

**Title:** Lower possession limits and shorter seasons directly reduce for-hire fishing effort in a multispecies marine recreational fishery

**Authors:** Ashley Trudeau<sup>1\*†, 2, 3</sup>, Eleanor A. Bochenek<sup>4</sup>, Abigail S. Golden<sup>2,3,5†</sup>, Michael C. Melnychuk<sup>5</sup>, Douglas R. Zemeckis<sup>6</sup>, Olaf P. Jensen<sup>1†, 2</sup>

**Affiliations:**

<sup>1</sup>Center for Limnology, University of Wisconsin, Madison, WI

<sup>2</sup>Department of Marine and Coastal Sciences, Rutgers University, New Brunswick, NJ

<sup>3</sup>Graduate Program in Ecology and Evolution, Rutgers University, New Brunswick, NJ

<sup>4</sup>Fisheries Cooperative Center, Haskin Shellfish Research Laboratory, Rutgers University, Cape May, NJ

<sup>5</sup>School of Aquatic and Fishery Sciences, University of Washington, Seattle, WA

<sup>6</sup>Department of Agriculture and Natural Resources, Rutgers University, New Brunswick, NJ

\*Corresponding author

†Current affiliation

## Abstract

Managers of recreational fisheries often rely on implicit and rarely-tested assumptions regarding how fishing effort will change in response to regulations. For instance, they assume that reduced seasons will directly reduce fishing effort without producing angler behavioral adaptations to maintain fishing opportunities and harvest. Vessel trip reports from a multispecies for-hire fishery in New Jersey, USA allowed us to empirically evaluate changes in fishing effort as overlapping seasons for four species became shorter and as possession limits decreased. We conducted focus groups with fishery stakeholders and then developed statistical models to evaluate hypotheses describing how anglers aboard for-hire vessels adapted to regulations. Fishing effort aboard charter boats remained consistent and primarily responded to the availability of “something” to harvest, suggesting that their customers are willing to substitute target species. Party boat anglers, in contrast, responded to the possession limits of black sea bass (*Centropristis striata*), and summer flounder (*Paralichthys dentatus*). Because party anglers were less willing to substitute target species, party vessel operators are likely particularly vulnerable to reductions in fishing opportunity and harvest potential.

**Keywords:** Recreational fishery, fisheries management, fishing effort, vessel trip reports, ARIMA model

## 1 **Introduction**

2           Recreational fisheries management worldwide struggles to limit harvest while  
3 concurrently meeting biological and socioeconomic objectives (Cox et al. 2002; Post et al. 2002;  
4 Abbott et al. 2018). Fisheries managers set and tune regulations such as season length,  
5 possession limits, and size limits to meet recreational harvest quotas, but angler response to these  
6 management changes is poorly understood. Anglers may adjust their behavior to compensate for  
7 new restrictions (e.g. Beaudreau et al. 2018; Gentner 2004; Powers and Anson 2018), or they  
8 may choose to leave the fishery (e.g. Holzer and McConnell 2017; Mackay et al. 2020;  
9 Whitehead et al. 2015). Restrictive regulations may not result in the expected reduction in  
10 harvest in the presence of compensatory behavior. Conversely, declining participation in the  
11 fishery can harm coastal communities that rely on income from the recreational fishing industry  
12 (Chan et al. 2018; Murray et al. 2010; NMFS 2018). Further complicating this calculation, in  
13 multispecies fisheries, anglers may switch targets when regulations are no longer acceptable to  
14 them (Beaudreau et al. 2018). This may be a desirable outcome if it relieves pressure on  
15 threatened stocks, but these alternative targets may then be subject to enough harvest pressure to  
16 become depleted (Abbott et al. 2018). Whereas fisheries managers can frequently monitor  
17 commercial harvest throughout the season (e.g. Gerritsen and Lordan 2011; Lee et al. 2010),  
18 recreational fisheries managers generally have few options for monitoring harvest or making  
19 changes mid-season (Pereira and Hansen 2003). An empirical understanding of the link between  
20 fishing regulations and resulting fishing effort is therefore needed to better inform fisheries  
21 management choices.

22           Regulations have the potential to reduce the utility that anglers receive from fishing, but  
23 their effects depend on individual preferences. Recreational anglers place value on catch (i.e. the

24 number of fish kept and released), harvest (i.e. the number of fish kept), and the overall fishing  
25 experience (Hunt et al. 2019). Throughout this paper, we will use “catch” to indicate all fish  
26 caught, including those kept and released, while “harvest” refers only to fish that are caught and  
27 kept. In a utility-maximizing approach to understanding angler decisions, the choice of whether  
28 or not to fish will depend on whether the expected fishing experience, catch, and harvest provide  
29 enough utility to outweigh the cost in time and money incurred by taking the trip (e.g. McFadden  
30 1974). Reductions in season length do not necessarily reduce the value of fishing trips, but they  
31 narrow the window of opportunity for anglers to schedule their fishing trips. This loss of  
32 opportunity potentially results in the loss of benefits related to the overall fishing experience if,  
33 for instance, inclement weather cancellations are proportionally more common. Reduced  
34 possession limits, in contrast, may reduce the benefits anglers receive from harvest itself.  
35 Anglers may still catch a lot of fish, which may still be satisfying to individuals who are highly  
36 catch-oriented (e.g. Schroeder and Fulton 2013). However, the lower possession limit places a  
37 ceiling on the harvest that anglers can take home, meaning that anglers who primarily fish for  
38 food may no longer decide to take the trip. Since the experience of the fishing trip is still valued  
39 by many anglers regardless of catch, however, fishing effort can remain highly elastic to  
40 regulations, depending on angler preferences (e.g. Beardmore et al. 2011a). When anglers do  
41 leave the fishery as a result of benefit loss associated with restrictive regulations, coastal  
42 communities experience negative economic effects as vessel operators and other businesses  
43 associated with the recreational fishery lose revenue (NMFS 2018). Understanding these  
44 potential angler responses therefore allows fisheries managers to weigh tradeoffs in the  
45 biological, social, and economic outcomes of their decisions.

46 Much uncertainty therefore exists when predicting how recreational fishing effort, and  
47 therefore harvest, will respond to changes in regulations. This uncertainty arises in part from  
48 unknowns associated with angler behavior, motivations, and preferences (e.g. Brinson and  
49 Wallmo 2017; Johnston et al. 2010). While anglers tend to express preferences for longer open  
50 seasons (Holzer and McConnell 2017; Young et al. 2019; Melnychuk et al. 2021), shorter  
51 seasons do not necessarily cause anglers to reduce their fishing effort. For example, during  
52 extreme reductions in season length for the red snapper (*Lutjanus campechanus*) fishery in the  
53 Gulf of Mexico, daily angler effort substantially increased, leading to a “derby style” fishery  
54 where private anglers, who own or rent their own boats, attempted to fish as much possible  
55 during their allotted time (Powers and Anson 2018, 2016). Shorter seasons therefore still  
56 corresponded to lower harvest across the season, but not in proportion to the change in season  
57 length. Because the functional response of fishing effort to shorter seasons is not often quantified  
58 and likely varies widely by fishery, this “effort compression” effect complicates managers’  
59 predictions of the response of harvest to changes in regulations. Further, reductions in possession  
60 limits can reduce the attractiveness of fishing opportunities to anglers (Whitehead et al. 2015),  
61 but angling effort in different fisheries may show different degrees of elasticity to changes in  
62 these regulations (Beard et al. 2003; Beardmore et al. 2011a) and may therefore not substantially  
63 affect overall harvest (van Poorten et al. 2013). In fisheries where open seasons overlap for  
64 multiple species, predicting angler response is further complicated. For example, in the  
65 multispecies for-hire recreational fisheries in Alaska, increased restrictions on harvest of Pacific  
66 halibut (*Hippoglossus stenolepis*) has been associated with increased harvest of less restricted  
67 species (Beaudreau et al. 2018). This substitution behavior can lead to a continuous “spiraling”  
68 effect of regulations where managers implement increasingly strict limits on an increasing

69 variety of species, and anglers continue to adapt by diversifying their targets in order to maintain  
70 their harvest (Abbott et al. 2018; Beaudreau et al. 2018). The effects of regulations on fishing  
71 effort may therefore depend on how anglers and operators of for-hire vessels respond to fishing  
72 opportunity (i.e. season length), harvest potential per trip (i.e. possession limit or variety of  
73 species available), and preferences for specific species (e.g. the popularity of species among  
74 harvest- or trophy-oriented anglers).

75 Because of this uncertainty in angler response to regulation, managers of open-access  
76 fisheries have not always successfully kept removals below sustainable harvest limits (Coleman  
77 et al. 2004; Cooke and Cowx 2004; Cox et al. 2002; Post et al. 2002; NEFSC 2019). This  
78 inconsistency in constraining recreational harvest points to a need for empirically understanding  
79 the effects of regulations on fishing effort in a multispecies context. Forecasting and  
80 “nowcasting” techniques have already been successfully used to predict landings in the Gulf of  
81 Mexico recreational fishery for individual species (Carter et al. 2015; Farmer and Froeschke  
82 2015), but not to infer the effects of multiple species’ regulations on fishing effort. By  
83 understanding the dynamics of both catch and effort in response to regulations, managers can  
84 reduce the uncertainty around how changes in season length of multiple (or individual) seasons  
85 in multi-species fisheries will affect fishing effort.

