

Wintertime CO₂, CH₄ and CO emissions estimation for the Washington DC / Baltimore metropolitan area using an inverse modeling technique

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Abstract

2
3 Since greenhouse gas mitigation efforts are being mostly implemented in cities, the
4 ability to quantify emission trends for urban environments is of paramount importance.
5 However, previous aircraft work has indicated large daily variability in the results. Here
6 we use measurements of CO₂, CH₄ and CO from aircraft over five days within an in-
7 verse model to estimate emissions from the D.C./Baltimore region. Results show good

8 agreement with previous estimates in the area for all three gases. However, aliasing
9 caused by irregular spatiotemporal sampling of emissions is shown to significantly im-
10 pact both the emissions estimates and their variability. Extensive sensitivity tests allow
11 us to quantify the contributions of different sources of variability and indicate that daily
12 variability in posterior emissions estimates is larger than the uncertainty attributed to
13 the method itself (i.e. 17% for CO₂, 24% for CH₄ and 13% for CO). Analysis of hourly
14 reported emissions from power plants and traffic counts shows that 97% of the daily
15 variability in posterior emissions estimates is explained by accounting for the sampling
16 in time and space of sources that have large hourly variability and, thus, caution must
17 be taken in properly interpreting variability that is caused by irregular spatiotemporal
18 sampling conditions.

19 Introduction

20 As cities move toward mitigating their carbon footprints, estimating their emissions using
21 atmospheric observations is a valuable way to assess the efficacy of mitigation policies. Recent
22 work¹⁻⁷ has already demonstrated the capability of top-down (atmospheric measurement-
23 based) estimation methods to inform bottom-up inventory methods for some greenhouse
24 gases (GHGs). On regional and urban scales, top-down methods have been shown to be
25 effective at estimating emissions using either tower-based or aircraft-based concentration
26 measurements.⁸⁻¹²

27 Atmospheric trace gas concentration measurements from airborne platforms have been
28 used extensively to estimate emissions from a region. Both oil and gas basins and urban
29 regions have been studied using mass balance methods,¹³⁻¹⁷ including the Washington D.C./
30 Baltimore metropolitan area.^{11,12} Researchers have also used aircraft observations with trans-
31 port models in an inversion framework to estimate emissions at regional,¹⁸⁻²¹ urban^{22,23} and
32 local scales.²⁴

33 Several studies have investigated the source of daily variability in aircraft-based top-down

34 emissions estimates for a given region. Variability in estimated emission rates has previously
35 been attributed to uncertainty in the mass balance methodology, which would confound or
36 obscure real emissions changes.^{25,26} More recent work using airborne measurements over oil
37 and gas fields has shown that temporal variability in emissions must be considered when
38 interpreting estimates from single-day flights, however. Lavoie et al.²⁷ found significant
39 temporal variability in single source emissions of methane (CH_4) from the Eagle Ford oil and
40 gas production basin in Texas, while Schwietzke et al.²⁸ investigated the effect of episodic
41 CH_4 emissions from natural gas facilities on the regional mass balance estimates in the
42 Fayetteville Shale.

43 In this study, we use observations collected during five aircraft flights over a two-week
44 period in February 2016 within a Bayesian inversion framework to: 1) estimate emissions
45 of CO_2 , CH_4 and CO from the cities of Washington D.C. and Baltimore, MD, (Fig. 1), 2)
46 quantify the uncertainty, and its sources, in each day's emissions estimate and, 3) explain
47 the cause for the observed daily variability in the estimated emissions.

48 To this end, we use an ensemble of inversions where prior emissions, transport model
49 and observation dataset were varied. Ensemble spread and correlations between six trans-
50 port models were used to construct the full model-data mismatch covariance matrix, and
51 the background mole fraction was first estimated by using sensitivities to nearby outside
52 sources and then further optimized within the inversion. Additionally, sensitivity tests were
53 conducted investigating the impacts of background choice, omitting correlations in the trans-
54 port error covariance matrix and changing the magnitude of the prior emission errors. We
55 use the inversion ensemble and sensitivity tests to quantify the different sources of variabil-
56 ity and, thus, understand the uncertainty inherent in the inverse methodology. We then
57 investigate daily variability in estimated emissions and to what extent this variability can
58 be explained by aliasing caused by irregular sampling of spatial and temporal variability in
59 large sources within the study domain.

