

1 **Competition-driven growth of Atka mackerel in the Aleutian Islands ecosystem revealed by**
2 **an otolith biochronology**

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5 Mary Elizabeth Matta^{1*}, Kimberly M. Rand², Morgan B. Arrington³, Bryan A. Black⁴

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7 ¹Resource Ecology and Fisheries Management Division, Alaska Fisheries Science Center,

8 National Marine Fisheries Service, National Oceanic and Atmospheric Administration, 7600

9 Sand Point Way NE, Seattle, WA 98115, USA

10 ²Lynker Technologies, LLC, 202 Church Street SE #536, Leesburg, VA 20175, USA

11 ³School of Aquatic and Fishery Sciences, University of Washington, 1122 Boat Street, Seattle,

12 WA 98195, USA

13 ⁴Laboratory of Tree-Ring Research, University of Arizona, 1215 E. Lowell Street, Tucson, AZ

14 85721, USA

15

16 *corresponding author

17

18 **Abstract**

19

20 Fish otolith increment biochronologies can reveal factors contributing to growth. These annually
21 resolved, multi-decadal biochronologies have been used to identify temporal and spatial patterns
22 in growth variability. Here, we used a linear mixed modeling approach to determine the relative
23 importance of intrinsic factors (e.g., sex, age) and extrinsic factors (e.g., temperature, abundance
24 of con-specifics and competitors) to growth of Atka mackerel (*Pleurogrammus monopterygius*),
25 a commercially important groundfish dominant in the Aleutian Islands ecosystem. A yearly
26 alternating pattern of wide and narrow increments was observed, and was negatively correlated
27 with abundance of pink salmon (*Oncorhynchus gorbuscha*), suggesting possible competition
28 over shared food resources such as large copepods (which were positively correlated with the
29 otolith biochronology). There was no detectable effect of temperature on otolith growth, which
30 could be a result of relative stability in water temperatures at the study site. We also contrast the
31 otolith biochronology with a body condition index to examine the relationship between otolith
32 and somatic growth. This represents the first otolith biochronology developed in the Aleutian
33 Islands and provides insight into potential species interactions and their impacts on growth
34 within this highly dynamic and productive ecosystem.

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38 **Keywords:** Sclerochronology, Growth, Aleutian Islands, Body condition, Competition, Otolith

39 **1. Introduction**

40

41 Describing growth of individuals is essential for understanding fish population dynamics,
42 particularly as a rising number of studies have recognized the potential for growth to vary over
43 time (Stawitz et al., 2015; Lorenzen, 2016). Variability in growth is driven by a suite of inherent
44 factors such as ontogenetic development and extrinsic factors, including environmental
45 conditions, food availability, and competition for shared resources, all of which can vary across
46 space and time (Walters and Wilderbuer, 2000; Sibley et al., 2015; Morais and Bellwood, 2018).
47 Understanding past relationships between these factors and growth may aid with predicting
48 future effects on growth resulting from changes in the external physical and biological
49 environments to which a group of individuals is exposed. Given the level of climate change
50 projected to occur in many marine environments and subsequent cascading impacts on ecological
51 communities (Wang et al., 2010; Oliver et al., 2019), such information is highly relevant to
52 fisheries managers.

53 An increasing number of studies have used increment widths formed in the hard
54 structures of marine organisms, such as fish otoliths, to develop biochronologies as proxies for
55 somatic growth over time (Black et al., 2008; Morrongiello et al., 2012). Otoliths grow
56 continuously over the life of a fish, and the relative widths of their annual increments are often
57 related to variability in oceanographic and biotic factors (Black et al., 2011; Doubleday et al.,
58 2015; Ong et al., 2018). The dendrochronology (tree-ring science) technique of crossdating
59 ensures that increments have been assigned the correct year of formation and thereby eliminates
60 even modest dating errors that can severely mute environmental signals, especially in long-lived
61 species (Black et al., 2016). Crossdating becomes more challenging in shorter-lived species (< 15

62 yr) due to the lack of overlap among samples, but can still be highly effective when interannual
63 variability in growth is high (Matta et al., 2010). Once increments are dated and measured,
64 mixed effects modeling offers a flexible approach for directly evaluating the magnitude of
65 intrinsic effects, such as sex or fish age, and extrinsic effects, such as temperature or prey
66 abundance, on otolith growth (Helser et al., 2012; Morrongiello and Thresher, 2015). Moreover,
67 otolith biochronologies often predate instrumental records, allowing an extended window into
68 past environmental conditions (van der Sleen et al., 2016). As biochronology techniques are
69 more widely implemented, a global library of datasets is emerging that can inform us of
70 covariability across instrumental records, diverse taxa, and trophic levels that may inform
71 species' responses to projected climate change (Barrow et al., 2018; Black et al., 2019).

72 Particularly well-suited to biochronology studies are species that utilize localized habitats
73 and that possess clear growth increments. Atka mackerel (*Pleurogrammus monopterygius*) is a
74 commercially and ecologically important species endemic to the North Pacific Ocean for which
75 archived otoliths from survey and fishery collections exist (Anderl, 2012). Long-term tagging
76 studies indicate that Atka mackerel have relatively small home ranges, rarely migrating more
77 than 50 km outside their local aggregations (McDermott et al., 2005, 2016). Although Atka
78 mackerel are found in small aggregations elsewhere, the population epicenter remains focused
79 within the U.S. Aleutian Islands (Lauth et al., 2007b), likely due to the region's unusual
80 oceanographic dynamics. The Aleutian Islands, an active volcanic chain, act as a natural barrier
81 that separates the Bering Sea to the north and the Pacific Ocean to the south (Fig. 1). The
82 dominant climatic feature of the Aleutian Islands is the Aleutian Low, a low-pressure system that
83 is the primary driver of temperature variation in the region, varies on interannual to interdecadal
84 timescales, and is teleconnected to the El Niño Southern Oscillation (Rodionov et al., 2005;

85 O'Reilly, 2018). In the central U.S. Aleutian Islands lies Seguam Pass (Fig. 1), the site of one of
86 the most consistently dense aggregations of Atka mackerel, likely due to its strong vertical
87 mixing and swift bottom currents (Rand et al., 2019). Compared to other large marine
88 ecosystems in Alaska such as the Bering Sea and Gulf of Alaska, the Aleutian Islands are not as
89 well-studied, and constructing an otolith biochronology from a key species in this region may
90 allow insight into underlying environmental and biological dynamics.

91 The objective of our study was to identify key drivers of growth of Atka mackerel that
92 involve not only physical variables, but also biological indicators of ecosystem functioning. To
93 accomplish this, we developed an otolith growth time series based on Atka mackerel caught in
94 Seguam Pass, the portion of the Aleutian Islands where this species is most locally abundant. We
95 used a mixed modeling approach to explicitly determine the relative impacts on growth of
96 intrinsic factors such as sex or age, as well as extrinsic factors such as environmental variability,
97 competition, and food resources. We had the unique opportunity to include not only indices of
98 climate but also abundance time series of zooplankton, a major prey group, and pink salmon
99 (*Oncorhynchus gorbuscha*), a potential competitor, to provide an ecosystem-level perspective.
100 Finally, we compared anomalies in otolith growth to independent estimates of body condition to
101 test the otolith biochronology's viability as a proxy for somatic growth, an aspect that has rarely
102 been examined in previous biochronology studies. This Atka mackerel otolith biochronology
103 represents the first such index developed from the Aleutian Islands region, with the overarching
104 goal of providing information regarding relationships among oceanographic and biological
105 variables within this highly dynamic and productive large marine ecosystem.