86 The Marine Recreational Information Program (MRIP) produces estimates of recreational  
87 catch and effort for most coastal states. Estimates are aggregated by two-month “waves” or by  
88 year. More granular estimates of fishing effort, however can be difficult and expensive to obtain  
89 in recreational fisheries (McCluskey and Lewison 2008), but Vessel Trip Report (VTR) data  
90 provide a daily census count of recreational fishing effort aboard federally-permitted for-hire  
91 vessels in the Greater Atlantic Region. We then empirically evaluated the response of weekly

92 fishing effort to changes in possession limits, season length, and season overlap in the New  
93 Jersey (NJ), USA, for-hire sector of the bottom fishery using this VTR data. To do this, we fit  
94 statistical models incorporating effects of four species' overlapping open seasons, their season  
95 lengths, and the number of "blackout" days during which none of the four species are available  
96 to harvest to a time series of weekly fishing effort. Guided by hypotheses formulated through  
97 focus-group interviews with stakeholders, a model selection process allowed us to infer the  
98 dominant mechanisms by which changes in possession limits, season length, and species  
99 availability could have influenced overall fishing effort in the NJ for-hire bottom fishery.  
100 Differences in overall preferences between anglers participating in the charter and party boat  
101 fisheries were inferred by fitting these models to time series separately for each sector.

102         The New Jersey bottom fishery is primarily harvest-motivated (e.g. Bochenek et al.  
103 2012), so we hypothesized that lower possession limits for popular species would be associated  
104 with a reduction in angler trips in a given week. While lower possession limits reduce the harvest  
105 potential of single fishing trips, shorter and more fragmented fishing seasons instead limit angler  
106 access to the fishery. During closed seasons, no targeting of any affected species is permitted,  
107 even for catch and release angling. Shorter seasons therefore leave fewer days available to fish  
108 for a given species each year, and reduced overlap of these seasons may limit the variety of fish  
109 that an angler is allowed to catch and harvest. Reductions in fishery access through shortened  
110 seasons has historically been assumed to have a direct effect on fishing effort, where angling  
111 trips that would have taken place during the now-closed season simply do not occur. We  
112 hypothesized that reductions in fishing effort associated with shorter seasons may instead be  
113 lower or higher than expected depending on whether anglers tended to respond to benefit loss  
114 associated with regulatory change by either 1) compensating for reduced fishing opportunity or

115 2) reducing their participation in the fishery. Of course, angler response to these changes in  
116 regulations will be heterogeneous because their responses depend on motivations and  
117 preferences that vary among anglers (e.g. Beardmore et al. 2011b). If particular responses  
118 dominate angler effort dynamics, however, the overall effect on all angler effort will be useful in  
119 a broad-scale policymaking context. We conducted a time series analysis of weekly total angler  
120 trips from the recreational for-hire sector in NJ to test the following hypotheses derived from  
121 focus group data describing how anglers may have adapted to changes in fishing opportunity:

- 122 1) *Species availability hypothesis*: Anglers switch between preferred species to maintain their  
123 opportunities to go fishing.
- 124 2) *Season length hypothesis*: Anglers intensify their fishing effort during shorter open seasons to  
125 maintain their preferred harvest levels.
- 126 3) *Blackout effect hypothesis*: In response to an increasing number of “blackout” days, where  
127 neither of these four bottomfish are available for harvest, anglers will either a.) increase their  
128 fishing effort during the remaining open seasons or b.) begin to exit the fishery.

## 129 **Methods**

### 130 *Study system*

131 The NJ marine recreational fishery is socioeconomically important, ranking fourth in the  
132 nation in state sales revenue generated by the recreational fishing industry (NMFS 2018). NJ  
133 anglers are also responsible for substantial removals, ranking second among US states in pounds  
134 of recreational harvest and fourth in release numbers (NMFS 2020). The for-hire sector makes  
135 up between 5 and 20% of total recreational catch, depending on the species, while the remaining  
136 catch is made up by shore-based anglers and private anglers who own or rent their own boats  
137 (ASMFC 2017; MAFMC and ASMFC 2020). The for-hire fleet is made up of party boats (also



138 called head boats), where anglers pay between \$30 and \$90 “per head” for a 4-8 hour guided trip  
139 shared with up to 100 other anglers, and charter boats, where a smaller group of anglers  
140 (typically 6 or fewer) pays more, currently between \$400 and \$1000, for a more personalized  
141 guided fishing trip on a smaller vessel (Steinback and Brinson, 2013). For-hire fishing vessels  
142 are highly accessible. Anglers may borrow or rent fishing gear, and no additional licensing or  
143 registration is required to participate. Spending by out-of-state anglers is particularly impactful in  
144 the for-hire fishing industry, and fishing effort by these anglers in this sector is sensitive to  
145 changes in fares (Li et al. 2019; Steinback 1999). As overhead costs (e.g. fuel, bait, boat  
146 maintenance) increase among for-hire operators as a result of fuel prices and reduced season  
147 lengths, businesses and communities relying on revenue from this sector are increasingly  
148 vulnerable to volatility in angler numbers which could result from regulatory changes (Murray et  
149 al. 2010).

150 As fisheries managers have struggled to limit harvest in order to maintain or rebuild fish  
151 stocks, the NJ marine recreational fishery has experienced marked changes in possession limits  
152 and season lengths for summer flounder (*Paralichthys dentatus*), black sea bass (*Centropristis*  
153 *striata*), scup (*Stenotomus chrisops*), and tautog (*Tautoga onitis*) (Fig. 1, Tables S1-S4). In spite  
154 of these changes, black sea bass recreational harvest in recent years (2013-2017) has exceeded  
155 harvest limits by an average of 41% (MAFMC 2018), and tautog continues to be classified as  
156 overfished (ASMFC 2007; ASMFC 2017). Although summer flounder was not overfished as of  
157 the latest stock assessment (NEFSC 2019), changes in distribution, reductions in recruitment,  
158 upward corrections of previous years’ harvest estimates, and a strict fisheries management plan  
159 have led to the continuation of stringent harvest regulations (ASMFC 2018; Terceiro 2018).  
160 Summer flounder is a highly popular target species in the NJ marine recreational fishery, and the

161 resulting short and fragmented seasons in the face of perceived improvement in the summer  
162 flounder stock have led to widespread frustration among stakeholders (Terceiro 2018). Tautog  
163 season lengths were reduced in 2008 in response to overfishing in the recreational sector  
164 (ASMFC 2007). In spite of the rebuilding plan implemented at this time, tautog spawning stock  
165 biomass remains low, and the stock is classified as overfished (ASMFC 2017). In contrast, a  
166 fisheries management plan for scup that was implemented in 1998 and amended in 2007 was  
167 successful in reducing harvest, and the stock was declared recovered in 2009 (MAFMC and  
168 NMFS 2007; Northeast Data Poor Stocks Working Group 2009).

### 169 *Focus groups*

170 Four focus groups were conducted across a north-south transect of the NJ coast in the  
171 towns of Atlantic Highlands, Toms River, Tuckerton, and Cape May in the winter and spring of  
172 2019. Participants were identified through purposive sampling in which researchers consulted  
173 with NJ state agency staff, extension agents, and industry representatives to identify  
174 knowledgeable, experienced, and collaborative recreational fishing industry stakeholders. Two to  
175 four stakeholders from each of four industry segments (party boats, charter boats, private anglers  
176 who own their own boats or fish from shore, and associated businesses) in each of the four  
177 regions were identified, for a potential maximum of 16 participants per focus group. Of these, 44  
178 stakeholders were successfully contacted and invited, and 37 attended. The focus groups ranged  
179 from 8 to 11 participants, plus two note takers and a moderator, and they lasted between two and  
180 two and a half hours. Focus group participants were asked open-ended questions about their  
181 process for choosing bottomfish target species and how those decisions are influenced by  
182 management regulations and their clients' or their own personal preferences. All focus groups  
183 were audio recorded, transcribed, and coded for common themes, following the standard analysis

184 guidelines for qualitative research in Creswell and Poth (2016) and Roller and Lavrakas (2015).  
185 The focus group procedure was approved by the Rutgers Institutional Review Board (Protocol  
186 #E18-112).

187 Results from the focus groups were used to develop alternative hypotheses to be tested in  
188 the analysis of VTR data. Overall, recreational industry representatives expressed strong  
189 dissatisfaction with current regulations, especially season length and timing. As one focus group  
190 participant said, “What I’ve observed here is just absolute, total frustration, bordering on anger.  
191 And I keep saying to myself, these regulations are going to turn a lot of local fishermen to  
192 pirates.” Of particular concern to stakeholders were the partitioning of open seasons into shorter  
193 periods and the loss of overlapping seasons for different species (Table S5). A common point  
194 stakeholders discussed was that the loss of overlap between different species’ open seasons was  
195 leading anglers to intensively harvest whatever species remained open at a given time. Two  
196 possible mechanisms for this change in behavior were incorporated into the hypotheses for our  
197 model selection: 1) anglers maintain harvest potential by compressing fishing effort into shorter  
198 seasons to maintain harvest of particular species or 2) anglers switch target species in order to  
199 continue fishing on a consistent basis.