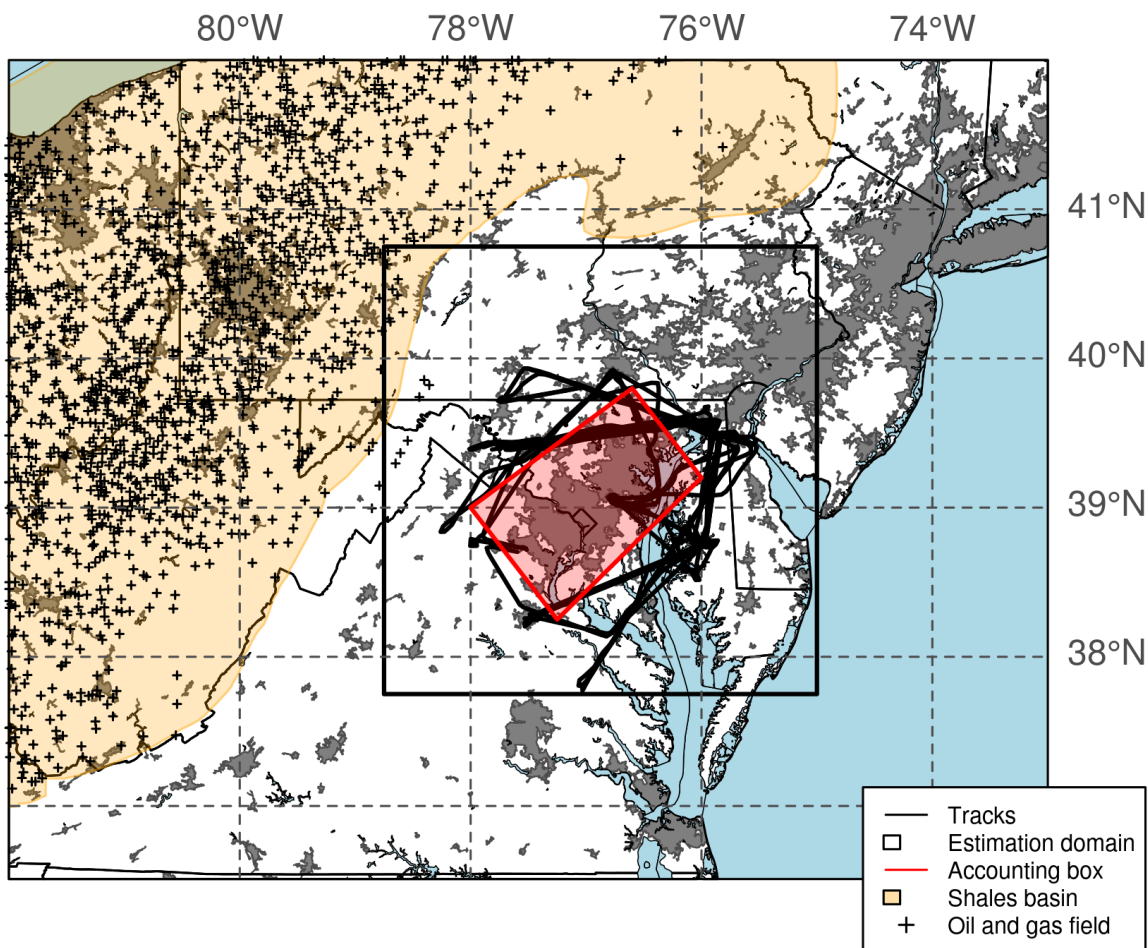


Figure 1: Computational domain (0.03° resolution) showing the inversion domain (black rectangle) and the outer domain (entire map) used to account for nearby outside sources. Flight tracks, Census-designated urban areas (gray shaded regions), the Marcellus, Devonian (Ohio) and Utica shale plays in the Appalachian basin and locations of the geometric center (centroid) of the oil and gas fields are also shown.²⁹ Total emissions are reported here within the accounting box (red polygon) defined by the corners: (39.80°N , 76.60°W), (39.00°N , 78.00°W), (38.25°N , 77.25°W) and (39.20°N , 76.00°W).

60 Methods

61 Observations

62 Trace gas observations from two airborne platforms were used in this study: Purdue Univer-
 63 sity's Beechcraft Duchess, housing the Airborne Laboratory for Atmospheric Research, or
 64 ALAR, (Purdue) and the University of Maryland's Cessna 402B research aircraft (UMD).

65 The two aircraft flew simultaneously for 5 days, mostly during afternoon hours, collecting
 66 trace gas mole fraction and meteorological data along transects at different altitudes that
 67 covered the full depth of the PBL (Fig. 1 and SI for further details). To determine the effect
 68 of withholding observations from the inversion system, we alternatively used CO₂ and CH₄
 69 observations from both aircraft, the UMD aircraft alone, or the Purdue aircraft alone, as
 70 part of the ensemble of inversions. Purdue did not measure CO, thus the CO inversions used
 71 only UMD observations.

72 Bayesian Inversion Framework

73 We estimate trace gas emissions using a Bayesian inverse analysis^{30,31} as in Lopez-Coto et
 74 al.³² Optimum posterior estimates of fluxes are obtained by minimizing the cost function J :

$$J(\mathbf{x}) = \frac{1}{2} \left[(\mathbf{x} - \mathbf{x}_b)^T \mathbf{P}_b^{-1} (\mathbf{x} - \mathbf{x}_b) + (\mathbf{H}\mathbf{x} - \mathbf{y})^T \mathbf{R}^{-1} (\mathbf{H}\mathbf{x} - \mathbf{y}) \right] \quad (1)$$

75 where \mathbf{x}_b is the first guess or a priori state vector, \mathbf{P}_b the *a priori* error covariance
 76 matrix which represents the uncertainties in our *a priori* knowledge about the fluxes and \mathbf{R}
 77 the error covariance matrix, which represents the uncertainties in the observation operator
 78 \mathbf{H} and the observations \mathbf{y} , also known as model-data mismatch. The observation operator
 79 \mathbf{H} is constructed using the sensitivity of observations to surface fluxes, or footprints (units:
 80 ppm $\mu\text{mol}^{-1} \text{m}^2 \text{s}$) generated with a transport model. Here we modify the formulation to
 81 include optimization of the background in the inversion (see SI for details).

