1 2	Density-dependent effects of eastern Kamchatka pink salmon ( <i>Oncorhynchus gorbuscha</i> ) and Japanese chum salmon ( <i>O. keta</i> ) on age-specific growth of western Alaska chum
$\frac{2}{3}$	salmon
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5	Running title: Competition among pink and chum salmon populations
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7	Tessa J. Frost <sup>1,4</sup> , Ellen M. Yasumiishi <sup>2</sup> , Beverly A. Agler <sup>3</sup> , Milo D. Adkison <sup>1</sup> , and Megan V.
8	McPhee <sup>1,5</sup>
9	
10	
11	<sup>1</sup> College of Fisheries and Ocean Sciences, University of Alaska Fairbanks, Juneau, Alaska, USA
12	
13	<sup>2</sup> Alaska Fisheries Science Center, National Marine Fisheries Service, National Oceanic and
14	Atmospheric Administration, Juneau, Alaska, USA
15	
16	Alaska Department of Fish and Game, Mark, Tag, and Age Lab, Juneau, Alaska, USA
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<sup>&</sup>lt;sup>4</sup> Present address: Southern Southeast Regional Aquaculture Association, Ketchikan, Alaska, USA <sup>5</sup> Corresponding author; mvmcphee@alaska.edu

#### 33 ABSTRACT

34 Hatcheries release > 4.5 billion juvenile Pacific salmon (*Oncorhynchus* spp.) into the North 35 Pacific Ocean annually, raising concerns about competition with wild salmon populations. We 36 used retrospective scale analysis to investigate how the growth of chum salmon (O. keta) from 37 western Alaska is affected by the abundance of chum salmon from Japanese hatcheries and wild 38 pink salmon (O. gorbuscha) from the Russian Far East. Over nearly five decades, the growth of 39 Kuskokwim River chum salmon was negatively correlated with the abundance of Japanese 40 hatchery chum salmon after accounting for the effects of sex and spring/summer sea surface 41 temperature in the Bering Sea. An effect of wild eastern Kamchatka pink salmon abundance on 42 the growth of Kuskokwim River salmon was detectable but modest compared to the intraspecific 43 competitive effect. A decrease in Japanese hatchery chum salmon releases in 2011-2013 was not 44 associated with increased growth of Bering Sea chum salmon. However, the abundance of wild 45 chum salmon from the Russian Far East increased during that time, possibly obscuring reduced 46 competition with hatchery chum salmon. Our results support previous evidence that chum 47 salmon are affected by intraspecific competition, and to a lesser extent interspecific competition, 48 in the North Pacific, underscoring that the effects of salmon hatchery production transcend 49 national boundaries. 50

51

52 Keywords (7): aquaculture; Bering Sea; competition; growth, hatcheries, North Pacific;

- 53 retrospective analysis
- 54

55 1 | INTRODUCTION

56 57 Abundance of Pacific salmon (Oncorhynchus spp.) has been increasing in North America 58 and Asia since the late 1970s, with record catches of pink salmon (O. gorbuscha) and chum 59 salmon (O. keta) beginning in the 1990s (Ruggerone & Irvine, 2018). Increases in salmon 60 abundance are attributed to favorable ocean growing conditions and increased hatchery output, 61 resulting in nearly twice as many salmon in the North Pacific Ocean during 1990-2005 compared 62 to 1952-1975. Despite favorable ocean conditions, decreases in growth rate and delayed maturity 63 have been observed throughout the chum salmon's range (Ishida, Kaeriyama, McKinnell, & 64 Nagasawa 1993; Ruggerone, Agler, & Nielsen 2011). Highly abundant Asian pink and chum 65 salmon in the Bering Sea may be driving density-dependent effects on salmon populations from 66 around the Pacific Rim (Cline, Ohlberger, & Schindler, 2019). 67 Chum salmon are widely distributed in the North Pacific Ocean regardless of stock of 68 origin (Seeb, Crane, Kondzela, et al. 2004; Myers, Klovach, Gritsenko, et al. 2007; Myers, 69 Walker, Davis, et al., 2009; Sato, Moriya, Azumaya, & Nagoya 2009). Chum salmon head to sea 70 soon after emergence and exhibit variable age at maturity, returning to spawn between two and 71 five years of age (Salo, 1991). Western Alaska chum salmon migrate between summer feeding 72 grounds in the Bering Sea and overwintering grounds in the Gulf of Alaska (Urawa, Sato, Crane, 73 Agler, et al. 2009). Japanese chum salmon exhibit a similar migration pattern, with the exception 74 of their first winter at sea which is spent in the western North Pacific Ocean (Urawa, 2004).

These migration patterns may lead to competition between Japanese chum salmon and western
Alaska chum salmon (after their first year at sea) when prey resources are limited (Ishida et al.,
1993).

78 Pink salmon are thought to be strong competitors due to their abundance (accounting for 79 ~ 60% of anadromous Pacific salmon; Heard, 1991; Ruggerone, Zimmerman, Myers, Nielsen, & 80 Rogers, 2003) and high prey consumption rates necessary to sustain rapid growth (Ruggerone & 81 Neilsen, 2005). Indirect evidence, based on correlative analysis and strong even/odd-year 82 abundance differences associated with the biennial life cycle of pink salmon, has been taken to 83 suggest that they can drive variation in zooplankton abundance and species composition (Batten, 84 Ruggerone, & Ortiz, 2018), depress the growth of sockeye salmon O. nerka (Ruggerone et al., 85 2003; Ruggerone, Agler, Connors, et al., 2016; but see McKinnell & Reichardt, 2012), and 86 influence reproductive success in seabirds (Springer & van Vliet, 2014). Direct estimates of pink 87 salmon abundance have been linked to reduced growth of chum salmon (Agler, Ruggerone,

88 Wilson, & Mueter, 2013) and productivity of sockeye salmon (Ruggerone & Connors, 2015;

89 Ruggerone et al., 2016). Chum salmon, the second-most abundant species of Pacific salmon, also

90 appear to exert intraspecific competitive effects on growth within (Ishida et al., 1993) and among

91 (Agler et al., 2013; Ruggerone et al., 2011) populations. But with less inter-annual contrast in

92 abundance, the competitive effects of chum salmon are more difficult to detect and have received

93 less study compared to pink salmon.

