

1 **Tori lines mitigate seabird bycatch in a pelagic longline fishery**

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17 **Abstract**

18 Albatross bycatch has been increasing over the past decade in the US central North Pacific tuna longline
19 fishery. A controlled field experiment was used to assess the efficacy of bird scaring or tori lines as a
20 seabird bycatch mitigation measure for this fishery in a 3-factor sampling design with other mitigation
21 methods (blue-dyed bait, offal discharge). A multilevel geospatial Bayesian regression modeling
22 approach was used to assess 3 albatross-gear interaction metrics (attempted contacts, contacts,
23 captures) recorded for each longline set using an electronic monitoring system. We found albatross
24 contacts with baited hooks were ca. 3 times (95% highest posterior density interval [HDI]: 1-7) less likely
25 for sets equipped with tori lines rather than without tori lines. Attempts to contact baited hooks were ca. 2
26 times (95% HDI: 1-4) less likely for tori line-equipped sets. Albatrosses were also less likely to be
27 captured in tori line sets but captures were too few to support strong inference compared with the contact
28 rates. Tori lines were therefore found to be an effective management measure to mitigate albatross
29 interactions with this fishery. Offal discharge during setting, however, was associated with higher seabird
30 interactions — but that inference was not strong since offal discharge and blue-dyed bait were
31 confounded treatments in some sets. Nonetheless, it was apparent that neither offal discharge nor blue-
32 dyed bait was helpful in reducing albatross interactions in this trial and so the efficacy of those measures
33 warrants further experimental investigation.

34
35 **Key words:** Albatross; Bycatch; Electronic monitoring; Longline fisheries; Tori line; Bayesian modelling;
36 Gaussian process

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44 competing interests that influenced the work reported in this paper.

45 **Availability of data:** The fisheries electronic monitoring data used in this study are owned by and are
46 available from the U.S. government agency NOAA Fisheries and restrictions apply to their availability.
47 Under the terms of a nondisclosure agreement with U.S. NOAA that the authors who analyzed the data
48 had to execute, and under sections 1905 and 201-209 of Title 18 of the United States Code (referred to
49 as the Trade Secrets Laws and Conflict of Interests Laws, respectively), the authors are prevented from
50 making the U.S. government data publicly available.

51 **Software availability:** All statistical modeling software used in this study are cited in the Methods section.

54 **1. INTRODUCTION**

55 Incidental capture or bycatch in commercial fisheries is considered a major threat to the global
56 conservation of pelagic seabirds (Anderson et al. 2011; Gray and Kennelly 2018). This is particularly the

57 case for albatrosses and petrels, which are two of the three most threatened groups of seabirds (Dias et
58 al. 2019). Of 29 albatross and large petrel species listed under the Agreement on the Conservation of
59 Albatrosses and Petrels (ACAP), 19 are categorized as threatened by the IUCN (Phillips et al. 2016;
60 IUCN 2021). It has been estimated recently that at least 160,000 seabirds are killed annually in pelagic
61 longline fisheries, of which a large proportion are albatrosses and the large petrel species (Anderson et
62 al. 2011). Hence a range of gear technologies have been proposed to help mitigate seabird bycatch in
63 pelagic longline fisheries such as night setting, branchline weighting, bird curtains, side setting, blue-dyed
64 bait, strategic offal discharge, underwater setting devices, hook shielding devices, and bird-scaring or tori
65 lines (Gilman et al. 2005; ACAP 2019).

66 A tori line, initially developed by Japanese tuna longline fishers (Brothers et al. 1999), is a line
67 with streamers that is towed from the stern of the vessel as crew set baited hooks. The forward
68 movement of the fishing vessel creates drag on the streamer line, so that a section of the line is in the air
69 above the sea surface. This aerial portion of the streamer line, as the name suggests, can have
70 streamers attached at various intervals to contribute to discouraging seabirds from attacking baited
71 hooks.

72 The five tuna regional fisheries management organizations (RFMOs) have measures in place that
73 either require or include tori lines as a seabird bycatch mitigation option in designated areas (ICCAT
74 2011; IATTC 2012; IOTC 2012; WCPFC 2018; CCSBT 2020). Tori lines are also prescribed for use in
75 combination with other measures by ACAP (2019). Some nations such as Japan, Uruguay, New Zealand
76 and South Africa, require tori line use in pelagic longline fisheries (Japan Ministry of Agriculture, Forestry
77 and Fisheries 2008; Uruguay Direccion Nacional de Recursos Acuaticos 2015; South Africa Department
78 of Agriculture, Forestry and Fisheries 2019; New Zealand Ministry for Primary Industries 2020).

79 Seabird bycatch of the US central North Pacific tuna longline fishery comprised mainly Laysan
80 (*Phoebastria immutabilis*) and black-footed albatrosses (*P. nigripes*), which are categorized as Near
81 Threatened with stable and increasing population trends, respectively (IUCN 2021). Albatross bycatch
82 has been increasing over the past decade in this fishery (Gilman et al. 2016; WPRFMC 2020). This trend
83 is attributable to increasing fishing effort and an increase in the number of albatrosses attending the
84 vessels. The latter is possibly in response to variability in ocean productivity linked to inter-annual and
85 decadal climate cycles, and to the increasing population trend of black-footed albatrosses (Gilman et al.
86 2016; IUCN 2021). Currently, a range of seabird bycatch mitigation measures, including the use of blue-
87 dyed fish bait and strategic offal and bait discharge, are regulatory requirements in this fishery (Gilman et
88 al. 2016).

89 Management authorities prioritized investigating whether tori lines would reduce seabird bycatch
90 in this fishery, but the efficacy of this measure has not been rigorously evaluated for this high-value deep-
91 set tuna fishery (WPRFMC 2020). Although tori lines are a commonly promoted seabird bycatch
92 mitigation measure in pelagic longline fisheries around the world, there have been surprisingly few
93 rigorous experimental assessments (Yokota et al. 2011; Melvin et al. 2013; Sato et al. 2016; Domingo et
94 al. 2017; Jimenez et al. 2020). So, an experimental field trial using commercial pelagic longline vessels
95 and a video-based electronic monitoring system was undertaken to evaluate whether tori lines might be
96 an effective seabird mitigation measure for the US central North Pacific tuna longline fishery.

