1	Highly accurate species identification of eastern Pacific rockfishes (Sebastes spp.) with
2	high-throughput DNA sequencing
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20 Abstract

21 Genetic species identification is often necessary for species flocks, such as rockfishes in the genus Sebastes (Teleostei, Scorpaenidae). Traditional visual identification 22 23 methods are challenged by the presence of many sympatric rockfish species with morphologically similar juveniles. Here we present a straightforward approach for 24 25 species identification in rockfishes using 96 nuclear microhaplotype loci that can be 26 efficiently genotyped using high-throughput DNA sequencing. Self-assignment of nearly 27 1 000 samples from 54 species resulted in >99% accurate species identification at a 95% confidence threshold. Phylogenetic relationships of Sebastes uncovered with these 28 29 same loci were highly concordant with relationships previously derived primarily with 30 mitochondrial DNA. We also assessed ascertainment bias and consequent reduced 31 nucleotide diversity and heterozygosity in non-ascertainment species to understand the 32 potential utility of these markers for those species. The data and protocol presented here will be useful for research and management of rockfishes in the northeastern 33 34 Pacific Ocean.

35 Introduction

36 Species identification is necessary when taxa that are the subject of study have closely 37 related and morphologically similar congeners. Generally, visual identification is the first priority, as it is typically low-cost and rapid. However, it can be inaccurate, particularly 38 for juvenile life stages, which often lack the morphological characteristics - especially 39 40 color and pattern - used to identify adults. Genetic identification has emerged as a 41 compelling alternative, by exploiting fixed nucleotide differences in specific gene regions between potential species of interest. Genetic identification approaches are commonly 42 used both for ecological studies and for management of fisheries and wildlife. Some 43 44 approaches further identify sampled individuals to population or family group. Genetic 45 analyses provide such identifications through methods such as categorical assignment 46 tests, mixed stock analysis, and parentage analysis (Pella and Milner 1987; Pella and 47 Masuda 2001; Jones et al. 2003). These methods utilize nucleotide variation within species, employing differences in allele frequencies among distinct populations or 48 49 stocks, and through segregating polymorphisms within populations (e.g., Abadía-Cardoso et al. 2013; Clemento et al. 2014). 50

In genetic studies of wild populations, it is possible to select a set of markers with the capacity to both distinguish the species identity of an individual, as well as its population or family of origin. In this scenario, the same genetic markers need to simultaneously contain fixed, or large frequency, differences between species and variation within a species, thereby providing inference at both levels of identification (e.g., Baetscher et al. 2019). However, ascertainment bias impacts the ability to use

genetic markers for this type of multi-level identification. Ascertainment bias occurs 57 because the initial assessment of genetic variation within a small number of samples 58 means that more common SNPs are more likely - and rare SNPs are less likely - to be 59 identified (Nielsen et al. 2004; Clark et al. 2005). When the ascertainment samples 60 61 consist of a small subset of the total number of species analyzed, ascertainment bias 62 suggests that only some of the variation from the initial SNP discovery samples will be shared across species due to different demographic histories and rates of mutation (Li 63 and Kimmel 2013). One manifestation of this bias can be reduced variation in the 64 65 genetic markers commensurate with the evolutionary genetic distance between the taxa used for marker discovery and other species of interest (Wakeley et al. 2001; Vowles 66 and Amos 2006). Another outcome is that markers may not amplify phylogenetically 67 distant species because of uncharacterized variation in the primer sites. 68

69 In marine fishes, few groups are as speciose as the rockfishes of the genus Sebastes, which includes over 100 species globally, almost all of which are found 70 71 exclusively in the North Pacific Ocean (Love et al. 2002). Nearshore species are 72 abundant in kelp forests and are prominent in studies of ecology and community 73 structure along the west coast of North America (Carr 1991). Rockfishes also support 74 important commercial and recreational fisheries throughout their Northeastern Pacific 75 range from Alaska to Mexico, where some regulations do not differentiate among 76 species and others apply to species complexes, both because many species co-occur 77 and also to alleviate the need to identify each fish (e.g., the "Other Rockfish" stock complex in the Gulf of Alaska; Tribuzio et al. 2017). Despite this inconsistent regulatory 78 79 framework, adult rockfishes can be accurately identified in most cases based on

morphometric characteristics; however, juveniles and cryptic species are frequently
misidentified (Butler et al. 2012).