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107

108 2. Materials and methods

109

110 2.1 Study species, collection area, and sample selection

111 Atka mackerel is a member of the greenling family (Hexagrammidae) and is one of the
112 most abundant groundfish species in the U.S. Aleutian Islands. Atka mackerel form large, dense
113 aggregations often in areas of high current velocity (McDermott et al., 2005; Lauth et al., 2007b)
114 and undergo diel vertical migrations, rising in the water column during daytime hours,
115 presumably to feed on pelagic zooplankton (Nichol and Somerton, 2002). Atka mackerel are
116 considered primarily zooplanktivores (Yang, 1999; Rand et al., 2010); however, fish
117 consumption (mostly smelts) has been shown to increase during the summertime months,
118 especially in the central portion of the Aleutian Islands (Rand et al., 2010). Although Atka
119 mackerel diets can be quite diverse, in general, euphausiids and pelagic calanoid copepods are
120 the dominant components (Yang, 1999; Rand et al., 2010). Atka mackerel are considered
121 important prey to several fish species in the Aleutian Islands, including Pacific cod (*Gadus*
122 *macrocephalus*), Pacific halibut (*Hippoglossus stenolepis*), and arrowtooth flounder (*Atheresthes*
123 *stomias*) (Yang, 1999), as well as a large marine mammal, the Steller sea lion (*Eumetopias*
124 *jubatus*) (Sinclair et al., 2005). Similar to other hexagrammids, their reproductive strategy is
125 somewhat unique in that males guard nests of eggs on the ocean floor for extended periods of
126 time (up to three months) (Lauth et al., 2007b). The spawning season lasts from late June to mid-
127 October, peaking in September (McDermott and Lowe, 1997; Lauth et al., 2007a). Atka
128 mackerel populations reach 50% sexual maturity at 3.6 years (McDermott and Lowe, 1997;
129 Cooper et al., 2010) and have a maximum observed age of 15 years (Anderl, 2012).

130 Seguam Pass, located in the central Aleutian Islands, is relatively shallow, dominated by
131 waters from the Alaskan Stream current and is ecologically important (Ladd et al., 2005; Mordy
132 et al., 2005; Zador and Ortiz, 2018). Within the vicinity of Seguam Pass, one of the largest and
133 most persistent aggregations of Atka mackerel supports a large commercial fishery (Lowe et al.,
134 2017) and is the primary prey for Steller sea lion (Sinclair et al., 2005, 2013). Some unique
135 features of Seguam Pass are a north-south temperature dipole, strong flow dominated by tidal
136 currents, and strong vertical mixing that transports nutrients to the upper water column (Stabeno
137 et al., 2005). The northern portion of Seguam Pass is characterized by deepwater upwelling and
138 is often the site of anomalously high chlorophyll-a measurements (Mordy et al., 2005), which in
139 turn correlate with abundance of oceanic zooplankton such as *Neocalanus* and *Eucalanus*
140 copepods and the euphausiid *Euphausia pacifica* (Coyle, 2005), species common in the Atka
141 mackerel diet (Yang, 1999, 2003).

142 Otoliths were collected by fishery observers on commercial vessels actively targeting
143 Atka mackerel in the vicinity of Seguam Pass (within NMFS Statistical Area 541) in the years
144 1996, 1998, 2000, 2002, 2005, 2008, 2010, and 2015 (Fig. 1). Otoliths were randomly selected
145 for growth-increment analysis from fish aged 5 years or older (von Biela et al., 2015). Only
146 otoliths with sufficiently clear growth patterns were retained to ensure precision of growth-
147 increment measurements. A minimum of 30 otoliths per collection year were included in the
148 growth-increment analysis, with collection years spaced no longer than 5 years apart, to ensure
149 adequate replication and overlap of otolith growth-increment time series.

150

151 2.2 Otolith analysis

152 Otoliths were snapped in half along the transverse plane and gently burned over an
153 alcohol flame to enhance the contrast between opaque and translucent growth zones (Anderl,
154 2012). Otolith annuli were enumerated during routine production age determination according to
155 protocols established by the Alaska Fisheries Science Center (Anderl, 2012). Atka mackerel do
156 not form a visible annulus during the first year of life (Anderl et al., 1996), so one year was
157 added to all age estimates. Otoliths were illuminated with reflected light from a dual gooseneck
158 LED light source and photographed with a digital camera attached to a Leica MZ9.5
159 stereomicroscope. Starting from the otolith edge and working towards the core, each increment
160 was dated with its year of formation based on capture date and age. Using Image-Pro Premier
161 image analysis software (Media Cybernetics), a measurement axis was drawn perpendicular to
162 each growth zone through the ventral side of the sectioned otolith and increments were
163 delineated distal to each translucent growth zone (Fig. 1). The location of the measurement axis
164 was consistent among specimens. The final increment at the otolith edge was not measured if it
165 was still in the process of forming. Because the first visible increment forms during the second
166 year of life, the portion of the otolith prior to the first increment was not measured. Otolith
167 growth increment measurements from individual fish were then plotted separately relative to fish
168 age at the time of increment formation and to calendar year of formation to evaluate patterns in
169 growth with respect to ontogenetic development and to produce raw growth-increment time
170 series (Fig. 2).

171 Individual growth-increment time series were too short to perform true crossdating.
172 Crossdating is a common practice in biochronology studies whereby synchronous patterns of
173 anomalously wide or narrow increments are matched among individuals (Douglass, 1920). This

174 procedure ensures the annual resolution of each measurement time series and that the final
 175 growth-increment dataset is annually resolved and preserves environmental signals. High-
 176 frequency variability is especially vulnerable to loss in the absence of crossdating. While an
 177 alternating pattern of wide and narrow increments was noted in many individuals (Fig. 1), Atka
 178 mackerel are too short-lived to robustly apply these crossdating techniques given that individuals
 179 lacked the overlap to crossmatch among one another. Therefore, some unknown dating error
 180 remains. The biennial pattern of wide and narrow increments observed across many individuals
 181 may have helped reduce some of this dating error, as it was often used to resolve decisions
 182 regarding growth observed at the otolith edge (i.e., whether edge growth should be attributed to
 183 the year of collection or the previous calendar year).

184

185 *2.3 Biochronology development (mixed modeling)*

186 Linear mixed models were used to assess various potential predictors of otolith growth,
 187 similar to methods described in Morrongiello and Thresher (2015) and Matta et al. (2018).
 188 Otolith increment widths are generally an allometric function of age at formation (Helser et al.,
 189 2012), such that otolith growth can be modeled as

190

$$191 \quad W_{ij} = a_j A_{ij}^{b_j},$$

192

193 where W is the width of the i^{th} increment formed at age A of the j^{th} fish. Increment width and
 194 increment age were natural log-transformed to linearize the data and stabilize the variance
 195 structure, such that $y_{ij} = \ln(W_{ij})$, $\alpha_j = \ln(a_j)$, $\beta_j = b_j$, and $x_{ij} = \ln(A_{ij})$. Models with random intercepts
 196 and age slopes for each specimen were considered to allow each fish to have its own growth

197 trajectory. These models were extended to consider potential effects of sex and age at capture
 198 (intrinsic fixed factors), calendar year of formation and cohort (random factors), and finally,
 199 effects of both environmental and biological extrinsic fixed factors (Morrongiello and Thresher,
 200 2015; Matta et al., 2018). The most complex possible model took the form:

201

$$202 \quad y_{ijkl} = (\bar{\alpha} + \alpha_i^S + \alpha_k^Y + \alpha_l^C) + (\bar{\beta} + \beta_{ij}^S + \beta_{jk}^Y + \beta_{jl}^C)x_{ijkl} + f(\cdot) + \varepsilon_{ijkl},$$

203

204 where y_{ijkl} is the width of the annual growth increment relative the age at formation (x_{ijkl}) for fish
 205 i at age j in year k and cohort l , $\bar{\alpha}$ and $\bar{\beta}$ are the fixed mean intercept and slope that describe the
 206 population-wide decline in increment width as a function of age, α_i^S and β_{ij}^S are random effects
 207 describing the growth trajectories of individual specimens with respect to age, α_k^Y and β_{jk}^Y are the
 208 random year intercept and the corresponding slope with respect to age, α_l^C and β_{jl}^C are the
 209 random cohort intercept and the corresponding slope with respect to age, and $f(\cdot)$ represents
 210 additional fixed factors, (e.g., intrinsic factors sex and age at capture, and extrinsic factors
 211 described in detail in Section 2.4). Continuous variables were mean-centered prior to model
 212 fitting to aid convergence. A compound symmetry covariance structure was assumed (i.e., all
 213 variances were considered homogenous). Relationships between extrinsic fixed effects were
 214 examined for collinearity using pairwise plots, correlation coefficients, and variance inflation
 215 factors (Zuur et al., 2009).

216 Models were fitted using the *lme4* package (Bates et al., 2015b) in R statistical computing
 217 software (R Core Team, 2019). First, restricted maximum likelihood (REML) was used to fit
 218 models to identify the optimal structure of the random components (specimen, cohort, and year).
 219 This was accomplished using the maximal intrinsic fixed effect parameterization of age * sex +

220 age at capture (Morrongiello and Thresher, 2015). Models were considered that allowed the
221 random effects to vary with age. Models with various random effects structures were assessed
222 for singularity, an indication of possible over-parameterization, using the `isSingular()` function in
223 R statistical computing software (R Core Team, 2019). We also performed a principal
224 components analysis of the random effects structure (`rePCA` function) to determine the
225 proportion of variance explained by each random effect, eliminating any random slopes that were
226 not supported by the data (Bates et al., 2015a). Next, maximum likelihood (ML) was used to fit
227 models to identify the optimal structure of the intrinsic fixed components. Finally, the model
228 with the optimal random and fixed intrinsic structure was extended to add extrinsic fixed factors
229 (see Section 2.4). The most parsimonious model was refitted to the data using REML to provide
230 unbiased parameter estimates (Zuur et al., 2009). Model selection was based on Akaike's
231 Information Criterion corrected for small sample sizes (AIC_c) (Burnham and Anderson, 2002).
232 Models were ranked by AIC_c differences, or $\Delta_i = AIC_{ci} - AIC_{cmin}$, where AIC_{ci} is the AIC_c of the
233 i th model and AIC_{cmin} is the minimum AIC_c value of the models compared. The model with $\Delta_i =$
234 0 is considered the most parsimonious, and models with $\Delta_i > 10$ have essentially no support
235 (Burnham and Anderson, 2002). We also calculated the Akaike weight (AIC_{wt}), a measure of the
236 relative likelihood (probability ranging from 0 to 1) of the model given the data (Burnham and
237 Anderson, 2002). Finally, for each model we calculated the marginal R^2 (R_m^2) and the conditional
238 R^2 (R_c^2), which explain the proportion of the variance explained by the fixed factors alone and the
239 fixed and random factors combined, respectively (Nakagawa and Schielzeth, 2013).

240 Prior to the addition of extrinsic fixed factors, the random year effects from the most
241 parsimonious model were extracted. The year intercepts encapsulate all random environmental
242 variation after age-related declines in growth and other intrinsic factors have been accounted for,

243 and can be considered analogous to the “master chronology” produced by the regional curve
244 standardization technique used in dendrochronology (Helsler et al., 2012). The random year
245 intercepts are reported here for comparison with previously published biochronologies, and with
246 time series that were too short to explicitly include in our mixed effects models (indices of body
247 condition and zooplankton abundance; see Section 2.4).

248

249 *2.4 Environmental and biological datasets*

250 Extrinsic oceanographic and biological factors hypothesized to have effects on otolith
251 growth were considered in the linear mixed effects models described in Section 2.3. With respect
252 to climate variables, gridded Hadley ISST 1° monthly sea surface temperatures (Rayner, 2003)
253 were obtained from the KNMI Climate Explorer (<http://climexp.knmi.nl>, retrieved 12/3/2018) and
254 averaged over the area 51-53°N, 172-174°W to calculate a regional mean-annual temperature
255 specific to Seguam Pass. Due to strong vertical mixing, sea surface temperature likely represents
256 relative interannual variation in water temperature throughout the water column (Stabeno et al.,
257 2005). Annual temperatures were used recognizing that otolith growth, while peaking in the
258 summer and minimized in the winter, occurs over a large portion of the year and actual timing of
259 deposition may differ at different life stages or on an individual basis (Kimura et al., 2007).
260 However, we noted that annual mean temperatures were well-correlated with seasonal mean
261 temperatures across the years 1980-2017, with Pearson correlation coefficients ranging from
262 0.62 in winter to 0.89 in summer. Therefore, annual mean temperatures likely well represent
263 overall interannual variability in water temperatures at Seguam Pass.

264 We also included an index of pink salmon abundance (Irvine and Ruggerone, 2016). Pink
265 salmon, which due to their semelparous 2-year spawning cycle are typically far more abundant in

266 odd years in the North Pacific (Ruggerone et al., 2010), are also significant predators of
267 zooplankton (Davis et al., 2005; Karpenko et al., 2007). The large Eastern Kamchatka population
268 of pink salmon, which occupies the Central and Eastern Aleutians during its ocean-going phase,
269 is known to induce trophic cascades (Batten et al., 2018) and compete with other species for
270 resources (Ruggerone and Nielsen, 2004; Springer and van Vliet, 2014). Finally, we included
271 estimates of age-3+ biomass (ln-transformed) from the Aleutian Islands Atka mackerel stock
272 assessment (Lowe et al., 2017) to account for any density-dependent, intraspecific competition
273 between adult fish that may affect growth.

274 While we wanted to include additional variables such as prey abundance explicitly in the
275 mixed effects models, those time series were not sufficiently long. Including them would require
276 truncating the growth increment data and result in lost information with respect to other
277 explanatory variables. However, these indices can still be compared against the random year
278 intercepts from the most parsimonious intrinsic model (i.e., the otolith biochronology, see
279 Section 2.3) for informational purposes. Pearson correlation coefficients were calculated between
280 the otolith biochronology and annual means of each of the following biological metrics.