82 Transport Models

83 In order to generate an ensemble of transport models and therefore better represent the
 84 uncertainties, NOAA's Hybrid Single-Particle Lagrangian Integrated Trajectory dispersion
 85 model (HYSPLIT)³³ was driven with 5 different meteorological products: the High Resolu-
 86 tion Rapid Refresh (HRRR) NOAA operational forecast product³⁴ and 4 configurations of

87 the Weather Research and Forecasting model (WRF³⁵) provided by the National Center for
88 Atmospheric Research (NCAR) that included 4 different PBL parametrizations, 2 sources of
89 initial and boundary conditions and the inclusion of the Building Energy Parameterization
90 (BEP) urban canopy model in one of the configurations. In addition, the vertical mixing
91 option in HYSPLIT also varied (Table S1 and SI for details).

92 **Emissions Inventories**

93 Nine CO₂ emissions inventories were used in the inversion to investigate the resultant vari-
94 ability in the posterior emissions (Table S2). Four of them (Vulcan (VU³⁶), ODIAC (OD³⁷),
95 FFDAS (FF³⁸) and ACES (AC³⁹) are existing anthropogenic CO₂ inventories but for a
96 different year; one provided only on-road emissions (DARTE (DA⁴⁰)); one is the mean of
97 the previous five (EB); and the rest (flat (FL) and simple (SP³²)) are constructed here to
98 complement the ensemble of prior fluxes. In addition, we use the ACES mean for February
99 between 12 - 19 EST (AC2). Since DARTE only provided on-road emissions, a simple calcu-
100 lation of urban emissions was used to complement it. CH₄ prior emissions were represented
101 using EPA's gridded inventory (EP) for 2012,⁴¹ EDGAR v4.3.2⁴² for 2012 (EG), the mean
102 of the previous two (EB), and a flat prior (FL). For CO we use EDGAR v4.3.2 (EG),⁴³
103 the National Emissions Inventory (NEI) for 2011 from EPA (NI,⁴⁴), the annual mean ACES
104 inventory (AC as in the CO₂ case) scaled using the mean observed $\Delta\text{CO}:\Delta\text{CO}_2$ ratio (6.18
105 ppb/ppm) and, again, a flat prior (FL).

106 **Background Determination**

107 Properly accounting for the background is critical for the inversion as the flux correction is
108 based on the observed enhancements above the background value. The impact of upwind
109 sources can be important especially in areas such as the one under study here, where multiple
110 sources exist in the surroundings (Fig. 1). Thus, we estimated the contribution from outside
111 the domain using a Lagrangian approach by convolving footprints from a reduced set of our

112 ensemble of transport models and prior fluxes. We extended the domain to the full extent
113 shown in Fig. 1. The full background was then represented as the ensemble mean of the
114 contribution from outside of the domain of interest (\mathbf{y}_{oc} , time-varying along the track) plus
115 the long-range background (\mathbf{y}_{lr} , constant for a given flight). This methodology provided a
116 time varying *a priori* background that included uncertainties that was then further optimized
117 in the inversion (SI).

118 **Error Covariances**

119 **i) Prior Flux Error Covariance**

120 The prior flux error covariance represents the uncertainties in the prior estimation of the
121 fluxes. Although bottom-up CO₂ emissions estimates are made on global and national scales
122 with small uncertainties, considerable errors are introduced when the emissions are disaggre-
123 gated to grid cells, due to the usage of proxies to spatially distribute emissions.⁴⁵ Reported
124 errors at grid cell levels range from 4% to more than 190%, averaging about 120%.⁴⁶ For CH₄
125 and CO it is likely that the errors at grid cell levels are even larger than for CO₂ because
126 of the less well-known characteristics of these species' sources. Given these reported uncer-
127 tainties at grid cell levels, we use a value of 100% of the grid cell emissions as uncertainty in
128 this work for all the prior inventories and gases with the exception of FFDAS where we use
129 a scaled up version of the provided uncertainties and the EB case for CO₂ where we use the
130 standard deviation of the ensemble at each pixel to represent the uncertainties. In all cases,
131 a covariance exponential model in space was assumed. (See SI for details)

132 **ii) Outside Contribution (background) Prior Error Covariance**

133 We consider a double exponential model, in space and time, to represent the error covari-
134 ance of the outside contribution (\mathbf{y}_{oc}) along the track. The diagonal is populated with the
135 uncertainty of the initial guess outside contribution based on the variance from the different
136 transport models and prior fluxes (SI).

137 **iii) Model Error**

138 The model-data mismatch error covariance was assumed to have three independent con-
139 tributions: 1) uncertainty in the observations, 2) uncertainty in the long-range background
140 concentration and 3) uncertainty in the transport model representation. The uncertainties in
141 the observations have their origin in the measurement uncertainties and the representativity
142 of the assigned mean to the averaging period (one minute in our case). This contribution is
143 not correlated and thus the covariance was considered diagonal. The long-range background
144 (\mathbf{y}_{lr}) determination also introduces uncertainty into the system. This contribution was also
145 assumed to be uncorrelated. Lastly, the transport model uncertainty is complex with several
146 previously published methods for its determination. Here we tested two methods, both based
147 on the ensemble of transport models. First, we tested a diagonal covariance populated with
148 the inter-model variance simulated using the same surface fluxes (the prior emissions in each
149 inversion case) in all the transport models similar to Engelen et al.⁴⁷ and Desroziers et al.⁴⁸
150 As stated in Engelen et al.,⁴⁷ this estimate can be too large for some models and too small
151 for other models, thus, in order to better represent the fidelity of each model and for each ob-
152 servation, we weighted the inter-model standard deviation with the relative error computed
153 by using the wind measurements from the aircraft. This definition of the transport model
154 error covariance assumes there are no correlations in space and time which is unlikely to be
155 true. Therefore, for the second method, which was used in the main ensemble of inversions,
156 we computed the correlations between the different transport models and included them in
157 the covariance matrix, leaving the first method as a sensitivity test (see SI for details).