Mitochondrial DNA (mtDNA) data suggest that Sebastes arose during the middle 82 Miocene in the Northwest Pacific and quickly diversified and dispersed into habitats 83 produced by high-latitude cooling and upwelling systems throughout the North Pacific 84 85 (Hyde and Vetter 2007). Originally, phylogenetic relationships among rockfishes were 86 defined by morphologic and meristic characters, with genetic data - specifically mitochondrial DNA - incorporated by the early 2000s (see Kendall 2000 for a 87 88 comprehensive review). Closely related species have been the subject of recent genetic 89 studies, which have identified cryptic species where adult specimens are morphologically similar and sometimes indistinguishable (Orr and Blackburn 2004: 90 91 Gharrett et al. 2005; Burford and Bernardi 2008, Orr and Hawkins 2008; Hyde and 92 Vetter 2009, Hess et al. 2013; Frable et al. 2015). These discoveries of cryptic species have coincided with increased genetic monitoring of rockfish populations for commercial 93 94 and recreational groundfish fisheries and population assessments (Orr and Blackburn 2004: Berntson and Moran 2009). Previous research used mitochondrial and nuclear 95 loci to genetically identify rockfish species (Hyde and Vetter 2007; Pearse et al. 2007). 96 97 In this study, we describe a new protocol for genetic species identification of rockfishes, including almost all lineages found in the California Current and Gulf of Alaska Large 98 Marine Ecosystems. We efficiently assay 96 nuclear microhaplotype loci using high-99 100 throughput DNA sequencing of amplicons and provide almost perfect species 101 identification in this group of fishes. The genetic markers (described in Baetscher et al. 102 2018; 2019) were discovered using double-digest restriction-site associated DNA

103 sequencing (ddRADseq; Peterson et al. 2012) of *S. atrovirens* (kelp rockfish) samples. 104 Initially, these markers were selected for the high heterozygosity necessary for identifying family relationships within S. atrovirens and its sympatric close relatives S. chrysomelas 105 106 (black-and-yellow rockfish) and S. carnatus (gopher rockfish). Additionally, the markers were designed for multiplexed analysis using next-generation DNA sequencing of 107 108 amplicons, which allows researchers to generate genotype data for hundreds-to-109 thousands of fish in a single sequencing reaction. Given that some collection techniques 110 employed to sample juvenile rockfishes can capture hundreds of fish in a single sampling event (Ammann 2004), a high-throughput method is particularly useful." In the 111 112 approach we describe here, we conducted species identification using genetic 113 assignment tests. Such assignment tests are employed to determine the likelihood that 114 a sample originates from one or more populations based on allele frequencies derived 115 from reference samples taken from those populations (Paetkau et al. 1995; Rannala 116 and Mountain 1997). Self-assignment provides a metric of how well a particular set of 117 genetic markers can differentiate among taxa when the identity of the true taxon is 118 known. Intuitively, the accuracy of assignment tests is limited by the biology and life 119 history of the organism – species with high gene flow have populations that are more 120 difficult to differentiate, and require high-resolution genetic data, whereas species with 121 almost no gene flow have populations that are typically easily discriminated using a 122 sufficient number of polymorphic loci. Given that our study involved classifying species 123 rather than populations, we anticipated identifying samples to true species with high 124 accuracy, assuming little-to-no ongoing gene flow among species.

125 Genotype data generated for testing species assignment allowed us to estimate 126 phylogenetic relationships of more than 50 rockfish taxa using nuclear DNA markers 127 and compare these results with a previously published phylogeny for Sebastes based 128 on seven mtDNA and two nuclear genes (Hyde and Vetter 2007). Depending on the 129 evolutionary history of the organism, nuclear and mtDNA genes can produce discrepant 130 signals of diversification (Shaw 2002; Chan and Levin 2005) and, thus, comparing the nuclear phylogeny against patterns derived in large part from mtDNA highlights areas 131 132 where the two marker types depict inconsistent relationships. A recent phylogenetic 133 study of six rockfish species using nuclear markers provides us with a comparison for the subgenus Sebastosomus (Wallace et al. 2022). Furthermore, we describe 134 135 phylogenetic relationships for a recently described cryptic species relevant to our 136 geographic study region.

137 Phylogenetic relationships help to contextualize the low levels of heterozygosity and nucleotide diversity for species not included in our marker ascertainment process 138 139 and allow us to assess this ascertainment bias based on evolutionary genetic distance. 140 Reduced heterozygosity diminishes the utility of these markers for intraspecific genetic analyses, including population structure and pedigree inference, even for species within 141 142 the same subgenus as the ascertainment species. This work describes a valuable 143 analysis tool for research of rockfishes when confident species identity is required, an examination of phylogenetic relationships across the genus, and insight into how 144 145 nucleotide diversity rapidly declines in species not included in the marker discovery 146 process.

147 Methods

148 Samples

149 Samples from adults of 54 species of rockfishes (Sebastes) and cabezon

150 (Scorpaenichthys marmoratus), the sister species of the genus Sebastes, were

151 obtained by trawl and hook-and-line fishing. Rockfishes were identified morphologically

152 by experts from the NOAA Southwest Fisheries Science Center or researchers at the

153 University of California Santa Cruz. For the majority of samples, DNA was extracted

154 from fin tissue using DNeasy 96 Blood & Tissue kits on a BioRobot 3000 (Qiagen, Inc.),

eluted into 200 μL, with extracts stored at 4° C. For species with few adult samples

available, DNA was extracted from juvenile samples as described. A small number of

157 samples were received as previously extracted DNA and stored at 4° C prior to

158 sequencing library preparation.

