

# Appendices for “Multistage Hierarchical Capture-Recapture Models”

## Appendix A

To fit the homogeneous CR model using a single-stage MCMC algorithm, we consider the joint posterior distribution for  $p$  and  $\psi$ . Under the PX-DA framework, this posterior distribution can be written as

$$[p, \psi | \mathbf{y}_{1:n}, \mathbf{y}_{(n+1):M}, n] \propto [\mathbf{y}_{(n+1):M} | p, \psi, \mathbf{y}_{1:n}, n] [p, \psi | \mathbf{y}_{1:n}, n], \quad (35)$$

$$\propto [\mathbf{y}_{(n+1):M} | p, \psi, \mathbf{y}_{1:n}, n] [\mathbf{y}_{1:n} | p, n] [n | p, \psi] [p] [\psi], \quad (36)$$

$$\propto [\mathbf{y}_{1:n} | p, n] [n | p, \psi] [p] [\psi], \quad (37)$$

where the full-conditional distribution of  $\mathbf{y}_{(n+1):M}$  is proportional to one when conditioned on  $n$  (and hence drops out of the right hand side) and the conditional distribution of  $\mathbf{y}_{1:n}$  is proportional to the product of zero-truncated binomials

$$[\mathbf{y}_{1:n} | p, n] \propto \frac{\prod_{i=1}^n [y_i | p]}{(1 - (1 - p)^J)^n}. \quad (38)$$

For a given joint proposal distribution  $[p, \psi]^*$ , the associated Metropolis-Hastings ratio to update  $p$  and  $\psi$  jointly is

$$r = \frac{[\mathbf{y}_{1:n} | p^{(*)}, n] [n | p^{(*)}, \psi^{(*)}] [p^{(*)}] [\psi^{(*)}] [p^{(k-1)}, \psi^{(k-1)}]^*}{[\mathbf{y}_{1:n} | p^{(k-1)}, n] [n | p^{(k-1)}, \psi^{(k-1)}] [p^{(k-1)}] [\psi^{(k-1)}] [p^{(*)}, \psi^{(*)}]^*}. \quad (39)$$

To implement the model using PPRB following Hooten et al. (2021), we obtain an initial MCMC sample for  $p$  and  $\psi$  by fitting the CR model to the observed data while conditioning

on fixed and known  $n$ . The posterior distribution for the first stage is proportional to

$$[\mathbf{y}_{1:n}|p, n][p][\psi], \quad (40)$$

with respect to  $p$  and  $\psi$ , and the associated first-stage Metropolis-Hastings ratio is

$$r = \frac{[\mathbf{y}_{1:n}|p^{(*)}, n][p^{(*)}][\psi^{(*)}][p^{(k-1)}, \psi^{(k-1)}]^*}{[\mathbf{y}_{1:n}|p^{(k-1)}, n][p^{(k-1)}][\psi^{(k-1)}][p^{(*)}, \psi^{(*)}]^*}. \quad (41)$$

At this first stage, we use a temporary proposal distribution  $[p, \psi]^*$  that is convenient.

For the second stage of the PPRB implementation, we assume that the proposal distribution is

$$[p, \psi]^* \propto [\mathbf{y}_{1:n}|p, n][p][\psi], \quad (42)$$

which is equivalent to the first-stage posterior, and randomly sample (with replacement) joint first-stage MCMC realizations to use as proposals in the second stage Metropolis-Hastings updates. The resulting second-stage Metropolis-Hastings ratio becomes

$$r = \frac{[\mathbf{y}_{1:n}|p^{(*)}, n][n|p^{(*)}, \psi^{(*)}][p^{(*)}][\psi^{(*)}][p^{(k-1)}, \psi^{(k-1)}]^*}{[\mathbf{y}_{1:n}|p^{(k-1)}, n][n|p^{(k-1)}, \psi^{(k-1)}][p^{(k-1)}][\psi^{(k-1)}][p^{(*)}, \psi^{(*)}]^*}, \quad (43)$$

$$= \frac{[n|p^{(*)}, \psi^{(*)}]}{[n|p^{(k-1)}, \psi^{(k-1)}]}, \quad (44)$$

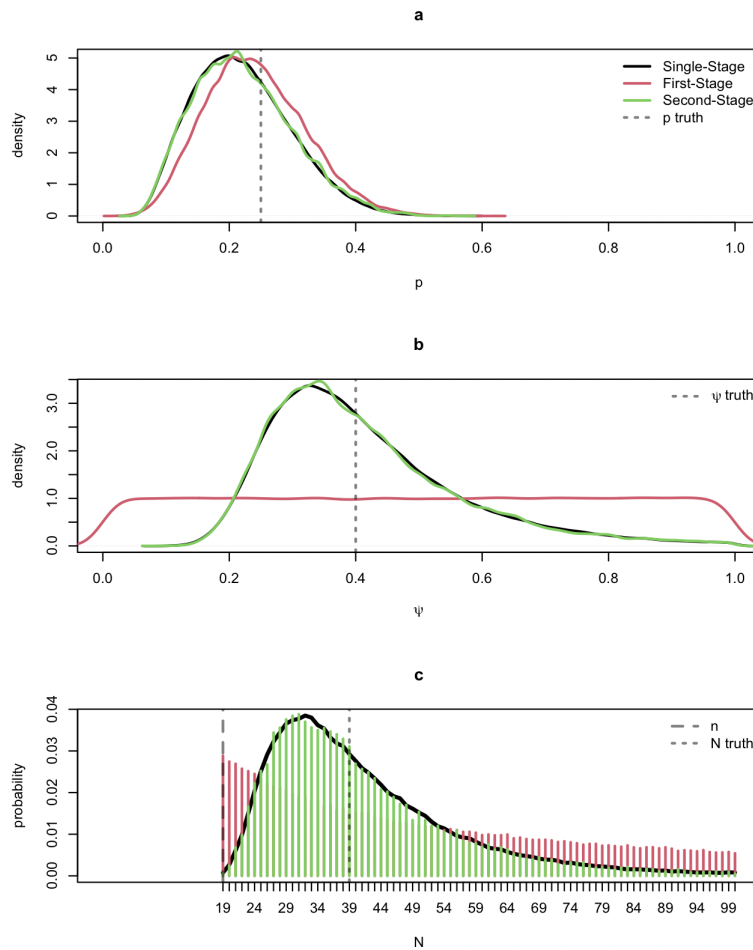
because the proposal cancels with the data model and priors. Thus, the second-stage Metropolis-Hastings ratio is merely a quotient involving the conditional model for  $n$  and can be evaluated easily using the first-stage MCMC sample.