#### 200 *Effort, catch, and management data*

201 Vessel Trip Report (VTR) data from for-hire vessels between 2001 and 2017 were  
202 obtained from the NOAA VTR database for the Greater Atlantic region. VTRs are a census of  
203 vessels with federal permits for black sea bass, summer flounder, or scup where operators report  
204 the number of anglers aboard and enumerate their catch and harvest. VTR data from 2018 and  
205 2019 were not included in the analysis because the 2018 switch to mandated electronic reporting

206 may have resulted in a systematic change in reporting compliance. VTRs do not report target  
207 species, so data were filtered according to the vessels' port state and the species they reported  
208 catching in order to capture NJ bottom fishing effort. Reports listing capture of bottom fish  
209 (defined in Table S6) and a port of departure in NJ were retained. Many more angler trips were  
210 reported aboard party vessels, so fishing effort was evaluated separately for party and charter  
211 vessels to avoid dominance of fishing effort dynamics by party operators. We first investigated  
212 how the for-hire fleet changed during this time period. To do this, we compiled annual counts of  
213 reporting vessels, the mean number of anglers per trip, and the mean number of trips per week  
214 for party and charter vessels. Next, to build our time series for testing our hypotheses of angler  
215 response to regulations, we compiled a weekly time series of fishing effort by summing the total  
216 number of angler trips reported by all vessels for each week. This process produced two time  
217 series of weekly counts of angler trips on charter and party boats.

218 Fishing effort can also respond to fishing quality (e.g. Wilson et al. 2020), so we included  
219 species-specific catch rates as predictors in our models. Although catch by species is reported in  
220 VTRs, reports of catch (number of fish caught) and harvest (number of fish retained, i.e., catch  
221 minus fish caught and released) after a trip are prone to recall bias (Bochenek et al., 2012). Catch  
222 rates to be used as predictor variables were therefore obtained instead from Marine Recreational  
223 Information Program (MRIP) access point intercept survey data (NOAA Fisheries 2021). These  
224 surveys take place at ocean access points that are selected within a stratified random sampling  
225 regimen. Among other data, respondents report their total species-specific catch, which includes  
226 both kept and released fish. Using the procedure described in the MRIP Survey Design and  
227 Statistical Methods documentation (Papacostas and Foster, 2021), we calculated the mean catch  
228 per trip for each of our four focal species for each two month survey "wave" in our time series

229 (Fig. S1). Missing values were imputed using linear interpolation for black sea bass and tautog  
230 catch rates. Scup catch rates were not included as predictors because of the high number of  
231 missing values. Summer flounder missing values occurred in winter months when the stock has  
232 migrated off-shore. These missing values were therefore replaced with zero. Average catch rates,  
233 rather than spawning stock biomass (SSB), were used to estimate the effects of fishing quality  
234 because fish species associated with bottom structure likely exhibit catch rate hyperstability (e.g.  
235 Dassow et al. 2019; Erisman et al. 2011). In addition, these catch rates could be calculated for  
236 every two months of the time series, while SSB estimates are only available on an annual basis.

237 NJ fishing regulations for summer flounder, black sea bass, tautog, and scup were  
238 collected for the years between 2001 and 2017. Open seasons and possession limits were  
239 obtained from annual releases of NJ recreational fishing regulations. Mid-season closures were  
240 found by searching the Federal and NJ Registers for rule changes impacting fisheries of the  
241 Northeastern United States. State and Federal Registers document rule changes for the federal and NJ  
242 state government. The Federal Register can be accessed at <https://www.federalregister.gov/> and  
243 the NJ Register at <https://www.state.nj.us/oal/rules/accessp/>. In cases where changes to  
244 regulations occurred mid-season, we included only the final regulations in the analysis.

#### 245 *Statistical Analysis*

#### 246 *Base ARMA model*

247 We used autoregressive-moving average (ARMA) models to quantify how implementing  
248 or changing a management measure affected fishing effort while accounting for autocorrelation  
249 and seasonal trends. Fitting a time series model at this granular scale allowed us to detect  
250 average effects of changes in regulations within and between years using external regressors.

251 Simultaneously, additional unexplained variation (i.e. variation in angler trips attributable to  
252 weather, changes in trip price, etc) is accounted for implicitly by seasonal and ARMA  
253 components. ARMA models account for short-term temporal autocorrelation in time series data  
254 by fitting autoregressive (AR) terms to lagged observations and moving average (MA) terms to  
255 lagged residuals (Box et al. 2008; Box and Jenkins 1970). Weekly time series have a long and  
256 non-integer period (52.14 weeks/year), but seasonal models are periodic, being at the same state  
257 as one year previous and repeating. To better align these weekly data with the model's seasonal  
258 component, a dynamic harmonic regression approach (Hyndman and Athanasopoulos 2018;  
259 Young et al. 1999) was used to fit an appropriate number of Fourier sine-cosine pairs to each  
260 time series of fishing effort data. Open seasons for our focal species are highly correlated with  
261 seasonality (Fig. 1), so by fitting an identical seasonal trend to each year, we were able to  
262 examine how differences in possession limits and season length (e.g. the loss of early and late  
263 summer for the summer flounder fishing season) influenced weekly fishing effort in the weeks  
264 that did experience differences in regulations among years. Following this approach, increasing  
265 numbers of sine-cosine pairs were generated using the forecast package in R v.4.1.0 (Hyndman  
266 et al. 2020; R Core Team 2021), and for each of these model fits, the auto.arima function of the  
267 forecast package was used to find the best fitting ARMA components. The best fitting  
268 combination of ARMA components and Fourier sine-cosine pairs was then chosen based on its  
269 AICc score. We tested for serial autocorrelation using the Breusch-Godfrey test.

### 270 *Candidate model construction*

271 In addition to the aforementioned ARMA and seasonality components, we included as  
272 predictors the regulations and catch rate variables relevant to the candidate model's hypothesis  
273 (Table 1, Appendix 1). Considerable variation in catch rates both within and between years were

274 evident (Fig. S1). To indicate possession limits and closed seasons for summer flounder, black  
275 sea bass, and tautog, an integer predictor indicated the possession limit in that week. A  
276 possession limit of 0 indicated a week where targeting the species was not permitted. Scup  
277 possession limits during open season remained at 50 for the entire time series, so an indicator  
278 variable was used instead to indicate whether week was open (1) or closed (0) for scup fishing  
279 (Table 1). An additional dummy variable ('Something open') was used to indicate whether at  
280 least one of the four bottomfish species was available for harvest during the week (i.e. a 0 during  
281 a blackout period, 1 otherwise).

282         The models did not include year as a covariate but instead attempted to explain annual  
283 variation in fishing effort through six co-variates that described fishing opportunities in each  
284 year. Four continuous variables specified the length of each species' season in days for each  
285 year. Two additional continuous variables indicated the total number of blackout days in each  
286 year as well as the number of open species available each week. In most years, regulations are  
287 announced in late spring of their effective year (i.e. shortly before the start of peak summer  
288 season). In 2010, 2011, and 2013, however, season lengths were adjusted mid-season for summer  
289 flounder and/or black sea bass. In years when regulations were changed mid-season, the final  
290 effective season length was used as a predictor. To correspond with the approximate date of the  
291 release of new regulations, annual variables, which included season lengths and annual harvest  
292 days, were updated annually on May 1.

293         The *null model* incorporated the assumption that anglers do not compensate for changes  
294 in season length and overlap by changing their behavior. This model therefore included only the  
295 focal species' possession limits and their catch rates (Table 1). The other three candidate models  
296 included additional predictor variables and interaction effects that tested three hypotheses for

297 how anglers may compensate for regulatory changes (Tables 1 and 2). The *blackout effect model*  
298 added the ‘Something open’ predictor to test the hypothesis that open seasons for any of the four  
299 focal species would attract fishing effort, regardless of which species were open. In addition, the  
300 annual number of blackout days and its interaction with ‘Something open’ was included to test  
301 for anglers’ response to an increasing number of blackout days on the calendar (Table 2). A  
302 positive interaction effect would suggest that anglers increased their fishing effort during the  
303 remaining open days in response to a higher number of blackout days, and a negative interaction  
304 effect would indicate that anglers instead tended to stop fishing in response to these changes. To  
305 illustrate, as the number of blackout days in a year increases from 0 to 30, an intensification of  
306 fishing effort during the open season would be indicated by a positive parameter value for the  
307 two-way interaction of the number of blackout days and the ‘Something open’ indicator. A week  
308 where at least one species is open during a year with 30 blackout days would then have a higher  
309 predicted fishing effort than that of a week in a year with 0 blackout days. On a blackout week,  
310 however, the ‘Something open’ indicator is zero, negating the interaction effect.

311         The *season length model*, in contrast, allowed anglers to display different responses to  
312 changes in the season length of different species. The model incorporated this behavior by  
313 including season length as a predictor conditional on the corresponding species’ open season (i.e.  
314 the possession limit is greater than 0). Our hypothesis that anglers would compensate by  
315 increasing their fishing effort during the remaining open season would be supported by a  
316 negative interaction effect between species-specific season length and the corresponding open  
317 season indicator. The *species availability model* accounted for specific substitution patterns used  
318 by anglers to maintain their fishing opportunities as the overlap between different species’ open  
319 seasons was reduced. A negative interaction effect between species-specific open seasons and



320 the number of species that were available would indicate a non-additive response of fishing  
321 effort to new open seasons. In other words, adding an additional open species to a given week  
322 would result in a lower increase in fishing effort than expected because many of those anglers  
323 were already fishing.