97 The findings from this study (1) contribute directly to the adoption of proven mitigation measures
98 for the US central North Pacific pelagic longline fishery to support compliance with the US seabird
99 bycatch regulatory requirements, and (2) demonstrate the capability of electronic monitoring systems for
100 fisheries bycatch risk assessment.

101 102 103 **2. METHODS**

104 105 **2.1. Experimental Design**

106 The experimental field trial comprised 175 sets deployed during 17 trips from 4 commercial longline
107 vessels undertaken between February and July 2020. Fig. S1 shows the central North Pacific study area
108 overlaid on the estimated non-breeding range of the Laysan and black-footed albatrosses. Branchline
109 weighting designs (45 g lead-centered swivels attached about 0.6 m from the hook), leader material
110 (wire), bait (saury, *Cololabis saira*), hook type (15/0 circle hooks), time-of-day of setting (during the
111 daytime) and the deck location of setting (from the conventional position at the stern, and not from the
112 side) were the same for all sets.

113 Here the longline set is the fundamental sampling unit or blocking factor (Bergh et al. 1990;
114 Jensen et al. 2018) that is nested within trip, which is itself nested within vessel. So, the sampling design
115 comprises three crossed random-effects: set, trip and vessel. This multilevel or hierarchical random-
116 effects structure needs to be accounted for in any statistical modeling of the estimated tori line effect on
117 albatross bycatch rates. The trial also comprised two other co-applied seabird bycatch mitigation
118 measures or treatments: (1) strategic discharge of offal and spent bait, and (2) blue-dyed fish bait (Gilman
119 et al., 2016). The treatment used for the first set of a trip was randomly selected. The trial therefore
120 comprised three treatments: (1) tori line, (2) offal/spent bait discharge and (3) blue-dyed bait. The 3-
121 treatment experimental design is summarized in Fig. 1, which shows the $2^3 = 8$ possible treatment arms if
122 this was a fully factorial study design.

123 Two of the eight arms were not possible here because, due to regulatory requirements, offal only
124 occurred in conjunction with the blue-dyed bait treatment (Fig. 1), resulting in a partially nested or partially
125 clustered experimental design (Baldwin et al., 2011; Candlish et al., 2018). The partially nested design
126 adds a complication that needs to be accounted for in any subsequent statistical modeling of the
127 estimated tori line effect on seabird bycatch rates. This was done here by creating a 3-category factor
128 cluster as follows: 1 = [offal=="no" & blue=="no"], 2 = [offal=="no" & blue=="yes"], 3 = [offal=="yes" &
129 blue=="yes"]. There was no level that included offal/bait discharge but not blue-dyed bait. The data
130 derived from this experiment were captured using the video-based electronic monitoring system.

131 132 **2.2. Tori Line Deployment and Design**

133 Figs. 2 and 3 show the tori line design used in the experiment. A 50 m-long, red, 3 mm diameter, 12-
134 strand, single-braid Dyneema (AmSteel AS-78 Dyneema, thermoplastic polyethylene) rope was used for
135 the aerial section of the tori line. Two 50 cm-long streamers were attached every 1 m along the aerial
136 section, with the first streamer attached at 2.5 m from the point of attachment to the tori pole on the
137 vessel deck, and the last streamer attached at 30.5 m, for a total of 28 streamers. The aerial section was
138 attached 5 m above the sea surface on schedule 40, marine grade, stainless steel poles. Streamers were
139 50 cm-long, 5 cm-wide, 6 mm thick, black polyethylene sheeting. One 100 cm-long strip was spliced
140 through the Dyneema rope to create the two 50 cm-long streamers (Fig. 3). The drag section of the tori
141 line was 55 m-long, and made of 6 mm diameter, 'Blue Steel' 12-strand polyolefin fiber rope. A weak link,
142 using monofilament nylon (polyamide), was in between the aerial section of the tori line and the tori pole
143 (see Pierre et al. 2016; Goad 2017; Goad and Debski 2017). Safety lines were incorporated to retain the
144 tori line if the weak link broke. When a vessel sets at 11.1 km/hr (6 knots), this would produce ca. 6.2 kg
145 of drag, resulting in a 50 m-long aerial section. The aerial and drag sections were connected by splicing
146 them onto a 6 mm stainless steel swivel, which was covered first with heat shrink wrap and then with
147 electrical tape.

148 This tori line design was selected to meet the minimum tori line design requirements of the two
149 Pacific tuna RFMOs in whose convention areas the US central North Pacific tuna longline fishery occurs
150 (IATTC 2012; WCPFC 2018). Materials were selected based on preferences expressed by Hawaii
151 longline fishers, local availability, cost and lessons learnt from tori line trials in New Zealand and Japan
152 (Goad and Debski 2017; Katsumata et al. 2019). The selected tori line design was also based on the
153 estimated distance astern that Laysan and black-footed albatrosses are able to access baited hooks.

154 155 **2.3. Electronic Monitoring System**

156 The electronic monitoring (EM) system included an 8-megapixel super low lux camera (GeoVision model
157 GV-ABL8712) positioned to provide a field of view off the vessel stern. The camera was set to record at
158 20 frames per second and imagery at 1440p resolution. Cameras were mounted inside of a setting shack
159 from the roof or otherwise from another structure at the stern if available. The low lux sensors optimize
160 the cameras' recording in low light conditions. The cameras had waterproof housings and a marine
161 sealant was used to augment waterproofing.

162 Using that video-based system, the EM analyst collected the data fields and applied the data
163 collection methods summarized in Table S1. The data fields are organized by categories of: seabird
164 bycatch mitigation method, other gear and other fishing methods that potentially have significant effects
165 on seabird catch, response variables of seabird attempts and contacts during the set, seabird scan
166 counts, and environmental variables that potentially have significant effects on seabird catch. Covariates
167 or predictors that explain seabird catch and survival risk were adapted from Gilman et al. (2014, 2016).
168 The protocol for recording seabird attempts and contacts with baited hooks was adapted from Boggs

169 (2001) and Gilman et al. (2003), and these data fields are defined in Table S1. The open-source EM
170 reviewing software *Review* (Chordata 2019) was used to process the EM data.