94 Understanding the effects of competition on growth and life history of Pacific salmon is 95 an unresolved but consequential issue, particularly as hatcheries around the Pacific Rim continue 96 to release billions of juvenile salmon into the North Pacific every year (NPAFC, 2020). Here, we 97 focus on potential competitive effects on chum salmon stocks originating from western Alaska, 98 including the watersheds of Norton Sound and the Yukon and Kuskokwim Rivers. Unexpected 99 declines in Chinook (O. tshawytscha) and chum salmon populations in western Alaska between 100 1997-2002 prompted Alaska to declare a "fisheries disaster" (Krueger, Zimmerman, & Spaeder, 101 2009). Periodic reductions in harvests of salmon in western Alaska continue to have detrimental 102 economic and cultural consequences for rural Alaskan communities (Loring & Gerlach, 2010). It 103 is therefore important to understand how competition might affect regional chum salmon 104 populations.

105 We used retrospective analysis of scale growth to investigate the relationship between 106 growth of chum salmon, known to originate from the Kuskokwim River in western Alaska or to 107 rear in the Bering Sea, and the abundance of Asian pink and chum salmon, specifically two 108 major stocks: Japanese hatchery chum salmon, and pink salmon from eastern Kamchatka. The 109 correlation between scale radius and fish length allowed us to examine growth during defined 110 periods of ocean residence (Yasumiishi, Criddle, Helle, Hillgruber, & Mueter, 2016) and 111 therefore make specific predictions about the effects of competition on growth patterns. First, we 112 predicted that competition should manifest in reduced growth of Kuskokwim River chum salmon 113 from the second year of growth at sea onwards, but not during the first year of growth at sea 114 when these populations would have minimal overlap with Asian populations (Agler et al., 2013). Second, we examined growth of chum salmon in the Bering Sea in relation to reduced abundance 115 116 of Japanese chum salmon associated with the Tohoku earthquake and tsunami of 2011. This 117 disaster destroyed a number of hatcheries along the Pacific coast, reducing the numbers and

118 subsequent survival of hatchery chum salmon released that spring (Watanabe, Sasaki, Saito, & 119 Ogawa, 2015). Hatchery chum salmon releases were reduced in 2012 and 2013 as well (NPAFC, 120 2020). We hypothesized that this decrease in hatchery chum salmon production would result in 121 increased growth of chum salmon in the Bering Sea during years 2012-2016. In both cases, we 122 modeled competition effects while accounting for covariates expected to influence growth in 123 chum salmon: sex and age at maturity (Morita, Morita, Fukuwaka, & Matsuda, 2005) and sea-124 surface temperatures (SST) in spring (Wechter, Beckman, Andrews, Beaudreau, & McPhee, 125 2017) and/or summer (Mueter, Peterman, & Pyper 2002).

126 The two-pronged approach we took provided additional insight into competition among 127 salmon in the North Pacific. First, by treating the first year of growth of Kuskokwim River chum 128 salmon as a 'counterfactual' (i.e., little to no competition with Asian pink and chum salmon), we 129 were better able to assess the possibility that spurious statistical associations between growth and 130 Asian pink and chum salmon abundance indices arose from coinciding time series, which can 131 confound inferences about causal effects of competition (e.g., Cunningham, Adkison, & 132 Westley, 2018). Second, the sharp decline in Japanese hatchery chum salmon beginning with the 133 2011 earthquake provided a way to test for density dependence in the Bering Sea against a 134 backdrop of limited contrast in hatchery releases of chum salmon over the previous three decades 135 (NPAFC, 2020), which has hampered efforts to understand the competitive effects of hatchery 136 chum salmon.

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- 139 **2** | **METHODS**
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# 141 **2.1** | Study populations142

We analyzed two collections ('Kuskokwim River' and 'Bering Sea') of scales from chum salmon that rear in the Bering Sea. The Kuskokwim River samples were collected by the Alaska Department of Fish and Game (ADF&G) from mature chum salmon captured in commercial and test fisheries, using set and drift gillnets with mesh sizes  $\leq 15$  cm (Bue, 2005), in district W-1 near Bethel, Alaska, in 1963 and 1968 – 2010 (Figure 1a). This dataset represented fish that survived their ocean residence and provided stock-specific information on temporal trends in growth. We targeted 25 samples for each sex and dominant age (4 or 5 years old) combination. 150 The Bering Sea dataset was composed of immature and maturing chum salmon caught 151 incidentally in the Bering Sea commercial trawl fishery for walleye pollock Gadus 152 chalcogrammus. Scales were collected by fishery observers deployed by the National Oceanic 153 and Atmospheric Administration (NOAA) on vessels and processing plants (Chilton, 2016). The 154 Bering Sea dataset included chum salmon captured within 52°N-60°N, 160°W-175°W (Figure 155 1a) from 16-31 July of each year, 2001-2016. This collection included chum salmon from an 156 uncharacterized mixture of populations (including hatchery origin) that were intercepted in the 157 Bering Sea prior to their homeward migration. Bering Sea sampling was opportunistic, with 158 sample sizes depending on the number of salmon intercepted in the pollock fishery.

