishing experiments and other
370 recommendations for the analysis of the survey data. Canadian Science Advisory

- 371 Secretariat Research. Document 2006/008. <http://waves-vagues.dfo->
372 [mpo.gc.ca/Library/331162.pdf](http://waves-vagues.dfo-mpo.gc.ca/Library/331162.pdf)
- 373 Brodie, S., Thorson, J. T., Carroll, G., Hazen, E. L., Bograd, S., Haltuch, M. A., Holsman, K. K.,
374 Kotwicki, S., Samhuri, J. F., Willis-Norton, E., & Selden, R. L. (2020). Trade-offs in
375 covariate selection for species distribution models: a methodological comparison.
376 *Ecography*, 43(1), 11–24. <https://doi.org/10.1111/ecog.04707>
- 377 Burnham, K. P., & Anderson, D. R. (2002). Model selection and multimodel inference: A
378 practical information-theoretic approach (2nd ed.). New York: Springer
- 379 Cao, J., Thorson, J. T., Richards, R. A., & Chen, Y. (2017). Spatiotemporal index standardization
380 improves the stock assessment of northern shrimp in the Gulf of Maine. *Canadian Journal*
381 *of Fisheries and Aquatic Sciences*, 74(11), 1781–1793. <https://doi.org/10.1139/cjfas->
382 [2016-0137](https://doi.org/10.1139/cjfas-2016-0137)
- 383 Cleveland, W. S. (1979). Robust locally weighted regression and smoothing scatterplots. *Journal*
384 *of the American Statistical Association*, 74(368), 829–836.
385 <https://doi.org/10.1080/01621459.1979.10481038>
- 386 Cox, S. P., Ianelli, J., & Mangel, M. (2013). Reports of the IPHC scientific review board.
387 International Pacific Halibut Commission, Report of Assessment and Research Activities
388 2013, 218–225.
- 389 Fernandes, P. G., & Fallon, N. G. (2020). Fish distributions reveal discrepancies between zonal
390 attachment and quota allocations. *Conservation Letters*, 13(3), e12702.
391 <https://doi.org/10.1111/conl.12702>
- 392 Grüss, A., & Thorson, J. T. (2019). Developing spatio-temporal models using multiple data types
393 for evaluating population trends and habitat usage. *ICES Journal of Marine Science*,
394 76(6), 1748–1761. <https://doi.org/10.1093/icesjms/fsz075>
- 395 Grüss, A., Thorson, J. T., Sagarese, S. R., Babcock, E. A., Karnauskas, M., Walter III, J. F.,
396 Drexler, M., 2017. Ontogenetic spatial distributions of red grouper (*Epinephelus morio*)
397 and gag grouper (*Mycteroperca microlepis*) in the U.S. Gulf of Mexico. *Fisheries*
398 *Research*, 193, 129–142. <https://doi.org/10.1016/j.fishres.2017.04.006>
- 399 Grüss, A., Walter III, J. F., Babcock, E. A., Forrestal, F. C., Thorson, J. T., Laretta, M. V., &
400 Schirripa, M. J. (2019). Evaluation of the impacts of different treatments of spatio-