171 172 **2.4. EM-Derived Data**

173 Nearly all recorded seabird interactions were for either the black-footed albatross or the Laysan albatross
174 exposed to this pelagic longline fishery — shearwaters accounted for ca. 1% of attempts to contact baited
175 hooks. The two albatross species records were combined into a single albatross category as there were
176 insufficient data to estimate species-specific effects.

177 The 2 albatross-gear response metrics were: (1) the recorded number of albatross attempts to
178 attack the gear/bait for each set, and (2) the recorded number of albatross contacts with the gear/bait for
179 each set. There were too few attempts or contacts > 2 or 3 to model meaningfully so these response
180 metrics were more appropriately restructured as a binary or Bernoulli response (0,1) variable with the
181 attempt rate being recoded as either 0 for no attempts and 1 for one or more attempts. The same
182 procedure was applied to the albatross contacts data.

183 There were only 5 Laysan and 5 black-footed albatross captures, based on the number observed
184 retrieved during the gear haulback, which were too few to warrant a more comprehensive statistical
185 analysis other than a modelled statistical summary. So, for this metric the median posterior albatross
186 capture rate and highest posterior density interval (HDI) were summarised by sampling from a binomial
187 likelihood with a Bayes-Laplace prior (Tuyl et al. 2008) using the `binom` R package (Dorai-Raj 2014) —
188 rather than merely using the raw or naïve capture summaries.

189 190 **2.5. Statistical Modeling Approach**

191 The statistical modeling approach for the two contact rate metrics was based on a Bayesian inference
192 workflow (Gabry et al. 2019) using spatially-explicit generalized additive mixed regression structured
193 models (geoGAMM: Fahrmeir and Lang 2001; Kammann and Wand 2003) with an appropriate response-
194 specific likelihood for the various forms of interaction rate data. The response metrics were binary data
195 (for example: 0 = no attempts, 1 = at least one attempt to attack the gear/bait) and so were sampled from
196 a Bernoulli probability distribution and so appropriately modelled using a regression model with Bernoulli
197 likelihood — which is a special case of a binomial likelihood but now with a single trial (Congdon 2003).

198 Specifically, geoGAMMs with Bernoulli likelihood were fit to the albatross interaction data (attempt
199 = at least 1 albatross attempt to attack the gear, contact = at least 1 albatross contact with the gear) while
200 accounting for potentially informative predictors using the Stan computation engine (Carpenter et al.,
201 2017) via the `brms` interface (Bürkner 2017). All models were implemented using weakly informative
202 regularizing priors (Lemoine, 2019) and so prior predictive graphical summaries were used to assess the
203 adequacy of the priors used (Gabry et al. 2019).

204 The predictors included hooks per set (as a nonlinear effect rather than just as an offset: see
205 Davies and Jonsen 2011 for discussion of nonproportional effort scaling), wind speed, cloudiness
206 (overcast or not), `offal blue-dyed bait` cluster, nonlinear mean density of seabirds attending the vessel
207 and specific set geolocation. The offal-blue-dyed-bait cluster covariate helps account for the partially
208 nested design issues summarized in Fig. 1.

209 Cubic smoothing splines (Wood 2006) were used to account for possible nonlinear functional
210 form of the covariates such as hooks per set (the longline fishing effort metric). The structured spatial
211 effect of the individual set geolocations was estimated in the geoGAMMs using a 2D Gaussian Process
212 structure (Gelfand and Schliep 2016).

213 The random effect structures (intercepts-only) included in the geoGAMMs were the identity of the
214 17 trips and the identity of the 4 vessels to account for any correlated or trip- and/or vessel-specific
215 heterogeneity in the interactions rates not accounted for by the other predictors. The 3-category cluster
216 variable was included as either a fixed effect or a random effect — if as a random effect then this form
217 was used as suggested by Candlish et al. (2018): (0+tori|cluster), where tori indicates whether tori lines
218 were used or not.

219 Model selection was based on leave-one-out cross-validation metrics to estimate any
220 comparative difference in expected predictive accuracy between the various models fitted such as
221 whether to include an explicit spatial effect or not or whether including a vessel-specific random effect
222 was necessary (Vehtari et al. 2017). The weight of evidence in favor of one model over any other
223 candidate models was also assessed using Bayesian stacking, which is the Bayesian analogue of model
224 averaging (Yao et al. 2018).

225 The posterior samples for all models were sourced from 4 chains and 12000 iterations after a
226 warmup of 2000 iterations per chain. Therefore, the posterior for each estimate comprised 10,000
227 samples or draws that were used to derive the uncertainty intervals (HDIs or highest posterior density
228 intervals: Kruschke and Liddell 2018) using the `tidybayes` package for R (Kay 2020a). Convergence
229 diagnostics such as effective posterior sample size and the Gelman-Rubin statistic ($R_{hat} < 1.01$) reflected
230 convergence of all Bayesian models used here (Gelman and Hill, 2007). Further evaluation of the best-fit-
231 model was then assessed using graphical posterior predictive checks (Gabry et al. 2019). All inference
232 was then made using the best-fit model.

233 In any experimental setting it is important to be able conclude that there was an effect, when
234 there really was an effect. This can be done using indices of existence and significance in a Bayesian
235 setting (Makowski et al. 2019). A probability statement about the *existence* of a particular effect and its
236 direction, such as tori line effects, can be determined with those 10,000 draws using the probability of
237 direction metric proposed recently by Makowski et al. (2019) — also known as the maximum probability of
238 an effect.

239 The *significance* (rather than just existence) of any such effect (or parameter estimate) was then
240 assessed using the HDI+ROPE approach (Kruschke and Liddell 2018). The region of practical
241 equivalence (ROPE) has been proposed as a robust procedure to determine the significance of a
242 meaningful effect in a Bayesian setting using the posterior draws from the best-fit model along with the
243 calculated 95% highest posterior density interval of those draws (Kruschke and Liddell 2018). An
244 appropriate ROPE range or “null hypothesis” region for a regression model with Bernoulli likelihood has
245 been defined by Kruschke and Liddell (2018) as [-0.18, 0.18].