158 **Sensitivity Analysis**

159 As described in the previous sections, the main inversion ensemble was composed by different
160 prior emissions (9 for CO₂, 4 for CH₄ and 4 for CO), 6 transport models and 3 combinations of
161 the observations for 5 flight days, totaling 1,290 inversions (810 + 360 + 120). This inversion
162 ensemble was configured with the background and prior and transport error covariances

163 choices that are most reasonable for the analysis. However, in order to additionally test
164 the sensitivity of the posterior estimates to inversion setup choices that might not be as
165 appropriate, we also investigated the effects of changing the background determination, the
166 transport error covariance, and the prior flux error covariance, separately from the main
167 inversion ensemble. Specifically, for the background test, we performed the inversion 1)
168 without optimizing the Lagrangian background, 2) scaling the Lagrangian background, and
169 3) selecting a single constant value along the track as background defined by the 1st, 5th
170 or 10th percentile, to compare with our base case of optimizing the background (OBC1).
171 For the scaled background case, a single scaling factor for each flight was applied to the
172 background time series. This scaling factor was the ratio of posterior to prior emissions for
173 the inversion case where the background was not optimized or scaled. We also tested the
174 impact of using only a diagonal transport error covariance as well as reducing and increasing
175 the uncertainty in the prior fluxes (50%, 100% and 200%). This sensitivity test resulted in
176 a total of 12 cases with 15,480 individual inversions, (Table S3).

177 Both the main inversion ensemble and the sensitivity test were analyzed in the same
178 fashion, grouping by cases (prior, transport, day, observation dataset or sensitivity case)
179 and then computing the mean and quantiles as shown in Figs. 2, S7, S10, S13 and S16.
180 The variability associated with each grouping was then computed as the standard deviation
181 among each case's mean value.

182 **Normalized Observed Emissions**

183 We construct an analysis to investigate whether the hourly variability of the energy gen-
184 eration and traffic sectors' emissions, combined with the specific flight pattern on a given
185 day, can explain the daily variability in the posterior CO₂ estimates. Both of these sources
186 have publicly available data at the hourly level: Continuous Emissions Monitoring System
187 (CEMS⁴⁹) data for power plants and Travel Monitoring Analysis System (TMAS⁵⁰) data
188 for traffic counts. First, we sum all the power plant emissions and traffic counts within the

189 footprint (we use the ensemble mean footprint as a mask) of each observation used in the
190 inversion and within the defined accounting box. We match the hourly power plant emissions
191 and traffic counts with the observation time, accounting for transport time to the point of
192 the observation at hourly temporal resolution. Then we average this value (the sum of all
193 traffic counts or powerplant emissions within each footprint) over all observations in each
194 flight for each of the five flights. Using an average allows us to account for the difference
195 in the number of observations per flight. Because traffic counts and power plant emission
196 rates are in different units, we define the normalized observed emissions (nOE), allowing for
197 the combination of the two sectors. We normalize counts and power plant emissions each to
198 their respective campaign mean so that the campaign mean is equal to one. Furthermore,
199 we use the relative contribution of the different sectors in the ACES 2011 annual mean³⁹
200 within the defined accounting box to construct the normalized observed emissions (nOE) for
201 each flight as follows:

$$nOE_i = f_e \frac{CEMS_i}{\langle CEMS \rangle} + f_r \frac{TMAS_i}{\langle TMAS \rangle} + 1 - (f_e + f_r) \quad (2)$$

202 In the above definition, i is the index indicating the flight, f_e is the contribution of the
203 electricity production sector (16%) and f_r is the contribution of the traffic emissions (46%)
204 in ACES. The last term of Eq. 2 represents the remainder of anthropogenic CO₂ emission
205 sectors. By this construction, the mean nOE for the campaign is also equal to 1.

206 Results

207 In the following subsections we present the main results of the analysis and discuss the
208 variability and uncertainty of the emissions estimates. In this context, the terms variability
209 and uncertainty are not used as synonyms. Rather, we use the term variability to describe
210 how a property (posterior total emissions for the most part) changes (varies) with respect
211 to different variables like time, space or model choices. The term uncertainty refers to the

212 ability of the inverse method to represent the measurand, and it combines all sources of
213 variability for a single day's estimate.

214 Emissions Rates

215 Our mean estimates for the defined accounting box are 87 ± 28 kmol s⁻¹ for CO₂, 0.42 ± 0.12
216 kmol s⁻¹ for CH₄ and 0.59 ± 0.16 kmol s⁻¹ for CO (mean $\pm 1\text{-}\sigma$) where the bounds presented
217 here represent the posteriors' daily variability. Ren et al.¹¹ using a mass balance method,
218 estimated emission rates of 96 kmol s⁻¹ for CO₂, 0.57 ± 0.28 kmol s⁻¹ for CH₄ and $0.55 \pm$
219 0.27 kmol s⁻¹ for CO using the same flight observations as this study. In addition, Salmon et
220 al.¹² estimated a CO emission rate (also using a mass balance method) of 0.54 ± 0.47 kmol
221 s⁻¹ in February 2015. Our estimates are consistent with these within $1\text{-}\sigma$ uncertainties for
222 both methods.