- 401 temporal variation in catch-per-unit-effort standardization models. *Fisheries Research*,
402 213, 75–93. <https://doi.org/10.1016/j.fishres.2019.01.008>
- 403 Guan, L., Chen, Y., Staples, K. W., Cao, J., & Li, B. (2017). The influence of complex structure
404 on the spatial dynamics of Atlantic cod (*Gadus morhua*) in the Gulf of Maine. *ICES*
405 *Journal of Marine Science*, 74(9), 2379–2378. <https://doi.org/10.1093/icesjms/fsx064>
- 406 Gullestad, P., Sundby, S. & Kjesbu, O. S. (2020). Management of transboundary and straddling
407 fish stocks in the Northeast Atlantic in view of climate-induced shifts in spatial
408 distribution. *Fish and Fisheries*, 21(5), 1008–1026. <https://doi.org/10.1111/faf.12485>
- 409 Hodgdon, C. T., Tanaka, K. R., Runnebaum, J., Cao, J., & Chen, Y. 2020. A framework to
410 incorporate environmental effects into stock assessments informed by fishery-
411 independent surveys: a case study with American lobster (*Homarus americanus*).
412 *Canadian Journal of Fisheries and Aquatic Sciences* 77(10), 1700–1710.
413 <https://doi.org/10.1139/cjfas-2020-0076>
- 414 Illian, J.B., Sørbye, S.H., & Rue, H. (2012). A toolbox for fitting complex spatial point process
415 models using integrated nested Laplace approximation (INLA). *Annals of Applied*
416 *Statistics*, 6(4), 1499–1530. <http://dx.doi.org/10.1214/11-AOAS530>
- 417 International Court of Justice. (1984). Case concerning delimitation of the of the maritime
418 boundary in the Gulf of Maine area (Canada/United States of America). *International*
419 *Court of Justice Reports* 246. Retrieved from <https://www.icj-cij.org/en/case/67>
- 420 Johnson, K. F., Thorson, J. T., & Punt, A. E. (2019). Investigating the value of including depth
421 during spatiotemporal index standardization. *Fisheries Research*, 216, 126–137.
422 <https://doi.org/10.1016/j.fishres.2019.04.004>
- 423 Johnston, R., & Sosebee, K. (2014). History of the United States bottom trawl surveys, NAFO
424 subareas 4–7. Northwest Atlantic Fisheries Organization SCR Document 14-024.
425 <https://www.nafo.int/Portals/0/PDFs/sc/2014/scr14-024.pdf>
- 426 Kristensen, K., Nielsen, A., Berg, C. W., Skaug, H., & Bell, B. M. (2016). TMB: Automatic
427 differentiation and Laplace approximation. *Journal of Statistical Software*, 70(5), 1–21.
428 <http://dx.doi.org/10.18637/jss.v070.i05>
- 429 Lidskog, R., Ugglå, Y., & Soneryd, L. (2011). Making transboundary risks governable: reducing
430 complexity, constructing spatial identity, and ascribing capabilities. *AMBIO*, 40, 111–
431 120. <https://doi.org/10.1007/s13280-010-0123-3>

- 432 Lynch, A. J., Cooke, S. J., Deines, A. M., Bower, S. D., Bunnell, D. B., Cowx, I. G., Nguyen, V.
433 M., Nohner, J., Phouthavong, K., Riley, B., Rogers, M. W., Taylor, W. W., Woelmer, W.,
434 Youn, S.-J., & Beard, T. D. Jr. (2016). The social, economic, and environmental
435 importance of inland fish and fisheries. *Environmental Reviews*, 24(2), 115–121.
436 <https://doi.org/10.1139/er-2015-0064>
- 437 Miller, T. J., Das, C., Politis, P. J., Miller, A. S., Lucey, S. M., Legault, C. M., Brown, R. W.,
438 Rago, P. J. 2010. Estimation of Albatross IV to Henry B. Bigelow calibration factors.
439 NEFSC Reference Document 10-05. <https://repository.library.noaa.gov/view/noaa/3726>
- 440 Murawski, S., & Gavaris, S. (2004). Computation of allocation shares for Canada and the USA
441 of the transboundary resources of Atlantic cod, haddock and yellowtail flounder on
442 Georges Bank. Transboundary Resource Assessment Committee Reference Document
443 2004/05. [https://www.bio.gc.ca/info/intercol/trac-](https://www.bio.gc.ca/info/intercol/trac-cert/documents/ref/TRD_2004_05_E.pdf)
444 [cert/documents/ref/TRD_2004_05_E.pdf](https://www.bio.gc.ca/info/intercol/trac-cert/documents/ref/TRD_2004_05_E.pdf)
- 445 Perretti, C. T., & Thorson, J. T. (2019). Spatio-temporal dynamics of summer flounder
446 (*Paralichthys dentatus*) on the Northeast US shelf. *Fisheries Research*, 215, 62–68.
447 <https://doi.org/10.1016/j.fishres.2019.03.006>
- 448 R Core Team (2019). R: A language and environment for statistical computing. R Foundation for
449 Statistical Computing, Vienna, Austria. Retrieved from <https://www.R-project.org/>
- 450 Runnebaum, J., Guan, L., Cao, J., O'Brien, L., & Chen, Y. (2018). Habitat suitability modeling
451 based on a spatiotemporal model: an example for cusk in the Gulf of Maine. *Canadian*
452 *Journal of Fisheries and Aquatic Sciences*, 75(11), 1784–1797.
453 <http://dx.doi.org/10.1139/cjfas-2017-0316>
- 454 Runnebaum, J., Tanaka, K. R., Guan, L., Cao, J., O'Brien, L., & Chen, Y. (2020). Predicting
455 bycatch hotspots based on suitable habitat derived from fishery-independent data. *Marine*
456 *Ecology Progress Series*, 641, 159–175. <https://doi.org/10.3354/meps13302>
- 457 Song, A. M., Scholtens, J., Stephen, J., Bavinck, M., & Chuenpagdee, R. (2017). Transboundary
458 research in fisheries. *Marine Policy*, 76, 8–18.
459 <https://doi.org/10.1016/j.marpol.2016.10.023>
- 460 Stone, H. H., & Gross, W. E. (2012). Review of the Georges Bank research vessel survey
461 program, 1987–2011. Canadian Manuscript Report of Fisheries Aquatic Sciences 2988.
462 <http://waves-vagues.dfo-mpo.gc.ca/Library/345565.pdf>