246 The decision rule is that if the HDI lies entirely outside the ROPE then reject the “null hypothesis”
247 that samples are the same or equivalent (Kruschke and Liddell 2018). If the HDI is lies entirely within the
248 ROPE then accept the “null”. Otherwise, the decision to reject or accept is “undecided”. This is called the
249 HDI+ROPE decision rule (Kruschke and Liddell 2018). Kelter (2020) suggested recently that using the
250 entire posterior distribution (aka the full or 100% ROPE) could be used for a more robust decision. The
251 existence and significance metrics were derived here using the `BayestestR` package for R (Makowski et
252 al. 2019).

253 Finally, the estimated effects summaries based on the best-fit conditional regression models were
254 then adjusted for variable sample size of the treatments using the marginal means approach (Searle et al.,
255 1980; Lenth, 2016) and implemented using the `emmeans` package for R (Lenth 2020). These marginal
256 mean effects were summarised as the median and the highest posterior density intervals (80%, 95%) using
257 the `tidybayes` package for R (Kay 2020a) and `stat_halfeye()` function from the `ggdist` package for R
258 (Kay 2020b). The posterior ratio (and 95% HDI) based on the 10,000 posterior samples for the 2 densities
259 was then used to assess any apparent difference between the contact rate for sets deployed with or without
260 tori lines. The posterior ratio summary was also included in the summary plot. The `ggplots2` (Wickham
261 2016) and `colorspace` (Zeileis et al. 2020) packages for R were used for the summary graphics while the
262 `patchwork` package for R (Pedersen 2020) was used for the multi-panel arrangements.

263 264 265 **3. RESULTS**

266 Strong inference was possible using the Bayesian structured modeling workflow that comprised (1) prior
267 predictive checks to assess the adequacy of the priors used for (2) a robust statistical model accounting
268 for experimental design constraints and potential predictors of interaction rates other than the tori line
269 treatment effect and then followed by (3) posterior predictive checks of the adequacy of the statistical
270 model fitted to the interaction data. Of the potentially informative predictors for all the geoGAMMs fitted to
271 the albatross interaction data, there was a significant seabird density effect, where albatross interaction
272 rates increased nonlinearly with increasing seabird density, and model selection based on leave-one-out
273 cross-validation (LOOcv) and the Bayesian stacking suggest that the spatial effect was a significant effect
274 and had relevance for any model inference (Fig. S2). On the other hand, LOOcv and Bayesian stacking
275 metrics suggested that inclusion of the vessel-specific random effects was not necessary but that trip-
276 specific random effects were — but there was little difference in model fit using vessel, set and trip as
277 random effects but the best-fit model included trip-specific effects only. So, the best-fit GAMMs selected
278 for inference for either the attempt or contact rates excluded the vessel- and set- specific random-effects
279 and used only trip-specific random-effects. The best-fit geoGAMMs identified by the LOOcv and Bayesian

280 stacking metrics fitted the interaction data well as shown for example by the posterior predictive check
281 tests summarized in Supplemental Material Fig. S3. All inference was based on those best-fit geoGAMMs
282 and in all models the tori line treatment effect conditional on all other predictors was significant
283 statistically.

284 Figs. 4 and 5 show the existence and significance of the modelled conditional tori line effect
285 based on the posterior draws from the best-fit contact rate geoGAMM. Specifically, Fig. 4 shows that the
286 tori line effect had a 0.994 probability of being negative while Fig. 5 shows that the tori line effect can
287 indeed be considered as statistically significant using either the full (100%) HDI+ROPE or the 95%
288 HDI+ROPE metric, while the estimated offal-blue-dyed-bait cluster-specific effects were equivocal
289 although the specific `offal=yes + bluedye=yes` effect had a 0.969 existence probability of being positive
290 and hence associated with a higher contact rate rather than reducing the rate (Fig. 4).

291 For attempts (not shown), the tori line effect was similar to the results for contact rate, and had a
292 0.995 probability of being negative and was statistically significant using either the full (100%) HDI+ROPE
293 or the 95% HDI+ROPE metric. The estimated offal-blue-dyed-bait cluster-specific effect was equivocal for
294 the specific `offal=no + bluedye=yes` effect but significant for the `offal=yes + bluedye=yes` effect that
295 also had a 0.978 existence probability of being positive.

296 Fig. 6 shows the estimated marginal means for the tori line effect for the albatross contact rate
297 sourced from the posterior draws from the best-fit conditional geoGAMM. The top panel shows the
298 estimated marginal means density distribution for the sets with and without tori lines where it is apparent
299 that it was ca. 3 times (95% HDI: 1-7) less likely to have one or more albatross contact when tori lines
300 were deployed. The bottom panel summarizes the same predicted effect as in the top bottom but now the
301 summary is conditioned on the offal-blue-dyed-bait cluster level. Again, contact rate is predicted to be
302 lower on sets with tori lines deployed and that this effect difference increases with 1 or 2 co-applied
303 mitigation measures — contact rates were highest when the sets were deployed with both offal and blue-
304 dyed bait but that co-application of tori lines for those sets moderated that undesirable effect.

305 The estimated marginal means for the tori line effect for the albatross attempt rate sourced from
306 the posterior draws from the best-fit conditional geoGAMM, not shown, was similar to the contact rate
307 results. It was ca. 2 times (95% HDI: 1-4) less likely to have one or more albatross attempt when tori lines
308 were deployed. When conditioned on the offal-blue-dyed-bait cluster level, again, the attempt rate is
309 predicted to be lower on sets with tori lines deployed and this effect difference increases with 1 or 2 co-
310 applied mitigation measures — attempt rates were highest when the sets were deployed with both offal
311 and blue-dyed bait, but that co-application of tori lines for those sets moderated that undesirable effect.
312 Importantly, it was apparent that even deploying sets with blue-dyed bait and not offal discharge also
313 increased the albatross attempt rate that was again moderated by co-application of tori lines.

314 Albatrosses were less likely to be captured when tori-lines were deployed: albatross captures
315 were ca. 1.1 times (95% HDI: 0.3 - 2.8) less likely when using tori lines, but captures were too few to
316 support strong inference compared with the contact rate metrics.