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160 **2.2** | Scale measurements

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162 Annual growth increments (Figure 1b) were measured from Kuskokwim chum salmon 163 scales (n = 3,742; Supporting Information Table S1) by ADF&G personnel following methods 164 described by Hagen, Oxman, & Agler (2001) and Agler et al. (2013). Briefly, images of salmon 165 scales were scanned from acetate impressions using an Indus 4601-11 Screen Scan microfiche 166 reader and measured using Image-Pro Premier 9.0 software with a customized application. Two 167 readers read a majority (91%) of the Kuskokwim scales, and their measurements were similar. A 168 randomly selected set of Kuskokwim scales (n = 50) was re-measured by an independent reader 169 to test for a reader effect using multiple analysis of variance (MANOVA) of growth 170 measurements and reader identity.

171 Annual growth increments were measured from Bering Sea chum salmon scales (n =172 1,845; Supporting Information Table S2) as described by AFSC (2017). Acetate impressions 173 made of scales mounted on gum cards were scanned at 24x magnification using Z Scan 46-II 174 Image Scanner through NOAA's imaging lab and digitized using Image-Pro Plus 7.0 software 175 with a customized application. All Bering Sea scales were measured by a single reader. 176 We followed the same scale-quality criteria for both collections. A scale was only 177 measured if: 1) the scale came from the "preferred" zone of the fish (Hagen et al., 2001); 2) age, 178 sex, and length data could be matched to the scale; 3) the scale was not regenerated; 4) the annuli 179 were clearly visible and defined; and 5) the scale could be measured along its longest axis. 180 Individual fish lacking scales meeting these criteria were not included in the analysis. We 181 followed the growth-increment notation of Morita & Fukuwaka (2006) where  $g_1$  indicates growth during the first year at sea,  $g_2$  describes growth accrued during the second year at sea, and so on. The width of each growth increment was measured as the distance between adjacent annuli (Figure 1b).

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#### 2.3 | Indices of eastern Kamchatka pink salmon and Japanese chum salmon abundance

187 188 Analyses focused on interactions in the Bering Sea between western Alaska chum salmon 189 and the dominant stocks of pink and chum salmon during the study period: Japanese chum 190 salmon, assumed to be entirely of hatchery origin, and eastern Kamchatka pink salmon primarily 191 of wild origin (Urawa et al., 2004; Urawa et al., 2009; Myers et al., 2007; Irvine & Ruggerone, 192 2016). Abundance estimates (Figure 2a,b) were obtained from Irvine & Ruggerone (2016); see 193 also Ruggerone & Irvine (2018). Specifically, we used their estimates of the number of salmon 194 (including those harvested) returning to coastal waters for Japanese chum salmon and pink 195 salmon from the east coast of Kamchatka. These numbers corresponded to the 'Japan' column in 196 Table 9 and the 'EKam' column in Table 8, respectively, of Irvine & Ruggerone (2016). 197 Abundance indices were transformed by natural logarithm prior to analysis. Irvine & Ruggerone (2016) also provided estimates of total biomass (immature and maturing) of salmon by species 198 199 and region; however, these estimates required additional, untested assumptions and were highly 200 correlated with returning adult abundance estimates for both Japanese chum salmon and pink 201 salmon from eastern Kamchatka (r > 0.99), so we used returning adult abundance estimates.

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#### 2.4 | Environmental covariates

To account for environmental influences on chum salmon growth represented by SST, we developed a seasonal index for spring and summer SST in the Bering Sea (Figure 2c, d).

207 Temperature data were obtained from NOAA's Earth System Research Laboratory's Physical

208 Sciences Division website (available from https://www.esrl.noaa.gov/psd/cgi-

209 bin/data/timeseries/timeseries1.pl), based on National Centers for Environmental

210 Prediction/National Center for Atmospheric Research gridded reanalysis data (Kalnay et al.,

211 1996). Summer SST was averaged across July and August in the central Bering Sea (54.3°-

212 60.0°N, 170.6°-178.1°W; Figure 1a). This spans an area used by immature chum salmon during

their summer feeding months (Echave, Eagleton, Farley, & Orsi, 2012) and encompasses both

shelf and slope habitats. Spring SST was based on average SST during April and May at the M2
mooring buoy in the southeastern Bering Sea (56.87°N, 164.06°W; Figure 1a).

The timing of the spring phytoplankton bloom in the eastern Bering Sea is strongly correlated with sea-ice extent (Sigler, Stabeno, Eisner, Napp, & Mueter, 2014). NOAA's ice retreat index (IRI; available from <u>www.beringclimate.noaa.gov/data/index.php</u>) at the M2 buoy did not span our entire study period, so we used the Spring SST index as a proxy for the date of sea-ice retreat (justified by linear regression between IRI and mean April/May SST at M2,  $R^2 =$ 0.72, p < 0.001; see also Wechter et al., 2017).

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#### 224 **2.5** | Statistical analyses

226 Linear mixed-effects models were fit using the lme4 V1.1-12 package (Bates, Maechler, 227 Bolker, & Walker, 2015) in the statistical language R, v. 3.3.3 (R Core Team, 2017). Prior to 228 modeling, all variables were normalized (mean = 0, SD = 1) to facilitate comparison of model 229 coefficients. Each annual growth increment was normalized separately within each of the two 230 datasets (Kuskokwim River and Bering Sea). For objective 1 (growth of Kuskokwim River chum 231 salmon), we used the second-order Akaike information criterion (AICc) to compare full models 232 to models with all possible combinations of abundance indices and covariates, including null 233 models (random effects only). For objective 2, in which we tested for a signal of reduced 234 Japanese hatchery releases 2011-2013 on chum salmon growth in the Bering Sea, we used 235 likelihood-ratio tests to determine whether a model containing the reduction in releases fit the 236 data significantly better than the model without the reduction. Model comparison (for differing 237 fixed effects) was conducted on models fit with maximum likelihood; coefficients of selected 238 models were estimated with restricted maximum likelihood (Zuur, Ieno, Walker, Saveliev, & 239 Smith, 2009).