- 463 Teh, L. S. L., & Sumaila, U. R. (2015). Trends in global shared fisheries. *Marine Ecology*
 464 *Progress Series*, 530, 243–254. <https://doi.org/10.3354/meps11049>
- 465 Thorson, J. T. (2019). Guidance for decisions using the Vector Autoregressive Spatio-Temporal
 466 (VAST) package in stock, ecosystem, habitat and climate assessments. *Fisheries*
 467 *Research*, 210, 143–161. <https://doi.org/10.1016/j.fishres.2018.10.013>
- 468 Thorson, J.T., Adams, C. F., Brooks, E. N., Eisner, L. B., Kimmel, D. G., Legault, C. M.,
 469 Rogers, L., & Yasumiishi, E. (2020). Seasonal and interannual variation in spatio-
 470 temporal models for index standardization and phenology studies. *ICES Journal of*
 471 *Marine Science*, 77(5), 1879–1892. <https://doi.org/10.1093/icesjms/fsaa074>
- 472 Thorson, J. T., & Barnett, L. A. K. (2017). Comparing estimates of abundance trends and
 473 distribution shifts using single- and multispecies models of fishes and biogenic habitat.
 474 *ICES Journal of Marine Science*, 74(5), 1311–1321.
 475 <https://doi.org/10.1093/icesjms/fsw193>
- 476 Thorson, J.T., Shelton, A.O., Ward, E.J., & Skaug, H.J. (2015). Geostatistical delta-generalized
 477 linear mixed models improve precision for estimated abundance indices for West Coast
 478 groundfishes. *ICES Journal of Marine Science*, 72(5), 1297–1310.
 479 <https://doi.org/10.1093/icesjms/fsu243>

480

481 **Tables**

482

483 Table 1. Akaike information criterion (AIC) for VAST model runs for each of the three stocks,
 484 with all models based on 77 parameters. Runs are sorted by AIC, and Δ AIC relative to the model
 485 with the lowest AIC is also shown.

486

Cod			Haddock			Yellowtail		
Run	AIC	Δ AIC	Run	AIC	Δ AIC	Run	AIC	Δ AIC
2	32062.13	0	2	46100.37	0	4	18159.15	0
1	32062.78	0.64	4	46151.88	51.50	2	18216.15	56.99
3	32132.00	69.87	1	46195.46	95.09	3	18250.13	90.97
4	32163.11	100.98	3	46275.89	175.52	1	18306.41	147.26

Note:

Run 1: gamma error distribution; logit link for encounter probability

Run 2: gamma error distribution; Poisson link for encounter probability

Run 3: lognormal error distribution; logit link for encounter probability

Run 4: lognormal error distribution; Poisson link for encounter probability

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490 Table 2. Summary of 15 major decisions for VAST used in this analysis