## Appendix B

To demonstrate the recursive implementation of the hierarchical CR model, we fit the model to simulated data using the single-stage and two-stage approaches. To simulate CR data for this example, we set  $M = 100$  individuals in the superpopulation, membership probability  $\psi = 0.4$ , and detection probability  $p = 0.25$ . Then we used the hierarchical model in Section 2 as a generative process to simulate data based on  $J = 3$  occasions which resulted in  $N = 39$  individuals in our population, with  $n = 19$  individuals observed by our measurement process. The simulated observed data  $y_i$ , for  $i = 1, \dots, n$ , can be summarized by the values  $\mathbf{y} = (1, 1, 1, 1, 1, 1, 1, 2, 2, 1, 1, 1, 1, 2, 2, 1, 1, 2, 1)'$ .

In the PX-DA implementation, we assumed  $M - n = 81$  augmented individuals with all-zero capture histories. Thus, we assumed the same  $M$  as in our data simulation and this allows us to infer the true  $\psi$ . We note that the estimation of  $N$  is not constrained by  $M$  empirically in this example, therefore we could use larger values of  $M$  without influencing the inference for parameter  $p$ .

We fit the hierarchical CR from Section 2 to our simulated data using two approaches: 1) a standard single-stage MCMC algorithm for the hierarchical model and 2) a two-stage algorithm based on the recursive formulation of the same model. In both cases, we used  $K = 200000$  MCMC iterations. The results of our analyses are summarized in Figure 5. The posterior comparison shown in Figure 5 indicates that the two-stage PPRB approach yields the same inference as the conventional single-stage MCMC algorithm. Both approaches fit exactly the same model, but the recursive framework suggests that other specifications for the conditional model for  $n$  (e.g., Poisson) are straightforward to implement. Furthermore, in more complicated models, we can benefit from parallel evaluation of the PMF for  $n$  which can improve stability and facilitate computation for large data sets.



**Figure 5:** Marginal posterior distributions for a)  $p$ , b)  $\psi$ , c)  $N$ . Distributions shown are a result of the single-stage MCMC algorithm (black), first-stage of the two-stage MCMC algorithm (red), and second-stage of the two-stage MCMC algorithm (green). Subfigure c shows marginal posterior probability mass functions; black line shown for single-stage for comparison.



[46,] 450 -150 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[47,] 500 -150 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[48,] 550 -150 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[49,] 0 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[50,] 50 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[51,] 100 -200 0 0 0 0 0 0 0 1 0 0 0 0 0 0  
[52,] 150 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[53,] 200 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[54,] 250 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[55,] 300 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[56,] 350 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[57,] 400 -200 0 0 0 0 0 0 0 0 1 0 0 0 0 0  
[58,] 450 -200 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[59,] 500 -200 0 0 0 0 2 0 0 0 0 0 0 0 0 0  
[60,] 550 -200 0 0 0 0 0 0 0 0 0 2 0 0 0 0  
[61,] 0 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[62,] 50 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[63,] 100 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[64,] 150 -250 0 0 0 0 0 0 0 1 0 0 0 0 0 0  
[65,] 200 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[66,] 250 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[67,] 300 -250 1 0 0 0 0 0 0 0 0 0 0 0 0 0  
[68,] 350 -250 0 0 0 0 0 0 0 0 2 0 0 0 0 0  
[69,] 400 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[70,] 450 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[71,] 500 -250 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[72,] 550 -250 0 0 0 0 0 0 0 0 1 1 0 0 0 0  
[73,] 0 -300 0 0 0 0 0 0 0 0 0 0 0 0 1 0  
[74,] 50 -300 0 0 0 0 0 0 0 0 0 0 0 0 0 1  
[75,] 100 -300 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[76,] 150 -300 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[77,] 200 -300 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[78,] 250 -300 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[79,] 300 -300 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[80,] 350 -300 0 0 0 0 0 0 0 0 0 0 0 0 0 0  
[81,] 400 -300 0 0 0 0 0 0 0 1 0 0 0 0 0 0  
[82,] 450 -300 0 0 0 0 0 0 0 1 0 0 0 0 0 0  
[83,] 500 -300 0 0 0 0 0 0 0 0 0 1 0 0 0 0  
[84,] 550 -300 0 0 0 0 0 0 0 0 0 0 1 0 0 0