317
318

319 4. DISCUSSION AND CONCLUSIONS

320 In any experimental setting it is important to be able conclude that there was an effect, when there really
321 was an effect. And it is equally as important to be able to conclude that there was no effect, when there
322 was in fact no effect. So, we used a Bayesian structured workflow to support robust statistical inference
323 about the efficacy of tori lines as a seabird bycatch mitigation measure in an experimental setting. We
324 found that tori lines did result in substantial reductions in albatross-gear contact rates in our trial for the
325 US central North Pacific tuna longline fishery. In fact, we can be > 99% sure that this effect occurred in
326 our experimental trial and this is a general finding consistent with previous studies on the application of
327 tori lines in other pelagic longline fisheries (Yokota et al. 2011; Melvin et al. 2013; Sato et al. 2016;
328 Domingo et al. 2017).

329 We also found that discharging offal and spent bait during setting might exacerbate and not
330 mitigate seabird catch risk. This is consistent with findings from a demersal longline fishery study that
331 found that higher quantities of offal discharges correlated with higher white-chinned petrel catch rates
332 (Delord et al. 2005). However, it is unclear whether discharging caused the higher bird interaction rates or
333 vice versa. Crew may have discharged offal or spent bait in response to observing high seabird
334 interactions. Discharging offal from processed catch, spent bait and dead discards away from setting and
335 hauling operations may draw scavenging seabirds' attention away from where baited hooks are available

336 and reduce seabird catch rates during that fishing operation (Cherel et al. 1996; McNamara et al. 1999).
337 However, this might be a short-term effect. Based on research conducted in trawl fisheries, increased
338 time between offal discharge events and retention of offal reduced the number of seabirds attending
339 vessels (Abraham et al. 2009; Pierre et al. 2012). The lower the seabird density attending vessels, the
340 lower the seabird catch risk (Gilman et al. 2005; Abraham et al. 2009). Retention might also reduce
341 competitive seabird scavenging behavior and foraging intensity, reducing capture risk (Delord et al. 2005;
342 Gilman et al. 2016). The Hawaii longline fishery may be unique in being required to 'strategically'
343 discharge offal during setting. The tuna RFMOs include either no discharging or strategic discharging as
344 one option (ICCAT 2011; IATTC 2012; IOTC 2012; WCPFC 2018; CCSBT 2020), while CCAMLR (2018)
345 prohibits offal and discard discharging during setting in demersal longline fisheries, consistent with the
346 recommendations of Agreement on the Conservation of Albatrosses and Petrels (ACAP 2019). It is a
347 research priority, with policy implications both locally and internationally, to determine how alternative
348 offal management practices affect long-term seabird catch rates, including effects on factors that may
349 significantly explain seabird catch risk of long-term seabird density at fishing vessels and seabird
350 scavenging and competitive behavior.

351 The study demonstrated that an EM system can be designed to collect variables that significantly
352 explain seabird catch risk. This included the employment of some of the seabird bycatch mitigation
353 methods that were used during setting in the study (*tori* lines including streamer line position in relation to
354 baited hooks, blue-dyed bait, offal management), but not branchline weight amount or distance from the
355 hook (leader length), consistent with previous assessments of EM system capabilities (Ames et al. 2005;
356 Piasente et al. 2012; Gilman et al. 2020). The EM system was also capable of enumerating seabirds to
357 the species level during scan counts, which was achieved in some previous EM trials (McElderry et al.,
358 2011; Piasente et al. 2012) but not others (McElderry et al. 2004, 2011), and the environmental factors
359 Beaufort wind force scale and cloud cover. All sets were made during the daytime. Nighttime setting
360 might prevent the EM analyst from consistently or accurately collecting some of these variables (Ames et
361 al., 2005; Piasente et al. 2012). During night, seabird scan count estimates might be more accurate when
362 using thermal or infrared night-vision cameras (Gilman et al. 2019). We explored but determined it was
363 not feasible for the EM analyst to accurately estimate the relative hue, value and chroma of bait due to
364 variable lighting conditions. It was also not feasible for the EM analyst to estimate the duration that baits
365 soaked in blue dye prior to setting due to the camera field of view. After the project was finished, the
366 bullet cameras (GV-ABL8712) used to record the stern of the vessel were severely corroded. Dome
367 cameras with similar image capabilities may be more suitable.

368 Our findings on the efficacy of the *tori* line design evaluated for the US central North Pacific tuna
369 longline fishery might not be applicable to fisheries that encounter different seabird species complexes.
370 The ca. 50 m-long streamer line aerial section covered 99.7% of the attempts and contacts by Laysan
371 and black-footed albatrosses. Laysan and black-footed albatrosses have limited diving capacities,
372 typically only making body thrusts to reach prey within about half a meter of the surface (Kezama et al.
373 2019). Furthermore, unlike in other regions, secondary interactions, where relatively small species of
374 deep-diving seabirds access baited hooks at depth and bring the baited hook to the sea surface where
375 larger seabird species have a second opportunity to access the terminal tackle and become captured
376 (Jimenez et al. 2012; Melvin et al. 2014), is understood to not occur in the US central North Pacific tuna
377 longline fisheries (Gilman et al. 2016).

378 During the experiment, there was one incident of the *tori* line's safety line entangling with gear.
379 There were no incidents of *tori* lines breaking. No safety issues were raised by the fishers related to
380 deploying and retrieving the *tori* line. Additional research could identify changes to the *tori* line design that
381 improve its efficacy while remaining practical and safe to use. For instance, alternative colors and
382 materials for the streamers might increase seabird mitigation efficacy (Delord et al. 2005). With over half
383 of observed seabird interactions occurring within 10 m of the vessel stern, improved *tori* line protection of
384 bait hooks in this area should be explored, while also considering practicality and safety of alternative *tori*
385 line designs.