491

#	Description	Decision
1	Spatial domain used when calculating derived quantities	Eastern Georges Bank [†] for cod [‡] and haddock [§] Georges Bank [†] for yellowtail flounder [¶]
2	Which categories (species/sizes) to include	Cod [‡] , haddock [‡] and yellowtail flounder [§] ; each analyzed separately
3	Identify whether to analyze encounter, abundance, and/or biomass sampling data	Biomass
4	Including spatial and/or spatiotemporal variation	Both
5	Choosing the spatial smoother and resolution	Anisotropic Matérn correlation function
6	Choosing the number of spatial and spatiotemporal factors	On
7	Specifying temporal correlation on model components	Fixed effects
8	Including density covariates as a semi-parametric model	NA
9	Accounting for catchability covariates and confounding variables	NA
10	Treating area swept as a catchability covariate or offset	On
11	Including vessel effects as overdispersion	On
12	Choosing among distributions and link functions	Logit & Poisson for encounter probability link function Gamma & lognormal for observed error distribution
13	Derived quantities	Biomass index
14	Bias correction for derived quantities	On
15	Model selection	AIC

[†] See Figure 2 for geographic boundaries

[‡] *Gadus morhua*

[§] *Melanogrammus aeglefinus*

[¶] *Limanda ferruginea*

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494 Table 3. Haddock spring biomass (mt) in Northeast Fisheries Science Center strata 17–18 and
495 21–22, west (USA) and east (Canada = CAN) of the Hague Line, on eastern Georges Bank
496 (Figure 2) using the Transboundary Resource Assessment Committee (TRAC) allocation method
497 (left) and VAST (right). For the TRAC allocation method, “0” represents observed zeros, italics
498 denotes one to two tows, and NA indicates that no tows occurred. One significant digit is shown
499 for VAST to emphasize that the model did not estimate zero biomass in any substratum in any
500 year. Only years that overlap for the Georges Bank spring haddock data that comes with VAST
501 (1968–2015) and the data used in the analysis in the main body of the text (1985–2017) are
502 shown

503

504 Table is shown on next page

505

506

Year	TRAC								VAST							
	USA 17	CAN 17	USA 18	CAN 18	USA 21	CAN 21	USA 22	CAN 22	USA 17	CAN 17	USA 18	CAN 18	USA 21	CAN 21	USA 22	CAN 22
1985	14	99	0	18	NA	3696	NA	54	8.1	124.3	3.3	46.0	121.4	1046.8	55.4	165.9
1986	0	21	0	0	0	1297	NA	0	4.8	119.3	1.5	32.1	35.6	947.8	10.4	165.7
1987	0	101	0	17	NA	63	NA	69	1.7	88.2	0.5	23.7	86.6	861.9	26.5	120.6
1988	0	13	NA	0	0	310	0	0	2.1	78.9	0.6	20.2	28.0	440.2	9.4	74.0
1989	28	146	NA	79	0	751	NA	256	10.9	236.0	4.0	79.5	70.2	616.5	25.2	161.1
1990	0	64	NA	NA	33	1305	NA	21	4.2	210.6	1.3	57.7	52.2	856.4	19.4	156.0
1991	NA	37	NA	0	0	28	NA	0	2.3	140.1	0.7	29.5	15.3	230.0	5.5	33.2
1992	NA	80	NA	0	NA	376	NA	0	0.7	74.5	0.2	17.3	10.5	242.2	4.1	45.1
1993	NA	439	NA	0	NA	387	NA	154	1.5	128.9	0.5	38.8	39.6	557.7	14.9	102.4
1994	11	1	0	NA	6	5644	NA	0	1.1	43.7	0.4	14.5	22.8	1739.5	7.3	277.6
1995	NA	60	0	NA	NA	3356	NA	888	1.1	56.2	0.3	14.1	82.2	2180.1	29.5	404.4
1996	NA	32	NA	0	NA	972	31	0	1.3	46.8	0.4	13.5	98.3	914.9	36.4	161.7
1997	10	28	0	11	45	1239	NA	74	3.2	100.6	1.3	32.4	82.3	995.4	28.1	185.7
1998	3	84	NA	5	282	227	0	108	1.9	125.3	0.8	38.9	96.3	797.2	38.1	154.2
1999	0	1598	NA	0	42	366	37	38	4.8	319.4	1.7	83.5	65.7	637.3	34.0	117.6
2000	0	220	0	NA	522	151	NA	55	1.8	153.5	0.6	41.2	185.0	1409.9	64.2	258.1
2001	NA	446	NA	0	1214	4339	NA	15	10.7	266.5	4.4	92.8	267.8	1400.4	94.3	313.5
2002	0	332	NA	16	0	896	93	77	11.5	281.5	4.4	93.6	171.1	1318.3	70.2	242.3
2003	2	77	NA	0	1123	NA	19	NA	4.7	454.5	1.8	116.2	161.8	1648.1	40.6	312.6
2004	NA	977	NA	75	NA	669	NA	2	116.4	811.6	44.3	261.2	518.8	1402.6	145.9	232.3
2005	680	948	0	NA	NA	3945	132	484	90.2	557.8	40.7	208.0	386.6	2641.1	155.9	613.1
2006	5	323	0	97	143	4140	NA	40	21.2	290.8	8.6	112.8	221.2	2006.8	89.3	444.9
2007	7	64	0	90	295	795	NA	123	7.0	515.8	3.0	148.1	294.7	2223.2	91.9	297.4