Preliminary analysis indicated that spring and summer SSTs were moderately correlated in the Kuskokwim dataset (Pearson's r = 0.53) but more strongly correlated over the Bering Sea dataset (Pearson's r = 0.73). Consequently, we included only spring SST in the models for objective 2, as it showed considerable contrast over the study period (Figure 2) and might better capture environmental variation affecting chum salmon growth (Wechter et al., 2017). To determine if growth of western Alaska chum salmon varied with Asian pink and chum salmon abundance (objective 1; Kuskokwim dataset), linear mixed-effects models were fit separately for each growth increment  $g_2 - g_4$  to examine the influence of predictor variables on inter-annual variability in growth. The model (1) for each growth increment was:

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$$g_{iy} = \beta_1 Sex_i + \beta_2 Age_i + \beta_3 SprSST_y + \beta_4 SumSST_y + \beta_5 JpnChum_y + \beta_6 KamPink_y + \beta_7 (JpnChum*KamPink)_y + \alpha_y + \varepsilon_{iy}$$
(1)

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253 where  $g_{iy}$  was the size of the growth increment of individual *i* during calendar year *y* of growth, 254 Sexi was a factor (male vs. female), Agei (factor: 4 vs. 5 years old) was age at maturity of 255 individual *i*;  $SprSST_v$  was spring SST during year *y*,  $SumSST_v$  was summer SST during year *y*, 256 JpnChum<sub>v</sub> was the abundance of Japanese chum salmon during year y, KamPink<sub>v</sub> was the 257 abundance of eastern Kamchatka pink salmon in year y, and the interaction between JpnChumy 258 and KamPink<sub>v</sub> allowed for the competitive effect of one stock being dependent on the abundance 259 of the other stock. The random effect  $\alpha_v$  accounted for correlated growth among individuals 260 growing during the same calendar year y, and  $\varepsilon_{iv}$  represented the residual variance. Age was not 261 included in the full model for  $g_4$  (present only in age-5 fishes). The random intercepts  $a_v$  and the residuals  $\varepsilon_{iv}$  were assumed to be independent and normally distributed with means of zero and 262 variances  $\sigma_a{}^2$  and  $\sigma_\epsilon{}^2$ , respectively; these assumptions were checked with model diagnostic plots 263 264 (not shown). We repeated the same modeling approach using  $g_1$  as a counterfactual response 265 variable, hypothesizing no effect of competition with Asian pink and chum salmon.

To determine if reduced Japanese hatchery chum salmon releases in 2011-2013 resulted in increased growth of Bering Sea chum salmon (objective 2; Bering Sea dataset), linear mixedeffects models were fit separately to each growth increment  $g_2 - g_4$  to examine the influence of reduced hatchery output while accounting for other sources of variation (sex, age, and spring SST). Because the decrease in Japanese hatchery outputs occurred 2011-2013, the effect was anticipated during growth years 2012-2016. The model (2) for each growth increment was:

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$$g_{iy} = \beta_1 Period_y + \beta_2 Sex_i + \beta_3 Age_i + \beta_4 SprSST_y + \alpha_y + \varepsilon_{iy}$$
(2)

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275 where  $Period_{y}$  was a factor identifying years prior to versus during the period affected by 276 decreased hatchery chum salmon releases; Age represented age at capture (including immature 277 and maturing fish); and the other variables are as described for the previous model. Age was not 278 included in the full model for  $g_4$  (present only in age-5 fishes). Likelihood-ratio tests were used 279 to determine whether addition of the reduced-release period significantly improved fit over a 280 model containing only sex, age, and spring SST; tests were evaluated at  $\alpha = 0.05$ . Model 281 assumptions, including no temporal autocorrelation, were checked with diagnostic plots (not 282 shown).

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### 284 **3 | RESULTS**

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286 Sample size varied by population, sex, age class, and growth increment (Supporting 287 Information, tables S1 and S2). Target sample size was generally achieved for Kuskokwim River 288 chum salmon except for the 1963 and 1969 brood years, which were represented by < 20289 individuals each. No reader effect was found over the entire scale (MANOVA, p = 0.97) or for 290 single growth increments (ANOVA, p > 0.5 for all) for Kuskokwim River scales. Bering Sea 291 sample sizes ranged from 7 to 120 per sex, year, and growth increment, with  $g_4$  tending to have 292 fewer samples than  $g_2$  and  $g_3$ . Average sizes of age-specific growth increments showed different 293 relationships over time, with only  $g_i$  in the Kuskokwim River dataset showing an upward trend, 294 and  $g_4$  in the Kuskokwim River dataset and  $g_3$  in the Bering Sea dataset showing downward 295 trends (Supporting Information, figures S1 and S2).

In Kuskokwim River chum salmon, the top model for  $g_1$  (no expected competition) contained only a positive correlation with summer SST. Summer SST appeared in all of the bestsupported models for  $g_1$  (those within two AICc units of the top model; Table 1). Three of the six best-supported models included salmon abundance indices, but in all cases the coefficients were positive and small (Figure 3). The null model for  $g_1$  was >9 AICc units above that of the top model (Supplementary Information, Table S3).