#### 324 *Model selection*

325 We evaluated the ability of our four candidate models to explain weekly natural log-  
326 transformed total fishing effort for party boats and charter boats. Angler trip counts were log-  
327 transformed to account for the greater variance in fishing effort during peak fishing season. The  
328 fit of the competing models was compared using the corrected Akaike Information Criterion  
329 (AICc) and their associated Akaike weights calculated using the MuMIn package (Bartoń 2020).

330 To evaluate the relative effects of changes in black sea bass and summer flounder  
331 possession limits and season length, we produced annual predictions of angler trips for  
332 hypothetical years under different regulations. Tautog regulations were not evaluated in this way  
333 because possession limits primarily changed within rather than between years in order to protect  
334 tautog from excessive harvest during their summer spawning season (Table S3). For each of  
335 these predictions, the fishing effort associated with each week of an average year (i.e. the  
336 average value of each week's sine-cosine coefficient pairs across all years of the time series) at  
337 average catch rates (i.e. the average value of each week's CPUE for summer flounder, black sea  
338 bass, and tautog) were forecast using the best fitting ARMA model with the predict.Arima  
339 function (R Core Team, 2021). Only the species of interest was "opened" for the hypothetical  
340 forecasted year. A year of weekly predictions were forecasted for each combination of  
341 possession limits and season lengths. We then applied a bias correction to these predictions

342 based on a non-parametric smearing adjustment (Duan, 1983), and annual fishing effort was  
343 summed for each forecasted year. These predictions produced estimates of the annual fishing  
344 effort associated with different combinations of season length and possession limits for specific  
345 species. These forecasts are intended to illustrate the relative effects of possession limit and  
346 season length changes for different species on angler trips in past years. They are not intended to  
347 forecast out-of-sample future changes in fishing effort.

## 348 **Results**

### 349 *Fleet changes*

350 The NJ for-hire bottom fishing fleet has experienced a number of changes since 2001.  
351 The decline in charter vessels reporting each year since 2010 is particularly distinctive, declining  
352 from 119 to 57 reporting vessels (Fig. 2A). In spite of the decline in charter vessels, the number  
353 of charter boat anglers and the mean number of anglers per charter trip have remained largely  
354 constant (Fig. 2B and 2C). This consistency in angler numbers is explained by a near-doubling in  
355 the average number of trips taken per charter vessel between 2010 and 2015, from 17.6 to 28.6  
356 charter trips per year. In contrast, the number of party boats has shown a less extreme overall  
357 decline, with the exception of a period between 2001 and 2005, where the number of reporting  
358 party boats dropped by nearly half (Fig. 2A). This change in party boat numbers corresponds  
359 with simultaneous decline in party boat angler trips (Fig. 2B) and a reduction in the average  
360 number of trips taken by each vessel (Fig. 2D). Both party boat numbers and angler numbers  
361 largely recovered by 2010, but they remained lower than in the early 2000s.

### 362 *Model selection*

363 *Charter boat fishing effort*

364 The time series of charter boat and party boat fishing effort differed in their best fitting  
365 models (Tables 3 and 5), indicating that charter and party boat anglers responded differently to  
366 changes in regulations. The *blackout effect model* was unambiguously the best fit to charter boat  
367 fishing effort, receiving 100% of the Akaike weight (Table 3). The ARMA and seasonal  
368 components of the model effectively removed serial autocorrelation of the residuals according to  
369 the Breusch-Godfrey test (Table S7). Total fishing effort on charter boats was relatively  
370 consistent between years (Fig. 2B), and variation in weekly effort was driven mainly by  
371 seasonality rather than by open seasons of specific species (Tables 4 and S8). In spite of these  
372 species' popularity, neither black sea bass, summer flounder, or tautog possession limits, nor  
373 scup open seasons were associated with significant changes in fishing effort on their own (Table  
374 4). All else being equal, the opening of at least one of the four species was associated with an  
375 over 6-fold increase in angler trips (i.e.  $\exp(1.954)=7.06$ ), suggesting that charter anglers are  
376 flexible in their species preferences ( $p=0.008$ , Table 4). All else being equal, the availability of at  
377 least one of the four focal species was associated with an over 600% increase in fishing effort  
378 compared to a "blackout" day. The interaction of the 'Something open' indicator with the annual  
379 number of blackout days, however, was not significant ( $p=0.143$ , Table 4). Charter boat anglers  
380 therefore did not appear to leave the fishery in response to increasing numbers of blackout days,  
381 which would have been evident by a negative interaction. Nor did they appear to compensate for  
382 blackout days by increasing fishing effort, which would have been evident by a positive  
383 interaction. Fishing effort of charter angler trips also did not appear to respond to summer  
384 flounder, black sea bass, or tautog catch rates when aggregated at the two-month level.

385 *Party boat fishing effort*

386 The species availability model was unambiguously the best fit to the time series of party  
387 boat angler trips (Table 5). Summer flounder, tautog, and black sea bass possession limits were  
388 significant predictors of fishing effort, where an increase in limit of 1 fish was respectively  
389 associated with a 26%, 15%, and 5% increase in angler trips. The opening of multiple species  
390 simultaneously, however, did not have an additive effect on fishing trips. The negative  
391 interaction between summer flounder, tautog, and black sea bass open seasons and the number of  
392 open species suggests that a subset of the anglers fishing for summer flounder, for example, were  
393 already previously fishing for black sea bass or tautog before the flounder season opened. Weeks  
394 where all three species are open for harvest therefore experienced fewer angler trips than would  
395 be predicted by only the species-specific possession limits. Scup open seasons, in contrast, were  
396 associated with increased angler trips in combination with the availability of additional target  
397 species. In our dataset, scup only occurred in combination with at least one other species ( $N$   
398  $\text{species}=2$ ). During these combined seasons of two species (usually tautog and scup), scup is  
399 associated with only a small decrease in mean angler trips (i.e.  $\exp(-0.944+1*0.99)=0.996$ , or a  
400 0.4% decrease in angler trips). In combination with two or three other species, however, scup  
401 season is respectively associated with a 63% or 268% increase in fishing trips. These overlaps  
402 typically occurred in the peak summer season, when scup is available inshore. During the rest of  
403 its winter open season, scup has migrated offshore to deeper water, where it is more difficult to  
404 target (NMFS 1999). Fishing effort did not obviously respond to black sea bass or summer  
405 flounder catch rates, but angler trips did increase by 3% in correspondence with an increase in  
406 tautog catch rates of 1 fish per trip ( $p = 0.015$ , Table 6).

407 Multicollinearity was detected between certain predictor variables for the species  
408 availability model, with variance inflation factors (VIF) as high as 9.9 (Table S11). To test the

409 sensitivity of the parameter estimates to this collinearity, we completed a supplemental analysis  
410 by re-fitting the model without the most highly correlated predictors (i.e. summer flounder  
411 possession limits and catch rates). Coefficient estimates were effectively the same, except for the  
412 interaction effect of tautog possession limits with the number of open species (Tables S12 and  
413 S13). When summer flounder-associated predictors were removed from the model, this  
414 interaction was no longer significant.

415 Black sea bass seasons experienced substantial variation in both possession limits and  
416 season lengths among years (Fig. 3). Only modest increases in annual angler trips relative to  
417 closed season were associated with the possession limits of two or three fish that were  
418 implemented in peak summer fishing seasons starting in 2014 (Fig. 1). Higher possession limits,  
419 in contrast, were associated with tens of thousands more angler trips per year. Summer flounder  
420 season lengths experienced less change among years in our analysis (Fig. 4A). In spite of this  
421 limited variation of season lengths, distinct changes in annual fishing effort were detected. As  
422 possession limits were lowered, however, the response of annual fishing effort to season length  
423 became less distinct (Fig 4B).

## 424 **Discussion**

425 Previous survey-based studies of recreational anglers' stated preferences have highlighted  
426 the importance of preserving fishing opportunities in the form of open fishing seasons in order to  
427 maintain angler satisfaction (Brinson and Wallmo 2017; Young et al. 2019). The use of VTR  
428 data allowed us to investigate the empirical response of anglers aboard for-hire vessels to  
429 reduced fishing opportunity. We found evidence of substantial reductions in annual fishing effort  
430 within the party and charter boat fisheries as a result of reduced possession limits and, to a lesser

431 extent, contracting season lengths. These results support the concerns expressed by focus group  
432 participants regarding reduced profitability of for-hire fishing vessels in the face of these  
433 increased restrictions. Fishing effort dynamics within the charter boat fishery were best  
434 explained by the *blackout effect model*, where the ability to harvest any one of the four species  
435 was a more important predictor of fishing effort than the availability of any specific species.  
436 Fishing effort in the party boat fishery, in contrast, was best explained by the *species availability*  
437 *model*, and angler trips specifically responded positively to summer flounder and black sea bass  
438 open seasons. The non-additive effects of additional open seasons suggested a significant degree  
439 of substitution behavior occurring among party boat angler trips as species open and close  
440 throughout the season. The interaction effect of tautog open season with species availability was  
441 non-significant in the sensitivity model fit that eliminated summer flounder predictors.  
442 Substitution behavior may therefore be less common among tautog anglers. Among charter boat  
443 anglers, however, substitution behavior appears to be even more prevalent, as indicated by the  
444 strong positive effect of the “Something open” predictor.