2008	2	135	NA	164	484	151	NA	204	5.0	185.1	1.8	62.3	230.2	784.7	113.6	140.1
2009	100	279	0	42	7452	7085	0	22	82.1	411.6	30.4	156.3	1186.6	3599.1	400.7	699.4
2010	105	96	0	168	1553	3379	41	125	47.9	382.1	19.5	147.1	616.6	2464.1	294.2	436.9
2011	19	978	0	179	415	3008	NA	843	36.9	711.4	13.9	248.8	405.0	3105.1	198.5	701.4
2012	17	2321	1	358	NA	4138	NA	115	31.0	1166.8	13.1	429.4	514.4	2352.5	226.2	414.3
2013	0	634	0	335	NA	20964	140	371	15.4	526.9	6.3	230.8	584.5	6623.6	221.2	1031.0
2014	151	2872	0	565	798	2737	NA	1401	113.6	1572.1	61.9	641.2	1308.3	6758.4	485.5	1454.5
2015	417	1071	67	858	NA	6472	NA	796	285.9	1988.0	150.4	914.2	2195.4	7934.6	903.0	1834.8

VAST spatial allocation

510 **Figure captions**

511

512 Figure 1. Northeast Fisheries Science Center (NEFSC) bottom trawl survey offshore strata, with
513 Georges Bank (GB) strata 13–22 in black.

514

515 Figure 2. Detail of Northeast Fisheries Science Center (NEFSC) strata 13–21 used by the
516 Transboundary Resource Assessment Committee (TRAC) for the assessment of Georges Bank
517 (GB) yellowtail flounder (upper); detail of NEFSC strata 16–22 used by the TRAC for the
518 assessment of eastern Georges Bank (EGB) Atlantic cod and haddock (middle). Fisheries and
519 Oceans Canada bottom trawl survey strata 5Z1–5Z4 overlaid on the NEFSC strata for
520 comparison (lower). The Hague Line separating USA and Canada jurisdictions is shown as a
521 dotted line in all plots.

522

523 Figure 3. Biomass (mt) and coefficient of variation (CV) for the four VAST model runs for each
524 stock. Run 1: gamma error distribution; logit link for encounter probability. Run 2: gamma error
525 distribution; Poisson link for encounter probability. Run 3: lognormal error distribution; logit
526 link for encounter probability. Run 4: lognormal error distribution; Poisson link for encounter
527 probability

528

529 Figure 4. Comparison of Transboundary Resource Assessment Committee (TRAC) allocation
530 method swept area biomass (solid line) and VAST estimates of biomass (dashed line) for USA
531 (upper) and Canada (middle). Shaded region is the 95% confidence interval for the VAST index.
532 Lower panel shows the proportion USA for TRAC (solid line) and VAST estimates (dashed).
533 Blue line is the loess smooth for the TRAC proportion USA. Shaded region is the 95%
534 confidence interval for the proportion USA computed from the VAST index.