302 After the first year at sea, Kuskokwim River chum salmon growth decreased with 303 increasing abundance of Japanese chum salmon (Figure 4); additionally, faster growth was 304 associated with males and earlier-maturing individuals (Table 1). All of the best-supported 305 models for growth increments  $g_2 - g_4$  included a negative correlation with Japanese chum salmon 306 abundance (Table 1). The coefficients for Japanese chum salmon abundance were consistently 307 more negative than those for Kamchatka pink salmon abundance, and null models were > 100308 AICc units away from the top models (Supporting Information, Table S3). Coefficients for 309 Kamchatka pink salmon abundance and its interaction with Japanese chum salmon abundance 310 differed among the growth increments (Figure 3). For  $g_2$ , the coefficient for Kamchatka pink 311 salmon abundance was slightly positive (included only in the fourth-ranked model), and no 312 interaction between Japanese chum salmon and Kamchatka pink salmon abundances appeared in 313 any of the best-supported models. In contrast, the coefficient for pink salmon abundance was 314 consistently negative for both  $g_3$  and  $g_4$ , and a positive interaction between the two salmon 315 abundance indices appeared in the three highest-ranking models for g<sub>3</sub> and in the fourth-ranked 316 model for  $g_4$ .

317 Sea-surface temperature appeared in many of the best-supported models for increments 318  $g_2 - g_4$  in Kuskokwim River chum salmon, but differed among increments (Supplementary 319 Information, Table S3). For  $g_2$ , the top model did not contain SST, but a negative correlation 320 with spring SST was included in the second-ranked model, a positive correlation with summer 321 SST was in the third-ranked model, and the fifth-ranked model contained both spring and 322 summer SST. Positive coefficients for spring SST appeared in the top models for both  $g_3$  and  $g_4$ , 323 while summer SST (with positive coefficients) appeared only in the third-ranked model for  $g_3$ 324 and the sixth- and seventh-ranked models for  $g_4$ .

Estimated parameters from the models for growth increments  $g_2 - g_4$  in Bering Sea chum salmon with and without accounting for decreased hatchery releases in 2011-2013 are shown in Table 2. For  $g_2$  and  $g_4$ , the simpler model without the reduced-release period was preferred (likelihood-ratio tests, P > 0.5). For  $g_3$ , the model with the reduced-release period was preferred (P = 0.04), but its coefficient was negative, opposite of the hypothesized effect. As in the Kuskokwim dataset, faster growth was associated with males and younger fish. Growth during  $g_2$  $-g_4$  was positively related to spring SST, but coefficients were < 0.20 (Table 2).

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## 333 4 | DISCUSSION

334

Based on hypothesized competitive effects of wild Kamchatka populations of pink salmon and chum salmon released by Japanese hatcheries, we expected reduced growth of 337 western Alaska chum salmon after their first year at sea during periods of high abundance of 338 these Asian salmon stocks. We found strong support that chum salmon returning to the 339 Kuskokwim River compete with Japanese chum salmon, as evidenced by the negative correlation between the size of hatchery releases and marine growth  $(g_2 - g_4)$ , and the absence of 340 341 such a correlation with  $g_1$ , in chum salmon returning to the Kuskokwim River 1963-2010. Despite the hypothesis that pink salmon are strong competitors with profound consequences on 342 343 marine food webs (Ruggerone & Nielsen, 2004; Springer et al., 2018), the effect of eastern 344 Kamchatka pink salmon on Kuskokwim River chum salmon growth was modest compared to the 345 effect of Japanese chum salmon. Agler et al. (2013) examined  $g_1$  and  $g_3$  in chum salmon 346 returning to the Yukon River, to the north of the Kuskokwim River, and to Bristol Bay to the 347 south. They similarly detected a greater competitive effect of Asian chum salmon than Asian 348 pink salmon on  $g_3$  (and no effect on  $g_1$ ). Our results suggest that these findings might be broadly 349 applicable to chum salmon populations from across western Alaska.

350 We detected no negative effect of eastern Kamchatka pink salmon abundance on growth 351 of Kuskokwim River chum salmon during their second year at sea, but we did find evidence for competitive effects on later marine growth. This was particularly apparent in  $g_3$ , for which three 352 353 of the four best-supported models included a positive interaction between eastern Kamchatka 354 pink salmon abundance and the magnitude of Japanese hatchery chum salmon releases. The 355 interaction indicated that Kuskokwim River chum salmon g<sub>3</sub> growth was particularly poor in 356 years when the combined abundance of these two stocks was particularly high. Chum salmon are 357 able to mitigate competition with pink salmon by feeding on gelatinous zooplankton not targeted 358 by other salmon (Tadokoro, Ishima, Davis, Ueyenagi, & Sugimoto, 1996). However, as chum 359 salmon grow they consume more fish (Karpenko, Volkov, & Koval, 2007), so perhaps by their 360 third year at sea Kuskokwim River chum salmon are less able to switch diets in order to avoid 361 competition with pink salmon, particularly in years when Japanese chum salmon is also 362 abundant. Chum salmon might adjust their oceanic distribution southward to avoid pink salmon 363 (Azumaya & Ishida, 2000), but perhaps during later growth years, homeward migration 364 pathways prevent Kuskokwim River chum salmon from doing so.