159 Genotyping and analysis

160 Samples were genotyped with a set of 96 microhaplotype markers ascertained in S. 161 atrovirens, S. carnatus and S. chrysomelas using the Genotyping-in-Thousands by 162 sequencing (GT-seq; Campbell et al. 2015) protocol, as modified by Baetscher et al. 163 (2018). The amplicon-sequencing library preparation includes an initial multiplex PCR 164 step to amplify target loci and a second PCR to add sequencing adapters and barcodes 165 for identifying samples. Normalized libraries were sequenced using 2 x 75 bp paired-166 end approach on a MiSeq instrument (Illumina, Inc.). Raw sequence reads were sorted 167 by individual barcode using the MiSeq Analysis Software (Illumina), and then paired

168 reads were combined and mapped to a reference using the bioinformatic workflow in 169 Baetscher et al. (2018). Variants were called across samples using FREEBAYES 170 (Garrison and Marth 2012) and the output variant call format (VCF) files were filtered for 171 guality (minQ = 30; minDP = 10) and merged using VCFTOOLS (Danecek et al. 2011). 172 In microhaplotypes, multiple single nucleotide polymorphisms (SNPs) segregate 173 together within a single sequencing read and do not require statistical phasing (Stephens and Donnelly 2003), which makes it relatively straightforward to call 174 individual haplotypes from mapped data files and the combined VCF file using the 175 176 software program MICROHAPLOT (Ng and Anderson 2016). Resulting genotypes were 177 filtered in R (R Core Team 2016) using a minimum threshold of 20 reads per 178 individual/locus and a minimum read depth ratio of 0.4, which applies to heterozygotes 179 and is a measure of the number of reads of the second most common allele divided by the read depth of the most common allele. Loci with high rates of missing data or 180 deviations from Hardy-Weinberg equilibrium (HWE) were removed and then samples 181 182 with missing data at more than 25 of the remaining loci were dropped from further 183 analysis. This missing data threshold was intentionally liberal to avoid removing 184 samples of species in which a larger proportion of loci failed to amplify due to uncharacterized variation in the primer sites (Fig. S1). Such variation is more common 185 186 in genetic markers applied to species that are phylogenetically distant from the 187 ascertainment species due to different demographic histories and rates or directionalities of mutation. 188

Since juvenile rockfishes are commonly misidentified, only genotypes for adults
were included, except for species in which we had fewer than five adult samples and

samples from juveniles were available. The veracity of the species identity for these
juvenile samples was evaluated by the self-assignment analysis (see below). A
maximum of 32 individuals per species was included, when available, to generate a
dataset with a representative estimate of assignment accuracy across the genus. The
data set was tested for deviations from HWE using the R package PEGAS (Paradis
2010) and pairwise F_{ST} was calculated with heterozygosity weighted by group size, also
in R using HIERFSTAT (Goudet 2005).

198 Genetic assignment

199 Genetic self-assignment was conducted in the R package RUBIAS (Moran and 200 Anderson 2018) using the leave-one-out self-assignment function with default allele 201 frequency prior. Leave-one-out procedures remove the gene copies for each sample 202 from the allele counts of its known population/taxon of origin before calculating the 203 likelihood that the sample came from that population, in order to avoid overestimating 204 assignment accuracy. RUBIAS provides a likelihood for each sample assigning to every 205 reference population and a z-statistic for each sample assignment. The z-statistic is the 206 difference (in number of standard deviations) between the observed log-probability of an 207 individual's genotype given it came from a specific population, and the log-probability 208 expected for an individual from that population. The mean and standard deviation of the 209 expected log-probability values are computed by RUBIAS using the locus-specific allele 210 frequencies and the assumption of HWE. When the probability of assignment is high for 211 a given reference population but the z-statistic is outside the expected range (<-3 or 212 >3), this can be an indication that the sample belongs to a population that is not

included in the reference dataset. In an effort to ensure that only samples that were
confidently identified to true species were included, any samples that were assigned to
a reference taxon with a z-statistic <-3 or >3 were excluded from the final dataset.

216 Phylogenetic analyses

217 Samples verified by self-assignment were used to construct phylogenetic trees. To 218 generate consensus sequence data for building trees, species-specific VCF files were 219 produced by FREEBAYES and then a consensus FASTA file for each species was 220 created using VCFTOOLS (Danecek et al. 2011). A member of the sister genus to Sebastes, Scorpaenichthys marmoratus (cabezon) was used to root the phylogenetic 221 222 trees. Loci in each species-consensus FASTA file were concatenated with the GENEIOUS software program (v 7.1.7; Kearse et al. 2012) before export to MUSCLE 223 224 (Edgar 2004) with alignments output in ClustalW format. These were then used as input 225 for MEGA (v. 7.0.26; Kumar et al. 2016) to generate maximum-likelihood trees using the General Time Reversible (GTR) model (Nei and Kumar 2000) with 1 000 bootstrap 226 replicates, which was consistent with the model used by Hyde and Vetter (2007) for 227 228 their Sebastes phylogeny. A similar analysis was performed to generate an unrooted 229 maximum-likelihood tree, without cabezon, also using the GTR model and 1 0 0 0 230 bootstrap replicates.