223 The applied inversion methodology corrected the prior inventories (Fig. 2a,c,e) by quite
224 different amounts leading to consistent results in the posterior emissions, with variability due
225 to choice of prior of 11%, 13% and 6% (or 9.6, 0.055 and 0.035 kmol s⁻¹) for CO₂, CH₄ and
226 CO respectively ($1\text{-}\sigma$), significantly lower than the variability of the prior values themselves
227 (flat prior included), 41%, 65% and 87% (or 20.8, 0.097 and 0.38 kmol s⁻¹). The flat (FL)
228 prior led to the largest range and IQR for all of the three gases due to the loose constraint it
229 imposed on the inversion. For CO₂, the FFDAS³⁸ prior (FF) resulted in the lowest posterior
230 estimates as well as the lowest range and IQR due to the low prior uncertainty assigned,
231 making it hard for the inversion to deviate from the prior values. For CH₄, the inversions
232 using the 2012 EPA gridded inventory⁴¹ (EP) as a prior provided the lowest estimates,
233 probably due to the lower prior emissions allocated into the urban areas and, therefore,
234 lower prior uncertainties, making it harder to correct those areas. For CO, the scaled ACES
235 inventory (AC) led to the lowest estimates. Variability due to transport model choice was
236 15% for CO₂, 13% for CH₄ and 16% for CO ($1\text{-}\sigma$), (Figs. S7c, S10c and S13c). We note
237 that HR and MY2 provided the highest and lowest estimates respectively, while MY and BL

238 had the most variable results. The observation dataset choice impacted the results the least,
239 with only a 6 % standard deviation of the mean for CO₂ and 10% for CH₄ with very similar
240 range and IQR for each of the three cases (Figs. S7d, S10d). In contrast to the relatively
241 small effect of varying these three model choices (prior, transport model, and observation
242 dataset), the daily variability of the estimates was 33% for CO₂ and 28% for CH₄ and CO
243 (1- σ) (Figs. 2b,d,f). The mean estimates for each day do not overlap with the IQR of the
244 other days and while the CO₂ and CO estimates follow a very similar pattern (as they have
245 similar sources), they differ from that of CH₄. In addition, the coefficient of determination
246 between the daily emission estimates for the three gases is $r^2=0.90$ for CO vs CO₂, $r^2=0.40$
247 for CO₂ vs CH₄ and $r^2=0.19$ for CO vs CH₄. This suggests that the inversion is actually
248 providing different estimates for each day, and that the posterior differences between days
249 are not only the result of choices in the model set up.

250 The spatial distribution of the averaged CO₂ posterior emissions for each prior case
251 shows that most of the emissions are coming from the urban areas, even in the flat prior
252 case (Fig. S8). The results show that the roads (traffic emissions) and fine spatial scale
253 features are only resolved in modeling results when high resolution inventories are used as
254 the prior emissions. The inversion was able to spatially differentiate between the cities of
255 Baltimore and Washington DC, correcting their emissions differently (Fig. S9): emissions
256 from Washington, DC were adjusted upward in all cases while those from Baltimore were
257 corrected downward in the cases of AC, AC2 and VU and only slightly upward for the rest.

258 The spatial distribution of the averaged CH₄ posterior emissions (Fig. S11) indicates that
259 while some emissions are from urban areas, significant emissions occur NNE and NNW of the
260 Washington - Baltimore metropolitan area as well, which is different than for CO₂. All the
261 CH₄ priors were corrected upwards indicating an overall underestimation of emissions in the
262 existing inventories (Fig. S12), with the strongest corrections applied to point sources outside
263 urban areas. However, the urban areas were also corrected upward, with this correction being
264 larger for EP than for EG or EB cases. Our posterior mean ratio to the 2012 EPA gridded