While Kuskokwim River chum salmon appear to compete with Japanese hatchery chum salmon, a steady increase in hatchery production through the 1990s followed by little variation has provided limited contrast for quantifying the extent of competition. We used a rapid decrease 368 in Japanese hatchery chum salmon production in 2011-2013 to contrast growth of chum salmon 369 intercepted in the Bering Sea during periods of normal (2001-2011) and low (2012-2016) 370 Japanese hatchery chum salmon releases. Contrary to our expectations and to our findings with 371 the Kuskokwim River dataset over a longer period, chum salmon in the Bering Sea did not 372 exhibit increased growth following the reduction in Japanese hatchery chum salmon production. 373 This contradiction could be resolved in several ways. First, the reduction in Japanese hatchery 374 chum salmon outputs may have been too small to cause a detectable effect on growth of chum 375 salmon in the Bering Sea. Second, our Bering Sea samples came from the Alaskan walleye 376 pollock trawl fishery. If these samples were size selective, changes in growth might have gone 377 undetected with this method of sampling. Finally, we observed a substantial increase in wild 378 chum salmon abundance over the study period for the Bering Sea dataset (2001-2016). Wild 379 Russian chum salmon increased by 264% from 2005-2015, reducing the proportion of hatchery 380 chum salmon from 62% in 2005 to 45% of total chum salmon abundance in 2015 (Ruggerone & 381 Irvine, 2018). Russian chum salmon follow an ocean migration pattern similar to that of Japanese 382 and western Alaska chum salmon (Urawa et al., 2009), and thus in the years that western Alaska 383 chum salmon were hypothesized to have experienced relief from competition with Japanese 384 hatchery chum salmon, they might have experienced increased competition with Russian chum 385 salmon. Future analyses of chum salmon growth should account for the abundance of Russian 386 chum salmon populations in addition to Japanese hatchery chum salmon.

387 The relationship between chum salmon growth and spring and summer SST in the Bering 388 Sea was smaller than expected. However, previous research has shown that the effects of SST on 389 chum salmon growth are variable. In southeast Alaska, Yasumiishi et al. (2016) observed a 390 positive relationship between g<sub>2</sub> growth and cooler summer/fall SST, but after accounting for 391 density-dependent effects (growth was more strongly correlated with population abundance), 392 growth was positively associated with SST. In contrast, Agler et al. (2013) observed a negative 393 relationship between western Alaska chum salmon growth and Gulf of Alaska SST. Analysis of 394 Kwiniuk River chum salmon in Norton Sound, western Alaska, detected no correlation between 395  $g_2$  -  $g_4$  scale growth in four- and five-year old chum salmon and SST in the North Pacific Ocean 396 during winter, spring, or summer (Ruggerone & Agler, 2008). These contradictory findings 397 suggest that the relationship between salmon growth and SST is complex and may not be linear 398 across the range of variation experienced by chum salmon populations across years. Our

399 understanding might be limited by lack of information on population-specific ocean

- 400 distributions. Furthermore, the diving behavior of chum salmon (Azumaya & Nagasawa, 2009)
- 401 suggests that temperature measured at the ocean's surface may not be the best proxy for
- 402 environmental conditions affecting salmon growth and age/size trends (G. Brown, Department of
- 403 Fisheries and Oceans Canada, pers. comm.).

404 Both Kuskokwim River and Bering Sea chum salmon experienced decreases in growth 405 during a time of maximum releases of chum salmon by Japanese hatcheries. Similar size declines 406 have been observed in chum salmon populations throughout the Pacific Rim, including Russia, 407 Japan, Canada, and the United States (e.g., Ishida et al., 1993; Bigler, Welch, & Helle, 1996; 408 Zavolokin, Zavolokina, & Khokhlov, 2009; Ruggerone et al., 2011; Yasumiishi et al., 2016), 409 pointing to shared causes in the North Pacific. Density-dependent effects of Asian hatchery chum 410 salmon have been detected on both North American (Helle, Martinson, Eggers, & Gritsenko, 411 2007; Ruggerone et al., 2011; Agler et al., 2013) and Asian (Ishida et al., 1993; Zaporozhets & 412 Zaporozhets, 2004) chum salmon populations. In western Alaska, increased production of Asian 413 hatchery chum salmon was significantly related to reduced adult length at age, productivity, and 414 delayed age at maturity of Norton Sound chum salmon (Ruggerone et al., 2011). The reduced 415 growth of Kuskokwim chum salmon we observed during  $g_2$  may contribute to increases in chum 416 salmon age at maturity, as slower growing salmon tend to mature at an older age (Morita et al., 417 2005; Siegel, McPhee, & Adkison, 2017). Unfortunately, the abundance and age data collected 418 from chum salmon returning to the Kuskokwim River were of insufficient quality for estimating 419 age composition by brood year (H. Hamazaki, ADF&G, pers. comm.), so we were unable to 420 directly test this hypothesis.

421 Testing for density dependence in the open ocean is difficult, but our analyses provide 422 insight into the potentially adverse consequences of increasing hatchery salmon outputs into the 423 North Pacific ecosystem. Over the longer time series afforded by the Kuskokwim River chum 424 salmon scale collection, we found strong support for competitive effects of Japanese hatchery 425 chum salmon and to a lesser extent, competitive effects of wild pink salmon from eastern 426 Kamchatka. That we did not observe increased growth of Bering Sea chum salmon when the 427 production of Japanese hatchery chum salmon was reduced in 2011-2013 may be because 428 decreases in hatchery outputs were too small to detect a difference in growth, or because the 429 recent increase in wild Russian chum salmon abundance overwhelmed any reduction in Japanese

430 chum salmon abundance. Our study was limited to examining competition with two major Asian 431 stocks of salmon, Japanese hatchery chum salmon and eastern Kamchatka pink salmon, which 432 collectively have comprised up to 52% of Asian pink, chum, and sockeye salmon, and up to 46% 433 of these species in the Bering Sea (based on data in Ruggerone & Irvine, 2018). Western Alaska 434 chum salmon could also compete with other stocks in the Bering Sea, for example, wild pink and 435 sockeye salmon from western Alaska. Future work could incorporate the combined effects of 436 competition from multiple stocks, although including North American stocks would complicate 437 hypotheses about growth during the first year at sea. Despite the limited focus, our study adds to 438 the cumulative evidence indicating that Pacific salmon can experience density dependence in the 439 North Pacific and Bering Sea ecosystems, suggesting that hatchery production cannot continue to

- 440 increase without some detrimental consequences for wild salmon populations.
- 441

442 Author Contributions: TJF led data compilation and analysis, interpretation, and drafting of

443 initial manuscript. MVM conceived the initial study design and contributed to analyses,

444 interpretation and writing of manuscript. EMY, BAA, and MDA assisted with analyses,

445 interpretation, and manuscript preparation.