445         Although substitution behavior appears to vary between charter and party boat anglers,  
446 our ability to infer specific angler behaviors is limited because the number of angler trips in a  
447 week also depends on the availability of trips for hire. Responses of angler trips to regulations  
448 may therefore indicate differences in operator behavior rather than angler preferences. The  
449 corresponding decline in federally permitted charter vessels and increase in annual trips per  
450 vessel, for example, suggest that the demand for charter trips may exceed the supply. If the  
451 remaining operators are allowed to target bottom fish on a given day, they will most likely be  
452 able to reserve enough customers to fill their vessel. The response of charter angler trips to the  
453 availability of “something” may therefore be an indication of operator behaviors. Angler trips

454 aboard party vessels, however, appeared to show more room for variation. Similar to charter  
455 trips, the number of weekly party angler trips can be limited by the availability of spots aboard  
456 party vessels. Conversely, at very low demand, party vessels will cancel trips if the number of  
457 spots sold do not recoup costs. However, considerably more variation is possible in the number  
458 of anglers aboard large party boats once this threshold of profitability is reached, suggesting to us  
459 that party boat fishing effort dynamics primarily reflect angler preferences. In particular, the  
460 large negative effects of reduced possession limits on the number of weekly angler trips suggest  
461 that many anglers have quit bottom fishing on party vessels in response to these changes.  
462 Because substitution behaviors do not appear to be as strong in the party fishery as in the charter  
463 sector, party vessel operators probably could not rely on angler substitution of less popular  
464 bottom species to maintain their profits. Party vessel operators may therefore be particularly  
465 vulnerable to the negative economic effects of increased restrictions on bottomfish harvest.

466         Considerable additional variation existed in angler trips that was not explained by  
467 changes in regulations. For example, a nearly 50% drop in angler trips occurred between 2005  
468 and 2010 (Fig. 5), which did not correspond to any specific changes in regulations. This time  
469 period does, however, roughly correspond with a period of conflict over reductions in the  
470 acceptable biological catch (ABC) for summer flounder, the implementation of conservation  
471 equivalency among states, and the stock assessment methods used by fisheries scientists  
472 (Terceiro 2011). The rebound in party boat angler numbers in 2010 is also coincident with a new  
473 stock assessment indicating that the summer flounder stock was not overfished and did not  
474 experience overfishing between 2008 and 2010 (Terceiro 2018). As a new control rule was  
475 implemented after the 2011 season, however, the ABC was reduced, leading to another round of  
476 conflict between scientists, managers, and stakeholders (Terceiro 2018). At the seasonal level,

477 these changes in annual fishing effort stem from a reduction in “peak” fishing effort for summer  
478 flounder during the summer months of May through August (Fig. 5A). Although black sea bass  
479 availability is also associated with higher fishing effort aboard party boats, similar patterns in  
480 monthly fishing effort are evident during years with and without year-round black sea bass  
481 seasons (Fig. 5B). Therefore, although fishing regulations influenced the number of angler trips  
482 each week, we speculate that trust in management and public perceptions of summer flounder  
483 stock health are potentially important predictors of fishing effort.

484 Vessel Trip Report data represent a large and mostly untapped resource for studying  
485 marine recreational fishing effort dynamics. However, they also present several challenges. First,  
486 only vessels with federal permits are required to submit VTRs. Federal permit are required for  
487 summer flounder, black sea bass, and scup fishers, but not for tautog. Charter vessels in  
488 particular may be underreported in the VTR data if they do not target either of these three  
489 species. In addition, VTRs report catch but not target species. We therefore defined  
490 bottomfishing trips based on the reported capture of at least one of nine bottom-associated  
491 species, which may have excluded some bottomfishing trips where nothing was caught.  
492 However, fishing trips with no reported catch made up only 1.5% of all fishing reports, so we  
493 believe that any effects of their elimination should be minimal. By filtering data by catch, we  
494 may also have included some trips targeting non-bottomfish species, such as striped bass  
495 (*Morone saxatilis*) or bluefish (*Pomatomus saltatrix*), during which bottomfish were caught  
496 incidentally. Both of these species remained open during the “blackout” periods recorded in our  
497 time series. The distinctively reduced weekly effort aboard charter vessels evident during these  
498 blackout periods suggests, however, that our filtering was largely successful at removing these  
499 trips. In addition, minimum length limits are important issues for fishery stakeholders (Table S5),



500 but they were not included as predictor variables because of excessive collinearity with  
501 possession limits. Minimum length limits tended to increase as possession limits decreased, so  
502 some of the effects of minimum length limits on fishing effort were explained in our model fits  
503 by changes in possession limits. Lastly, although VTRs provide a census count of anglers aboard  
504 federally permitted for-hire vessels, operators targeting tautog are not required to acquire a  
505 federal permit. We expected that operators targeting tautog would also target other highly  
506 popular bottomfish that do require federal permits, but we may have missed vessels specializing  
507 in tautog fishing, particularly among charter vessels.

508         The apparent willingness of anglers to substitute target species aboard charter boats, and  
509 to a lesser extent aboard party boats, has a number of implications for management of marine  
510 recreational fisheries. In particular, the relatively stable fishing effort in the charter sector  
511 regardless of individual species' closures suggests that discards may be high for closed species  
512 that are caught incidentally when anglers target other bottom fish. In other fisheries where  
513 anglers show high willingness to substitute target species, discard mortality has been  
514 demonstrated to reduce the effectiveness of seasonal closures (Chagaris et al., 2019). This  
515 phenomenon highlights the importance of understanding angler motivations for maintaining  
516 fishing opportunities and/or harvest. The relative importance of preserving fishing opportunity  
517 versus harvest capacity has been investigated in a variety of systems (e.g. Melnychuk et al.,  
518 2021; Young et al., 2019) and angler response to these changes appears to depend in part on  
519 anglers' willingness to re-allocate fishing effort to other time periods or alternative species. In  
520 other harvest-oriented fisheries, anglers express strong preferences for higher possession limits  
521 (e.g. Mackay et al. 2020). Reductions in possession limits and complete closures reduce anglers'  
522 harvest capacity and therefore their expected satisfaction, resulting in reduced fishing effort

523 overall if anglers are unwilling to substitute less-restricted species (Powell et al., 2010).  
524 Redirected fishing effort can lead to increased harvest of substitute species (Beaudreau et al.,  
525 2018) or increased discard mortality when closed or restricted species are caught and released  
526 (Chagaris et al., 2019). Although we investigated only the response of for-hire recreational  
527 fishing effort, the effect of regulation change on total recreational fishing effort also depends on  
528 the response of private boat anglers. These anglers do not rely on the availability of spots aboard  
529 for-hire vessels, suggesting that they have more ability to respond to closures by re-allocating  
530 fishing effort to different times of year. This response was observed in the Gulf of Mexico red  
531 snapper fishery when season length was drastically reduced (Chagaris et al., 2019; Powers and  
532 Anson, 2018, 2016). In less extreme instances of season reductions, however, private anglers  
533 may instead target alternative species to maintain their level of harvest or opportunities to fish,  
534 leading to a more stable pattern of fishing effort similar to our observations of charter vessels.  
535 Alternatively, the costs of maintaining a private vessel may drive some private anglers to leave  
536 the fishery when regulations become more stringent. If this choice is widespread, fishing effort,  
537 harvest, and discards would decline, but coastal communities would also experience the negative  
538 economic impacts associated with reduced angler participation. Responses to regulations among  
539 both private and for-hire anglers are therefore important to understand when evaluating the  
540 effects of new regulations on fishing effort, harvest, and discard mortality. An ongoing project  
541 by this team is using stated preference methods to investigate these potential responses among  
542 private and for-hire anglers.

543 Fisheries managers constantly consider tradeoffs in ecological, social, and economic  
544 objectives with the goal of maintaining stocks above safe harvest limits, maintaining public  
545 access to the fishery, and supporting the economies of coastal communities (e.g. Punt 2017). In

546 addition to wrestling with uncertainties in population dynamics of important stocks, considerable  
547 uncertainty surrounds the response of fishers to changes in regulations and ecological conditions  
548 (Fulton et al. 2011). Accounting for the responses of human stakeholders with heterogeneous and  
549 often competing preferences is vital for enacting proactive management decisions (Johnston et  
550 al. 2010). For-hire vessels make up one of these heterogeneous stakeholder groups and provide  
551 relatively low-cost access to fish stocks for recreational anglers globally. Recreational fisheries  
552 are also a major source of fishing mortality (Coleman et al. 2004; Cooke and Cowx 2004), and  
553 many of the costs of reduced harvest are borne by for-hire vessels, their customers, and the  
554 coastal communities relying on their economic contributions. In recent years, for example, fleet  
555 diversity of the recreational fishery in the Mid-Atlantic has declined as more anglers switch to  
556 shore-based modes of fishing and away from for-hire vessels (NEFSC 2021). Between  
557 uncertainty surrounding new regulations each year and reduced participation of anglers in the  
558 for-hire sector, for-hire operators are left in a precarious economic position. Illustrating this  
559 concern, one focus group participant stated, “Name me one industry besides fishing [...] where  
560 we can’t go year to year and we can’t budget, we can’t forecast, we can’t predict. And you show  
561 me one industry where you have that every year, year after year, and still stay in business.”  
562 Fisheries managers are therefore left in the difficult position of being accountable for keeping  
563 recreational harvest within imposed limits while also balancing the biological, social, and  
564 economic objectives of stakeholders, including these for-hire operators. Uncertainty associated  
565 with angler responses to changes in fishing regulations is an important limitation in managers’  
566 ability to constrain recreational harvest. Further investigations of angler behavioral responses to  
567 regulation should continue to help managers to enact regulations that prevent overharvest while  
568 meeting the economic needs of coastal communities.