For the Bayesian analysis, FASTA alignments were converted to Nexus format using PGDSpider (v. 2.1.1.5; Lischer and Excoffier 2012), and then used as input for MRBAYES (v. 3.2; Huelsenbeck and Ronquist 2001). Parameters included a GTR substitution model and one million generations, where generation time was increased

experimentally until the standard deviation of split frequencies dipped below 0.01 and
the Potential Scale Reduction Factor (PSRF) converged to 1. This included a uniform
Dirichlet prior (1,1,1,1) and 25% burn-in with sampling from the posterior every 5000
generations. Phylogenetic trees generated by this analysis were visualized using
FigTree (v 1.4.3; Rambaut 2016).

240 Because the marker set was designed using data from S. atrovirens, S. 241 chrysomelas, and S. carnatus based on the variability in those species, the amount of variation in other species was expected to be affected due to ascertainment bias. This 242 243 bias was quantified as the decrease in mean internal heterozygosity and nucleotide 244 diversity for each species with increasing genetic distance from *S. atrovirens*. Genetic 245 distance was calculated in MEGA using a variety of model settings to determine the 246 extent to which estimates of genetic distance in these data are sensitive to model 247 choice (Fig. S2). Nucleotide diversity was calculated per variant site for each species in VCFTOOLS and then the sum of all sites within a species was divided by the total 248 249 number of bases in the 96 loci to account for invariant sites.

250

251 **Results**

252 <u>Genotyping and data analysis</u>

A total of 997 rockfish samples from 54 species were genotyped and analyzed with a
VCF file that had previously been generated from 1 690 rockfish samples and contained
4,322 variant sites from all species (Baetscher 2019; Table S1). Five loci (Sat_914,

Sat_934, Sat_1399, Sat_1871, Sat_2513) with large amounts of missing data across > 35% of species and one locus (Sat_1166) with three or more haplotypes per individual in some species, suggestive of a paralogous locus, were removed. Only genotypes that passed filtering thresholds for read depth, allelic ratio, and missing data were retained for analyses. In three species (*S. reedi, S. wilsoni*, and *S. crameri*) with fewer than two adult samples available, genotypes from juveniles were included. The number of samples per species ranged from two (*S. rufinanus*) to 32 (*S. atrovirens*; Table 1).

The majority of species-by-locus combinations conformed to HWE; however, the six species with the greatest number of deviations (> 10 loci out of HWE), were *S. rosaceus* (18 loci), *S. carnatus* (15 loci), *S. chrysomelas* (13 loci), *S. ensifer* (13 loci), *S. diaconus* (11 loci), and *S. mystinus* (10 loci). Thirty percent of the loci were out of HWE in four of the 54 species, and three loci, Sat_770, Sat_875, and Sat_2178, were out of HWE in 8-13 species. Pairwise F_{ST} ranged from 0.015 between *S. carnatus* and *S. chrysomelas* to 0.746 between *S. levis* and *S. entomelas* (mean $F_{ST} = 0.45$, s.d. = 0.13).

270 <u>Self-assignment</u>

Self-assignment resulted in 98.3% accuracy at a scaled-likelihood value of 0.95, and all
mis-assigned individuals at > 50% likelihood were either *S. carnatus* assigning to *S. chrysomelas*, or vice versa. These assignment results indicated that this set of genetic
markers cannot consistently distinguish between *S. carnatus* and *S. chrysomelas* and
that a single genetic reporting group would be appropriate for assignment.

The self-assignment analysis was performed again after creating a single *S. carnatus/chrysomelas* reporting group and 100% of samples were correctly assigned at a 50% scaled-likelihood threshold. At the 95% confidence level, assignment accuracy was 99.2% and all lower confidence assignments were *S. carnatus* or *S. chrysomelas* samples that assigned to the joint reporting group, but at a scaled-likelihood below 95%.