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## FOG\_12505 Figure legends

**FIGURE 1** Study area (a) and image of chum salmon scale (b). On (a), location of Kuskokwim River chum salmon sampling is shown in black oval, and region of Bering Sea chum salmon bycatch sampling is shown in grey oval. The M2 buoy where spring SST was measured is shown by an asterisk (56.87°N, 164.06°W), and the rectangle shows the region over which summer SST in the central Bering Sea (54.3°-60.0°N, 170.6°-178.1°W) was averaged. On (b) the annual growth increments are shown, depicting the first ( $g_1$ ), second ( $g_2$ ), and third ( $g_3$ ) years of growth at sea; annuli are marked by black bars.

**FIGURE 2** Annual variation, 1965-2013, in variables used to model Kuskokwim River chum salmon growth: a) Japanese hatchery chum salmon abundance; b) eastern Kamchatka pink salmon abundance; c) spring SST at M2 buoy; and d) summer SST in the Central Bering Sea.

**FIGURE 3** Coefficients and standard errors for salmon abundance indices in best-supported models ( $\Delta$ AICc < 2; see Table 1) for the size of Kuskokwim River chum salmon growth increments  $g_1 - g_4$ . c In each pane, coefficients are arranged in descending order of model support, with the top model ( $\Delta$ AICc = 0) at the top of each plot pane. Jpn. = Japanese, E. Kam = eastern Kamchatka. The top model for  $g_1$  contained no salmon abundance index.

**FIGURE 4** Scale growth of Kuskokwim River chum salmon, averaged over both males and females, versus normalized abundance of Japanese chum salmon by growth increment and age. Solid line and filled circles, four-year olds; dashed line and open circles, five-year olds.











Japanese chum abundance (normalized)

**TABLE 1** Coefficients (and their standard errors) from top models for scale growth increments in Kuskokwim River chum salmon. A "—" indicates that a variable was considered in the full candidate model but not included in top model. Variances within ( $\sigma^2_W$ ) and among ( $\sigma^2_A$ ) calendar year at sea are also reported.

	Growth increment								
	$g_1$	$g_2$	$g_3$	$g_4$					
Intercept	-0.03 (0.06)	0.29 (0.07)	0.35 (0.07)	-0.12 (0.06)					
Japanese chum		-0.27 (0.07)	-0.18 (0.07)	-0.17 (0.06)					
Kamchatka pink			-0.09 (0.07)	-0.07 (0.05)					
Chum * pink			0.23 (0.10)						
Sex (male)		0.12 (0.03)	0.38 (0.03)	0.45 (0.04)					
Age (5)		-0.57 (0.03)	-0.97 (0.03)	n/a					
Spring SST			0.09 (0.06)	0.08 (0.05)					
Summer SST	0.20 (0.06)								
$\sigma^2 w$	0.139	0.197	0.148	0.089					
$\sigma^2_A$	0.857	0.684	0.601	0.844					

**TABLE 2** Coefficients (and standard errors) from models for growth increments  $g_2 - g_4$  in Bering Sea chum salmon, as well as number of individuals (n) and results of likelihood-ratio test (LRT) comparing the model that included the period of reduced hatchery releases (2011-2013; 'with') to the model that did not ('without'). Variances within ( $\sigma^2_W$ ) and among ( $\sigma^2_A$ ) calendar year at sea are also reported.

Increment/	$g_2(n = 1,733)$		$g_{3}(n =$	1,451)	$g_4 (n = 316)$		
model	with	without	with	without	with	without	
Intercept	0.84 (0.14)	0.82 (0.13)	1.32 (0.19)	1.36 (0.19)	-0.18 (0.14)	-0.08 (0.13)	
Reduced release period	0.10 (0.11)		-0.20 (0.10)		-0.30 (0.17)		
Sex (male)	0.09 (0.05)	0.09 (0.05)	0.21 (0.05)	0.21 (0.05)	0.16 (0.11)	0.16 (0.11)	
Age (5)	-0.27 (0.04)	-0.27 (0.04)	-0.46 (0.05)	-0.47 (0.05)			
Spring SST	0.16 (0.08)	0.16 (0.07)	0.13 (0.07)	0.13 (0.08)	0.11 (0.12)	0.11 (0.13)	
$\sigma^2 W$	0.060	0.058	0.050	0.064	0.126	0.152	
$\sigma^2_A$	0.915	0.915	0.865	0.865			
LRT ( $\chi^2$ , df, P)	1.02, 1, 1	P = 0.31	4.40, 1, 1	P = 0.04	3.43, 1, 1	P = 0.06	