386 Our findings contribute to evidence-informed seabird conservation management policy - finding
387 *tori* lines to be an effective management measure to mitigate albatross interactions with this central North
388 Pacific tuna longline fishery, and in highlighting the need for further experimental investigation of
389 'strategic' offal discharge and blue-dyed bait. More primary, controlled field experiments, as well as
390 syntheses of accumulated research, are needed to disentangle the relative efficacy of the various seabird
391 bycatch mitigation methods provided as options by the tuna RFMOs and recommended by others (ICCAT

392 2011; IATTC 2012; IOTC 2012; WCPFC 2018; ACAP 2019; CCSBT 2020). Quantitative meta-analytic
393 modelling with significant overall expected effects provides the strongest and most generalizable
394 evidence (Sutton et al. 2000; Evans 2003; Chalmers 2007). Unlike in properly designed experimental
395 studies, observational studies do not experimentally manipulate specific variables and control for others.
396 As a result, estimated effects of individual variables from analyses of observer data are always
397 confounded by other variables. However, during commercial operations, fishers' implementation of
398 bycatch mitigation methods can differ substantially from their implementation during research, which can
399 result in large differences in efficacy (Gilman et al. 2005; Cox et al. 2007). Consequently, in addition to
400 experimental studies, properly designed analyses of observer data that standardize effort by explicitly
401 accounting for potentially significant explanatory variables are also needed to determine whether and why
402 application of mitigation methods perform differently during commercial use so that interventions can be
403 pursued to address any deviations in prescribed application. Meta-analytic syntheses of accumulated
404 research estimate an overall or pooled effect, and if effects vary across studies, they can also identify
405 reasons for between-study heterogeneity. Synthesis research also identifies knowledge gaps, and
406 conversely identifies areas where additional studies are not needed, guiding future research (Chalmers et
407 al. 2014). Due to the larger sample size plus the number of independent studies, correctly designed meta-
408 analytic assessments can provide estimates with increased accuracy over estimates from single studies,
409 with increased statistical power to detect a real effect. By synthesizing estimates from a mixture of
410 independent, small and context-specific studies, the overall estimated effect from meta-analyses is
411 generalizable and relevant over diverse settings (Nakagawa et al. 2015).

412
413

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620

621 **FIGURE CAPTIONS**

622
623 Fig. 1. Experimental design to evaluate the effect of tori lines as a seabird bycatch mitigation measure in
624 Hawaii's tuna longline fishery given the co-application of two other mitigation measures (offal discard,
625 blue-dyed bait). Terminal nodes show the number of sets completed for each of the $2^3 = 8$ potential
626 treatment combinations. Nodes with zero sets shows those treatment arms that did not occur in this
627 study, due to regulatory requirements, and hence leads to a partially nested rather than a fully factorial
628 study design. The seabird bycatch mitigation measure of branchline weighting was used in all treatments.
629

630 Fig. 2. Bird-scaring *tori* line design, with 50 m-long aerial section, 55 m-long drag section, tori line attached
631 5 m above the sea surface, with two 50 cm-long streamers attached every 1 m starting at 2.5 m from the
632 point of attachment of the aerial section to a pole located on the vessel deck to 30.5 m along the aerial
633 section.
634

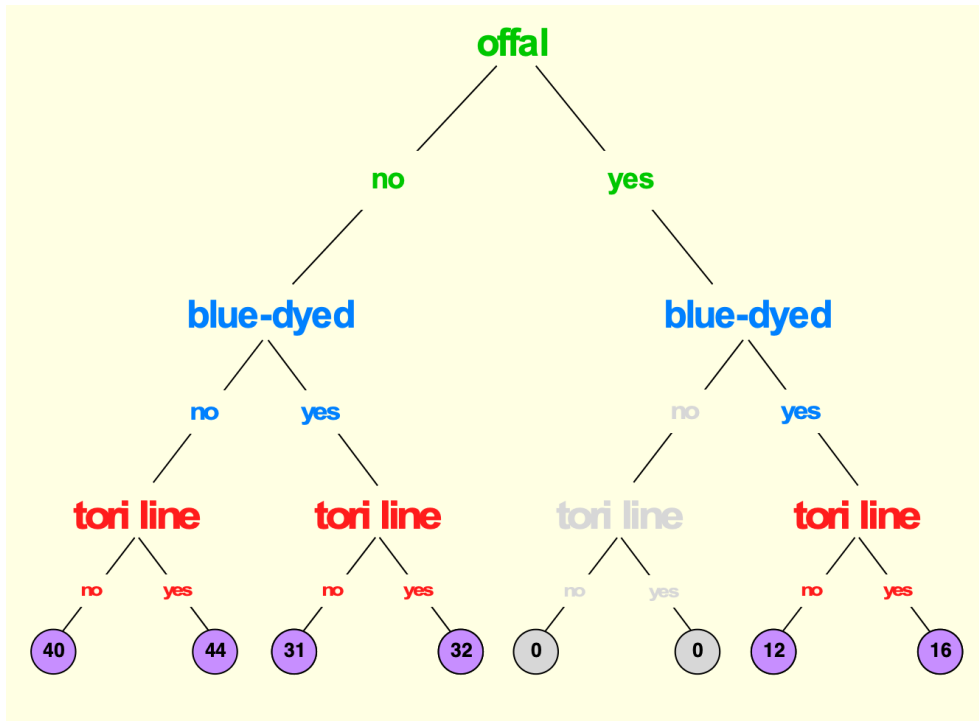
635 Fig. 3. A 100 cm-long plastic strip was spliced through a Dyneema rope to create two 50 cm-long streamers
636 on the aerial section of the *tori* line.
637

638 Fig. 4. Probability of direction plot for selected parameters estimated from the best-fit GAMM for the
639 albatross bait/gear contact rate. Polygons show the density summary of the posterior draws and colored
640 given the estimated direction (positive or negative) of the effect or parameter. The proportion of the
641 polygon that does not include zero is a statement about the probability of the proposed direction of the
642 effect.
643

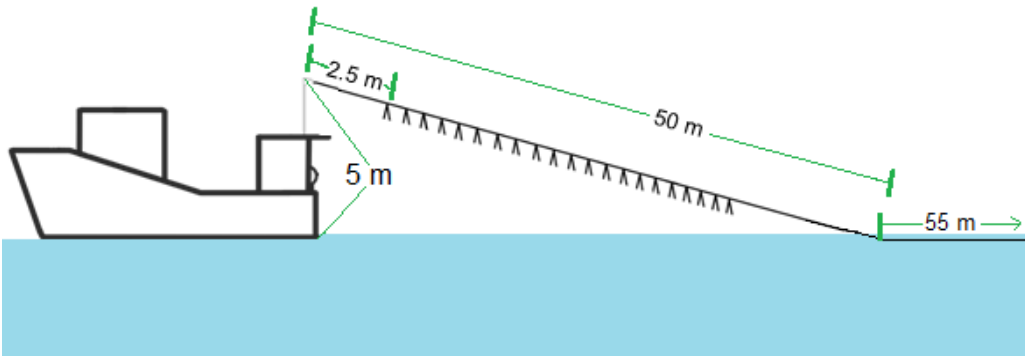
644 Fig. 5. ROPE-based summary of the significance of the tori line and offal-blue dyed bait cluster effects
645 derived from the best-fit GAMM for the albatross bait/gear contact rate. Panel A shows the effects given a
646 full ROPE based on a 100% highest posterior density interval. Panel B shows the effects given a ROPE
647 based on a 95% highest posterior density interval. Green polygon indicates a significant effect.
648