281 Zooplankton comprise a significant portion of the Atka mackerel diet. Therefore we
282 compared indices of copepod and euphausiid abundance with the otolith biochronology.
283 Abundance estimates (number per sample) of copepods and euphausiids were obtained from
284 Continuous Plankton Recorder (CPR) tows (Batten et al., 2003; Batten et al., 2018). Copepods
285 were separated into “large” (> 2 mm) and “small” (< 2 mm) categories with the understanding
286 that the two size groups may occupy different roles in terms of trophic energy transfer (Mazur et
287 al., 2007). The CPR data were collected continuously around the clock from 2000-2014 in the
288 Southern Bering Sea closest in proximity to the study area, the same region in which the Eastern

289 Kamchatka population of pink salmon has been shown to impact plankton communities (Batten
290 et al., 2018). Connectivity between this region and Seguam Pass has been demonstrated by
291 Stabeno et al. (2005); strong tidal currents within and adjacent to Seguam Pass flow north,
292 providing an important source of nutrients to the Southern Bering Sea. Data were averaged
293 across multiple tows during the summer months (June-August) for each year of collection to
294 estimate the mean abundance of each zooplankton category. Only summer data were used
295 because they allowed for the best time series continuity of the CPR sampling dataset and to
296 maintain consistency with Batten et al. (2018). Further details regarding CPR sampling and
297 analysis are described in Batten et al. (2003).

298 The otolith biochronology was also compared to fish body condition outside of the mixed
299 effects models. This metric was based on length and weight observations from 13,013 fish
300 collected by observers aboard commercial fishery vessels in NMFS Statistical Area 541 from
301 1990-2015. We calculated the relative condition factor K_n (Blackwell et al., 2000), where K_n is
302 the weight of an individual fish divided by the length-specific mean weight. (Length-specific
303 mean weight was calculated separately for each sex and month of collection, to control for
304 potential sex-specific differences and seasonal variation related to reproduction.) Therefore, a
305 fish in average condition for its length/sex/month category would have a K_n of 1, and those
306 above or below average condition would have a K_n of > 1 or < 1 , respectively. All K_n values
307 were then averaged by year to generate a relative annual index of condition representative of the
308 population at Seguam Pass for comparison with the otolith biochronology.

309

310

311 3. Results

312

313 Otoliths from 409 Atka mackerel collected between 1996 and 2015 were examined
314 during this study. Of these, 289 fish (71%) had increment boundaries that were deemed
315 sufficiently clear for measurement (Table 1, Fig. 2). There were 172 females and 117 males in
316 the otolith measurement sample, and fish ranged from 5 to 12 years in age at capture. A total of
317 1,426 growth increments were measured, producing a 28-year biochronology spanning 1988-
318 2015.

319 In our model selection process (based on Δ_i), the model with the optimal random effects
320 structure included a random intercept and slope for each specimen with respect to age at
321 increment formation, thereby allowing each specimen to have its own growth trajectory, as well
322 as random year and cohort intercepts. Of the possible combinations of intrinsic fixed effects,
323 only age at increment formation and age at capture were found to contribute to variation in
324 growth (sex was not a factor in the most parsimonious model). We then extended the optimal
325 intrinsic model to include the extrinsic fixed effects of Atka mackerel abundance, pink salmon
326 abundance, and water temperature. None of the correlation coefficients between extrinsic fixed
327 effects were significant, and variance inflation factors were all below 2.0, indicating there was no
328 collinearity among these explanatory variables (Zuur et al., 2007).

329 The most parsimonious full model ($\Delta_i = 0$) included an effect of pink salmon abundance
330 on variation in growth of Atka mackerel growth increments (Table 2; Table A.1). The Δ_i of the
331 next best model, which included terms for temperature and pink salmon abundance, indicated
332 some support; however, comparison of the AIC_{wt} of the two models suggested the pink salmon
333 abundance-only model was ~ 2.5 times more likely to be the model best representing the data

334 (Wagenmakers and Farrell, 2004). The other candidate extrinsic models had little to no support
335 based on Δ_i and AIC_{wt} , indicating low probability for effects of temperature or Atka mackerel
336 abundance on variability in otolith growth (Table A.1).

337 In the best intrinsic model (not including effects of pink salmon abundance), the otolith
338 biochronology was characterized by a biennial pattern of lower than average growth in odd years
339 followed by higher than average growth in even years, with an overall upward trend during the
340 most recent decade (Fig. 3a). Additionally, growth varied by cohort, with some cohorts
341 outperforming others; this relationship also appeared somewhat cyclical albeit on a longer
342 timeframe than the year effect (Fig. 3b). The magnitude of the year effect was greater than that
343 of the cohort effect.

344 In the full model (including all intrinsic and extrinsic effects), we observed an age-related
345 decline in growth-increment width (Fig. 4a). We also observed a very slight increase in growth
346 with respect to the age the fish was at the time of its capture (Fig. 4b), indicating that fish that
347 were older at capture were generally faster growers than those that were younger at capture.
348 Finally, we observed a negative relationship between otolith growth of Atka mackerel and
349 abundance of Eastern Kamchatka pink salmon (Fig. 4c). The full model explained a high
350 proportion of the total variance in otolith growth ($R_c^2 = 0.806$). The R_m^2 value of this model
351 (0.694), the proportion of the variance explained by fixed factors only, indicated that inclusion of
352 pink salmon abundance explained slightly more of the variance than the intrinsic model (0.680).

353 The otolith biochronology (the random year intercepts from the best intrinsic model) was
354 significantly correlated with several biological indices (Table 3). The otolith biochronology was
355 significantly and positively related to abundance of large copepods in the Southern Bering Sea (r
356 = 0.58, $p = 0.023$). This relationship was strongest over the years 2000-2010, with a correlation

357 coefficient of 0.87 ($p < 0.001$). However, after 2011 this relationship became somewhat
358 uncoupled, particularly in the year 2013 where the two time series varied inversely (Fig. 5a). The
359 otolith biochronology was negatively related to abundance of small copepods (Fig. 5b) and
360 abundance of pink salmon (Fig. 5c). The relationship between the otolith biochronology and pink
361 salmon abundance was strongest between the years 1990 and 2010 ($r = -0.74$, $p < 0.001$), and
362 similar to the relationship with large copepods, the relationship between the otolith
363 biochronology and pink salmon became uncoupled in 2013 due the anomalously low abundance
364 of pink salmon in that year. The otolith biochronology was not significantly related with
365 euphausiid abundance (Table 3). There was no significant relationship between the Atka
366 mackerel biochronology and body condition, or between body condition and any of the other
367 biological time series (Table 3).

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369

370 **4. Discussion**

371

372 In this study, we generated a 28-year otolith biochronology for Atka mackerel, one of the
373 most abundant groundfish species in the Aleutian Islands, Alaska. To the best of our knowledge,
374 this is the first published biochronology from this important large marine ecosystem, as well as
375 one of the first biochronologies anywhere to highlight the importance of upper-trophic level
376 processes relative to otolith growth. This work suggests that biotic interactions are among the
377 most important extrinsic controls on growth of Atka mackerel at Seguam Pass in the Aleutian
378 Islands ecosystem. In particular, we found evidence of competition between pink salmon and
379 Atka mackerel for shared food resources. Both species are zooplanktivores and forage within the

380 top 50 m of the water column (Nichol and Somerton, 2002; Armstrong et al., 2005), with
381 calanoid copepods and euphausiids comprising a major part of their diets (Yang, 1999; Davis et
382 al., 2005; Karpenko et al., 2007; Armstrong et al., 2008; Rand et al., 2010).