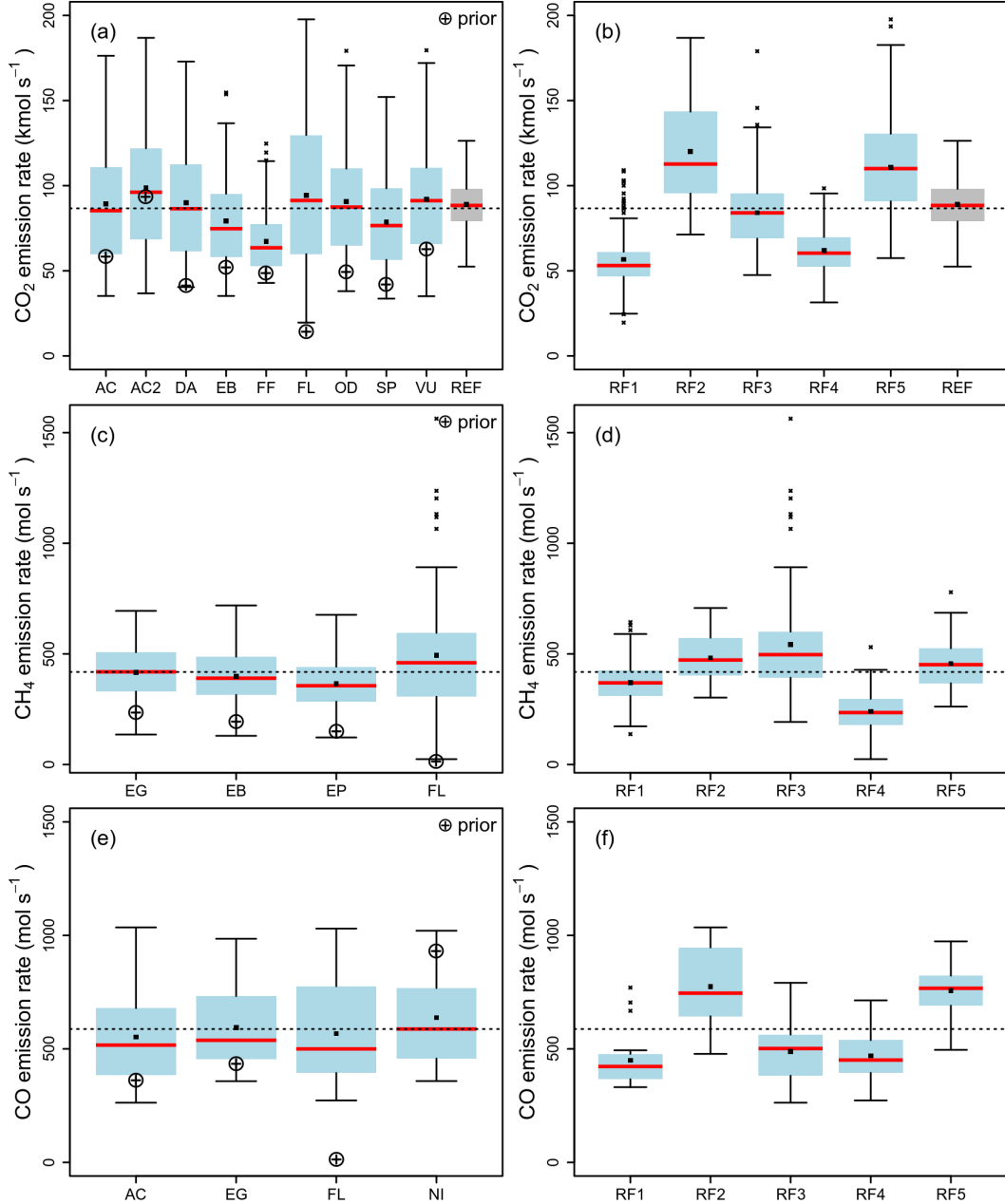


Figure 2: Boxplots of the total CO_2 , CH_4 and CO estimated emission rate within the accounting box grouped by: (a,c,e) the different inventories used as priors and (b,d,f) the different research flights. The grey bar in panels (a,b) are the values provided by ACES, scaled to totals of 2016, for February between 12 - 19 EST (referred as REF). Blue bars indicate the 25th to 75th range, whiskers the range up to 1.5 times the IQR, x's the outliers ($> 1.5 \times \text{IQR}$), red line the median, square markers the mean and the dotted line the posterior mean. The circled pluses in panel (a,c,e) represent each prior's total emissions. (See methods section and Tables S1 and S2 for abbreviations)

265 inventory (EP),⁴¹ 2.73 ± 0.76 , is in very good agreement with Ren et al.'s¹¹ estimate of 2.8
266 times the EPA values for the same region.

267 The spatial distribution of the mean posterior CO fluxes (Fig. S14) indicates that the
268 CO emissions largely originate in the urban areas, as they do for CO₂. In addition, the
269 correction (Fig. S15) is mostly applied in the urban cores, increasing the fluxes for AC, FL
270 and EG while strongly decreasing the emissions for NI case. Due to the construction of AC
271 for CO (using the ACES CO₂ inventory scaled by mean observed $\Delta\text{CO}:\Delta\text{CO}_2$ ratio), power
272 plant emissions were present in the prior, while we expect the power plants ratio to be small
273 compared to other sources. The inversion was able to correct down at least a few of them
274 (blue dots in Fig. S15a). The NEI CO prior case was strongly corrected down over all urban
275 areas, even in Philadelphia, indicating that the inversion is able to correct underestimation
276 as well as overestimation in the prior. The NEI CO overestimation has been extensively
277 reported in the literature;^{22,23,51} specifically in the DC/Baltimore region a close to 50 %
278 overestimation of the NEI CO inventory has been reported,^{11,12} similar to our result of 58%.

279 Sensitivity Analysis

280 For all three gases, the diagonal model-data mismatch error covariance (EDC1) provided
281 larger emissions estimates than the equivalent full covariance case (C1) (Fig. S16). In addi-
282 tion, the range and IQR within each case was larger with the diagonal covariance indicating
283 that the off-diagonal terms played an important role in limiting the number of possible so-
284 lutions. The background selection impacted both the mean estimates and the range and
285 IQR indicating that incorrect background specification can bias the estimation results. The
286 prior flux error sensitivity test showed that posterior emissions estimates were larger when
287 prior uncertainties were doubled, and the range and IQR within each case was also larger
288 indicating a potential over-fitting problem. When prior uncertainties were halved from the
289 base case, the estimates were lower and less variable, indicating the solutions were more
290 constrained by the prior fluxes than by the observations. This effect was similar to the FF-

291 DAS prior case for CO₂, for which the prior uncertainties were likely too small. Despite the
292 differences described above, the variability of the mean across the sensitivity analysis cases
293 remained relatively low, at 11% for CO₂, 17% for CH₄ and 8% for CO.

294 **Special Case: Flat Prior**

295 The inversions using a spatially flat prior (FL) were able to provide mean totals close to those
296 in which an inventory prior was used for all three gases (Fig. 2). This result demonstrates the
297 potential for using aircraft measurements to estimate an overall city-wide emission rate for
298 a location where a spatially explicit inventory or other emissions information is unavailable.
299 However, we have also shown that the range and IQR in the flat prior case was the greatest
300 among all the prior cases, implying that when using a flat prior sampling more time periods
301 (i.e. using more observations) is required to provide confidence in the estimates. The spatial
302 distribution of the campaign-averaged posterior fluxes for CO₂, CH₄ and CO (Fig. S17) is
303 consistent with the results obtained with the other priors as well. For example, CO₂ and CO
304 show very similar spatial distributions with most of the emissions originating in the urban
305 areas while CH₄ shows a broader spatial distribution. Note that these spatial patterns are a
306 result of a campaign of 5 days with winds coming from different directions (Fig. S1), leading
307 to a good triangulation of the source locations.