281 Phylogenetic trees

282 Species relationships were elucidated with maximum-likelihood and Bayesian trees. 283 Both rooted trees (Fig. 1, Fig. S3) and an unrooted tree (Fig. S4) recovered very similar 284 phylogenetic relationships. Branch support on the Bayesian tree was generally higher 285 than for the maximum-likelihood trees, which had consistent bootstrap values, but with 286 slight differences at some of the deeper nodes. Some of the most confident 287 relationships in the Bayesian tree included the position of S. atrovirens clustered with 288 members of the *Pteropodus* subgenus, as well as that *S. saxicola* and *S. semicinctus* appeared proximate to *Pteropodus* and distant from other members of the subgenus 289 290 Allosebastes (Fig. 1). Monophyletic relationships among taxa within the subgenus 291 Sebastomus garnered strong support with the exception of S. rufus, which groups with 292 the subgenus *Acutomentum* (Fig. 1). While the branch support for these phylogenetic 293 positions varied between the maximum-likelihood and Bayesian analyses, the overall 294 pattern among these subgenera appeared consistent.

295 Ascertainment bias

296 Observed heterozygosity in most species declined substantially when compared to S. 297 atrovirens, with a smaller decrease in S. chrysomelas and S. carnatus (mean for S. 298 atrovirens, chrysomelas, carnatus = 0.423, overall mean = 0.130; range = 0.012-0.458; 299 Fig. 2). Nucleotide diversity sharply declined with genetic distance from S. atrovirens 300 (Fig. 3), with low levels of diversity even in species in the same subgenus as S. 301 *atrovirens*. Genetic distance was calculated as pairwise differences since a comparison 302 indicated that nucleotide substitution model does not substantially alter distance 303 estimates for this dataset (Fig. S2). Based on these results, over 80% of species 304 analyzed in this study contained less than half of the nucleotide diversity of S. atrovirens 305 over a genetic distance of fewer than 0.04 base differences per site for 10 695 total 306 sites, excluding gaps and missing data.

307

308 Discussion

309 Here we demonstrate the high accuracy (>99% correct assignment) of a set of short 310 haplotypic markers for identifying 54 species of the genus Sebastes, including all of the species commonly found in the California Current Large Marine Ecosystem along the 311 312 Pacific coast of North America. Using these loci, we distinguish between closely related 313 and recently described cryptic species, describe phylogenetic relationships, and 314 quantify a decrease in the heterozygosity and nucleotide diversity of these genetic 315 markers in species with increasing evolutionary genetic distance from the ascertainment 316 species.

317 Ecological studies and management of fisheries require efficient methods to 318 conclusively identify sympatric marine species, particularly at the larval and juvenile 319 stages. In rockfishes, planktonic larvae from many species coexist during their pelagic 320 phase and remain challenging to identify morphologically as they recruit to settlement 321 habitats (Butler et al. 2012). Even as adults, the number of species present in 322 overlapping habitats, the presence of cryptic species (e.g., Frable et al. 2015), and 323 subtle differences in coloration or morphology (Ingram and Kai 2014) underscore the 324 need for genetic species identification. Previous marker types have been used for this 325 task; one such study included 33 species with 97.4% assignment accuracy (Pearse et 326 al. 2007), and the other, a much more complete survey of the genus, genotyped 103 327 individuals from 101 species at seven mitochondrial and two nuclear genes, but did not 328 test these data for genetic assignment accuracy (Hyde and Vetter 2007). Our method of 329 genotyping fewer than 100 multiplexed microhaplotype loci with high-throughput DNA sequencing is highly accurate, efficient for large sample sizes and can be coupled with 330 331 a reproducible analysis workflow based on the reference database for species 332 assignment generated by this study.

Self-assignment using genotype data from 90 retained microhaplotype markers accurately identified the true species identity of every sample for all 54 species, with the exception of two extremely closely related species. At a stringent likelihood threshold (> 95%), eight samples of *S. carnatus* and *S. chrysomelas* assigned to the combined *S. carnatus/chrysomelas* group at a lower level of confidence, but still above a 50% scaled-likelihood. Notably, these sister species have been the subject of ongoing research (Narum et al. 2004; Buonaccorsi et al. 2011) and our results from the self-

340 assignment demonstrate the challenge of separating the two groups with existing341 genetic markers and call into question their taxonomic status as two distinct species.

Coincidentally, S. carnatus/chrysomelas are also the most phylogenetically 342 proximate to the primary ascertainment species (S. atrovirens; Fig. 1; Fig. S3, Fig. S4), 343 and with nearly as much variation in these loci (Fig. 2, Fig. 3). And while these genetic 344 345 markers easily differentiate juvenile-stage cryptic species (e.g., S. mystinus/diaconus, 346 S. aleutianus/melanostictus) and those commonly misidentified even as adults (e.g., S. *flavidus/serranoides*), they underperform for *S. carnatus/chrysomelas*. This indicates 347 348 that these taxa are more genetically similar than every other pair of sister species 349 included in our dataset, at least in the portion of the genome surveyed with these loci, 350 consistent with the lowest pairwise F_{ST} value (0.015) in the study. Previous work on S. 351 carnatus and S. chrysomelas identified a single, highly diverged locus and concluded 352 that the pair is likely in the final stages of speciation, but with ongoing gene flow (Narum et al. 2004, Buonaccorsi et al. 2011). A more recent investigation using reduced-353 354 representation and whole genome resequencing found three distinct genomic regions 355 with elevated divergence and variation in genes pointing to the importance of coloration 356 and vision (Behrens et al. 2021). Results from these studies are consistent with the 357 general idea that speciation mechanisms in rockfishes likely involve both allopatric and sympatric processes, including habitat differentiation associated with depth gradients 358 359 (Ingram 2011) and mate choice reinforced by internal fertilization (Buonaccorsi et al. 360 2011).