383 Pink salmon is the dominant salmonid species in the North Pacific and is biennially
384 highly abundant (Ruggerone and Irvine, 2018). Pink salmon are voracious predators in the North
385 Pacific Ocean, impacting food webs through predation and competition, particularly the large
386 Eastern Kamchatka population, the range of which extends to 155°W during its ocean-going
387 phase (Ruggerone and Nielsen, 2004; Springer and van Vliet, 2014). During odd years when
388 they are highly abundant, pink salmon have been shown to depress populations of large
389 copepods, which in turn reduces grazing pressure on phytoplankton, allowing them to flourish
390 (Shiomoto et al., 1997; Kobari et al., 2003; Batten et al., 2018). Also during these odd years
391 when they are abundant, pink salmon have been shown to outcompete other North Pacific marine
392 predators for shared resources, reducing growth and survival of sockeye (*O. nerka*), chum (*O.*
393 *keta*), and Chinook salmon (*O. tshawytscha*) (Ruggerone et al., 2003; Ruggerone and Nielsen,
394 2004) and even altering hatch dates and reducing growth and breeding success in seabirds
395 (Springer and van Vliet, 2014). There is also evidence of competition between juvenile
396 groundfishes and pink salmon in the eastern Gulf of Alaska in years when resources are limited
397 (Daly et al., 2019).

398 Copepods in general dominate the zooplankton biomass in the Aleutian Islands (Zador
399 and Ortiz, 2018) and are an important component of the Atka mackerel diet. Indeed, we observed
400 a strong positive relationship between the abundance of large copepods and otolith growth.
401 However, there was some temporal variation in the strength of this relationship. The relative
402 abundance of pink salmon, as predicted by its 2-year life cycle and tendency to be highly

403 abundant in odd years, was unexpectedly low in 2013 and high in 2014, resulting in cascading
404 effects on zoo- and phytoplankton communities (Batten et al., 2018). During this same
405 timeframe, the relationships among the otolith biochronology, large copepods, and pink salmon
406 broke down, which could indicate additional unexplained effects on growth, or possible temporal
407 flexibility in the Atka mackerel diet at the study location. For example, in the northern portion of
408 Seguam Pass, fish have also been found to be an important component of the diet of Atka
409 mackerel, especially compared to the western Aleutian Islands (Rand et al., 2010). Still, the
410 overall relationships observed among the otolith biochronology, pink salmon index, and large
411 copepod index suggest the potential for competition between the two fish predators over this
412 important, energetically-rich prey resource. We hypothesize that in years when pink salmon
413 abundance is high, they overgraze large copepods, resulting in lower availability for Atka
414 mackerel, thereby reducing the capacity for growth. In years when pink salmon abundance is
415 low, Atka mackerel are able to capitalize on this food resource and experience overall higher
416 growth. To the best of our knowledge, direct impacts of pink salmon have not been described for
417 other groundfish species in the Aleutian Islands. Given our findings, such potential effects may
418 be worthy of further exploration.

419 Interestingly, the Atka mackerel otolith biochronology had a negative relationship with
420 abundance of small copepods, and no relationship with euphausiid abundance, despite
421 euphausiids being another significant part of the Atka mackerel diet (Yang, 1999, 2003). Though
422 the CPR is known to under-sample large euphausiids (Batten et al., 2018), the rate is assumed to
423 be consistent, and thus should capture relative year-to-year changes in abundance (S. Batten,
424 pers. comm., Marine Biological Association UK, Nanaimo, British Columbia, Canada). The
425 observed relationships between zooplankton abundance and the otolith biochronology may be a

426 result of the spatial distribution of zooplankton prey fields with respect to our sample collection
427 locations. Most of our specimens were collected in the southern portion of Seguam Pass, where
428 large oceanic copepods dominate the zooplankton community (Coyle, 2005). It could be,
429 therefore, within that small region of Seguam Pass, that large copepods are the main limiting
430 resource to Atka mackerel growth. Or, there may be additional unknown sources contributing to
431 variation in growth. Further studies of spatial and temporal variation in Atka mackerel diet
432 preferences may clarify some of these issues. Also worthy of further investigation are the
433 processes driving the observed relationships between the otolith biochronology and abundance of
434 large and small copepods. Detailed examination of copepod size spectra within Atka mackerel
435 diets could elucidate whether the observed biennial patterns are consistent with the relative size
436 composition of copepod prey consumed in odd and even years.

437 Unlike other previous otolith biochronology studies in the adjacent eastern Bering Sea
438 (Matta et al., 2010; van der Sleen et al., 2016) and Gulf of Alaska (von Biela et al., 2015; Matta
439 et al., 2018), temperature was not a predictor of growth. The Aleutian Islands are a unique
440 ecosystem with very different oceanographic and biological dynamics than the Bering Sea and
441 Gulf of Alaska. Seguam Pass in particular is characterized by strong currents and top-to-bottom
442 vertical mixing, resulting in little stratification (Ladd et al., 2005; Stabeno et al., 2005). This in
443 turn results in relatively stable water temperatures experienced by Atka mackerel at Seguam Pass
444 from year-to-year (Rand et al., 2010, 2019). Indeed, temperature has not been found to be a
445 significant predictor of Atka mackerel density (catch per unit effort) in the Aleutian Islands
446 (Rand et al., 2019). Additionally, bioenergetic modeling has demonstrated that diet composition
447 explains far more variation in body growth in Atka mackerel than thermal experience, even
448 across a longitudinal gradient where growth varies considerably (Lowe et al., 1998; Rand et al.,

449 2010). One other possibility is that Atka mackerel may not be as sensitive to fluctuations in
450 temperature as other Alaskan groundfish. Their small (< 50 km) home ranges (McDermott et al.,
451 2005, 2016) suggest that their thermal tolerance may be relatively high, as staying in place could
452 theoretically expose them to greater temperature variability than would migrating to remain in
453 thermally uniform water masses. That, coupled with the relative stability of interannual mean
454 temperature in Seguam Pass, may explain the lack of relationship between Atka mackerel growth
455 and temperature.

456 Otolith biochronology studies often assume but rarely test the relationship between
457 otolith and somatic growth (Black et al., 2013). Here, we did not detect a significant correlation
458 between the biochronology and body condition. One possible explanation is that body condition
459 in the form we used is a measure of cumulative growth that is dependent on weight and length
460 integrated over a fish's lifetime, and therefore is inherently subject to autocorrelation. In contrast,
461 the otolith biochronology represents the relative population-wide deviation in growth for any
462 given year. Furthermore, fish weight is a far more labile trait compared to either fish length or
463 otolith size, and therefore environmentally-driven changes in condition may occur on different
464 timescales than would be captured by the otolith biochronology, making a relationship more
465 difficult to detect. Another possibility is simply that the otolith biochronology is not a perfect
466 proxy for somatic growth. However, we failed to detect any significant relationships between
467 body condition and prey abundance, suggesting that other measures of fish health and fitness
468 should be explored. There is currently no available published information regarding whether K_n
469 is actually an appropriate indicator of growth and fitness in Atka mackerel. In general, there is
470 little agreement in the literature about the best way to assess fish body condition. While more
471 direct measures such as energy density are preferred, these are typically destructive, expensive,

472 and time-consuming. Methods based on length and weight are far more common, as these data
473 can be collected cheaply and rapidly and allow for the development of long-term datasets.
474 However, one drawback of these methods is they can change seasonally and based on life history
475 and reproductive state, and some may not reflect actual body composition (Wuenschel et al.,
476 2019) or relative fitness or survival (Cox et al., 2015). These methods suffer further from
477 difficulties in standardization and thus may induce biases in body condition estimates (Blackwell
478 et al., 2000; Wuenschel et al., 2019). Clearly, further work is necessary to determine the best
479 method of assessing body condition in Atka mackerel as well as other Alaska groundfish species.