308 **Discussion: Uncertainty and Sources of Variability**

309 **Method Combined Uncertainty**

310 We were able to disentangle and quantify the different sources of variability present in the
311 inversion-based emissions estimates and found that the largest source of variability in the
312 retrieved emissions is the daily variability. In the following analysis, we omit the daily
313 variability because the goal is to understand the uncertainty we expect in each day's estimate
314 and whether the daily variability is likely to be caused by general uncertainty in the method.

315 Here we assume each source of variability is independent of the others, so that the variances
316 can be summed to estimate the method uncertainty in each day’s estimate. We note that the
317 assumption of independence is not likely to be true and therefore this uncertainty estimation
318 might be biased due to not considering the correlations among them. In addition, the
319 ensemble construction (transport, priors, observation dataset, covariances and background)
320 might impact the ensemble spread and therefore might not be the true uncertainty in the
321 method but it does, however, provide an indication of the likely variability introduced by
322 the different model choices.

323 Three different cases are shown in Table 1 for estimating combined uncertainties. The
324 Combined Uncertainty 1 case considers all sources of variability tested in the inversion.
325 However, we believe that two transport models are outliers that suffer from improper mixing
326 and resulted in biased estimations. The highest retrieved fluxes are obtained consistently
327 using the HR configuration, suggesting that this configuration is too dispersive, although
328 more research is needed to be more certain. The lowest posterior estimates consistently
329 correspond to the configuration including the experimental vertical mixing parametrization
330 (MY2), indicating that this method may under-predict vertical mixing. Removing these two
331 outlier configurations reduces the variability due to transport model choice to 7% for CO₂,
332 10% for CH₄ and 8% for CO; these are used to calculate the Combined Uncertainty 2 case.
333 Because the flight tracks are different for each aircraft, the variability due to the observation
334 dataset may also be affected by the spatial and temporal distribution of the sources being
335 measured, so we remove this variability to also calculate the Combined Uncertainty 3 case.

336 **Daily Variability in Estimated Emissions: Aliasing**

337 The daily variability in our posterior emissions from the inversion ensemble was 33% for CO₂
338 and 28% for CH₄ and CO (Table 1). This variability is larger than each individual source of
339 variability as well as the three cases of the combined uncertainties as shown above, although
340 for CH₄ the two are comparable. In order to better understand the origin of this variability

Table 1: Sources of variability and combined uncertainty.

Source of uncertainty	ϵ CO ₂ (%)	ϵ CH ₄ (%)	ϵ CO (%)
daily	33	28	28
prior	11	13	6
transport	15	14	16
transport no outliers	7	10	8
observation dataset	6	10	6 ^a
sensitivity	11	17	8
Combined Uncertainty 1 (prior, transport, dataset and sensitivity)	22	27	20
Combined Uncertainty 2 (prior, transport no outliers, dataset and sensitivity)	18	26	14
Combined Uncertainty 3 (prior, transport no outliers and sensitivity)	17	24	13

^aCO variability due to the observation dataset is assumed to be the same as for CO₂.

341 in the estimates, we conducted an analysis of the temporal variability and spatial sampling
342 of the two largest sources of CO₂ in the accounting box, according to the ACES inventory:
343 energy generation and on-road traffic.³⁹ Thirteen power plants and 87 counting stations were
344 used within the accounting box (Fig. S18). Both of these sources have significant variability
345 throughout a single day, with traffic counts in the area varying by up to a factor of 20 between
346 night time and evening rush hour depending on the location (Fig. S19, S20), and individual
347 power plant reported emissions varying up to a factor of two within a single day, but even
348 more between days as they sometimes shut down completely (Fig. S21). If daily means
349 of these emissions are investigated, neither the average daily mean of powerplant emissions
350 nor the average daily mean of traffic counts correlates with the daily mean emissions from
351 our inversion posterior. However, the daily variability in the posterior estimates can indeed
352 be explained using an analysis that considers the hourly variability of these two sectors'
353 emissions, combined with the specific flight pattern on a given day. We define each day's
354 *normalized observed emissions* (nOE, Eq. 2) using powerplant and traffic count data to
355 conduct this analysis.