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583 **Competing interests**

584 The authors declare there are no competing interests.

585 **Contributors**

586 AT: Formal Analysis, investigation, visualization, writing—original draft, writing—review &  
587 editing

588 EB: Conceptualization, funding acquisition, writing—review & editing

589 AG: Conceptualization, funding acquisition, investigation, writing—original draft, writing—  
590 review & editing

591 MM: Methodology, writing—review & editing

592 DZ: Conceptualization, funding acquisition, investigation, writing—review & editing

593 OJ: Supervision, conceptualization, funding acquisition, writing—review & editing

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## 604 **Data availability**

605 VTR data are confidential, and access must be granted by the NOAA National Marine Fisheries  
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607 researchers who have been granted data access by NOAA.

608

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Table 1: Predictor variable descriptions for each of the four models. Checks mark the predictors that were included in each model, and asterisks indicate the predictors that were included as predictors conditional on the corresponding open season. Obelisks indicate predictors with two-way interactions with corresponding species' possession limits. Species-specific regulations and catch rates are indicated by the following abbreviations: black sea bass (BSB), summer flounder (SMF), tautog (TOG), and scup (SCP).

Predictors	Description	Variable type	Update frequency	Null model	Blackout effect model	Season length model	Species availability model
'BSB PL' 'SMF PL' 'TOG PL' 'SCP open'	Possession limits of each focal species, with 0 indicating closed season.  Scup possession limits did not change, so scup open season is a binary variable, with 1 indicating open season and 0 closed.	Integer	Weekly				
'CPUE BSB' 'CPUE SMF' 'CPUE TOG'	Mean catch per trip of each focal species	Continuous	Bimonthly	✓	✓	✓	✓
'Something open'	Indicator that at least one of the focal species' seasons is open (i.e. PL > 0)	Binary	Weekly		✓		
'N blackout days'	Annual number of blackout days, where none of the focal species are open	Integer	Annual		✓		

'Season length BSB'	Number of days of open season per year for each focal species	Integer	Annual	✓*	
'Season length SMF'					
'Season length TOG'					
'Season length SCP'					
'N species open'	Number of focal species open each week	Integer	Weekly		✓†

Table 2: Hypotheses about angler behavioral responses to changes in regulations for flounder (SMF), black sea bass (BSB), tautog (TOG), and scup (SCP). Hypotheses are illustrated with representative quotations from stakeholder focus groups and operationalized in candidate models with different interaction effects. Species-specific regulations and catch rates are indicated by the following abbreviations: black sea bass (BSB), summer flounder (SMF), tautog (TOG), and scup (SCP). Bolding indicates the interaction sign that supports the hypothesis represented by the selected quotation. Further detail on model parameterization can be found in Appendix 1 of the supplemental materials.

Model	Representative quotation	Interaction effects	Interpretation
0. Null	N/A	N/A	N/A
1. Blackout effects	<i>“We do see a tremendous setback that occurs because of the eight or ten-day closure during the end of June. The people just kind of stop coming when that happens and you go [snaps fingers] it’s over.”</i>	Something open * number of blackout days	Positive: Anglers compensate for fewer harvest days (or more blackout days) with higher weekly fishing effort  <b>Negative: Anglers leave the fishery on year with fewer harvest days/more blackout days.</b>
2. Season length	<i>“We used to fish through March. When you could do that, you didn’t need to all press in to whatever the next fishery was...Now everybody’s just trying to get every last day in that they can because there’s so few of them available.”</i>	Season length BSB : BSB open Season length SMF : SMF open Season length TOG : TOG open Season length SCP : SCP open	<b>Negative: Effort compression. Anglers compensate for shorter seasons of particular species by increasing weekly fishing effort while that species is open.</b>
3. Species availability	<i>“[Black] sea bass is the only thing open. [Tau]tog’s closed, fluke’s closed. All of</i>	BSB PL * N species open SMF PL * N species open TOG PL * N species open	<b>Negative: Anglers switch target species when the</b>

*the angler pressure is now on sea bass. Where it used to spread out and diversify and the anglers would do other things, no matter what it was. You have a very severe angler impact on a single species due to the way regulations are set up, leaving no other choice but to target specific species.”*

SCP Open \* N species open

**season for their initial target species closes.**

Table 3: Model fit and Akaike weights of all candidate models for the time series of charter boat fishing effort. Bolded values indicate the lowest AICc and highest weight.

Model	AICc	AICc weight	Log Likelihood	# Parameters
Null model	2192.76	0	-1052.24	42
<b>Blackout effect model</b>	<b>2173.31</b>	<b>1</b>	<b>-1063.02</b>	<b>23</b>
Season length model	2189.84	0	-1044.11	48
Species availability model	2193.47	0	-1048.16	46

Table 4: Coefficients of blackout effect model fit to charter boat fishing effort time series.

Coefficients of the autoregressive, moving average, and seasonal component can be found in Table S8. Species-specific regulations and catch rates are indicated by the following abbreviations: black sea bass (BSB), summer flounder (SMF), tautog (TOG), and scup (SCP). Bolded values are significant at the  $p < 0.05$  level. DF=870

Coefficient	Estimate	Standard error	T value	P value
<b>Intercept</b>	<b>2.203</b>	<b>0.760</b>	<b>2.897</b>	<b>0.004</b>
BSB PL	0.004	0.005	0.733	0.464
SMF PL	0.018	0.022	0.852	0.394
TOG PL	0.032	0.025	1.307	0.192
SCP Open	0.229	0.152	1.507	0.132
<b>Something open</b>	<b>1.954</b>	<b>0.740</b>	<b>2.641</b>	<b>0.008</b>
N blackout days	0.024	0.017	1.465	0.143
Something open * N blackout days	-0.019	0.017	-1.170	0.242
SMF CPUE	0.010	0.029	0.350	0.727
BSB CPUE	0.007	0.006	1.069	0.285
TOG CPUE	-0.017	0.019	-0.902	0.367



Table 5: Model fit and Akaike weights of all candidate models for the time series of party boat fishing effort. Bolded values indicate the lowest AICc and highest weight.

Model	AICc	AICc weight	Log Likelihood	# Parameters
Null model	1517.43	0	-726.55	31
Blackout effect model	1502.68	0	-715.94	34
Season length model	1510.27	0	-710.29	39
<b>Species availability model</b>	<b>1482.59</b>	<b>1</b>	<b>-704.82</b>	<b>35</b>

Table 6: Coefficients of species availability model fit to party boat fishing effort time series. Coefficients of the autoregressive, moving average, and seasonal component can be found in Table S10. Species-specific regulations and catch rates are indicated by the following abbreviations: black sea bass (BSB), summer flounder (SMF), tautog (TOG), and scup (SCP). Bolded values are significant at the  $p < 0.05$  level. DF=855

Coefficient	Estimate	Standard error	T value	P value
<b>Intercept</b>	<b>5.745</b>	<b>0.168</b>	<b>34.228</b>	<b>&lt;0.0001</b>
<b>BSB PL</b>	<b>0.050</b>	<b>0.010</b>	<b>5.050</b>	<b>&lt;0.0001</b>
<b>SMF PL</b>	<b>0.227</b>	<b>0.045</b>	<b>5.089</b>	<b>&lt;0.0001</b>
<b>TOG PL</b>	<b>0.141</b>	<b>0.033</b>	<b>4.283</b>	<b>&lt;0.0001</b>
<b>SCP Open</b>	<b>-0.994</b>	<b>0.268</b>	<b>-3.712</b>	<b>0.0002</b>
<b>BSB PL x N species available</b>	<b>-0.014</b>	<b>0.004</b>	<b>-3.880</b>	<b>0.0001</b>
<b>SMF PL x N species available</b>	<b>-0.047</b>	<b>0.014</b>	<b>-3.474</b>	<b>0.001</b>
<b>TOG PL x N species available</b>	<b>-0.034</b>	<b>0.013</b>	<b>-2.506</b>	<b>0.012</b>
<b>SCP Open x N species available</b>	<b>0.495</b>	<b>0.095</b>	<b>5.185</b>	<b>&lt;0.0001</b>
SMF CPUE	0.005	0.023	0.219	0.827
BSB CPUE	0.004	0.004	0.923	0.356
<b>TOG CPUE</b>	<b>0.030</b>	<b>0.012</b>	<b>2.444</b>	<b>0.015</b>