649 Fig. 6. Summary of the estimated marginal mean tori line effect derived from the best-fit GAMM for the
650 albatross bait/gear contact rate. Panel A shows the estimated tori line effect. Colored polygon shows the
651 density distribution summary, solid dot (+ numeric label) = median estimated of the density polygon, thick
652 horizontal line below each polygon shows the 80% highest posterior density interval for the density
653 polygon while the thin horizontal line is the 95% HDI. Panel B shows the estimated tori line effect
654 conditional on offal-blue-dyed-bait treatment cluster — solid dot = estimated median and vertical bar =
655 95% highest posterior density interval.
656

657 FIGURES
658



659
660 Fig. 1
661
662



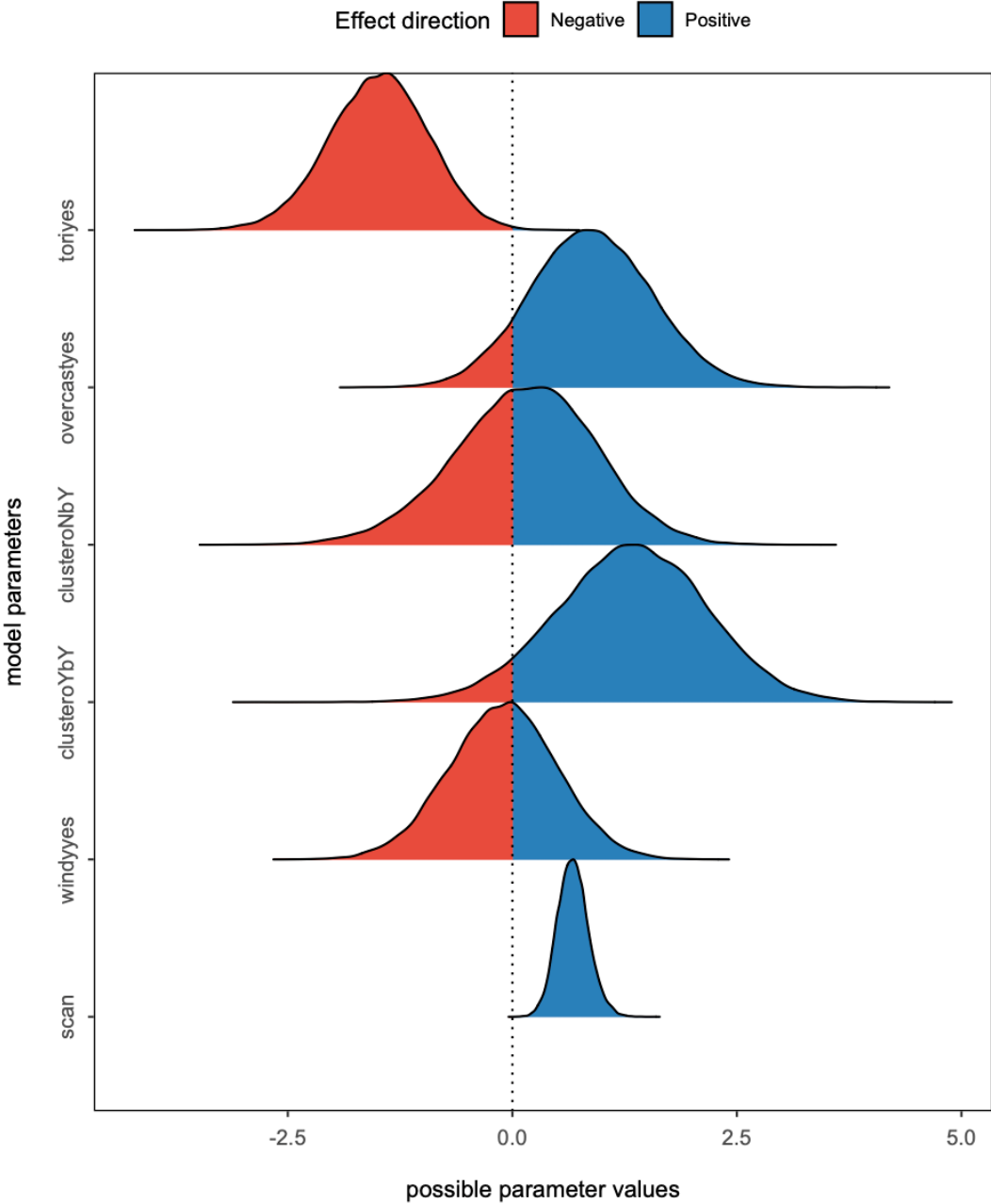
663
664 Fig. 2
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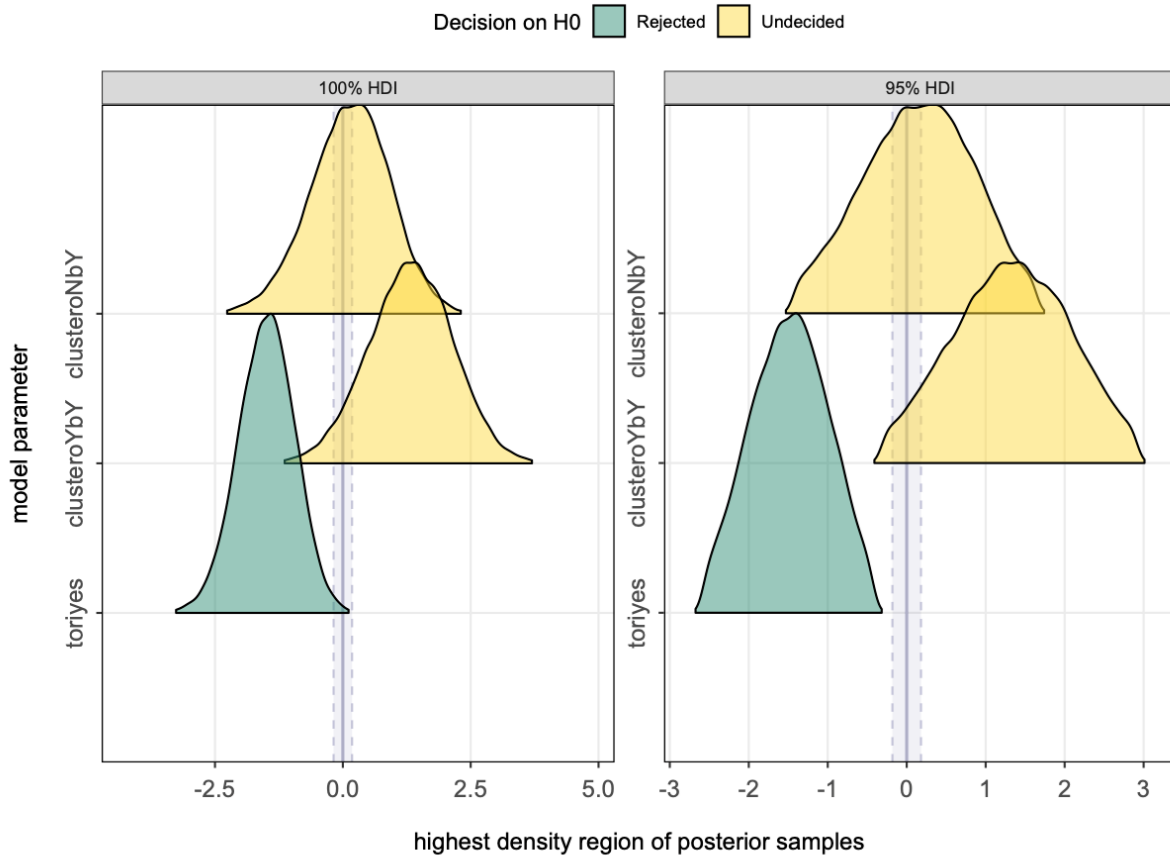
Fig. 3