361 Previously described rockfish species relationships relied heavily on 362 mitochondrial DNA data (Kai et al. 2003; Li et al. 2006; Hyde and Vetter 2007; Li et al. 363 2007), providing an opportunity to apply the nuclear markers from this study to estimate 364 phylogenetic relationships for comparison (Fig. 1, Fig. S3, Fig. S4). Rooted and 365 unrooted maximum-likelihood trees produced consistent topologies with very similar 366 branch support, although some deeper nodes in the unrooted tree garnered higher support, while other nodes were better supported in the rooted tree (Fig. 1, Fig. S4). 367 368 High confidence nodes in the Bayesian tree were generally well supported in the 369 maximum-likelihood tree, with most differences occurring at nodes with lower support, 370 such as the position of either S. alutus or S. borealis in a clade with S. melanostictus 371 and S. aleutianus (Fig. 1, Fig. S3). Few instances of well-supported Bayesian 372 relationships deviate from the maximum-likelihood tree, although S. polyspinis presents 373 one such case. The Bayesian tree topology from our data is the most appropriate for 374 comparison with the phylogeny in Hyde and Vetter (2007) since the analyses are 375 equivalent and, although Bayesian methods can overestimate node support, 376 bootstrapped maximum-likelihood values may be overly conservative (Douady et al. 2003). 377

Most relationships remain consistent between the microhaplotype tree topologies and the more complete *Sebastes* tree from Hyde and Vetter (2007). Although Hyde and Vetteranalyze species that are absent from our dataset, primarily from the northwest Pacific and North Atlantic, we analyze representatives from each major clade with the exception of the subgenera *Sebastocles* and *Mebarus*, whose constituents are exclusively in the northwest Pacific, with the exception of *S. atrovirens* which should

384 clearly be included in the *Pteropodus* subgenus. Generally, we find very high 385 concordance with Hyde and Vetter (2007) at the subgeneric level. Areas in which the 386 microhaplotype tree (Fig. 1) deviates from their tree include clade "D" nesting within 387 Pteropodus, and members of Eosebastes, S. aurora and S. diploproa, nesting within 388 Sebastichthys. At the species level, more variation exists. For example, both trees depict close phylogenetic relationships among S. atrovirens, S. carnatus, and S. 389 390 chrysomelas, with the microhaplotype tree placing S. maliger as a closer relative of the three species than S. caurinus, as in the mitochondrial tree. Other small differences in 391 392 the topologies include strong support that S. melanops is more closely related to S. 393 flavidus than S. serranoides (a relationship also identified by Wallace et al. 2022); and 394 that S. goodei is more closely related to S. paucispinis than to S. jordani. We also show 395 that S. diaconus and S. mystinus are easily distinguished and nearest neighbors in the 396 phylogeny, which is unsurprising since these species were only recently described as 397 separate taxa (Frable et al. 2015; Wallace et al. 2022).

398 Taxonomy of rockfishes, particularly of subgenera, has been and continues to be 399 dynamic, as highlighted by multiple revisions of subgeneric classifications (Love et al. 400 2002). For example, S. diploproa is part of the subgenus Sebastichthys in Kendall 401 (2000), who cites Eigenmann and Beeson (1894), but Li et al. (2006) designate S. 402 diploproa as a member of Allosebastes, attributed to Gilbert (1890). Phylogenetic 403 relationships described by the microhaplotype data are generally consistent with 404 mitochondrial data and support polyphyly of generally accepted subgenera, including 405 Acutomentum, Allosebastes, and Sebastosomus (Hyde and Vetter 2007; Li et al. 2007). 406 A formal re-description of these subgenera would alleviate some of the taxonomic

407 confusion but comprehensive taxonomic revision would require data from more species408 in the genus than are included in this study.