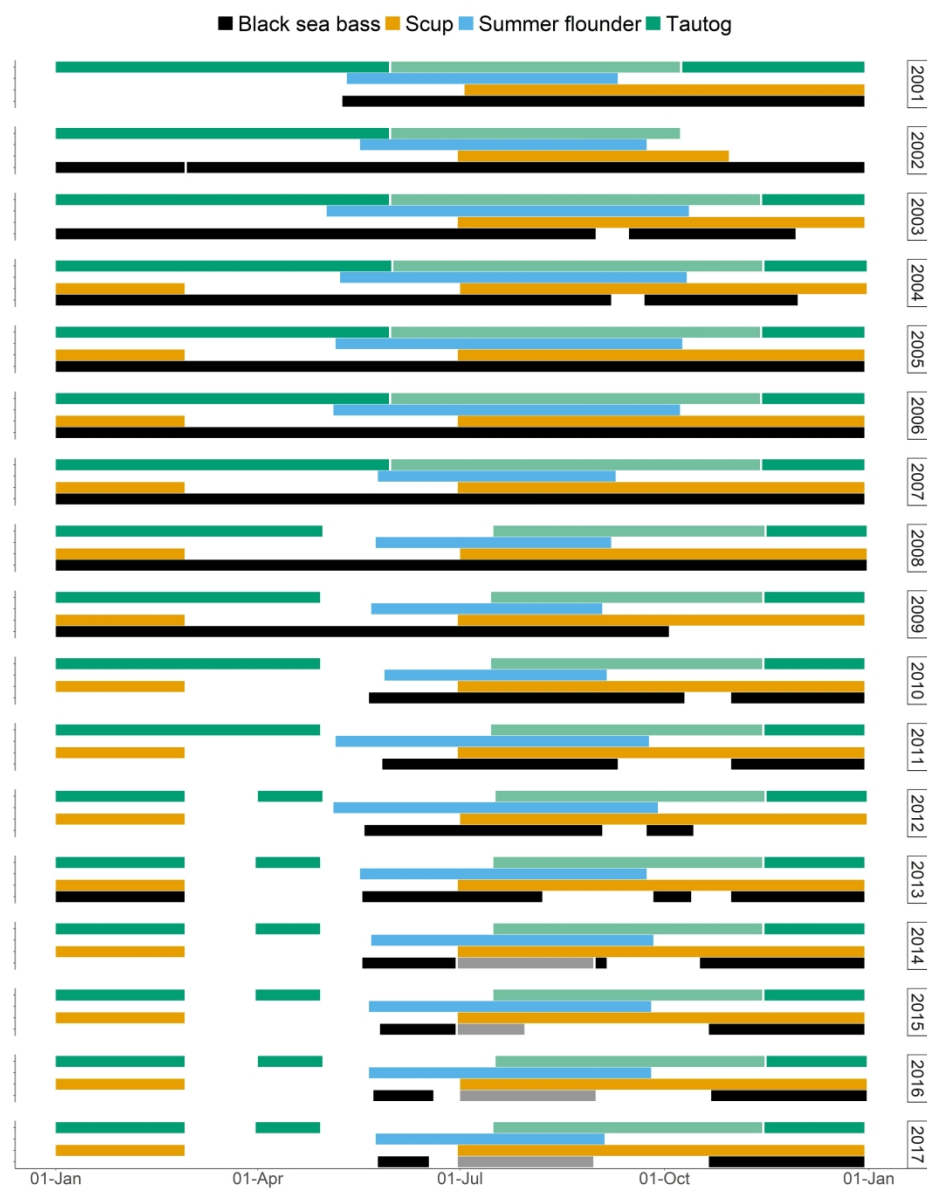


Figure 1: Changes in New Jersey season length and overlap for tautog (top, green), summer flounder (blue), scup (yellow), and black sea bass (bottom, black) between 2001 and 2017. Colored bars delineate open seasons for each of the four species. Light green bars for tautog indicate 1-fish possession limits during the summer and fall months. Gray bars starting in 2014 illustrate black sea bass summer seasons with 2 or 3 fish possession limits

1058x1375mm (72 x 72 DPI)

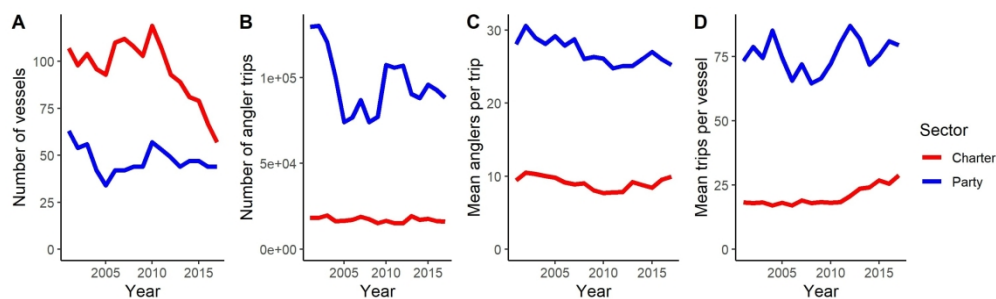


Figure 2: Changes in the annual number of vessels reporting from the for-hire bottom-fishing fleet (A), the number of angler trips reported (B), the mean number of anglers per trip reported (C), and the mean number of trips per vessel reported (D) between 2001 and 2017 in the NJ charter and party boat fleets.

1058x317mm (72 x 72 DPI)

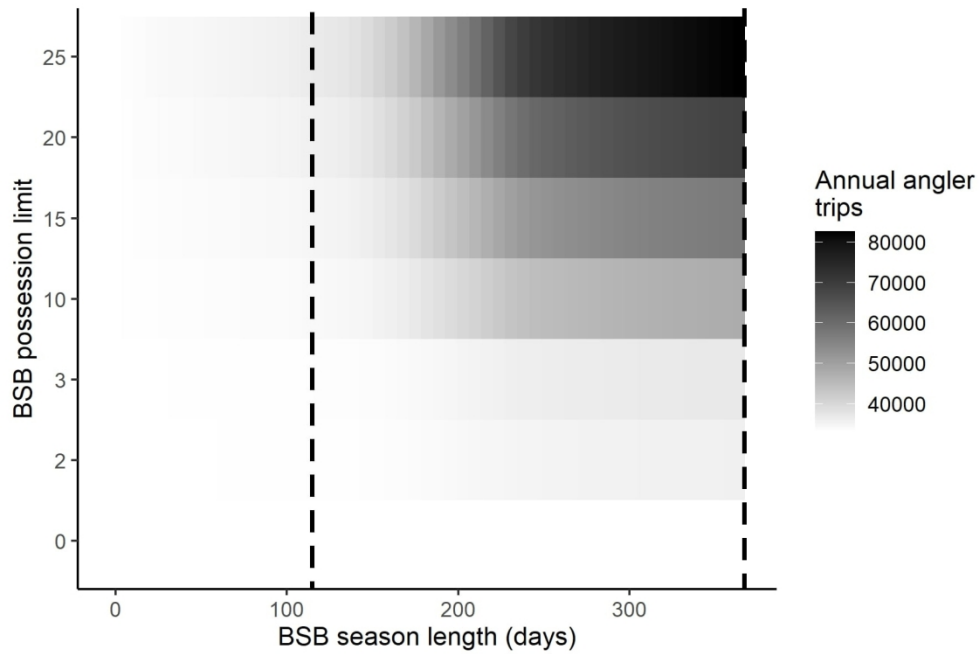


Figure 3: Annual party boat angler trips predicted across a range of season lengths and possession limits for black sea bass. In these model predictions forecasting effort from hypothetical regulations, only black sea bass season is open. The area between the two dashed lines indicates season lengths that are represented in the data.

635x423mm (72 x 72 DPI)

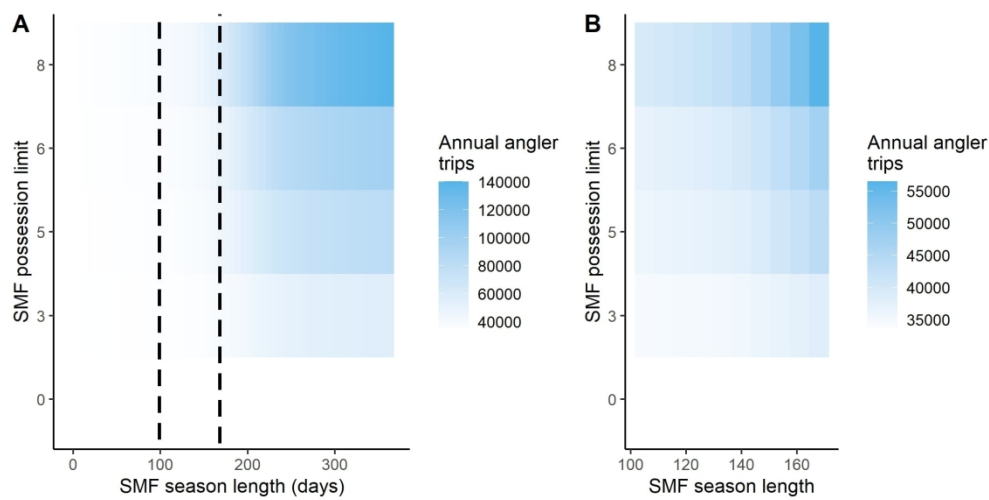


Figure 4: Annual party boat angler trips predicted across a full range of possession limits, hypothetical season lengths (A), and the season lengths represented in the data (B) for summer flounder (i.e. B is a subset of A). The area between the two dashed lines indicates season lengths that are represented in the data. In these model predictions forecasting effort from hypothetical regulations, only summer flounder season is open.