Brood year	Age 4			Age 5 To			
	Male	Female	Male	Female			
1963	5	1	6	3	15		
1968	15	17	18	19	69		
1969	2	5	1	4	12		
1970	18	16	16	1	51		
1971	18	22	13	13	66		
1972	22	24	19	15	80		
1973	21	18	5	19	63		
1974	23	23	5	19	70		
1976	25	21	25	25	96		
1977	25	25	14	19	83		
1978	21	18	24	24	87		
1979	23	24	21	23	91		
1980	24	25	29	28	106		
1981	25	28	24	25	102		
1982	26	25	25	24	100		
1983	25	31	25	25	106		
1984	25	25	25	25	100		
1985	25	25	25	23	98		
1986	25	25	24	25	99		
1987	23	24	25	25	97		
1988	25	25	25	24	99		
1989	22	25	24	25	96		
1990	25	25	25	24	99		
1991	25	25	24	25	99		
1992	25	25	25	25	100		
1993	25	25	25	25	100		
1994	23	24	23	23	93		
1995	22	16	18	23	79		
1996	15	22	22	25	84		
1997	23	21	20	19	83		
1998	18	22	22	24	86		
1999	25	24	24	22	95		
2000	25	25	25	17	92		
2001	25	25	25	25	100		
2002	25	26	6	2	59		
2003	22	25	25	25	97		
2004	25	25	25	26	101		
2005	25	25	28	34	112		
2006	25	25	31	26	107		
2007	25	25	35	25	110		
2008	25	28	13	17	83		
2009	25	26	33	35	119		
2010	33	25	0	0	58		

**TABLE S1** Number of Kuskokwim River chum salmon scales measured from 1963, 1968-2010 by brood year, age at maturity, and sex.

Brood			Femal	e				Male	e		Grand
year	$g_l$	$g_2$	$g_3$	$g_4$	Total	$g_l$	$g_2$	<b>g</b> 3	$g_4$	Total	Total
1996	7	7	7	7	7	8	8	8	8	8	15
1997	21	21	21	6	21	22	22	22	6	22	43
1998	24	24	20	5	24	26	26	23	8	26	50
1999	57	57	49	11	57	69	69	56	10	69	126
2000	29	29	24	4	29	56	56	51	11	56	85
2001	84	84	74	23	84	91	91	81	22	91	175
2002	67	67	48	13	67	49	49	40	9	49	116
2003	32	32	29	6	32	42	42	36	8	42	74
2004	31	31	27	16	31	37	37	33	15	37	68
2005	55	55	51	4	55	72	71	71	8	72	127
2006	54	54	36	19	54	64	64	47	24	64	118
2007	80	80	76	5	80	80	79	72	10	80	160
2008	29	29	14	6	29	41	41	17	10	41	70
2009	58	57	55	8	58	57	57	54	6	57	115
2010	34	34	31	5	34	53	53	50	8	53	87
2011	66	66	57	14	66	68	68	60	16	68	134
2012	104	104	74		104	120	120	95		120	224
2013	21	21			21	36	36			36	57
2014						1				1	1

**TABLE S2** Number of Bering Sea chum salmon scales measured by brood year, sex, and growth increment. Note that each growth increment is tallied separately, so a 5-year old individual would be represented in each of the columns  $g_1 - g_4$ .

**Table S3.** Model selection for counterfactual  $g_1$  and focal growth increments  $g_2 - g_4$ , for Kuskokwim River chum salmon. Models within 2  $\Delta$ AICc units are shown, as well as the null (random effect only) model. SprSST, spring SST; SumSST, summer SST; JpnChum, index of Japanese chum salmon abundance; KamPink, index of Kamchatka pink salmon abundance. Sign of coefficient is indicated by +/- for each of the modeled fixed effects.

Modeled variables	df	logLik	ΔAICc
gl (counterfactual)			
SumSST	4	-5065.24	0
SumSST + KamPink	5	-5064.41	0.35
SumSST + JpnChum	5	-5064.77	1.06
SumSST + JpnChum + KamPink	6	-5064.16	1.85
Sex + SumSST	5	-5064.23	1.98
-SprSST + SumSST	5	-5065.24	1.999
Null (random effect only)	3	-5070.90	9.32
<u>g2</u>			
Sex + Age - JpnChum	6	-4666.56	0
Sex + Age - SprSST - JpnChum	7	-4666.21	1.29
Sex + Age + SumSST - JpnChum	7	-4666.38	1.64
Sex + Age - JpnChum + KamPink	7	-4666.41	1.69
Sex + Age - SprSST + SumSST - JpnChum	8	-4666.43	1.74
Null (random effect only)	3	-4878.61	418
<u>g</u> <sub>3</sub>			
Sex + Age + SprSST + (-JpChum*-KmPink)	9	-4419.48	0
Sex + Age + (-JpnChum*-KmPink)	8	-4420.80	0.64
Sex + Age + SprSST + SumSST + (-JpnChum*-KamPink)	10	-4419.33	1.72
Sex + Age - JpnChum	6	-4423.35	1.73
Null (random effect only)	3	-5096.71	1342
<u>g4</u>			
Sex + SprSST - JpnChum - KamPink	7	-2422.70	0
Sex + SprSST - JpnChum	6	-2423.95	0.49
Sex - JpnChum - KamPink	6	-2424.09	0.77
Sex + SprSST + (-JpnChum*-KamPink)	8	-2422.10	0.84
Sex - JpnChum	5	-2425.52	1.61
Sex + SumSST - JpnChum - KamPink	7	-2423.53	1.67
Sex + SprSST + SumSST - JpnChum - KamPink	8	-2422.66	1.96
Null (random effect only)	3	-2480.56	108



**FIGURE S1.** Temporal trends in average growth by increment and age for Kuskokwim River chum salmon. a)  $g_1$ ; b)  $g_2$ ; c)  $g_3$ ; d)  $g_4$ . Solid line, four-year olds; dashed line, five-year olds.



**FIGURE S2.** Temporal trends in average growth by increment Bering Sea chum salmon. a)  $g_1$ ; b)  $g_2$ ; c)  $g_3$ ; d)  $g_4$ .