480 Most frequently, biochronology studies have been focused on long-lived fish, corals, and
481 molluscs (Black et al., 2005; Carilli et al., 2010; Peharda et al., 2016), but this method has also
482 been successfully applied to fishes with life spans similar in length to Atka mackerel
483 (Morrongiello et al., 2011; von Biela et al., 2015; van der Sleen et al., 2018). One drawback of
484 applying this approach to organisms with lower longevity is that the short lengths of the
485 individual growth time series used to build the biochronology preclude crossdating. The inability
486 to crossdate may therefore result in an unknown error rate that could mute environmental signals
487 (Black et al., 2016). Yet, a consistent alternating pattern of narrow and wide growth increments
488 was apparent in many individuals in our study, which likely reduced dating error by allowing for
489 crossmatching between clear individuals and those with potential interpretative differences with
490 respect to growth on the otolith edge. Moreover, dating errors in long-lived individuals induce
491 frame-shifts that increasingly diminish environmental signals back through time (Black et al.,
492 2016). For the Atka mackerel biochronology, any error rate, which cannot be quantified in the
493 absence of crossdating, should be constant over the length of the biochronology. A second
494 limitation inherent to biochronology studies is that precise measurements can only be taken from

495 clear specimens. However, Atka mackerel consistently have otolith patterns that are relatively
496 easy to interpret and precision among readers is high (Anderl, 2012), likely reducing the
497 potential for measurement and dating errors as well as sample-selection bias.

498 Using an otolith biochronology approach, we found evidence suggesting that growth of
499 Atka mackerel is driven by external biological factors. This study is the first of its kind located in
500 the Aleutian Islands region, to date a relatively poorly studied but large, diverse, and ecologically
501 rich marine ecosystem in the North Pacific Ocean. While previous biochronology studies in
502 Alaska waters have detected significant positive relationships between growth and temperature
503 for a variety of species including flatfish, rockfish, and greenlings (Matta et al., 2010; von Biela
504 et al., 2015; van der Sleen et al., 2016), this is the first to suggest direct linkages with biological
505 drivers, namely prey dynamics and interspecies competition. Future monitoring of these effects
506 on Atka mackerel growth within this dynamic ecosystem is warranted to determine whether
507 these interspecies relationships are truly stationary, especially given projected climate change
508 that may impact zooplankton and salmon populations.

509

510

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- 779 Zuur, A.F., Ieno, E.N., Walker, N.J., Saveliev, A.A., Smith, G.M., 2009. Mixed Effects Models
780 and Extensions in Ecology with R. Springer, New York, 574 pp.

781 Table 1. Summary of specimens used to develop Atka mackerel (*Pleurogrammus*
 782 *monopterygius*) otolith biochronology.

| Collection year | <i>n</i> fish | <i>n</i> increments | Fish length range (cm) | Fish age range (yr) |
|-----------------|---------------|---------------------|------------------------|---------------------|
| 1996 | 48 | 226 | 36-48 | 5-10 |
| 1998 | 18 | 99 | 41-43 | 6-10 |
| 2000 | 32 | 136 | 39-46 | 5-9 |
| 2002 | 33 | 160 | 41-47 | 5-10 |
| 2005 | 43 | 156 | 36-45 | 5-8 |
| 2008 | 28 | 146 | 39-43 | 6-10 |
| 2010 | 34 | 197 | 39-46 | 5-12 |
| 2015 | 53 | 306 | 39-47 | 5-12 |
| Total | 289 | 1426 | 36-48 | 5-12 |

783

784 Table 2. Parameter and variance estimates from most parsimonious full mixed model describing
 785 otolith growth of Atka mackerel (*Pleurogrammus monopterygius*).

| Fixed effects | Estimate | SE | t |
|----------------------|----------|-------|----------|
| Intercept | -2.721 | 0.021 | -132.279 |
| ln(Age at formation) | -1.065 | 0.022 | -47.515 |
| ln(Age at capture) | 0.095 | 0.045 | 2.120 |
| Pink salmon abund. | -0.001 | 0.000 | -3.467 |

| Random Effects | Variance | SD | Correlation |
|--------------------|----------|-------|-------------|
| Specimen intercept | 0.014 | 0.117 | 0.230 |
| Specimen slope | 0.025 | 0.159 | |
| Year | 0.006 | 0.075 | |
| Cohort | 0.002 | 0.046 | |
| Residual error | 0.036 | 0.190 | |

| Metric | Value |
|-------------------|---------|
| AIC _c | -235.96 |
| AIC _{wt} | 0.62 |
| R_m^2 | 0.69 |
| R_c^2 | 0.81 |

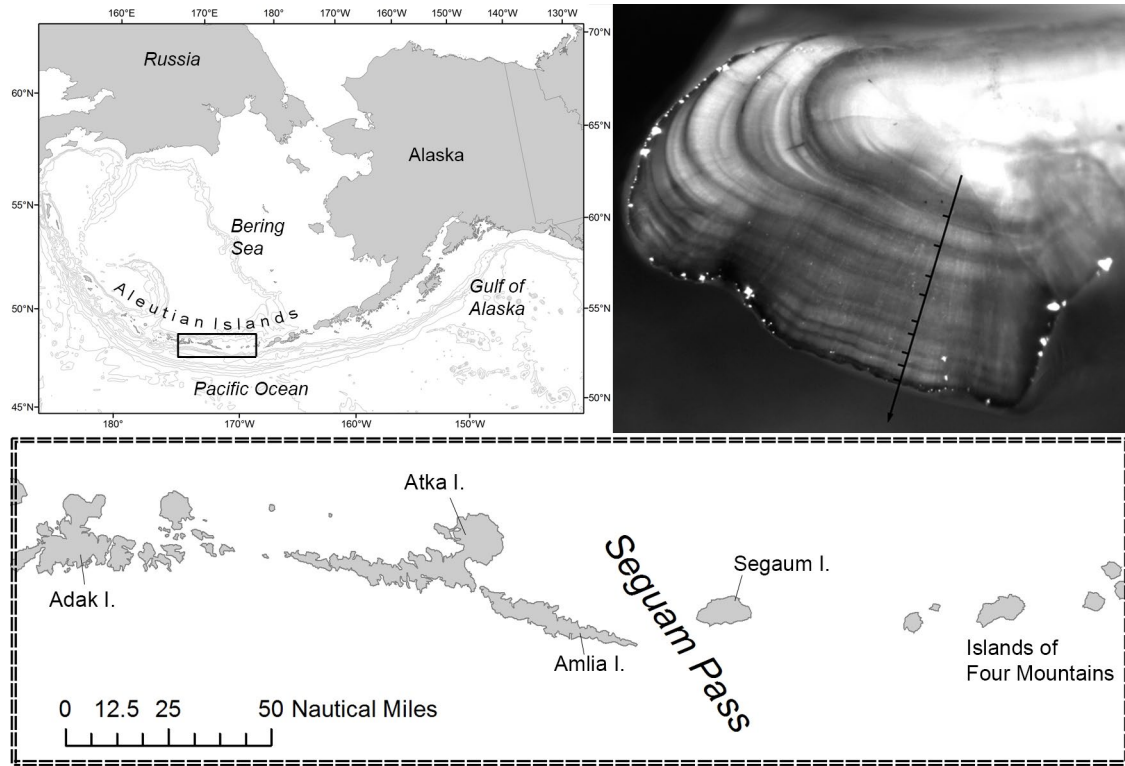
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788 Table 3. Correlations between Atka mackerel (*Pleurogrammus monopterygius*) otolith
 789 biochronology (intrinsic mixed model year effect) and biological variables. Significant
 790 correlations ($p < 0.05$) are in bold text. L. Copepods and S. Copepods are large and small
 791 copepods, respectively.