846x423mm (72 x 72 DPI)

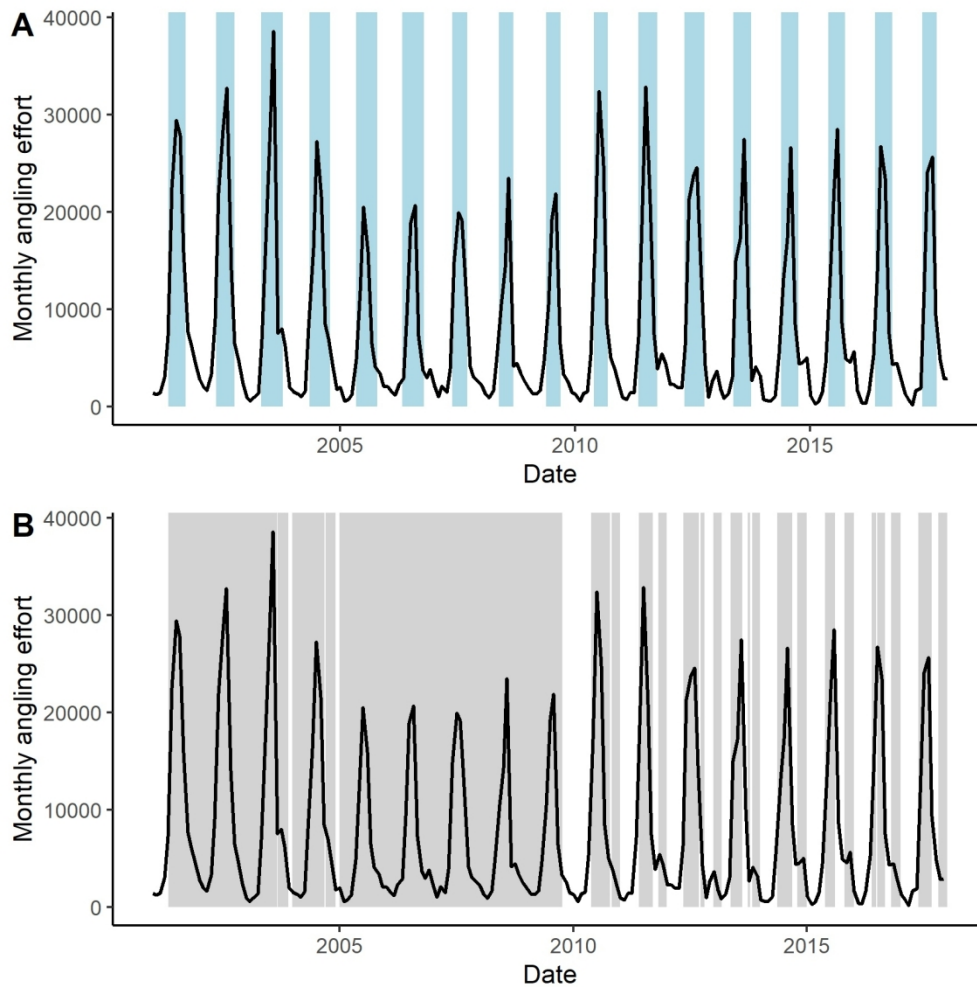


Figure 5: Monthly fishing effort in the party boat sector of the NJ for-hire recreational fishery. Summer flounder open seasons are highlighted in blue on plot A, and black sea bass seasons are in gray on plot B.

152x152mm (300 x 300 DPI)

## Appendix 1: Candidate models

Candidate models take the following form for the best fit number of sine-cosine pairs,  $k$ , autoregressive coefficients  $p$ , and moving average coefficients  $q$  at time  $t$ . Sine-cosine pairs are fit to observations at time  $t$  through the  $\omega$  coefficients. Error terms are indicated by  $\varepsilon$ . Catch per unit effort (CPUE) and regulation covariates are included for black sea bass (BSB), summer flounder (SMF), tautog (TOG), and scup (SCP).

### Null model

$$\begin{aligned}
 & \ln(1 + N \text{ angler trips}_t) \\
 &= \sum_{k=1}^K [\alpha_{1,k} \cos(\omega_{1k}t) + \alpha_{2,k} \sin(\omega_{2k}t)] \\
 &+ \sum_{p=1}^P \phi_p \ln(N \text{ anglers})_{t-p} \\
 &+ \sum_{q=1}^Q \theta_q \varepsilon_{t-q} \\
 &+ \beta_0 + \beta_1 \text{ Possession limit SMF}_t + \beta_2 \text{ Possession limit BSB}_t \\
 &+ \beta_3 \text{ TOG Possession limit}_t + \beta_4 \text{ Open season SCP}_t + \beta_5 \text{ CPUE SMF}_t \\
 &+ \beta_6 \text{ CPUE BSB}_t + \beta_7 \text{ CPUE TOG}_t + \varepsilon_t \\
 &\varepsilon_t \sim N(0, \sigma^2)
 \end{aligned}$$



## Blackout effect model

$$\ln(1 + N \text{ angler trips}_t)$$

$$\begin{aligned}
&= \sum_{k=1}^K [\alpha_{1,k} \cos(\omega_k t) + \alpha_{2,k} \sin(\omega_k t)] \\
&+ \sum_{p=1}^P \phi_p \ln(N \text{ anglers})_{t-p} \\
&+ \sum_{q=1}^Q \theta_q \varepsilon_{t-q} \\
&+ \beta_0 + \beta_1 \text{ Possession limit SMF}_t + \beta_2 \text{ Possession limit BSB}_t \\
&+ \beta_3 \text{ Possession limit TOG}_t + \beta_4 \text{ Open season SCP}_t + \beta_5 \text{ CPUE SMF}_t \\
&+ \beta_6 \text{ CPUE BSB}_t + \beta_7 \text{ CPUE TOG}_t + \beta_8 \text{ Something open}_t \\
&+ \beta_9 \text{ N blackout days}_t + \beta_{10} \text{ Something open}_t * \text{ N blackout days}_t + \varepsilon_t
\end{aligned}$$

$$\varepsilon_t \sim N(0, \sigma^2)$$

## Season length model

$$\ln(1 + N \text{ angler trips}_t)$$

$$= \sum_{k=1}^K [\alpha_{1,k} \cos(\omega_k t) + \alpha_{2,k} \sin(\omega_k t)]$$

$$+ \sum_{p=1}^P \phi_p \ln(N \text{ anglers})_{t-p}$$

$$+ \sum_{q=1}^Q \theta_q \varepsilon_{t-q}$$

$$+ \beta_0 + \beta_1 \text{ Possession limit SMF}_t + \beta_2 \text{ Possession limit BSB}_t$$

$$+ \beta_3 \text{ Possession limit TOG}_t + \beta_4 \text{ Open season SCP}_t + \beta_5 \text{ CPUE SMF}_t$$

$$+ \beta_6 \text{ CPUE BSB}_t + \beta_7 \text{ CPUE TOG}_t + \beta_8 \text{ Season length SMF}_t$$

$$+ \beta_9 \text{ Season length BSB}_t + \beta_{10} \text{ Season length TOG}_t$$

$$+ \beta_{11} \text{ Season length SCP}_t + \beta_{12} \text{ Open season SMF}_t * \text{ Season length SMF}_t$$

$$+ \beta_{13} \text{ Open season BSB}_t * \text{ Season length BSB}_t$$

$$+ \beta_{14} \text{ Open season TOG}_t * \text{ Season length TOG}_t$$

$$+ \beta_{15} \text{ Open season SCP}_t * \text{ Season length SCP}_t + \varepsilon_t$$

$$\varepsilon_t \sim N(0, \sigma^2)$$

## Species availability model

$$\ln(1 + N \text{ angler trips}_t)$$

$$\begin{aligned}
&= \sum_{k=1}^K [\alpha_{1,k} \cos(\omega_k t) + \alpha_{2,k} \sin(\omega_k t)] \\
&+ \sum_{p=1}^P \phi_p \ln(N \text{ anglers})_{t-p} \\
&+ \sum_{q=1}^Q \theta_q \varepsilon_{t-q} \\
&+ \beta_0 + \beta_1 \text{ Possession limit SMF}_t + \beta_2 \text{ Possession limit BSB}_t \\
&+ \beta_3 \text{ Possession limit TOG}_t + \beta_4 \text{ Open season SCP}_t + \beta_5 \text{ CPUE SMF}_t \\
&+ \beta_6 \text{ CPUE BSB}_t + \beta_7 \text{ CPUE TOG}_t \\
&+ \beta_8 \text{ Possession limit SMF}_t * N \text{ species open}_t \\
&+ \beta_9 \text{ Possession limit BSB}_t * N \text{ species open}_t \\
&+ \beta_{10} \text{ Possession limit TOG}_t * N \text{ species open}_t \\
&+ \beta_{11} \text{ Open season SCP}_t * N \text{ species open}_t + \varepsilon_t
\end{aligned}$$

$$\varepsilon_t \sim N(0, \sigma^2)$$