409 The set of nearly 100 microhaplotype loci target substantial variation in the 410 ascertainment species, S. atrovirens, S. carnatus, and S. chrysomelas (Baetscher et al. 411 2018; 2019) and contain a similar amount of variation in a closely related taxon (S. 412 *maliger*). However, variation declines rapidly with increasing genetic distance (Fig. 3), 413 even for members of the *Pteropodus* subgenus. Such reduced variation has been documented in studies of ascertainment bias in microsatellite loci across multiple 414 genera (Vowles and Amos 2006). Even so, the ascertainment bias we observe here is 415 416 even more significant than previously observed, with dramatically decreased nucleotide 417 diversity over relatively small evolutionary genetic distances, with only the most closely 418 related species to those included in the marker discovery process found to have 419 substantial variation (Fig. 3). The surprising amount of variation in S. rosaceus and S. 420 *ensifer*, despite their evolutionary distance from *Pteropodus*, might be explained by 421 cryptic structure in those species, as indicated by the relatively high number of loci that 422 deviated from HWE. However, selectively removing loci for individual species would be challenging with the > 50 species included in this analysis. 423

Although the relatively low observed heterozygosity found in this set of markers for the majority of species analyzed here suggest limited utility for purposes other than species identification (e.g., pedigree reconstruction), the amplicon library preparation protocol is highly flexible and enables researchers to add additional loci or swap out markers that would increase power for species of particular interest. Such an effort

could bolster this set of markers for population genetic structure or pedigree analyses in
additional species, and previous research has shown that genotyping samples with a
single set of genetic markers to both identify species and analyze pedigree relationships
is an economical approach (Baetscher et al. 2019).

Here, we describe an efficient method for genotyping and analyzing genetic data 433 434 to identify species of rockfishes, particularly for taxa commonly captured together as 435 juveniles. The genetic markers we employ, and our subsequent analytical workflow, provide highly accurate species identification and estimates of phylogenetic 436 relationships largely consistent with previous genetic data. In addition, we describe a 437 438 flexible protocol for modifying the set of target loci and accounting for ascertainment 439 bias to suit the specific needs of a variety of ecological studies and fisheries 440 management objectives.

441

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Fig. 1: Consensus tree estimated using the General Time Reversible (GTR) model and
 Bayesian posterior analysis for 54 *Sebastes* species and one member of the sister
 genus, *Scorpaenichthys*. Genetic data includes 96 nuclear markers. Bayesian posterior
 probabilities above 50 are indicated at nodes.



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Fig. 2: Heterozygosity for 54 *Sebastes* species genotyped with 96 nuclear genetic markers. Mean internal heterozygosity per species is indicated as the dark bar inside the box. Boxes represent the 25th and 75th percentiles (first and third quartiles) and whiskers extend to the smallest and largest values 1.5 times the distance between the first and third quartiles. Points beyond that distance are plotted individually. Fill colors indicate subgenera classification consistent with designations in Table 1.



Fig. 3: Genetic distance from *Sebastes atrovirens* and nucleotide diversity for 54 *Sebastes* species classified to subgenus. Genetic data includes 96 nuclear markers.
Genetic distance is measured as pairwise differences (base differences per site,

excluding gaps and missing data). Names for species with nucleotide diversity >0.003

- and all members of the *Pteropodus* subgenus are shown.
- 720

721 Tables

Table 1. Number of samples per Sebastes species included in the self-assignment and

723 phylogenetic analyses. Mean nucleotide diversity, mean internal heterozygosity and

nominal subgenera classification is included.

		nucleotide		
Species	samples	diversity	heterozygosity	subgenus
Sebastes atrovirens	32	0.00678	0.4578	Pteropodus
S. chrysomelas	32	0.00592	0.4086	Pteropodus
S. carnatus	32	0.00561	0.4032	Pteropodus
S. rosaceus	19	0.00467	0.2275	Sebastomus
S. maliger	16	0.00433	0.3285	Pteropodus

S. ensifer	19	0.00393	0.2362	Sebastomus
S. chlorostictus	15	0.00368	0.1829	Sebastomus
S. constellatus	16	0.00354	0.1473	Sebastomus
S. melanostictus	7	0.00353	0.1814	Zalopyr
S. caurinus	9	0.00343	0.2336	Pteropodus
S. umbrosus	18	0.00329	0.1788	Sebastomus
S. jordani	15	0.00308	0.0956	Sebastodes
S. dallii	6	0.00300	0.1667	Auctospina
S. zacentrus	6	0.00285	0.1350	Allosebastes
S. oculatus	18	0.00267	0.1163	Sebastomus
S. aleutianus	9	0.00262	0.1250	Zalopyr
S. diploproa	16	0.00259	0.1388	Sebastichthys
S. alutus	24	0.00256	0.1038	Acutomentum
S. wilsoni	16	0.00255	0.1345	Allosebastes
S. semicinctus	17	0.00236	0.1136	Allosebastes
S. nebulosus	21	0.00235	0.1548	Pteropodus
S. mystinus	25	0.00234	0.1068	Sebastosomus
S. diaconus	26	0.00234	0.0932	Sebastosomus
S. melanops	31	0.00223	0.1021	Sebastosomus
S. ruberrimus	13	0.00221	0.0958	Sebastopyr
S. miniatus	20	0.00218	0.0976	Rosicola
S. elongatus	16	0.00207	0.1149	Hispaniscus
S. hopkinsi	24	0.00195	0.0790	Acutomentum
S. aurora	14	0.00192	0.0770	Eosebastes
S. emphaeus	20	0.00192	0.1341	Allosebastes
S. flavidus	31	0.00189	0.1189	Sebastosomus
S. babcocki	11	0.00187	0.1146	Rosicola
S. rufus	11	0.00181	0.0970	Sebastomus
S. auriculatus	16	0.00173	0.0734	Auctospina
S. nigrocinctus	29	0.00173	0.1145	Sebastichthys
S. serriceps	5	0.00169	0.0915	Sebastocarus
S. proriger	31	0.00164	0.0790	Allosebastes
S. borealis	8	0.00163	0.0637	Zalopyr
S. melanostomus	20	0.00159	0.0726	Eosebastes
S. reedi	14	0.00158	0.0313	unclassified
S. rastrelliger	12	0.00156	0.0945	Pteropodus
S. ovalis	31	0.00153	0.0521	Acutomentum
S. entomelas	15	0.00153	0.0926	Acutomentum
S. pinniger	16	0.00147	0.0944	Rosicola
S. crameri	19	0.00144	0.0461	Eosebastes