| | Otolith | Condition | L. Copepods | S. Copepods | Euphausiids | Salmon |
|-------------|--------------|-----------|--------------|-------------|-------------|--------|
| Otolith | 1 | | | | | |
| Condition | -0.19 | 1 | | | | |
| L. Copepods | 0.58 | -0.27 | 1 | | | |
| S. Copepods | -0.58 | -0.11 | -0.28 | 1 | | |
| Euphausiids | 0.24 | 0.02 | -0.05 | -0.20 | 1 | |
| Salmon | -0.52 | -0.08 | -0.73 | 0.35 | 0.07 | 1 |

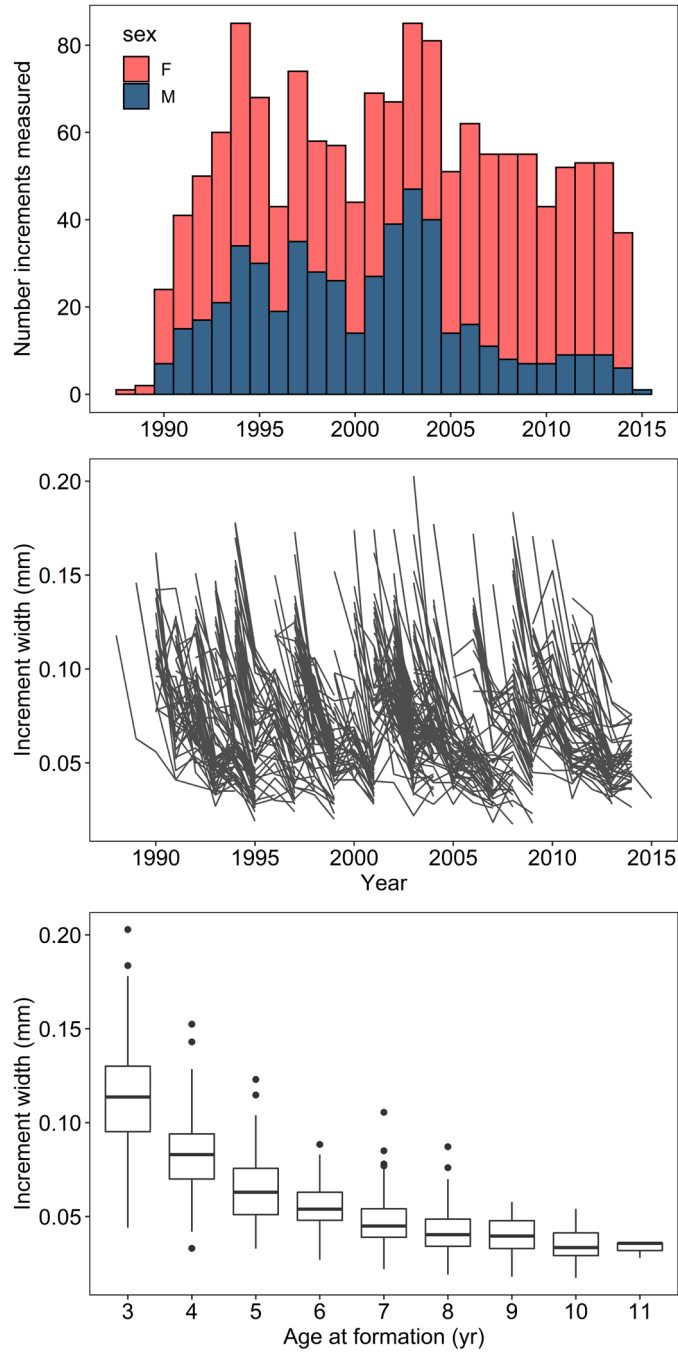
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794 Figure 1. Counter-clockwise from top left: Map of Aleutian Islands and surrounding waters,
 795 close-up of Seguam Pass study area (inside black box). Actual collection locations cannot be
 796 shown due to vessel confidentiality agreements. Transverse section of an Atka mackerel
 797 (*Pleurogrammus monopterygius*) otolith showing axis of measurement and demarcation of
 798 growth increments. Note alternating pattern of wide and narrow increments.

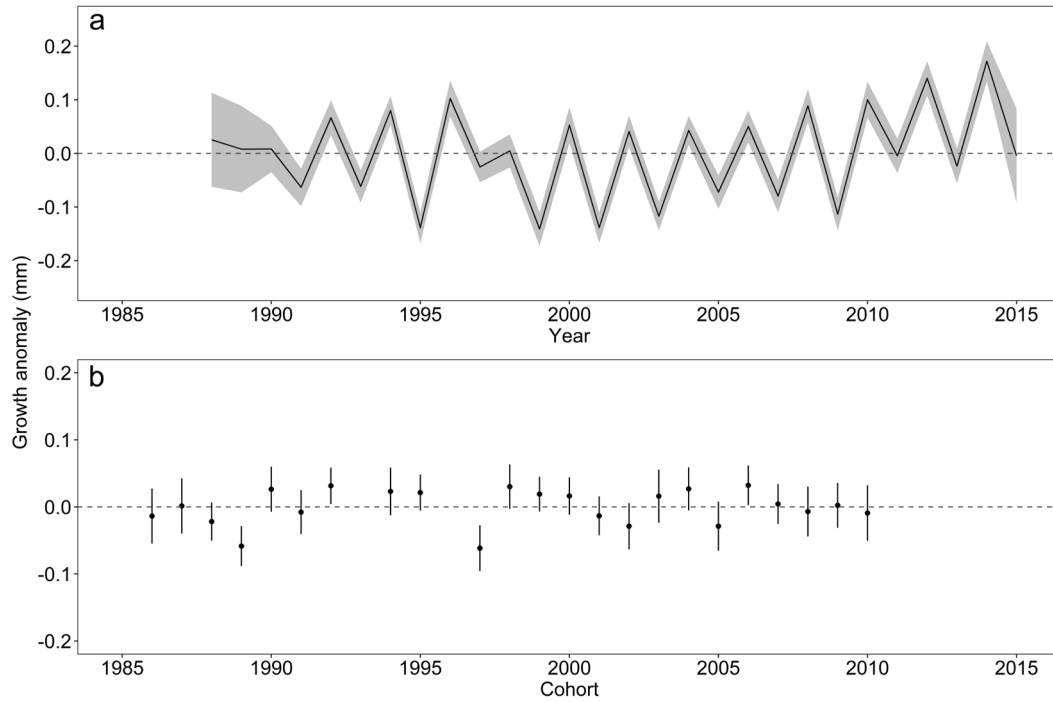
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801 Figure 2. Atka mackerel (*Pleurogrammus monoptyerygius*) otolith biochronology sample depth
 802 (top; males are in blue and females are in coral), raw growth increment measurements (middle),
 803 and decline in increment width relative to age at formation (bottom).