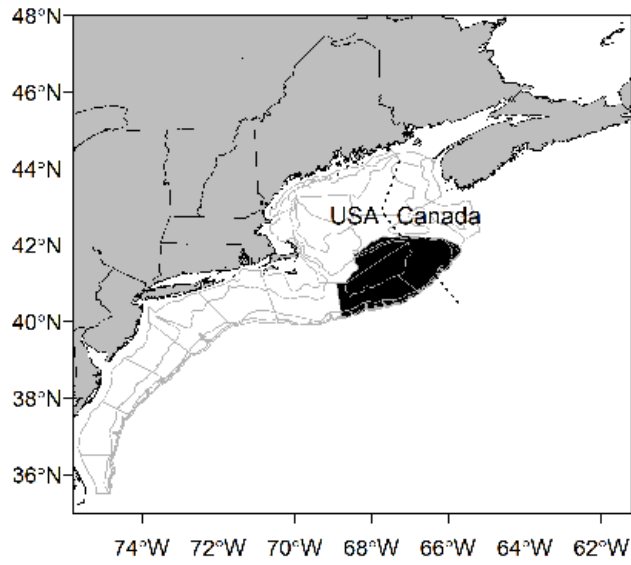
535

536 Figure 5. Comparison of the proportion USA for Transboundary Resource Assessment
537 Committee (TRAC) and VAST estimates of eastern Georges Bank (EGB) cod biomass. TRAC
538 proportion USA is shown for 1985–2018, with loess fits for 1985–2017 and 1986–2018. VAST
539 estimates of proportion USA are shown for 1985–2017 and 1986–2018. A Poisson link model

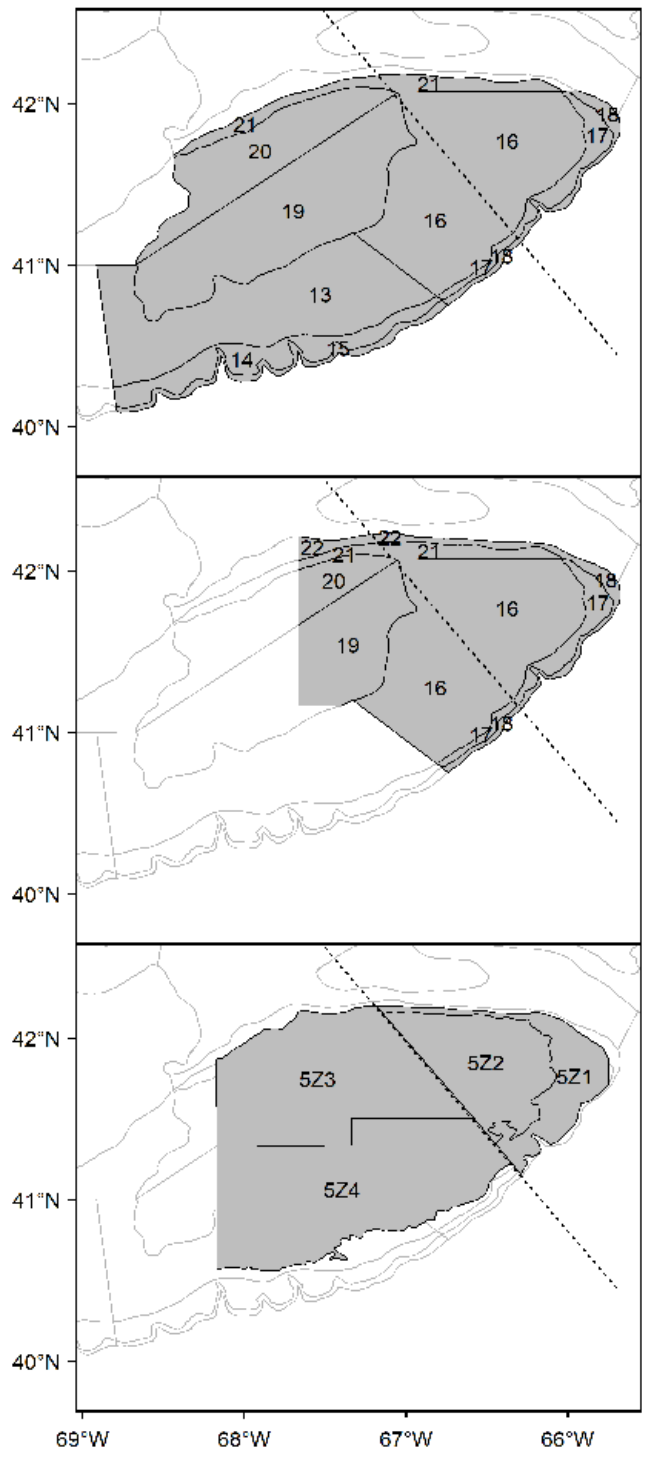
VAST spatial allocation

540 was used for both VAST runs. Proportion USA for VAST 1985–2017 is the same as that shown
541 in Figure 3.