Total samples	997			
S. rubrivinctus	19	0.00063	0.0120	Hispaniscus
S. serranoides	32	0.00100	0.0643	Sebastosomu
S. goodei	32	0.00118	0.0803	Sebastodes
S. rufinanus	2	0.00119	0.0899	Acutomentum
S. polyspinis	8	0.00122	0.0731	unclassified
S. moseri	3	0.00128	0.1031	Acutomentum
S. saxicola	16	0.00134	0.0904	Allosebastes
S. levis	32	0.00137	0.0325	Hispaniscus
S. paucispinis	32	0.00143	0.0410	Sebastodes

726 Supplemental Material

Table S1. Number of samples per *Sebastes* species included in the variant call format
(VCF) file used to call genotypes.

Species	samples
S. atrovirens	242
S. chrysomelas	113
S. goodei	65
S. levis	48
S. ovalis	36
S. serranoides	33
S. flavidus	32
S. hopkinsi	32
S. melanops	32
S. paucispinis	32
S. proriger	31
S. caurinus	30
S. carnatus	29
S. diaconus	26
S. emphaeus	26
S. ensifer	26
S. mystinus	25
S. alutus	24
S. miniatus	23
S. crameri	22
S. nebulosus	22
S. melanostomus	20
S. rosaceus	20
S. umbrosus	20
S. rubrivinctus	19
S. semicinctus	19
S. auriculatus	16
S. aurora	16
S. babcocki	16
S. constellatus	16
S. diploproa	16
S. elongatus	16
S. entomelas	16
S. jordani	16

Total	1440
S. rosenblatti	1
S. brevispinis	1
S. nigrocinctus	6
S. simulator	7
S. melanostictus	7
S. serriceps	8
S. polyspinis	8
S. borealis	8
S. rufinanus	9
S. helvomaculatus	9
S. aleutianus	9
S. moseri	10
S. rufus	12
S. ruberrimus	13
S. zacentrus	15
S. dalli	15
S. chlorostictus	15
S. wilsoni	16
S. saxicola	16
S. reedi	16
S. rastrelliger	16
S. pinniger	16
S. oculatus	16



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Fig. S1: Number of genetic markers (loci) that do not amplify in all samples for a given

731 Sebastes species out of 96 total markers. Species are ranked from fewest-missing to

732 most-missing loci. Color corresponds to subgenus. No loci are missing entirely from S.

733 atrovirens, S. chrysomelas, S. carnatus, S. caurinus, or S. maliger.



Fig. S2: Genetic distances are base substitutions per site between each Sebastes 735 736 species and S. atrovirens. Nucleotide substitution models are indicated by symbols. The 737 Jukes-Cantor (Jukes and Cantor 1969), Tajima-Nei (Tajima and Nei 1984), and Kimura 738 2-parameter (Kimura 1980) models included uniform rates among sites. The maximum composite like- lihood model (Tamura, Nei, and Kumar 2004) and Tajima-Nei model 739 740 with 1000 bootstraps included rate variation among sites modeled with a gamma distribution and differences in sequence composition (Tamura and Kumar 2002). The 741 742 dataset included 10,695 sites, excluding gaps and missing data and analyses were 743 conducted in MEGA v. 7.0.26.



Fig. S3: Rooted maximum-likelihood tree estimated with the General Time Reversible model and 1000 bootstrap replicates for 54 *Sebastes* species and one member of the sister genus, *Scorpaenichthys*. Genetic data includes 96 nuclear markers. Bootstrap values above 50 are indicated above nodes on relevant branches. Analyses included 11,622 sites and retained gaps and missing data.



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Fig. S4: Unrooted maximum-likelihood tree estimated using the General Time
 Reversible model with 1000 bootstrap replicates for 54 *Sebastes* species. Bootstrap

support above 0.50 is shown on branches. The analysis included 11,622 sites and

retained gaps and missing data.