probability of direction



670 Fig. 4
671
672

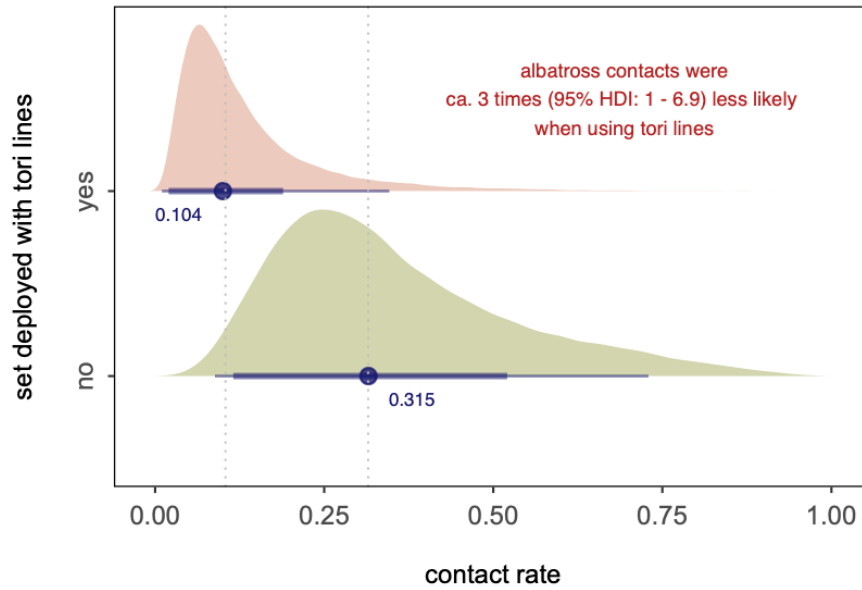
Region of Practical Equivalence (ROPE) parameter plot



673
674 Fig. 5
675

(A)

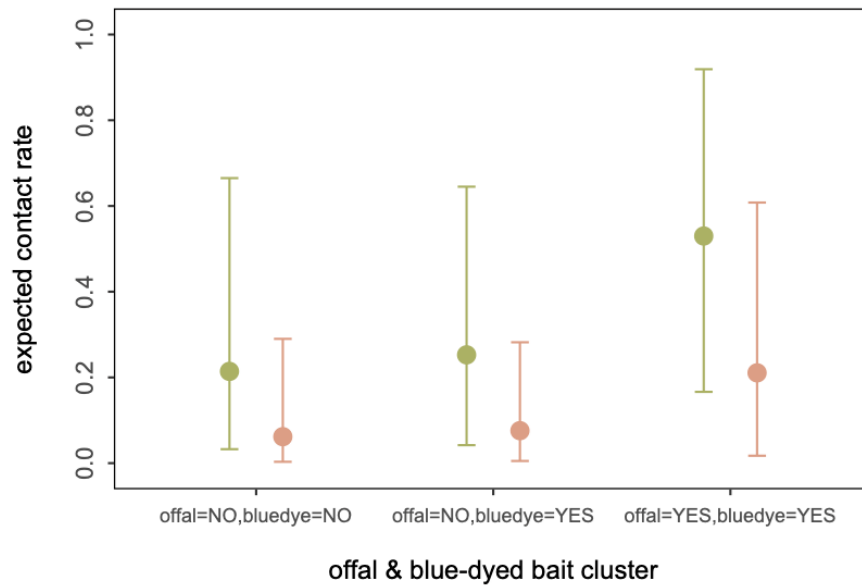
GAMM adjusted marginal tori line effect
density plots (with median and 80% & 95% HDI summaries)



(B)

GAMM adjusted marginal cluster-specific tori line effect
(with median & 95% HDI summaries)

tori lines deployed: ● no ● yes



676 Fig. 6
677