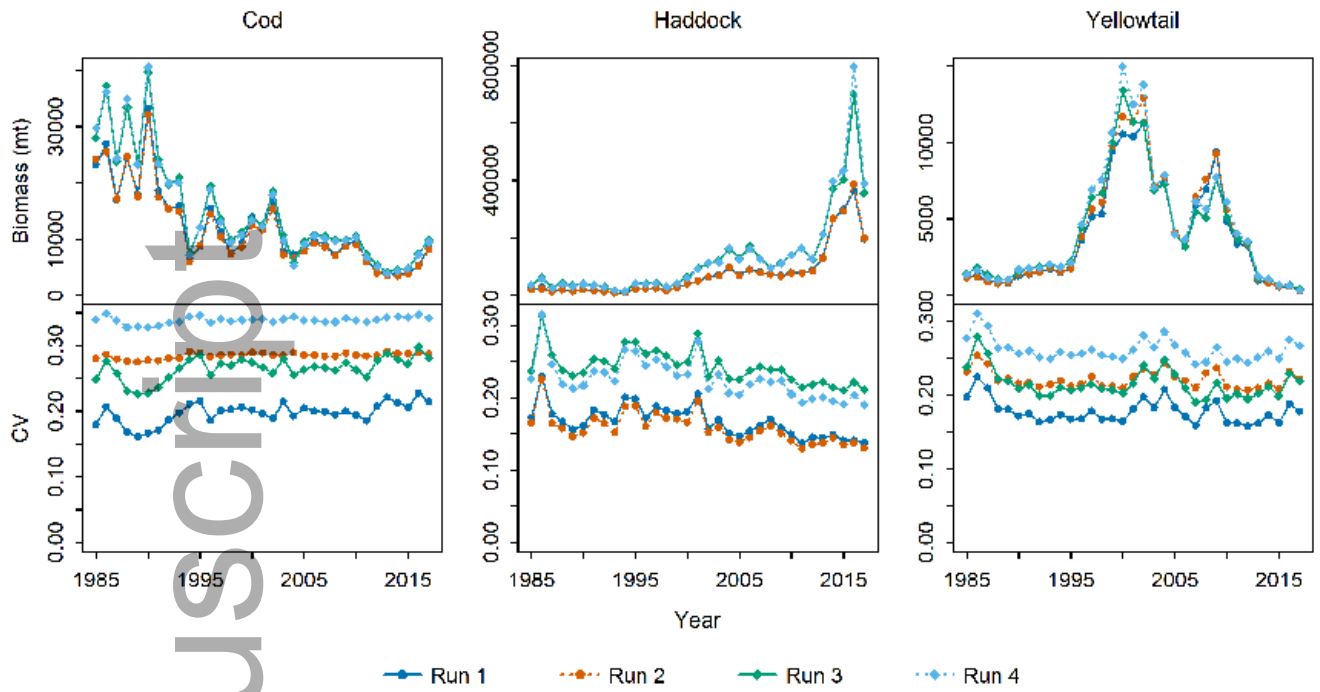
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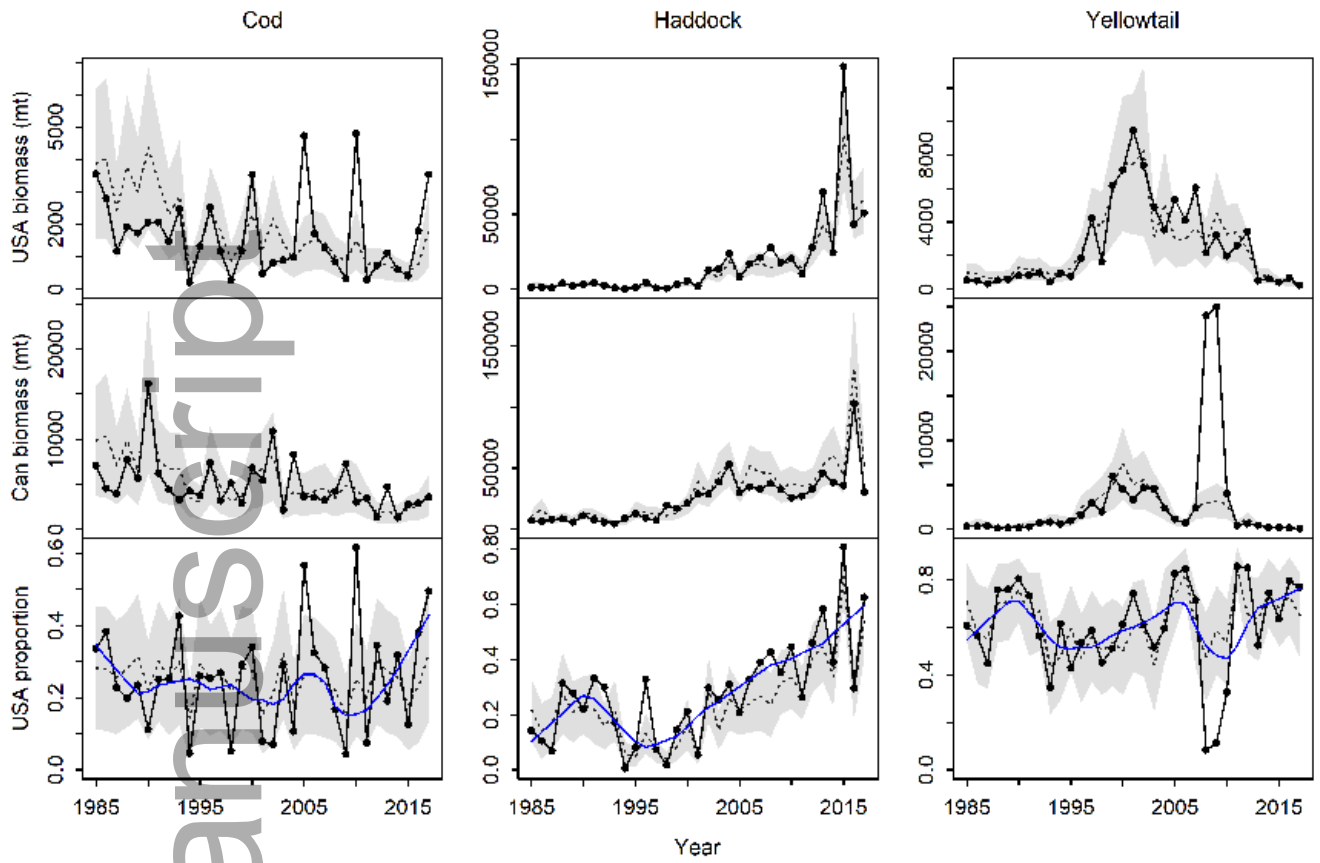
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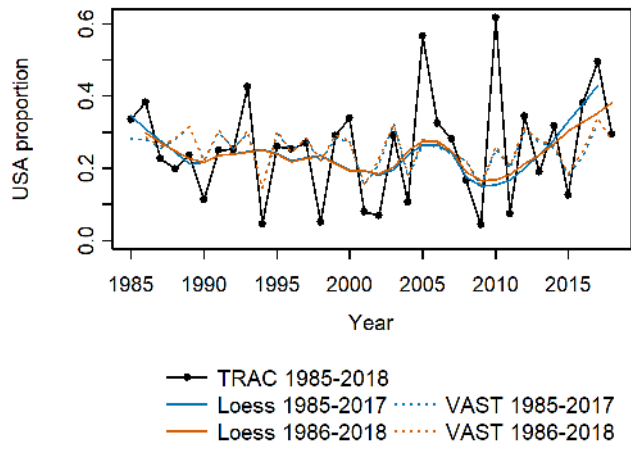
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