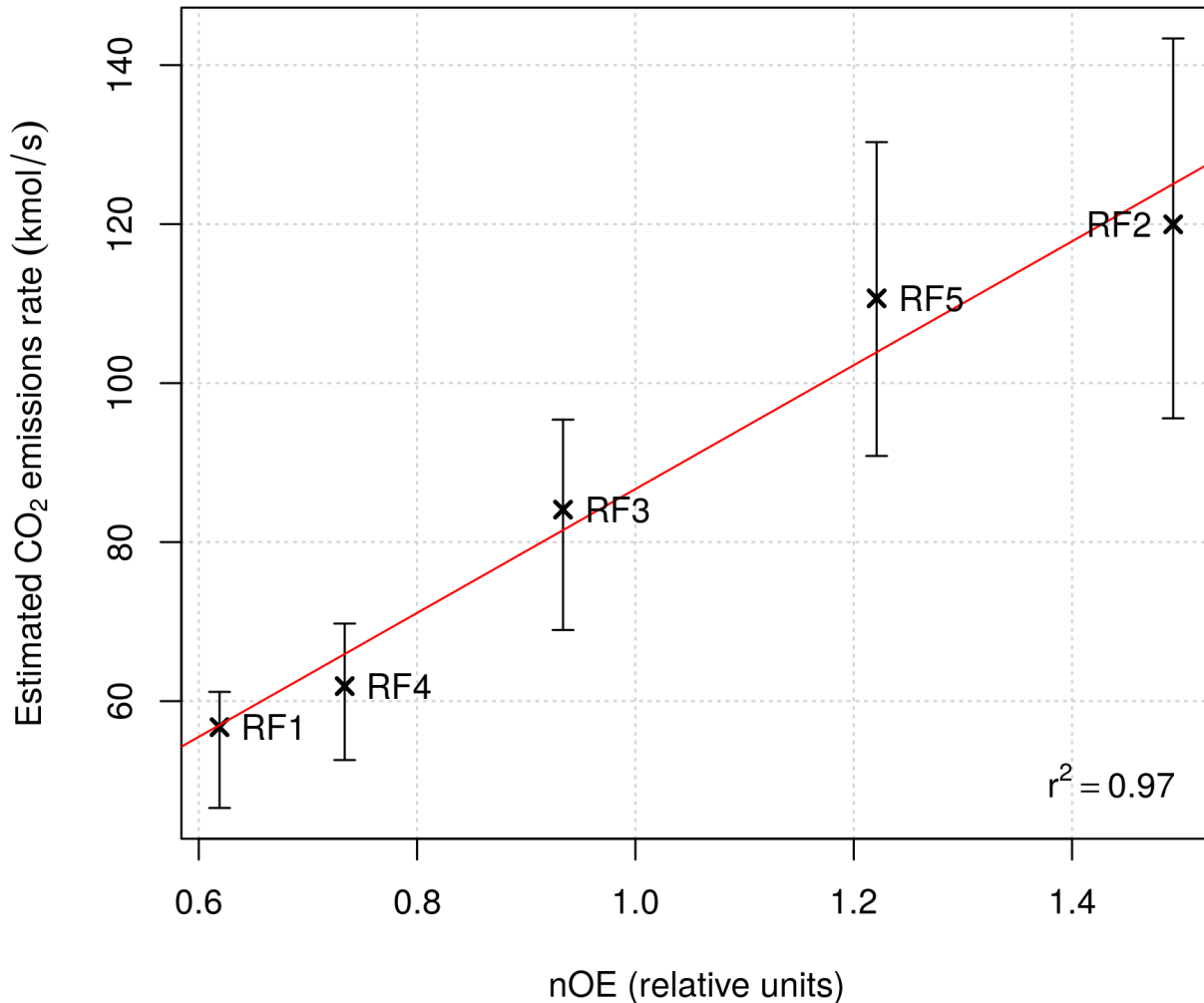


Figure 3: Estimated CO₂ emission rates (kmol s⁻¹) for each research flight as a function of the normalized observed emissions (nOE) computed using CEMS and TMAS hourly data. Errors bars correspond to 25th and 75th percentiles of the ensemble of inversions for each day. Red line indicates the linear fit.

356 Fig. 3 shows the daily mean estimated CO₂ emissions from the inversion as a function
 357 of the *normalized observed emissions* (nOE), with error bars representing the 25th and 75th
 358 percentiles of the ensemble of inversions for each day. The correlation between the two is
 359 nearly perfect ($r^2 = 0.97$), implying that the daily variability observed by the inversion is
 360 caused by irregular spatiotemporal sampling (aliasing) of the rapidly changing underlying
 361 emissions.

362 These results, showing that aliasing of large hourly variability in emissions from large

363 CO₂ sources can explain 97% of the variability in our CO₂ emissions estimate, suggest that
364 similar spatiotemporal variability in CO and CH₄ sources could explain the variability in our
365 estimates for those gases as well. We note that for CH₄ this is less clear due to the larger
366 estimated uncertainty in the posterior emissions, but it is plausible given that large temporal
367 variability in CH₄ source emissions has been reported in oil and gas production fields,^{27,52}
368 and likely exists in urban areas as well.

369 **Path Forward**

370 Flight campaigns are extremely useful for greenhouse gas (GHG) and pollutant emissions
371 estimation because of the fast deployment and large spatial coverage that is provided by
372 a moving platform. However, they are limited by the reduced temporal coverage as well
373 as the difficulty of measuring all the areas at the same time. We have shown that this
374 irregular sampling (in time and space) generates aliasing of the emissions impacting both
375 the emissions estimates and the variability of those estimates. Therefore, moving forward,
376 multiple flights over a region over different hours, days, months and seasons are recommended
377 as well as multiple aircraft flying together with well-coordinated flight plans based on forecast
378 back-trajectories so that the coverage of the cities can be maximized at all times. Addition
379 of measurements from every platform (surface, aerial or from space) available should also
380 help reduce the aliasing of emissions. This aliasing of emissions is likely not exclusive to
381 aircraft campaigns but rather a ubiquitous problem to all monitoring systems based on
382 spatiotemporally discrete sampling (aircraft, cars, polar orbiting satellites as well as sparse
383 tower networks) and it must be considered when designing the measurements and interpreting
384 the results.

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389 Supporting Information Available

- 390 • supplemental-information.pdf: Detailed methodology and supplemental tables and fig-
391 ures

392 This material is available free of charge via the Internet at <http://pubs.acs.org/>.

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609 Graphical TOC Entry

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