804



805

806 Figure 3. Otolith growth anomalies of Atka mackerel (*Pleurogrammus monopterygius*) predicted

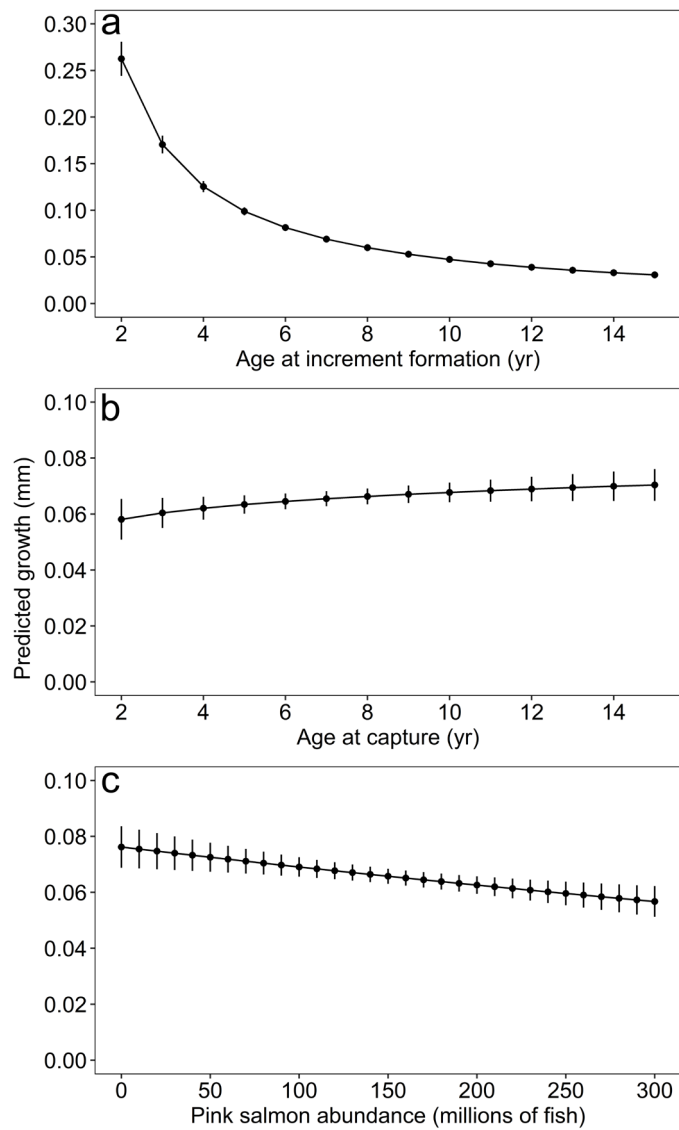
807 by the most parsimonious intrinsic mixed effect model. a) Annual variation in growth (year

808 effect or "biochronology", black line, with standard error represented by shaded area), and b)

809 cohort-specific growth variation (points, with standard error represented by bars), after effects of

810 fish age at increment formation and age at capture have been removed.

811



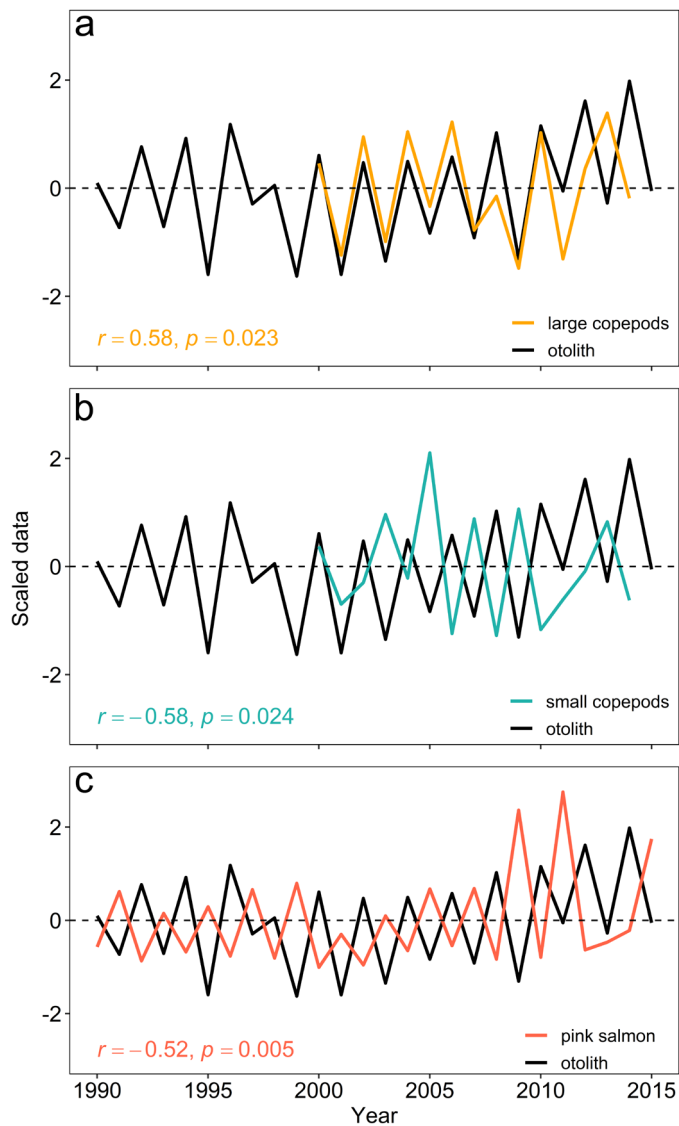
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813 Figure 4. Effects of a) fish age at increment formation, b) fish age at capture, and c) pink salmon

814 abundance on otolith growth of Atka mackerel (*Pleurogrammus monopterygius*). Points

815 represent means and bars represent 95% confidence intervals.

816



817
 818 Figure 5. Biological time series significantly correlated ($p < 0.05$) with the Atka mackerel
 819 (*Pleurogrammus monopterygius*) otolith biochronology (year effect from intrinsic mixed effects
 820 model): a) large copepod abundance, b) small copepod abundance, and c) pink salmon
 821 abundance. All data have been standardized to a mean of 0 and standard deviation of 1.
 822
 823

824 Appendix Table A.1. Results of model fitting for the optimal extrinsic fixed effects structure
 825 describing growth of Atka mackerel (*Pleurogrammus monopterygius*) otolith increments. All
 826 extrinsic models share the same random effects and intrinsic fixed effects structure (see main text
 827 for details). Extrinsic fixed effects considered in the model set are Atka mackerel abundance
 828 index (Atka), pink salmon (*Oncorhynchus gorbuscha*) abundance (Pink), and sea surface
 829 temperature (Temp). K = number of parameters, AIC_c = Akaike Information Criterion corrected
 830 for small sample sizes, Δ_i = difference between each model's AIC_c and the model with the
 831 minimum AIC_c , AIC_{wt} = Akaike weight, and LL = log-likelihood.

832

| Extrinsic fixed effects | K | AIC_c | Δ_i | AIC_{wt} | LL |
|-------------------------|-----------|----------------|------------|-------------|---------------|
| None | 9 | -228.17 | 7.79 | 0.01 | 123.15 |
| Atka | 10 | -226.37 | 9.59 | 0.01 | 123.26 |
| Pink | 10 | -235.96 | 0 | 0.62 | 128.06 |
| Temp | 10 | -226.44 | 9.51 | 0.01 | 123.30 |
| Atka + Temp | 11 | -224.76 | 11.20 | 0.00 | 123.47 |
| Pink + Temp | 11 | -234.16 | 1.80 | 0.25 | 128.17 |
| Atka + Pink + Temp | 12 | -232.47 | 3.49 | 0.11 | 128.34 |

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